

# ECON42720 Causal Inference and Policy Evaluation

## 2 Regression Recap

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# About this Lecture

**Linear regression** is by far the most important **estimation technique in causal inference**

Yes, I know, **machine learning is all the rage these days**

- ▶ But a linear approximation is often the best we can do
- ▶ It's already hard enough to get the linear approximation right
- ▶ Fancy techniques are not always better

# Resources

The material behind these slides can be found in any good econometrics textbook. For introductory econometrics, I recommend

- ▶ Wooldridge, J. Introductory Econometrics: A Modern Approach. 7th Edition. South-Western College Publishing, 2019.
- ▶ Stock, J. and M. Watson. Introduction to Econometrics. 3rd Edition. Pearson, 2017.

# Regression recap

Why do we use linear regression?

- ▶ What we want to approximate: the conditional expectation function (CEF)

How does regression work?

- ▶ Interpretation of coefficients with and without controls
- ▶ Estimation with OLS and Inference

# Undergrad Recap: Goal of Linear Regression

Quantify the **expected effect** of a **one unit change in  $X$  on  $Y$**

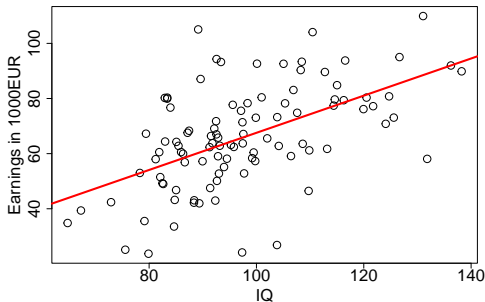
- ▶ If  $X$  goes up by one unit, by how many units does  $Y$  go up or down?
- ▶ **Causal interpretation:** If we/nature/an experimenter changes  $X$  by one unit, what is the expected effect on  $Y$ ?

This **effect** is equivalent to the **slope** coefficient  $\beta_1$  in a **linear regression model**

$$Y = \beta_0 + \beta_1 X + u$$

# Linear Regression

Regression analysis means that we **fit a straight line** through  $(X, Y)$  data points



In this example, the **regression line** is  $\text{Earnings} = -3000 + 700 \text{ IQ}$

- An increase in the IQ by one point increases earnings on average by 700 EUR

# What we are looking for: the conditional expectation function

In undergraduate econometrics, you probably learned about a **population regression model**

- ▶ the idea is that there is a **true relationship** between  $X$  and  $Y$  that we want to estimate
- ▶ we have a **sample** of  $n$  observations from this population and estimate  $\hat{\beta}_0, \hat{\beta}_1$  using OLS

But is the **population regression model (PRM)** really what we are looking for?

- ▶ Yes and no. The PRM is an approximation of our object of interest
- ▶ This object is called the **conditional expectation function (CEF)**

## Ingredients: Random Variables

**Random variables** are variables that take on different values with a certain probability

- ▶  $x$  is a random variable that takes on values  $x_1, x_2, \dots, x_n$  with probabilities  $f(x_1), f(x_2), \dots, f(x_n)$

**Expected value:** the average value of a random variable

$$\begin{aligned} E(x) &= x_1 f(x_1) + x_2 f(x_2) + \dots + x_n f(x_n) \\ &= \sum_{j=1}^n x_j f(x_j) \end{aligned}$$



## Expected Value: Example

$x \in \{-1, 0, 2\}$  with probabilities  $f(-1) = 0.3$ ,  $f(0) = 0.3$ ,  $f(2) = 0.4$ . The expected value of  $X$  is

$$\begin{aligned} E(x) &= (-1)(0.3) + (0)(0.3) + (2)(0.4) \\ &= 0.5 \end{aligned}$$

# Notation

We denote **random variables with lower case letters**  $x, y$ .

Typically, we **do not use indices when we talk about the population**. For example, the linear relationship between  $x$  and  $y$  in the population is

$$y = \beta_0 + \beta_1 x + u$$

Realisations of a random variable are denoted with **lower case letters with indices**  $x_i$  with  $i = 1, \dots, n$ . We also use indices when

- ▶ Talking about **relationships in the sample**
- ▶ Talking about particular realisations of a random variable:  $x_i = x$ , for example  $x_i = \textit{female}$  or  $x_i = 5$

This can be confusing at the start but you'll get used to it!

# The Conditional Expectation Function

We are interested in **explaining the relationship between  $x$  and  $y$**  in the population

A useful concept in this regard is the **Conditional Expectation Function (CEF)**:

$$E(y_i|x_i)$$

- ▶ what is the **population average of  $y_i$  for a given value of  $x_i$**
- ▶ i.e. what if  $x_i$  takes value  $x$ ?

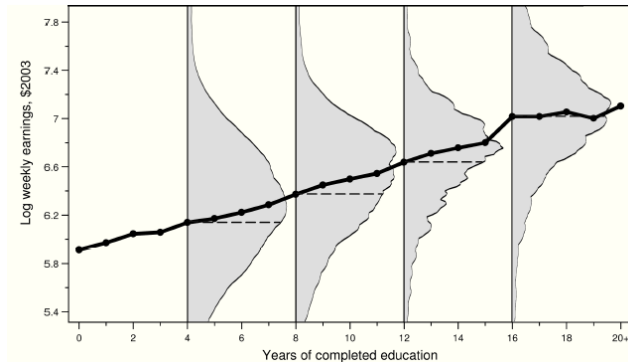
**Example:**  $x$  is a dummy that equals one if a person is female and zero otherwise.  $y$  is earnings.

The **CEF can take on two values:**

- ▶ Average earnings of women  $E(y_i|x_i = 1)$
- ▶ Average earnings of men/other  $E(y_i|x_i = 0)$

# A Continuous CEF

Education vs earnings (from Angrist & Pischke, MHE)



At every level of completed education, we have a different expected value of earnings

- ▶ At each value  $x_i$  we have a distribution of  $y_i$
- ▶ and the CEF comprises the averages of this distribution

# Regression is a Linear Approximation of the CEF

The **population regression model** (PRM) is a linear approximation of the CEF

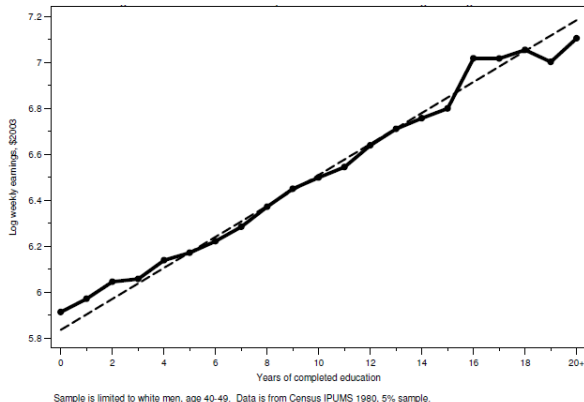
$$y = \beta_0 + \beta_1 x + u$$

- ▶ Our **ultimate goal** is to know the **CEF**
- ▶ But: with a linear regression, we can estimate the parameters of the PRM
- ▶ I.e. we **cannot estimate the CEF directly**

Why a **linear approximation is useful**:

- ▶ We typically have small sample sizes, so approximating a non-linear function is difficult
- ▶ We are often interested in marginal effects, so a linear approximation is often sufficient

## PRM: Linear Approximation of the CEF



The dashed line is the Population regression model  $y = \beta_0 + \beta_1 x + u$ . The solid line is the CEF  $E(y|x)$ .

It can be shown that the PRM is the **best linear approximation of the CEF** (see Angrist & Pischke, MHE, ch. 3).

## CEF and PRM: what's all this about?

The **CEF is the object of interest** in (most of) econometrics

- ▶ The PRM  $y = \beta_0 + \beta_1 x + u$  is a **linear approximation** of the CEF.
- ▶ But it is an approximation, so it can be wrong.
- ▶ With data, we can draw inference on the PRM but not on the CEF

**What does this mean for the empirical analysis?**

- ▶ We need to think about the **relationship between  $x$  and  $y$**  in the population
- ▶ Linear approximations are more innocuous when we consider small changes in  $x$

# The Sample Regression Function

Now suppose we have a **sample of size  $n$**  that was **randomly sampled from the population**,  $(y_1, x_1), \dots, (y_n, x_n)$

The **sample regression function** is

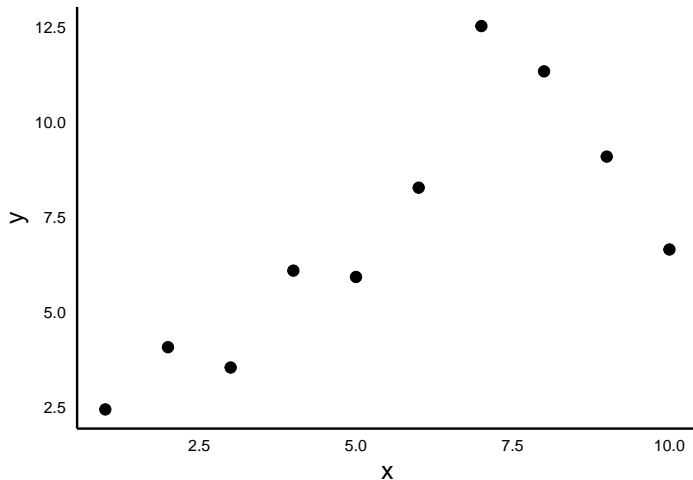
$$\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i \quad (1)$$

We can estimate the parameters  $\hat{\beta}_0$  and  $\hat{\beta}_1$  with Ordinary Least Squares (OLS)



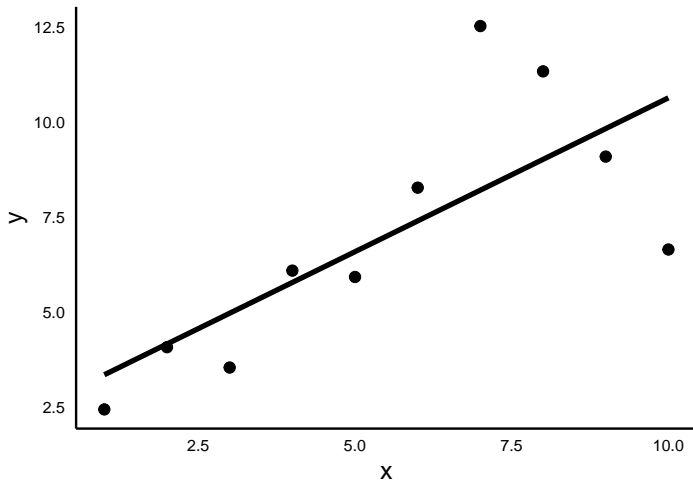
## OLS: Intuition

Let's start with some data points



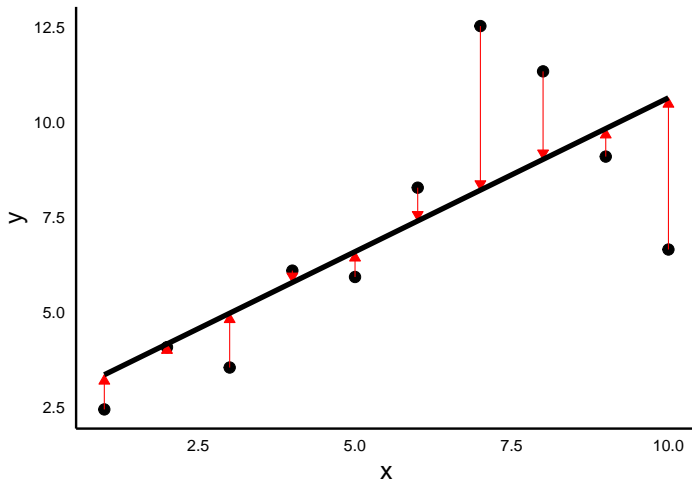
## OLS: Intuition

Goal: fit a regression line through those points



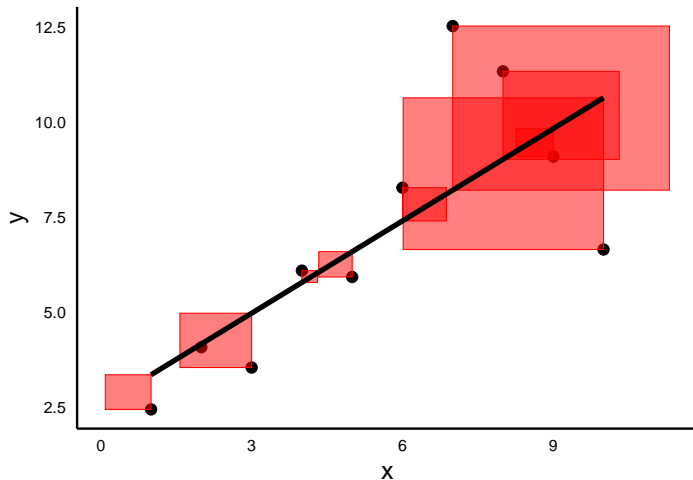
## OLS: Intuition

The key ingredient of OLS are the residuals  $\hat{u}_i = y_i - \hat{b}_0 - \hat{b}_1 x_i$



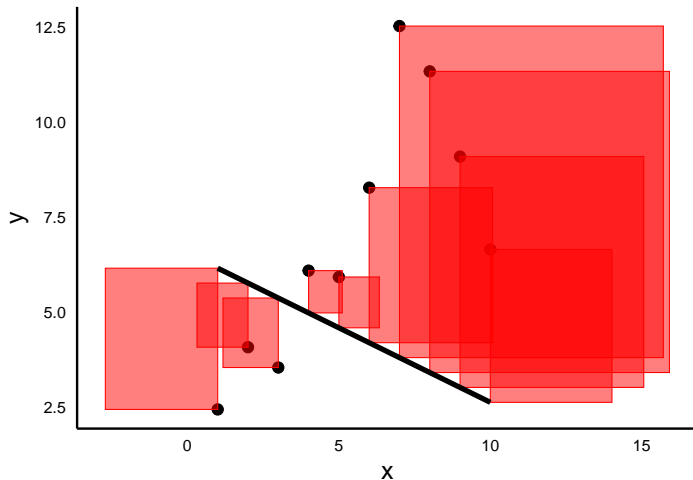
## OLS: Intuition

Now consider the square of each residual



## OLS: Intuition

Let's consider a different regression line: the squares are much larger!



# OLS: Intuition

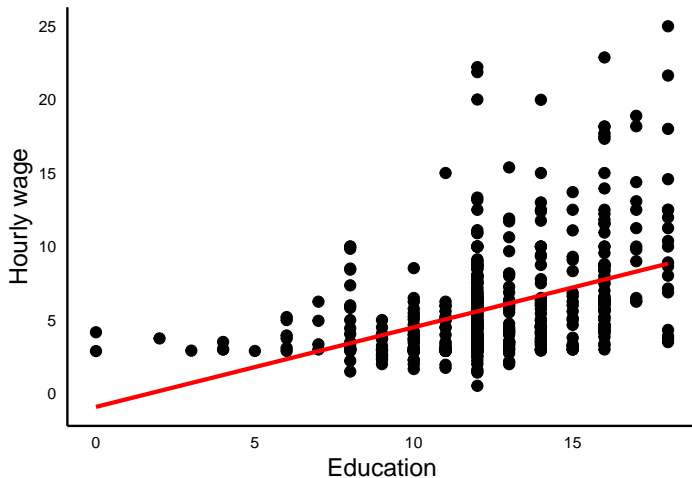
OLS **minimizes the average size of these squares**

It **minimizes the sum of squared residuals (SSR)**

The result is the **best-fitting line** that describes the relationship between  $x$  and  $y$  in the sample

## OLS: Data Example

Let's look at the relationship between education and wages with data from the U.S.



# Regression Output: Interpretation

Table 1: Effect of Education on Wages

|                         | <i>Dependent variable:</i>     |
|-------------------------|--------------------------------|
|                         | wage                           |
| educ                    | 0.541 ***<br>(0.053)           |
| Constant                | -0.905<br>(0.685)              |
| Observations            | 526                            |
| R <sup>2</sup>          | 0.165                          |
| Adjusted R <sup>2</sup> | 0.163                          |
| Residual Std. Error     | 3.378 (df = 524)               |
| F Statistic             | 103.363 *** (df = 1; 524)      |
| Note:                   | * p<0.1; ** p<0.05; *** p<0.01 |

A 1-year increase in education is associated with a 0.54 USD increase in hourly wages



## OLS: The Math

The Ordinary Least Squares (OLS) estimators  $\widehat{\beta}_0$  and  $\widehat{\beta}_1$  are derived through the minimization problem

$$(\widehat{\beta}_0, \widehat{\beta}_1) = \arg \min_{\widehat{b}_0, \widehat{b}_1} \sum_{i=1}^n [(y_i - \widehat{b}_0 - \widehat{b}_1 x_i)^2] \quad (2)$$

The sample means  $\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$  and  $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$  are the sample analogs of the population means  $E(y_i)$  and  $E(x_i)$

## OLS: The Math

The **residuals of the regression** are defined as  $\hat{u}_i = y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i$ .

When we “run an OLS regression”, we **minimize the sum of squared residuals (SSR)**,  $\sum_{i=1}^n \hat{u}_i^2$  and obtain values for  $\hat{\beta}_0$  and  $\hat{\beta}_1$

Solving the minimization problem (2) yields the **estimators**

$$\begin{aligned}\hat{\beta}_1 &= \frac{\frac{1}{n} \sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})}{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\widehat{Cov}(y_i, x_i)}{\widehat{V}(x_i)} \\ \hat{\beta}_0 &= \bar{y} - \hat{\beta}_1 \bar{x}\end{aligned}\tag{3}$$

# The Sampling Distribution of the OLS Estimator

To **draw inference about the population**, we need to know the **sampling distribution of the OLS estimator**

What we want to know:

- ▶ When will  $\hat{\beta}_1$  be **unbiased**?
- ▶ What is its **variance**?

To answer these questions, we need to make some **assumptions about the sample and population**

1. Population model is linear in parameters
2. Sample is randomly drawn from the population
3. Variation in  $x$
4. **Zero conditional mean assumption (ZCM)**

# OLS Assumptions

The four **OLS assumptions must be fulfilled** for the OLS estimator to be **unbiased and consistent**

**Unbiasedness:**  $E(\hat{\beta}_1) = \beta_1$

- ▶ across many random samples, the estimator gets it right on average

**Consistency:**  $\hat{\beta}_1 \xrightarrow{P} \beta_1$  as  $n \rightarrow \infty$

- ▶ if the sample size increases, the estimator converges to the true value
- ▶ this is a consequence of the Law of Large Numbers (LLN)
- ▶ As  $n$  gets larger, the sample becomes more representative of the population

# The Zero Conditional Mean (ZCM) Assumption

The **ZCM assumption is the most important assumption** in this module

- ▶ It is **not testable with the data** at hand
- ▶ It rarely holds in practice (except in randomised experiments)
- ▶ Causal inference techniques exploit scenarios where ZCM holds approximately

**Other names for the ZCM assumption:**

- ▶ **Conditional independence assumption (CIA)**
- ▶ **Exogeneity assumption**

# The Zero Conditional Mean (ZCM) Assumption

Consider the **population model**

$$y = \beta_0 + \beta_1 x + u$$

The **ZCM assumption** states that the **conditional mean of the error term is zero**

$$E(u|x) = E(u) = 0$$

What does this mean?

- ▶ The error term is **not systematically related** (speak: uncorrelated) with  $x$
- ▶ At any level of  $x$ , the average value of  $u$  is zero

## ZCM Assumption: Example

Does **higher education (causally) increase earnings?**

$$wage_i = \beta_0 + \beta_1 education_i + u_i$$

**What is the error term  $u_i$  here?**

- ▶ Any determinant of a person's wage that is **not education**
- ▶ E.g., innate ability, motivation, personality, etc.

## ZCM Assumption: Example

**Say  $u_i$  includes ability.** According to the ZCM assumption, the following must hold:

$$E(\text{ability} \mid \text{education} = 8) = E(\text{ability} \mid \text{education} = 12) = E(\text{ability} \mid \text{education} = 16)$$

So it **must hold that**:

- ▶ the average ability of people with 8 years of education is the same as
- ▶ the average ability of people with 12 years of education and
- ▶ the average ability of people with 16 years of education

This is hardly plausible  $\Rightarrow$  **ZCM assumption is violated**



## What if ZCM is Violated?

The OLS estimator of  $\beta_1$  is **biased and inconsistent**

**Bias:**  $E(\widehat{\beta}_1) \neq \beta_1$

- ▶ The expected value of the OLS estimator is not equal to the true value of  $\beta_1$
- ▶ Across many samples, the estimates are systematically too big or too small

**Inconsistency:**  $\widehat{\beta}_1$  does not converge to  $\beta_1$  as  $n \rightarrow \infty$

- ▶ Even if the sample size is very large, the OLS estimator does not converge to the true value of  $\beta_1$

# How to think about Violations of ZCM

The variable  $x$  is **typically a choice**

- ▶ People choose how much education to get, how often they go to the gym, how much they save, who they want to date, etc
- ▶ Firms choose how much to invest, how many workers to hire, how much to pollute, etc.
- ▶ Governments choose how much to spend on education, how much to tax, etc.

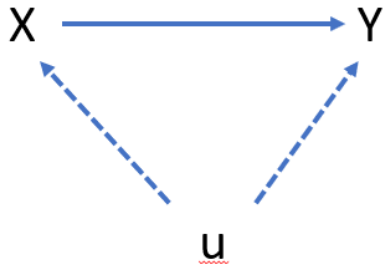
The **choice** of  $x$  is typically **influenced by other factors**  $u$

- ▶ Individual factors: ability, motivation, preferences, etc.
- ▶ Firm factors: technology, market conditions, etc.
- ▶ Government factors: ideology, political pressure, etc.

**Problem:**  $u$  **affects**  $y$  not just through  $x$  but also **through other paths** or directly

# How to think about Violations of ZCM

The error term  $u$  includes **one or more confounders**



Here  $u$  includes a confounder  $y$  directly and through  $x$

Example:  $x$  is education,  $y$  is earnings,  $u$  is ability

## Omitted Variable Bias

Suppose the **true model** is  $y = \beta_0 + \beta_1 x + \beta_2 s_1 + e$

However, we **estimate the model**  $y = \tilde{\beta}_0 + \tilde{\beta}_1 x + u$

It can be shown that the **OLS estimator is biased**

$$\tilde{\beta}_1 = \beta_1 + \underbrace{\beta_2 \frac{\text{Cov}(x, s_1)}{\text{Var}(x)}}_{OVB}$$

## So when does ZCM hold?

ZCM holds if  $x$  is **as good as randomly assigned** to individuals

- ▶ This is the case if  $x$  is assigned in a **randomised experiment**
- ▶ Or if  $x$  is assigned in a **quasi-experiment** that mimics random assignment
- ▶ Or if we can **control for all confounders** in the analysis

We should **always assume that ZCM is violated**. Researchers need to **think hard about confounders and how to eliminate them**.

# Controlling for Confounders: Multivariate Regression

We can include **confounders in the regression model** to control for them

$$y = \beta_0 + \beta_1 x + \mathbf{S}\boldsymbol{\gamma} + u$$

Here,  $\mathbf{S}$  is a vector of covariates  $\mathbf{S} = (s_1, s_2, \dots, s_k)$  and  $\boldsymbol{\gamma} = (\gamma_1, \gamma_2, \dots, \gamma_k)$  is a vector of coefficients,<sup>1</sup> i.e.

$$\mathbf{S}\boldsymbol{\gamma} = \gamma_1 s_1 + \gamma_2 s_2 + \dots + \gamma_k s_k$$

We are **only interested in  $\beta_1$ , the causal effect of  $x$  on  $y$**

- ▶ The other coefficients  $\gamma_1, \gamma_2, \dots, \gamma_k$  are not of interest (nuisance parameters)
- ▶ We include the covariates  $\mathbf{S}$  to control for confounders

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<sup>1</sup>Note: each element of  $\mathbf{S}$  is in itself an  $(n \times 1)$  vector, so  $\mathbf{S}$  is actually an  $(n \times k)$  matrix

# Interpretation of $\beta_1$ in Multivariate Regression

$\beta_1$  now has a **ceteris paribus interpretation**

- ▶ **Holding all other variables  $S$  constant**, a one unit increase in  $x$  leads to a  $\beta_1$  unit increase in  $y$

The **inclusion of  $S$**  allows for a **like-with-like comparison**

- ▶ We **compare units with the same values** of  $S$  but different values of  $x$
- ▶ But the like-with like comparison is only valid if  $S$  contains all confounders

# Conditional Mean Independence Assumption

The **Conditional Mean Independence Assumption (CMIA)** is a “light” version of the ZCM assumption.

$$E(u \mid x, \mathbf{S}) = E(u \mid \mathbf{S}) = 0$$

In plain English: as long as  **$\mathbf{S}$  is included, the error term  $u$  is uncorrelated with  $x$**

- ▶  $x$  is exogenous conditional on  **$\mathbf{S}$**
- ▶  $x$  is as good as random conditional on  **$\mathbf{S}$**



# Summary: What you need to understand for this module

## Logic of linear regression

- ▶ Why we use linear regression
- ▶ Why and how we use OLS to estimate the parameters of the linear regression model
- ▶ How to interpret the OLS estimator  $\hat{\beta}_1$

## Limitations of linear regression for causal inference

- ▶ The ZCM assumption is violated in most applications, leading to OVB
- ▶ How control variables can be used to control for confounders

# Appendix

# Regression with R

```
# Required packages (install if necessary)  
library(tidyverse)  
library(wooldridge)  
library(stargazer)
```

# Regression with R

This code shows how to estimate and present regressions with R

```
data('wage1') # load the data
df <- wage1
reg1 <- lm(wage ~ educ, data = df) # estimate simple regression
reg2 <- lm(wage ~ educ + exper, data = df) # estimate multivariate regression
stargazer(reg1, reg2, type = "text") # print regression results
```

## Regression outputs with Stargazer and R Markdown

You can generate nice regression tables with the `stargazer` package and R Markdown/Quarto. Here is a code chunk that gives you a nicely formatted latex table.

```
```{r, results='asis', echo=FALSE}  
stargazer(reg1, reg2,  
          header=FALSE,  
          type='latex',  
          title="Effect of Education on Wages")  
```
```

## Regression with R

We can also generate a scatter with a regression line

```
ggplot(df, aes(x = educ, y = wage)) + # generate scatter plot
  geom_point() + # add points
  geom_smooth(method = "lm", se = FALSE) + # add regression line
  theme_minimal()
```

# Contact

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