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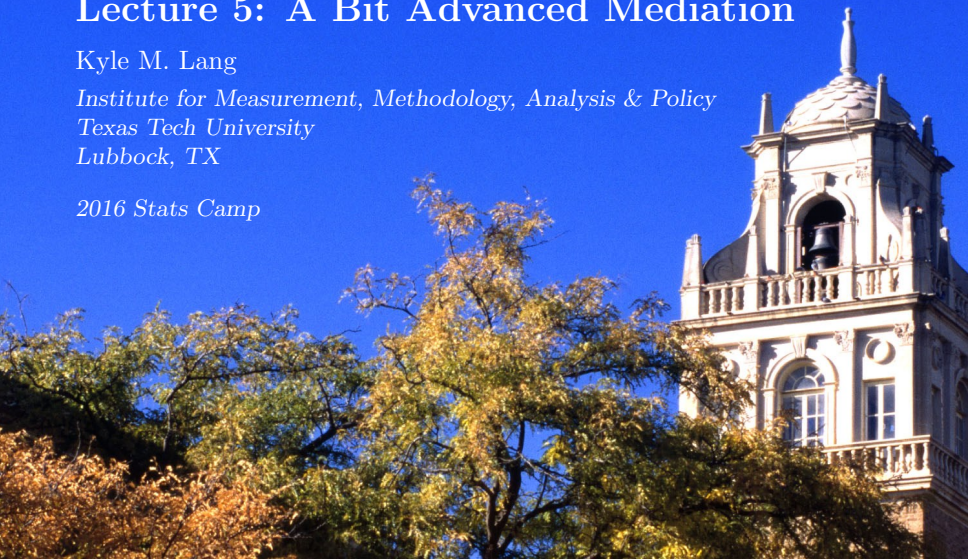


Lecture 5: A Bit Advanced Mediation

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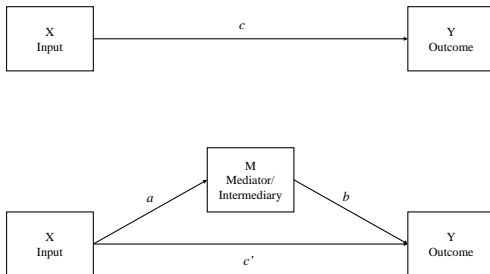
Institute for Measurement, Methodology, Analysis & Policy
Texas Tech University
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2016 Stats Camp



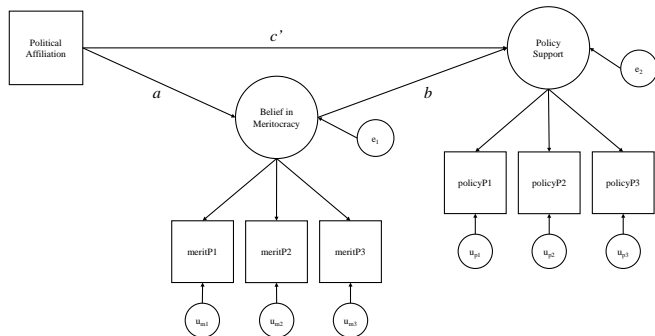
- Show how to test for indirect effects in latent variable models
- Discuss the interpretation of indirect effects
- Discuss effect size measures for indirect effects

So far, all of our models have been similar to:



But there is no reason that we need to restrict ourselves to mucking about with observed variables.

We can (and should) test for indirect effects using *latent variable models* such as:



Measurement error can be a big problem for mediation analysis, so latent variable modeling is highly recommended.

Example



```
library(lavaan)
dataDir <- "../data/"
dat1 <- readRDS(paste0(dataDir, "adamsKlpsData.rds"))
## Specify the CFA model:
mod1.1 <- "
merit =~ meritP1 + meritP2 + meritP3
policy =~ policyP1 + policyP2 + policyP3
"

## Fit the CFA and check model:
out1.1 <- cfa(mod1.1, data = dat1, std.lv = TRUE)
## Check model fit:
round(fitMeasures(out1.1)[c("chisq", "df", "pvalue", "cfi",
                             "tli", "rmsea", "srmr")], 4)
```

chisq	df	pvalue	cfi	tli	rmsea	srmr
16.8695	8.0000	0.0315	0.9215	0.8529	0.1129	0.0653

```
summary(out1.1)
```

Example



```
lavaan (0.5-20) converged normally after 22 iterations
```

Number of observations	87
Estimator	ML
Minimum Function Test Statistic	16.869
Degrees of freedom	8
P-value (Chi-square)	0.031

Parameter Estimates:

Information	Expected
Standard Errors	Standard

Latent Variables:

	Estimate	Std.Err	Z-value	P(> z)
merit =~				
meritP1	0.690	0.134	5.155	0.000
meritP2	0.968	0.142	6.830	0.000
meritP3	0.748	0.137	5.458	0.000
policy =~				
policyP1	0.851	0.186	4.570	0.000

Example



policyP2	0.996	0.167	5.967	0.000
policyP3	1.121	0.177	6.339	0.000

Covariances:

	Estimate	Std.Err	Z-value	P(> z)
merit ~ policy	-0.336	0.131	-2.563	0.010

Variances:

	Estimate	Std.Err	Z-value	P(> z)
meritP1	0.865	0.165	5.248	0.000
meritP2	0.445	0.201	2.211	0.027
meritP3	0.833	0.172	4.857	0.000
policyP1	1.836	0.324	5.671	0.000
policyP2	0.942	0.256	3.683	0.000
policyP3	0.857	0.297	2.882	0.004
merit	1.000			
policy	1.000			

Example



```
## Specify the structural model:
mod1.2 <- "
merit =~ meritP1 + meritP2 + meritP3
policy =~ policyP1 + policyP2 + policyP3

policy ~ b*merit + polAffil
merit ~ a*polAffil

ab := a*b
"

## Fit the structural model and test the indirect effect:
out1.2 <- sem(mod1.2, data = dat1, std.lv = TRUE,
              se = "boot", boot = 2500)
summary(out1.2)
```


Example



```
lavaan (0.5-20) converged normally after 24 iterations
```

```
Number of observations                        87
```

```
Estimator                                    ML
```

```
Minimum Function Test Statistic             20.665
```

```
Degrees of freedom                          12
```

```
P-value (Chi-square)                        0.056
```

```
Parameter Estimates:
```

```
Information                                Observed
```

```
Standard Errors                           Bootstrap
```

```
Number of requested bootstrap draws        2500
```

```
Number of successful bootstrap draws        2475
```

```
Latent Variables:
```

	Estimate	Std.Err	Z-value	P(> z)
merit =~				
meritP1	0.545	0.125	4.374	0.000
meritP2	0.858	0.150	5.737	0.000
meritP3	0.609	0.116	5.251	0.000

Example



policy =~				
policyP1	0.799	0.193	4.141	0.000
policyP2	0.924	2.899	0.319	0.750
policyP3	1.001	2.037	0.491	0.623

Regressions:

		Estimate	Std.Err	Z-value	P(> z)
policy ~					
merit	(b)	-0.195	0.202	-0.962	0.336
polAffil		0.169	0.139	1.217	0.224
merit ~					
polAffil	(a)	-0.411	0.100	-4.097	0.000

Variances:

	Estimate	Std.Err	Z-value	P(> z)
meritP1	0.922	0.181	5.109	0.000
meritP2	0.341	0.460	0.740	0.459
meritP3	0.869	0.185	4.700	0.000
policyP1	1.801	0.337	5.352	0.000
policyP2	0.918	173.253	0.005	0.996
policyP3	0.922	140.529	0.007	0.995
merit	1.000			
policy	1.000			

Example



Defined Parameters:

	Estimate	Std.Err	Z-value	P(> z)
ab	0.080	0.090	0.890	0.373

```
parameterEstimates(out1.2, boot = "bca.simple")[-c(1 : 3)]
```

	label	est	se	z	pvalue	ci.lower	ci.upper
1		0.545	0.125	4.374	0.000	0.337	0.831
2		0.858	0.150	5.737	0.000	0.630	1.119
3		0.609	0.116	5.251	0.000	0.383	0.833
4		0.799	0.193	4.141	0.000	0.422	1.146
5		0.924	2.899	0.319	0.750	0.662	1.424
6		1.001	2.037	0.491	0.623	0.575	1.783
7	b	-0.195	0.202	-0.962	0.336	-0.610	0.173
8		0.169	0.139	1.217	0.224	-0.089	0.455
9	a	-0.411	0.100	-4.097	0.000	-0.642	-0.235
10		0.922	0.181	5.109	0.000	0.600	1.304
11		0.341	0.460	0.740	0.459	-0.100	0.679
12		0.869	0.185	4.700	0.000	0.562	1.308
13		1.801	0.337	5.352	0.000	1.226	2.583
14		0.918	173.253	0.005	0.996	0.342	1.507

Example



15		0.922	140.529	0.007	0.995	-0.038	1.614
16		1.000	0.000	NA	NA	1.000	1.000
17		1.000	0.000	NA	NA	1.000	1.000
18		2.444	0.000	NA	NA	2.444	2.444
19	ab	0.080	0.090	0.890	0.373	-0.063	0.275

Interpretation of Indirect Effects

Although indirect effects are composed parameters, they have direct interpretations, independent of the interpretations of their constituent paths:

- The $X \rightarrow M \rightarrow Y$ indirect effect ab is interpreted as:
 - The expected change in Y for a unit change in X that is transmitted indirectly through M , or...
 - For a unit change in X , Y is expected to change by ab units, indirectly through M , or...
 - Participants who differ by one unit on X are expect to differ by ab units on Y as a results of the effect of X on M which, in turn, affects Y .
- The interpretation/scaling of the indirect effect is entirely defined by the input X and outcome Y
 - The scaling of the intermediary variable M does not affect the interpretation of the indirect effect.

Partially Standardized Indirect Effect



$$ab_{ps} = \frac{ab}{SD_Y}$$
$$c'_{ps} = \frac{c'}{SD_Y}$$
$$c_{ps} = \frac{c}{SD_Y} = ab_{ps} + c'_{ps}$$

- Simple
- Removes binding to the scale of Y
- Still scale-bound by X
- Not clear what constitutes a “large” effect

Completely Standardized Indirect Effect



$$ab_{cs} = \frac{SD_X ab}{SD_Y}$$

$$c'_{cs} = \frac{SD_X c'}{SD_Y}$$

$$c_{cs} = \frac{SD_X c}{SD_Y} = ab_{cs} + c'_{cs}$$

- Simple
- Removes all scale binding
- Not clear what constitutes a “large” effect

$$P_M = \frac{ab}{c} = \frac{ab}{c' + ab}$$

- Very simple
- Not bounded by 0 and 1
- Explodes toward $\pm\infty$ as $c \rightarrow 0$
- Very unstable
 - High between-sample variability
 - Requires $N \geq 500$

$$R_M = \frac{ab}{c'} = \frac{P_M}{1 - P_M}$$

- Very simple
- Not bounded by 0 and 1
- Explodes toward $\pm\infty$ as $c' \rightarrow 0$
- Very unstable
 - High between-sample variability
 - Requires $N \geq 2000$

Developed by Fairchild, MacKinnon, Taborga, and Taylor (2009).

- Given a non-zero total effect, represents the proportion of variance in Y accounted for by the indirect effect.

$$R_{med}^2 = r_{MY}^2 - (R_{Y.MX}^2 - r_{XY}^2)$$

- Mostly sensible interpretation
- Predicated on the assumption that $\beta_{YX} \neq 0$
- $|ab| > |c| \Rightarrow R_{med}^2 < 0$
 - Not a strict proportion

Developed by Preacher and Kelley (2011).

- Gives the proportion of the *maximum possible* indirect effect represented by ab .

$$\kappa^2 = \frac{ab}{\max(ab)}$$

- Bounded by 0 and 1
- Values closer to 1.0 indicate a bigger effect
- A bit of a pain to calculate.

Computing $\max(ab)$



$$a \in \left\{ \frac{\sigma_{YM}\sigma_{YX} \pm \sqrt{\sigma_M^2\sigma_Y^2 - \sigma_{YM}^2} \sqrt{\sigma_X^2\sigma_Y^2 - \sigma_{YX}^2}}{\sigma_X^2\sigma_Y^2} \right\} = [a_{low}, a_{high}],$$

$$b \in \left\{ \pm \frac{\sqrt{\sigma_X^2\sigma_Y^2 - \sigma_{YX}^2}}{\sqrt{\sigma_X^2\sigma_M^2 - \sigma_{MX}^2}} \right\} = [b_{low}, b_{high}],$$

$$\max(a) = \begin{cases} a_{high}, & \text{if } \hat{a} > 0 \\ a_{low}, & \text{if } \hat{a} < 0 \end{cases}, \quad \max(b) = \begin{cases} b_{high}, & \text{if } \hat{b} > 0 \\ b_{low}, & \text{if } \hat{b} < 0 \end{cases},$$

$$\max(ab) = \max(a)\max(b)$$

Example



```
## Specify the model:
mod2 ← "
policy ~ b*sysRac + cp*polAffil
sysRac ~ a*polAffil

ab := a*b
"

## Estimate the model:
out2 ← sem(mod2, data = dat1)
##
## Extract/compute the necessary quantities:
ab ← prod(coef(out2)[c("a", "b")])
ab
```

```
[1] 0.1015958
```

Example



```
cPrime ← coef(out2)["cp"]  
##  
sdY ← sd(dat1$policy)  
sdX ← sd(dat1$polAffil)  
##  
r2MY ← with(dat1, cor(policy, sysRac))^2  
r2XY ← with(dat1, cor(policy, polAffil))^2  
R2Y.MX ← inspect(out2, "r2")["policy"]
```

Example



```
## Partially Standardized:  
abPS ← ab / sdY  
abPS
```

```
[1] 0.08559454
```

```
cPrimePS ← cPrime / sdY  
cPrimePS
```

```
cp  
0.1138675
```

```
cPS ← abPS + cPrimePS  
cPS
```

```
cp  
0.199462
```

Example



```
## Completely Standardized:  
abCS ← (sdX * ab) / sdY  
abCS
```

```
[1] 0.1345859
```

```
cPrimeCS ← (sdX * cPrime) / sdY  
cPrimeCS
```

```
cp  
0.1790413
```

```
cCS ← abCS + cPrimeCS  
cCS
```

```
cp  
0.3136272
```


Example



```
## Proportions:  
pm ← ab / (cPrime + ab)  
pm
```

```
      cp  
0.429127
```

```
rm ← ab / cPrime  
rm
```

```
      cp  
0.7517031
```

```
## R2:  
R2med ← r2MY - (R2Y.MX - r2XY)  
R2med
```

```
      policy  
0.06905689
```

Compute κ^2



```
## Subset the data:
tmpData <- dat1[ , c("polAffil", "sysRac", "policy")]
colnames(tmpData) <- c("x", "m", "y")
##
## Extract pertinent variance/covariance elements:
cov1 <- cov(tmpData)
sYM <- cov1["x", "m"]
sYX <- cov1["y", "x"]
sMX <- cov1["m", "x"]
s2X <- cov1["x", "x"]
s2M <- cov1["m", "m"]
s2Y <- cov1["y", "y"]
```

Compute κ^2



```
## Possible range of a:
```

```
aMarg ← sqrt(s2M * s2Y - sYM^2) * sqrt(s2X * s2Y - sYX^2)
aInt ← c(
  (sYM * sYX - aMarg) / (s2X * s2Y),
  (sYM * sYX + aMarg) / (s2X * s2Y)
)
aInt
```

```
[1] -0.4378558  0.5793099
```

```
##
```

```
## Possible range of b:
```

```
bMarg ← sqrt(s2X * s2Y - sYX^2) / sqrt(s2X * s2M - sMX^2)
bInt ← c(-1 * bMarg, bMarg)
bInt
```

```
[1] -1.289996  1.289996
```

Compute κ^2



```
##  
## max(a):  
aMax ← ifelse(coef(out2)["a"] < 0,  
              aInt[1],  
              aInt[2])  
  
aMax
```

```
      a  
0.5793099
```

```
##  
## max(b)  
bMax ← ifelse(coef(out2)["b"] < 0,  
              bInt[1],  
              bInt[2])  
  
bMax
```

```
      b  
1.289996
```

Compute κ^2



```
##  
## max(ab)  
abMax ← aMax * bMax  
abMax
```

```
      a  
0.7473075
```

```
##  
## Kappa Squared:  
k2 ← ab / abMax  
k2
```

```
      a  
0.1359491
```

Suppose:

1. Σ is given by:

	x	m	y
x	1.5		
m	0.3	1.4	
y	0.6	0.45	1.55

2. The estimated paths are:

- $a = 0.2$
- $b = 0.246$
- $ab = 0.049$

Compute κ^2 for the estimated ab .

- Fairchild, A. J., MacKinnon, D. P., Taborga, M. P., & Taylor, A. B. (2009). R-squared effect-size measures for mediation analysis. *Behavior Research Methods*, 41(2), 486–498.
- Preacher, K. J., & Kelley, K. (2011). Effect size measures for mediation models: Quantitative strategies for communicating indirect effects. *Psychological Methods*, 16(2), 93.