

ASSOCIATIVE MEMORY AND BELIEF FORMATION^{*}

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Abstract

This paper experimentally studies the role of associative memory for expectation formation. Real-world information signals are often embedded in memorable contexts. Thus, today's news may cue the selective retrieval of similar past news and hence contribute to the widely documented pattern of expectation overreaction. Based on a stylized version of models of associative memory in the literature, we develop a simple and tightly controlled experimental setup in which participants observe sequences of news about the stock market value of hypothetical companies. Here, identical types of news are associated with identical stories and images. In this setup, participants' expectations strongly overreact to recent news, and we successfully verify the model's relatively nuanced predictions about how the magnitude of overreaction should depend on the history of news. We use our experimental data to estimate the model and further provide direct causal evidence that expectation overreaction indeed occurs because of imperfect and associative memory. These patterns hold both within an experimental session and over a time lag of three days. We also document that the effects of associative memory depend on the structure of the environment: once today's news are associated with the stories and images of previous opposite news, expectations systematically underreact.

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“One new narrative may remind of another that has been lying fairly dormant. . . there is cue-dependent forgetting.”

Robert Shiller, “Narrative Economics”, 2017

1 Introduction

This paper experimentally studies the role of imperfect and associative memory for the formation of expectations or beliefs. In the textbook model of Bayesian updating, memory plays no role: agents entertain a prior belief, update this belief upon receipt of information, and yesterday’s posterior equals today’s prior. Our paper starts from the premise that people do not constantly actively entertain beliefs about every potentially relevant state of the world. Rather, when people are prodded to act on or update their beliefs, they first need to reconstruct their prior knowledge and beliefs from memory. This observation raises the empirical question how exactly people retrieve prior information, and which features of news make it more or less likely for memory traces to get recollected.

The second observation that motivates our paper is that real-world information signals typically do not just consist of abstract information. Rather, information is often embedded in memorable contexts, including stories, images, emotions, or sounds. Moreover, *similar news are frequently embedded in similar contexts*. For example, when individuals receive negative feedback about their performance, these negative news are often associated with scolding, public shaming, and emotions of insufficiency. Similarly, when good news prevail in the stock market, people are disproportionately exposed to bulls, upward-sloping trend lines, and good-times stories (Shiller, 2017).

This motivates the question about the role of associative recall in expectation formation. The associative nature of memory has recently received increased attention in the theory literature (Mullainathan, 2002; Bordalo et al., 2017a). A central prediction that emerges from this body of work is that such asymmetric context-cued recall could lead to overreaction: after receipt of a piece of news, people reconstruct past knowledge from memory, yet predominantly remember those past news that appeared in similar contexts as today’s news. As a consequence, expectations might *look like* they overreact to recent news. However, even though modeling errors in belief formation is often argued to require knowledge about the underlying psychological micro-foundations (Fudenberg, 2006), direct empirical evidence on a potential link between overreaction and memory is scarce.

To make progress, we present laboratory experiments that are structured around the predictions of a simple formal framework that applies the idea of associative recall to belief formation, based on the formulations in Bordalo et al. (2017a) and Mullainathan (2002). In this model, decision-makers (i) have imperfect memory and (ii) are more

likely to recollect a piece of news from the past if the context in which it is experienced is similar to today’s context. This stylized version of existing models predicts overreaction in expectations, yet it also makes relatively nuanced predictions about how such overreaction depends on parameters such as the precise signal history, the correlation structure between signals and contexts, the imperfectness of memory, or the strength of associative recall. Each of our seven experimental treatments with a total of almost 500 lab subjects is designed around these predictions.

We propose a new experimental paradigm to investigate the role of associative memory for belief formation in an economic decision context. This paradigm builds a bridge between (i) the types of tightly-controlled, quantitative, and financially incentivized designs that dominate modern experimental economics research on bounded rationality and (ii) psychological research on cued recall. In our experiment, participants predict the stock market value of twelve hypothetical companies. We adopt this particular framing just to make the task intuitive for subjects, rather than because we think of our work as a finance application. The experiment comprises two distinct periods that we think of as “past” and “present.” Across both periods, a subject sequentially observes multiple pieces of news about a company on their decision screen, where each piece of news takes on the value +10 or −10. The value of a company is deterministic and given by 100 plus the sum of all news that were shown up to a given point in time. As in the motivating examples, the news are not communicated to subjects as mere abstract numbers but are embedded in a context, which consists of a story and an image that explain the piece of news. For example, for one company, a positive signal would be shown with a story about the company having launched a successful advertisement campaign with a celebrity, accompanied by a picture of that celebrity.

In the baseline version of the experiment, as in the motivating examples, identical news trigger identical contexts: there is a one-to-one mapping between {Company \times type of news} and context. That is, for each company, all positive news are communicated using the same context, and all negative news are communicated using the same context. However, the same context is never used for different types of news or for different companies. All of this is known to subjects.

In the first period of the experiment, a subject sequentially observes between zero and three pieces of news for a company and then states a first belief about the value of a company. This process is repeated for all twelve companies. After the first period of the experiment, subjects work on an unrelated real effort task for 15 minutes, which generates a time lag between the first and second period. In the second period, subjects observe up to one additional piece of news for a company and immediately after state their second-period belief about the value of that company. In addition, subjects explicitly indicate how many positive or negative signals they recall to have seen. Again, this

procedure is repeated for all twelve companies. As before, the true value of a company is given by 100 plus the sum of all signals that have been accumulated throughout the entire experiment, including in the first period.

In this setup, our interest lies in evaluating the extent to which second-period beliefs overreact with respect to the second-period signal. Because of the simple deterministic structure of the experiment, the prediction of both a rational (Bayesian) model and a model with imperfect but no associative recall is that the OLS coefficient in a regression of second-period beliefs on second-period signals equals one. In contrast, our formal framework that borrows from prior theoretical work predicts that (i) the OLS coefficient is larger than one, meaning that expectations overreact; (ii) overreaction increases in the number of first-period signals that take on the same realization as the second-period signal (because more news can be cued); (iii) overreaction disappears if memory is exogenously manipulated to be perfect; (iv) overreaction disappears if associative recall is exogenously shut down; and (v) expectations under- rather than overreact if recent news appear in a context that was previously associated with the opposite type of signal. We provide causal tests of each of these predictions. All of our experiments were pre-registered, including a pre-analysis plan.

We test predictions (i) and (ii) using the baseline treatment variation *Main* discussed above. We find that, in the experimental data, expectations strongly overreact: the aggregate OLS regression coefficient of the second-period signal is 1.12, substantially larger than its rational or imperfect-but-no-associative-recall benchmark of one. Moreover, as predicted, across tasks the magnitude of overreaction is strongly increasing in the number of first-period signals that take on the same value as the second-period signal. For instance, when subjects do not observe any first-period signals that match the second-period signal, their expectations do not overreact at all. In contrast, overreaction monotonically increases in whether one, two, or three first-period signals match the second-period signals (prediction (ii)). These patterns point to the importance of associative recall in our setup and cannot be explained by an account of recency effects.

We study heterogeneity in overreaction and its individual-level correlates. Within treatment *Main*, about 35% of subjects do not exhibit overreaction, while 65% overreact to the second-period signal to varying degrees. Using pre-registered heterogeneity analyses, we show that overreaction is stronger among subjects who score lower on a Raven matrices IQ test and exhibit more pronounced imperfect recall. In contrast, response times are only weakly related to overreaction.

To provide causal evidence for the role of imperfect and associative memory, we turn to testing predictions (iii) and (iv). To this effect, we exogenously manipulate the strength (or relevance) of both imperfect and associative memory. To show that imperfect memory is necessary in order for overreaction to arise in our setup, we introduce

treatment *Reminder*. This treatment follows exactly the same structure as condition *Main*, except that before subjects observe the second-period signal for a given company, they are reminded of their own first-period belief. Viewed through the lens of our formal framework, this treatment manipulation eliminates the imperfectness of memory, so that asymmetric recall and hence overreaction can no longer take place. We find that, in treatment *Reminder*, subjects' beliefs do indeed not overreact.

Having documented the role of imperfect memory for overreaction, we next study the role of associative recall. To show that associative memory generates overreaction in our experiments, we introduce treatment *NoCue*. This treatment follows the same structure as condition *Main*, except that each piece of news is communicated with a separate context. That is, subjects never observe the same story or image twice, even if they receive the same signal for a given company twice. Thus, viewed through the lens of our formal framework, this treatment manipulation eliminates (or at least substantially reduces) the extent to which associativeness can affect recall. The results show that overreaction disappears entirely in *NoCue*, and the treatment difference in overreaction between *Main* and *NoCue* is quantitatively large and statistically significant.

In all experiments reported above, types of news and contexts (stories and images) were connected through a one-to-one mapping. That is, all positive signals for a given company appeared with the same context, and all negative signals appeared with the same (yet different) context. In treatment *Underreaction*, we modify this correlation structure between signals contexts to test prediction (v) above. Specifically, in the second period of the experiment, positive signals are communicated along with the context that was associated with negative signals for that same company in the first period. Likewise, negative signals for a company appear in contexts that were previously associated with positive signals for that same company. In this treatment, our formal framework predicts that expectations should underreact to the last signal because it selectively cues the retrieval of signals that took on a different value than the current signal. To causally isolate the role of memory, we again implement a control condition *Underreaction reminder*, in which subjects were reminded of their first-period belief before receiving a second-period signal.

We find that expectations in *Underreaction* systematically underreact, also relative to treatment *Underreaction reminder*. Moreover, based on our simple model, we again derive and empirically verify relatively nuanced predictions about the particular signal histories that should lead to more or less underreaction, based on the logic of imperfect and associative recall. These results highlight that associative memory generates predictable patterns of over- and underreaction, purely depending on the precise ways in which contexts are linked to pieces of news.

All experiments summarized up to this point rely on a setup in which the time lag

between the first and second period is 15 minutes. In the final part of the paper, as a robustness check, we investigate the role of the length of the time lag for the relevance and quantitative importance of associative memory. We implement two additional pre-registered treatment variations. Condition *Time lag* follows the same procedures as treatment *Main*, except that the time lag between the first and second period is three days: participants complete the first part in the laboratory and then return to the lab three days later. To again isolate the causal role of memory, we also implemented treatment *Reminder time lag*, which is identical to *Time lag*, except that subjects are reminded of their first-period belief before they receive a second-period piece of news. The results are again in line with our simple model and very similar to the results in the baseline treatments.

All of our main results are derived from theoretically-motivated but ultimately reduced-form regression specifications. Thus, in supplementary analyses in the final part of the paper, we directly estimate our stylized model of memory and belief formation, in particular the parameters that govern the imperfectness of memory and the strength of associative recall. The results of these estimations suggest that associative recall plays a quantitatively large role in generating observed beliefs.

In summary, the central contribution of our paper is a theoretically-structured experimental analysis of the role of associative recall for belief formation. This paper hence fits into an emerging literature that has argued for the importance of associative memory for economics. [Mullainathan \(2002\)](#) and [Bordalo et al. \(2017a\)](#) present models of how cued recall shapes economic decision-making across a broad set of domains, including how it generates overreaction in belief formation. Related work has investigated the implications of memory in applied settings such as updating biases ([Gennaioli and Shleifer, 2010](#); [Wilson, 2014](#)), financial markets ([Bodoh-Creed, 2013](#); [Bordalo et al., 2017b, 2018](#); [Wachter and Kahana, 2019](#)), and self-esteem ([Koszegi et al., 2019](#)).

As much of the simple formalism that structures our experiments directly draws from this literature, we view our experiments as providing some of the first direct evidence from tightly structured and quantitative economic decision making tasks in favor of this emerging body of theoretical work. Psychological experiments on associative recall exhibit a different structure than the experiments that are presented here (see [Kahana, 2012](#), for an overview). These experiments typically consist of explicit cued recall problems (such as with words), rather than of the types of quantitative reasoning tasks that characterize modern experimental economics research. [Bordalo et al. \(2019\)](#) present an experiment on shape and color recognition that shows a link between associative memory and the representativeness heuristic.¹

¹More indirectly, our paper also relates to recent experimental work on the role of motivated (self-serving) memory, as in [Zimmermann \(forthcoming\)](#) and [Carlson et al. \(2018\)](#), although these experiments

More broadly, our paper fits into the recent experimental literature that has focused on unearthing the micro-foundations behind reduced-form behavioral biases (Enke and Zimmermann, 2019; Enke, 2017; Enke and Graeber, 2019; Esponda and Vespa, 2016; Martínez-Marquina et al., 2017; Dertwinkel-Kalt et al., 2017; Frydman and Jin, 2018). Finally, because we experimentally identify a mechanism behind overreaction, our work also relates to a literature that documents overreaction or excess movement in survey expectations about macroeconomic variables or geopolitical events (Augenblick and Rabin, 2018; Augenblick and Lazarus, 2018; Bordalo et al., 2018, 2017b).

The remainder of the paper proceeds as follows. Section 2 offers a stylized formal framework that motivates the experimental design and structures the empirical analysis. Section 3 describes the experimental design, implementation, and pre-registration. Sections 4 and 5 present the evidence on overreaction and the roles of imperfect and associative memory therein. Section 6 considers the case of underreaction, while Section 7 studies the role of the length of the time lag. Section 9 concludes.

2 Theoretical Framework

2.1 Setup

This section presents a stylized model to guide the design of the experiments and to structure the empirical analysis. The mechanics of the model directly build on some of the formulations in Mullainathan (2002) and Bordalo et al. (2017a). The framework rests on two key assumptions: (i) people do not permanently entertain a prior belief but may instead forget it over time, so that once they are prompted to update their beliefs they first need to reconstruct past knowledge from memory and (ii) this recollection process is subject to associative recall, meaning that news are more likely to get remembered if they were observed in a context that is similar to the context in which today’s signal is observed. For simplicity, we abstract away from additional behavioral assumptions that the literature on associative memory has incorporated, such as salience effects or rehearsal.

Consider a decision-maker (DM) who forms beliefs about the state of a time-varying stochastic variable θ_t with initial value ν . We consider two periods that we will think of as “past” and “present.” In any given period, the value of θ is given by its initial value plus the sum of all news n_j that have accumulated up to this point, where $n_j \in \{-q, q\}$. News are equally likely and i.i.d. We will use the terms “news” and “signal” interchangeably.

A piece of news n_j is associated with a memorable context $c_j \in \{L, H\}$. In the “past”, k news arrive, so that $\theta_1 = \nu + \sum_{x=1}^k n_x$. In $t = 1$, there is a one-to-one mapping between

are not concerned with associative recall.

type of news (positive or negative) and context (high or low): $n_x = n_y \Leftrightarrow c_x = c_y$.

In the “present” ($t = 2$), the DM observes one final piece of news n_{k+1} . Thus:

$$\theta_2 = v + \sum_{x=1}^k n_x + n_{k+1} \quad (1)$$

Just as in $t = 1$, the piece of news is associated with a context. We will consider two regimes. In the first, second-period news and contexts are associated in the same way as in the first period. In the second regime, the DM receives second-period news in a context opposite to what he was exposed to in the first period, meaning that he observes positive news in a “negative” context and vice versa. As a shorthand for this “correlation” between news and context, we define

$$\rho \equiv \begin{cases} 1 & \text{if } P(c_2 = H | n_{k+1} = q) = P(c_2 = L | n_{k+1} = -q) = 1 \\ -1 & \text{if } P(c_2 = L | n_{k+1} = q) = P(c_2 = H | n_{k+1} = -q) = 1 \end{cases}$$

2.2 Memory and Beliefs

In period 1, the DM observes a collection of news which deterministically pin down the true state. Thus, when prompted for their belief, the DM will state $b_1 = v + \sum_{x=1}^k n_x$.

Our object of interest is the extent to which the DM’s belief about θ responds to the latest piece of news n_{k+1} . A rational (or Bayesian, though there is no uncertainty here) DM would observe the collection of n_x and then correctly predict $\theta_2 = v + \sum_{x=1}^k n_x + n_{k+1}$.

Suppose instead that the DM potentially forgets some of the news between $t = 1$ and $t = 2$. That is, when our DM wakes up in $t = 2$, he has potentially forgotten some of the $n_1 \dots, n_k$. Thus, his belief (after observing n_{k+1}) is given by

$$b_1 = v + \sum_{x=1}^k m_x n_x + n_{k+1} \quad (2)$$

where $m_x \in \{0, 1\}$ denotes whether the DM remembers piece of news n_x .

Whether or not the DM remembers a piece of news is determined by both (i) imperfect and (ii) associative memory. First, by imperfect recall we mean that, irrespective of the piece of news, there is some probability that the DM will forget. The reduced-form assumption of imperfect recall is a shorthand for different mechanisms that have been highlighted in the psychological literature, in particular that of interference ([Kahana, 2012](#)). By the logic of interference, it is harder for people to recall a specific item if they have been exposed to many similar items in the past. In our experiments, we generate interference (and hence imperfect memory) by implementing the same type of judgment task multiple times with different sets of signal realizations and contexts.

Second, by associative recall we mean that the probability of recalling a piece of news from the past is higher if it is cued by today's signal. That is, a signal is more likely to get remembered if it occurred with the same context as today's signal.

We formalize the probability of remembering signal n_x as follows. First, the baseline probability of recall is $r < 1$. In addition, there is an increase in the probability of recalling $(1 - r)a$, $a < 1$, if the context c_{k+1} that is associated with n_{k+1} is the same as the context that is associated with news n_x . This formulation implies that associative memory matters more for DM with highly imperfect memory (low r). Formally:

$$m_x = \begin{cases} 1 & \text{with probability } r + (1 - r)a \mathbb{1}_{c_x = c_{k+1}} \\ 0 & \text{else} \end{cases} \quad (3)$$

We assume that the DM forms beliefs exclusively from what he recalls.²

Denote by $z \geq 0$ the number of news in $t = 1$ that were observed in the same context as n_{k+1} . The expected belief in period $t = 2$ is then given by:

$$\begin{aligned} E[b_2 | n_x, n_{k+1}] &= v + n_{k+1} + \sum_{x=1}^k m_x n_x \\ &= v + n_{k+1} + \sum_{x=1}^k r n_x + \sum_{x=1}^z (1 - r) a n_x \end{aligned} \quad (4)$$

$$= v + [1 + \rho z (1 - r) a] n_{k+1} + r \sum_{x=1}^k n_x \quad (5)$$

$$= v + n_{k+1} + [r + \rho (1 - r) a] \sum_{x=1}^z n_x + r \sum_{x=z+1}^k n_x \quad (6)$$

Equation (5) is the core expression that we subject to systematic experimental tests. It implies the rational (Bayesian) prediction that second-period beliefs will respond with a coefficient of one to variation in the second-period signal. On the other hand, viewed through the lens of imperfect and associative memory, equation (5) suggests that expectations will overreact. This is the central prediction of our framework. Moreover, equation (6) clarifies that this overreaction can equivalently be understood as increased sensitivity of beliefs to past news that were communicated in the same context as today's news (the third term), relative to news that were communicated in a different context (the fourth term in eq. (6)). Equation (5) suggests the following abstract hypotheses, which we concretize for our experimental implementation in Section 3:

²This implies naïveté as in [Mullainathan \(2002\)](#). In principle, naïveté could come in two facets: (i) the DM fails to realize that he sometimes forgets, i.e., that there are signals he does not recall; (ii) the DM realizes that he sometimes forgets, but he naïvely infers what he does not recall by failing to take associativeness into account.

Hypotheses.

1. *If the correlation between news and context is positive ($\rho = 1$), expectations overreact to today's news, on average. Put differently, expectations are more sensitive to past news that took on the same realization as today's news.*
2. *Overreaction increases in the number of past news that were communicated in the same context as today's news (z).*
3. *Overreaction increases in the imperfectness of memory ($1 - r$).*
4. *Overreaction increases in the strength of associative recall (a).*
5. *If the correlation between news and context is negative ($\rho = -1$), expectations underreact to today's news, on average.*
6. *This underreaction increases in the number of past news that there were communicated in the same context as today's news.*

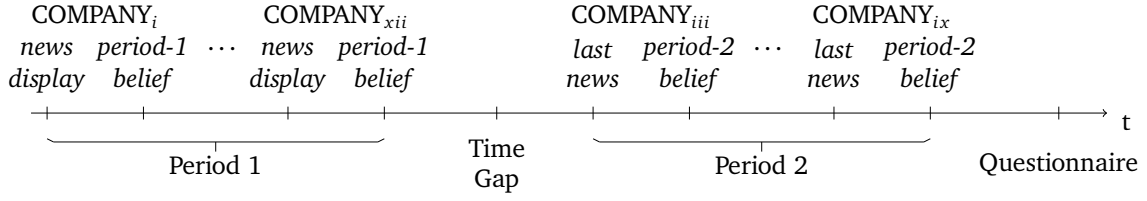
3 Experimental Design

An environment in which the role of memory for belief formation can be studied requires (i) a dynamic setup in which subjects state beliefs twice with a delay period in-between, such that previously formed posteriors do not mechanically translate into current priors; (ii) variation in signal histories that allow for nuanced predictions about when overreaction should be more or less pronounced; (iii) treatment variations that allow the exogenous manipulation of memory constraints, the role of associativeness, and the correlation structure between signals and contexts; and (iv) incentive-compatible belief elicitation. Our design was built to accommodate these features.

3.1 Experimental Setup

In order to isolate the role of memory, we implemented a simple deterministic decision environment in which, absent potential memory constraints, behaving rationally is trivial. Subjects were asked to guess the stock market value of twelve hypothetical companies at a given point in time. Continuing the notation from Section 2, the value of a company at time t , θ_t , was given by the baseline value, $v = 100$, plus the sum of all news about that company up to t . News could be positive, $n_x = 10$, or negative, $n_x = -10$. Positive and negative news were equally likely and were randomly and independently

Figure 1: Experimental Timeline



drawn by the computer. All of this was known to subjects. The value of a company at time t is given by:

$$\theta_t = 100 + \sum_{x=1}^k n_x. \quad (7)$$

where k is the number of signals that were shown up to t .

News were not only communicated as abstract numbers, but were shown on subjects' computer screens with what we refer to as a context: an image and a story. The written instructions clarified that these images and stories were supplied to "explain" to subjects why a particular piece of news for a company was observed. For instance, all stories that accompanied positive news gave some rationale for why the value of the company had gone up, such as a successful marketing campaign or a recent technological innovation. The content of the story and the picture were tailored to match each other. The signal, picture and story were displayed on subjects' computer screens for 15 seconds. The time was calibrated such that subjects would have sufficient time to process the news, as well as to fully grasp the content of the picture and the story. Appendix E contains examples of these images and stories (see Figures 10 and 11).

The experiment consisted of two periods, as summarized in Figure 1. In both periods, participants estimate the stock market value of hypothetical companies. It was made salient to subjects that information from the first period of the experiment would also be relevant for their estimates in the second period.

In period 1, subjects sequentially observed news and corresponding contexts for a particular company on their computer screens. Then, they were asked to estimate the company's current stock market value. This procedure was repeated for all twelve companies. Across companies, the number of signals varied between zero and three.³ Beliefs in period 1 allow us to verify whether subjects understand the basic information structure, had sufficient time to process the information, and are in principle able to form correct posteriors in our decision environment.

³All subjects saw three companies with three pieces of news, three with two pieces of news, three with one pieces of news and three with zero pieces of news.

After period 1, we implemented a time gap in which subjects went through a 15 minute real effort task. The real-effort task required subjects to type multiple combinations of letters and numbers into the keyboard. Subjects had 15 minutes to type in as many combinations as they could. For each correctly solved task, subjects received 5 cents. The purpose of the real effort task was to trigger memory constraints, such that previously formed beliefs (period 1) do not necessarily translate into second-period priors. While the memory literature contains many demonstrations that 15 minutes are sufficient to activate long-term memory and the corresponding memory constraints, in Section 7 we show the robustness of our findings when we increase the time lag to three days rather than 15 minutes. It is further worth mentioning that our procedure of working with twelve similar companies induces what psychologists call “interference”, which makes it harder for subjects to memorize and correctly attribute separate signals or beliefs. This effect likely contributes to our reduced-form assumption of imperfect memory, but ultimately our paper does not need to take a stand on where exactly imperfect memory comes from.

In the second period, for each company, subjects were shown up to one additional piece of news. Specifically, for ten companies, subjects received an additional piece of news, while for two companies, there were no additional news. We included two companies with no additional news because these allow us to directly assess whether subjects perfectly remember their first-period belief in the second period.

Immediately after observing the additional piece of news for a company, subjects were asked to state a second-period belief about the value of that company. Second-period beliefs constitute our main outcome of interest. In addition, on a subsequent decision screen, subjects were asked to recall the total number of positive and negative news that were shown to them in the course of the entire experiment for that company. These recall measures are not financially incentivized. Again, this procedure was repeated for all twelve companies.

The experimental instructions and comprehension questions saliently emphasized that first-period signals would also be relevant for the second period, so that subjects presumably attempted to memorize either the signals themselves or their own first-period belief as sufficient statistic for these signals. At the same time, the instructions do not suggest to subjects that they need to memorize the stories or images that accompany a signal.

The experiment was independently randomized across subjects across the following layers: (i) the order of companies in the first period; (ii) the order of companies in the second period; (iii) whether or not a company received a piece of news in the second period; and (iv) the actual signal realizations.

Beliefs were incentivized using a binarized scoring rule, which is incentive-compatible

regardless of subjects' risk attitudes (Hossain and Okui, 2013). Under this scoring rule, subjects could potentially earn a prize of 10 euros. The probability of receiving the prize was given by 100 minus the squared distance between a subject's belief and the true value of the asset. In order to avoid hedging motives, at the end of the experiment one of the 24 beliefs was randomly selected for payment. Since second-period beliefs are our main outcome measure, we incentivized them more heavily, in expectation: with 90% probability a second-period belief was randomly selected for payment, and with 10% probability a first-period belief. To avoid extreme outliers due to typing mistakes, the computer program restricted beliefs to be in $[50, 150]$.

3.2 Treatment Variations and Sources of Exogenous Variation

We conducted five treatments, referred to as *Main*, *Reminder*, *No Cue*, *underreaction* and *Underreaction reminder*. In combination, these treatments allow for causal tests of all of the abstract predictions laid out in Section 2. That is, the treatments were designed to identify (i) potential overreaction in expectations; (ii) the ways in which the quantitative magnitude of such overreaction depends on the precise signal history; and (iii) the causal roles of imperfect and associative memory for overreaction; and (iv) the role of the correlation between context and news.

Treatment *Main*. In treatment *Main*, there is a one-to-one mapping between type of news (positive or negative) for a given company and the context with which the signal is communicated. That is, every positive news for company A is communicated with the same context (image and story). Likewise, every negative news for company A is communicated with the same context (albeit a different one than the positive news). The same logic holds for all other companies. Thus, it can never happen that a context is communicated with news for different companies, or with both positive and negative news. A context deterministically identifies a piece of news.

Thus, treatment *Main* resembles our opening examples and implements a situation in which we hypothesize to observe overreaction. In addition, note that because the number and realizations of the signals vary across companies and subjects, the twelve tasks exhibit substantial variation in signal histories. We leverage this source of exogenous variation to test the within-treatment predictions derived in Section 2 about how the presence or quantitative magnitude of overreaction depends on the number of first-period signals that occurred in the same context (have the same realization) as the second-period signal. 80 subjects participated in this treatment.

Treatment *Reminder*. In treatment *Reminder*, we seek to remove subjects' memory constraints, holding everything else constant. Conceptually, we think of this treatment as exogenously setting the parameter $r = 1$ in the framework of Section 2 (meaning perfect memory). This allows to precisely identify the role of (imperfect) memory in our setup. The setup is exactly the same as in *Main*, except that at the beginning of the second period (i.e., before a subject observes the last signal for a company), subjects were reminded of their own first-period belief for that company. Thus, in contrast to treatment *Main*, we assist subjects in the recall of their first-period belief, so that they presumably no longer need to reconstruct their prior knowledge from memory. Comparing treatments *Main* and *Reminder* allows us to cleanly identify the role of memory imperfection r for overreaction. 50 subjects participated in this treatment.

Treatment *NoCue*. Treatment *NoCue* was designed to isolate the role of associative recall. Conceptually, we think of this treatment as setting the associative recall parameter $a = 0$ (meaning no associative recall). In terms of implementation, the setup in this treatment was exactly the same as in *Main*, except that each piece of news is communicated with a different context. That is, a given context (image and story) never appears twice, even if the company and type of news is identical. Thus, it is no longer the case that every positive news for a given company is communicated with the same context, and every negative news for a given company is communicated with the same context. As a consequence, stories and images can no longer trigger associative recall. At the same time, all other features of the environment remain unchanged. Comparing treatments *Main* and *NoCue* therefore allows us to cleanly identify the role of associative recall a . 80 subjects participated in this treatment.

Treatments *Underreaction* and *Underreaction reminder*. All treatments described above relied on a design in which the observation of a positive piece of news in the second period cues the asymmetric recollection of positive first-period news (and analogously for negative news), which corresponds to $\rho = 1$ in our formal framework. Treatments *Underreaction* and *Underreaction reminder* conceptually correspond to setting $\rho = -1$. In both treatments, the first period proceeded exactly as in treatment *Main*. In the second period, however, the remaining news were communicated on subjects' decision screens along with the *opposite* story and image, relative to the first period. That is, a positive piece of news for company A was communicated along with the story and image that were associated with negative news for company A in the first period of the experiment. Analogously, a negative piece of news for company A was communicated along with the story and the image that were associated with positive news for company A in the first period of the experiment. 80 subjects participated in this treatment.

To allow for causal inference, we again conducted a second treatment: condition *Underreaction reminder*. This treatment was identical to *Underreaction*, except that subjects were reminded of their own first-period belief right before they received the second-period signal for a company. In all other respects, these two treatments followed the same procedure as treatments *Main* and *Reminder*. 50 subjects took part in *Underreaction reminder*.

3.3 Predictions

Equation (5) in the conceptual framework directly suggests the following estimating equation for subject i 's second-period belief in task j :

$$b_2^{i,j} = \beta_1 n_{k+1}^j + \beta_2 \sum_{x=1}^k n_x^j + \epsilon \quad (8)$$

That is, we regress a subject's second-period belief on the value of the last signal as well as the first-period stock value (or the corresponding first-period belief). Appendix A.2 formally derives the properties of the OLS estimator $\hat{\beta}_1$ and shows that $E[\hat{\beta}_1] \approx 1 + \rho(1-r)a\bar{z}$, where \bar{z} is the average number of first-period signals that were observed in the same context as the second-period signal. By applying the abstract predictions derived in Section 2 to this experimental design and estimating equation, we are hence ready to state the following predictions:

Predictions.

1. In treatment *Main*, we observe overreaction: $\hat{\beta}_1 > 1$.
2. In treatment *Main*, overreaction increases in the number of first-period signals that were observed in the same context as the second-period signal.
3. Overreaction is stronger in treatment *Main* than in *Reminder*.
4. Overreaction is stronger in treatment *Main* than in *No Cue*.
5. In treatment *Underreaction*, we observe underreaction $\hat{\beta}_1 < 1$. Moreover, underreaction increases in the number of first-period signals that were observed in the same context as the second-period signal.
6. Underreaction is stronger in treatment *Underreaction* than in *Underreaction reminder*.

For clarity, Table 1 explicitly spells out which abstract model prediction from Section 2 maps into which specific experimental prediction, and which experimental treatments we use to test a given prediction.

Table 1: Mapping from model predictions to experimental predictions

Abstract model prediction	Treatments	Experimental prediction
1. Overreaction if news and context co-occur	<i>Main</i>	$\hat{\beta}_1^{Main} > 1$
2. Overreaction increases in # identical past contexts	<i>Main</i>	$\hat{\beta}_1^{Main}$ increases in z
3. Overreaction increases in imperfectness of memory	<i>Main</i> & <i>Reminder</i>	$\hat{\beta}_1^{Main} > \hat{\beta}_1^{Reminder}$
4. Overreaction increases in strength of associative recall	<i>Main</i> & <i>NoCue</i>	$\hat{\beta}_1^{Main} > \hat{\beta}_1^{NoCue}$
5. Underreaction if news and context negatively corr.	<i>Underreaction</i>	$\hat{\beta}_1^{Under} < 1$
6. Underreaction increases in imperfectness of memory	<i>Under.</i> & <i>Under. rem.</i>	$\hat{\beta}_1^{Under} < \hat{\beta}_1^{Under. rem.}$

3.4 Procedures and Logistics

Upon arrival in the lab, subjects received written instructions about the experiment. Subjects were given unlimited time to read the instructions and could ask questions at any point in time. After all subjects had indicated that they had finished the instructions, they completed a total of seven computerized control questions to verify adequate comprehension. Whenever a subject did not solve a control question correctly, a computer screen pointed out the mistake and explained the correct solution. As we pre-registered (see below), we exclude subjects from the analysis that answered more than one control question incorrectly (7% of potential participants).

Treatments *Main*, *Reminder*, and *NoCue* were conducted in the BonnEconLab of the University of Bonn and treatments *Underreaction* and *Underreaction reminder* in the University of Cologne’s Laboratory for Experimental Economics. Assignment to the relevant treatments was randomized within experimental sessions: *Baseline*, *Reminder*, and *NoCue* were all implemented in the same sessions, as were *Underreaction* and *Underreaction reminder*. We only compare treatments that were randomized within session. The experiments were computerized using Qualtrics. Experimental sessions lasted up to 90 minutes, and subjects earned an average of 15 euros.

3.5 Pre-Registration

All experiments in this paper were pre-registered in the AEA RCT registry, including a pre-analysis plan. The pre-registration includes (i) the design of the treatments described in Section 3.2; (ii) the design of the treatment discussed in Section 7; (iii) the heterogeneity analysis discussed in Section 4.3; (iv) the regression equation (8) through which we analyze all data; (v) all predictions outlined in Section 3.3; (vi) the sample size in each treatment; and (vii) that subjects would be dropped from the sample (and replaced) if they answer more than one control question incorrectly. The pre-registration is available at <https://www.socialscienceregistry.org/trials/4247>.

4 Baseline Results on Overreaction

4.1 Preliminaries

Before we present the results, we conduct two checks on our experimental data. First, we verify people’s understanding of the experimental setup by investigating the accuracy of people’s beliefs at the end of the first period, before memory constraints become relevant. Table 8 in Appendix C shows that beliefs almost perfectly correspond to the true value of a company, in each of the three treatments: in a regression of subjects’ beliefs on actual company values, the OLS coefficient is always almost exactly one and hence rational. The average percentage deviation between first-period beliefs and the truth is only 0.4%, while the median deviation is zero. This provides reassuring evidence that subjects appear to understand the decision task well.

Second, to provide credence to our assumption that subjects can no longer perfectly remember their first-period belief once the second period starts, we consider the relationship between subjects’ second-period and first-period beliefs in those tasks in which there was no second-period signal. Table 9 in Appendix C reports the results. In a regression of second-period on first-period beliefs, the OLS coefficient is only 0.56 and hence far away from the perfect memory benchmark of one. This suggests that memory is indeed imperfect in our setup, hence opening up a potential role for associative recall. Moreover, this regression coefficient is very similar in treatments *Main* and *No Cue*.

4.2 Treatment *Main*: Overreaction in Expectations

Throughout the empirical analysis, we present OLS regressions to test the hypotheses outlined in Section 3.3. Table 2 presents the results for treatment *Main*. In columns (1)–(3), we present three regression specifications. First, a regression in which we regress second-period beliefs on the value of the last signal (+10 or –10), controlling for the first-period belief. Second, an analogous regression in which we control for the objective first-period stock value as opposed to the first-period belief. Third, a comprehensive specification in which we control for experimental session fixed effects, first-period signal history fixed effects, company fixed effects, experimental order fixed effects, and subject fixed effects. In this third specification, controlling for first-period beliefs or stock values is redundant as these are implicitly accounted for by the first-period signal history fixed effects. In each regression specification, an observation corresponds to a subject-task, for a total of ten tasks per subject.⁴ Throughout, we cluster the standard errors at the subject level.

⁴Naturally, and as specified in the pre-analysis plan, we restrict attention to those tasks in which a subject indeed received a signal in the second period.

Table 2: Treatment *Main*

	Dependent variable: 2nd period belief						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
2nd period signal	1.10*** (0.02)	1.11*** (0.02)	1.11*** (0.03)	0.87*** (0.04)	0.87*** (0.04)	0.85*** (0.04)	0.87*** (0.04)
Belief in 1st period	0.75*** (0.03)					0.59*** (0.05)	
Stock price in 1st period		0.74*** (0.03)					
Value of cued 1st period signals				0.92*** (0.03)	0.90*** (0.03)		
Value of non-cued 1st period signals				0.59*** (0.05)	0.59*** (0.05)		
2nd period signal × # 1st period signals in same context						0.34*** (0.05)	0.31*** (0.05)
Session FE	No	No	Yes	No	Yes	No	Yes
1st period signal history FE	No	No	Yes	No	No	No	Yes
Company FE	No	No	Yes	No	Yes	No	Yes
Order FE	No	No	Yes	No	Yes	No	Yes
Subject FE	No	No	Yes	No	Yes	No	Yes
Observations	800	800	800	800	800	800	800
Adjusted R^2	0.80	0.80	0.80	0.81	0.81	0.81	0.81

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The sample includes all observations from treatment *Main* where subjects observed a second-period signal.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

For all coefficients that are reported in Table 2, the rational prediction is that they equal one. The simple framework outlined in Section 2 instead predicts that the coefficient of first-period beliefs or stock values is less than one (due to imperfect memory) and that the coefficient of the last signal is greater than one (due to imperfect and associative memory). This is indeed what we find, see columns (1)–(3). In terms of magnitude, the OLS coefficient suggests that expectations substantially overreact, by 11 percent relative to the rational prediction of one. The last row of Table 2 reports the p-value for the null hypothesis that the coefficient of the first-period signal equals one. We reject this rational null hypothesis at all conventional levels of significance.

As highlighted by equation (6) in the formal framework, our hypothesis is that such overreaction occurs because the first-period signals get recollected more successfully if they are cued, that is if they take on the same value as the second-period signal. To investigate this more explicitly, columns (4) and (5) of Table 2 include as separate regressors the overall value of those first-period signals that do (or do not) equal the second-period signal. The results show that expectations are much more responsive to the value of the

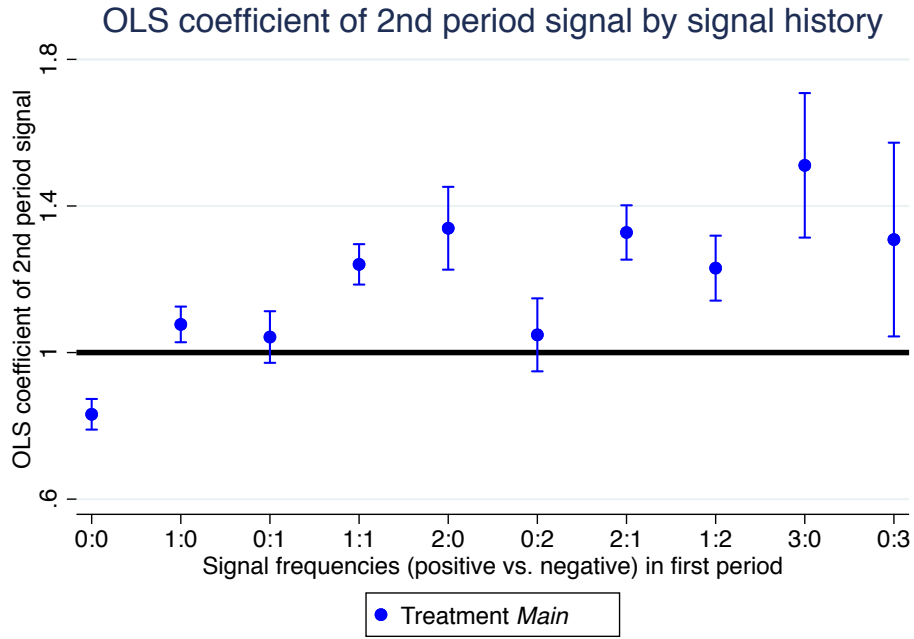


Figure 2: OLS coefficient in a regression of second-period beliefs on the last signal, separately for each set of signal frequencies in the first part. The regressions control for a subject’s first-period belief. Standard error bars are computed based on clustering at the subject level.

cued first-period signals. Here, the difference in regression coefficients is statistically significant at all conventional levels.

Figure 2 visualizes the results. For each set of possible signal frequencies in the first period of the experiment, we regress second-period beliefs on the value of the second-period signal, and then plot the OLS coefficient and corresponding standard error. The figure shows that this coefficient is almost always larger than one. At the same time, there appears to be significant variation in the quantitative magnitude of this effect. Visual inspection suggests that the coefficient is increasing in the number of first-period signals. This is intuitive because if there are no past signals that can be cued, then trivially associative memory and asymmetric recall cannot generate overreaction. Thus, it is reassuring that expectations do not at all overreact in the case of zero positive and zero negative first-period signals (at the very left of Figure 2). In fact, if anything, this coefficient is smaller than one, consistent with a large literature on belief updating that shows that in lab environments where the role of (associative) memory is shut down, people’s belief updating typically exhibits conservatism ([Benjamin, 2018](#)).

This discussion directly relates back to Prediction 2 in Section 3.3, which posits that the magnitude of the regression coefficient should be increasing in the number of first-period signals that take on the same value as the second-period signal. This is a direct test of the role of associative memory because with either perfect memory ($r = 1$ in the model) or imperfect but no associative memory ($a = 0$), this prediction would not

hold, compare equation (5). Moreover, while a particular form of recency bias could in principle generate the type of overreaction we observe in Table 2, this is not the case for the prediction that overreaction depends on the signal history in specific ways.

Columns (6)–(7) provide a formal statistical test of Prediction 2. Here, we interact the value of the second-period signal with the number of first-period signals that were communicated in the same context (z in the model). The regression table suppresses the corresponding raw term for brevity. The results show that the interaction term of interest is consistently positive and statistically highly significant. The magnitude suggests that each additional first-period signal increases the responsiveness to the second-period signal by about 30%, on average. Moreover, once the interaction term is accounted for, the regression coefficient of the second-period signal (which now econometrically corresponds to the case of no cued first-period signals) is less than one. This result is analogous to the discussion of Figure 2: when no signals are cued, expectations do not overreact but – if anything – even underreact ([Benjamin, 2018](#)).

The evidence in support of Prediction 2 suggests that associative memory does not only help people in recalling that they saw such a signal in the first period, but also *how often* such a signal occurred. To corroborate this result, we next turn to subjects’ direct recall data. Figure 3 shows average levels of reported recall of first-period signals in condition *Main*, as a function of whether these first-period signals were identical to or different from the second-period signal. That is, the figure shows how many signals subjects report to have recalled, as a function of whether those signals were cued or not. The figure shows that the recall of cued signals is very accurate, on average. In contrast, the recall of non-cued signals is more compressed: subjects overestimate the frequency of signals that did not appear at all and they substantially underestimate (under-recall) the frequency of signals that appeared often. This is indicative that associative memory helps not only with remembering *whether* a certain type of signal has appeared before, but also *how often* it appeared.

4.3 Heterogeneity Analysis

Next, we examine across-subject heterogeneity in overreaction. To estimate the presence of such heterogeneity, we require a measure of a subject’s overall overreaction “type.” To this effect, we run our standard regression of second-period beliefs on the second-period signal, but now separately for each subject. That is, we estimate a subject-specific responsiveness parameter that should be equal to one if a subject is either rational or has imperfect-but-no-associative-memory.

Figure 4 presents the distribution of types. While the beliefs of a notable fraction of subjects do not reflect associative recall (35% have a regression coefficient of at most

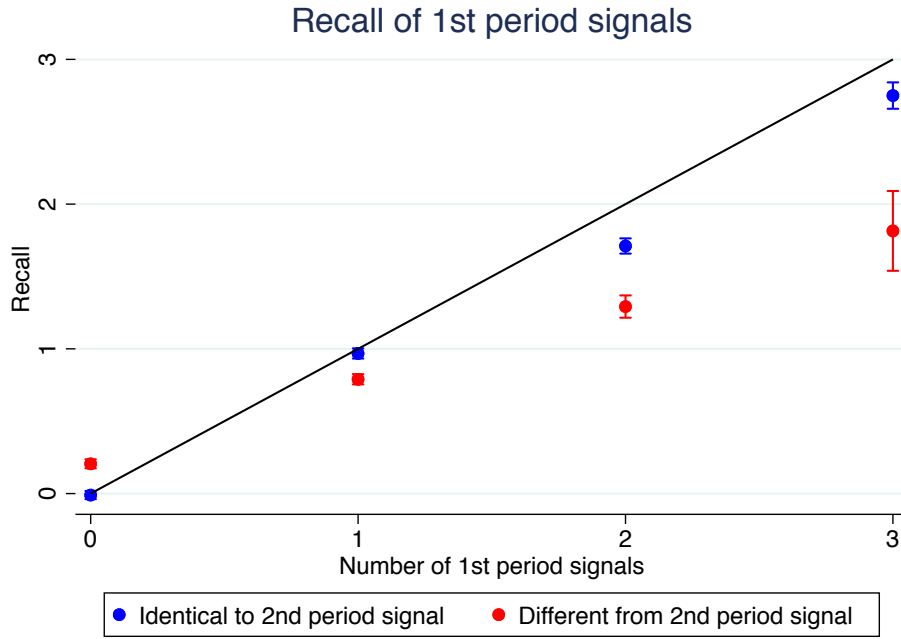


Figure 3: Recall of first-period signals in Treatment *Main*, depending on whether the second-period signal was identical to or different from the first-period signals. We construct the recall variables as follows. In the case of recall of signals that are different from the second-period signal, we use the reported recall quantity. In the case of recall of signals that are identical to the second-period signal, we use the reported recall minus one. That is, we make the arguably very plausible assumption that subjects always remember the value of the second-period signal that they just saw a few seconds ago.

one), the majority of participants exhibit overreaction to varying degrees.

A natural question is what explains this heterogeneity. To investigate this, we turn to three heterogeneity analyses, all of which were specified in our pre-analysis plan: (i) performance on a Raven matrices IQ test that was administered at the end of the experiment; (ii) a measure of the strength of memory that is estimated from the experimental recall data as a proxy for r : for each subject, we regress the reported recall of non-cued signals on the actual number of corresponding signals and save this regression coefficient as a measure of the strength of (non-cued) memory; and (iii) response times.

Table 3 reports the results. Here, we implement our standard regression methodology, yet additionally interact the value of the last signal with our pre-specified subject-specific variables. We find that subjects with higher Raven scores and better non-cued recall exhibit less overreaction. The relationship between overreaction and response times is negative, but not statistically significant.⁵

⁵Table 11 in Appendix C reports analogous analyses for the recall data.

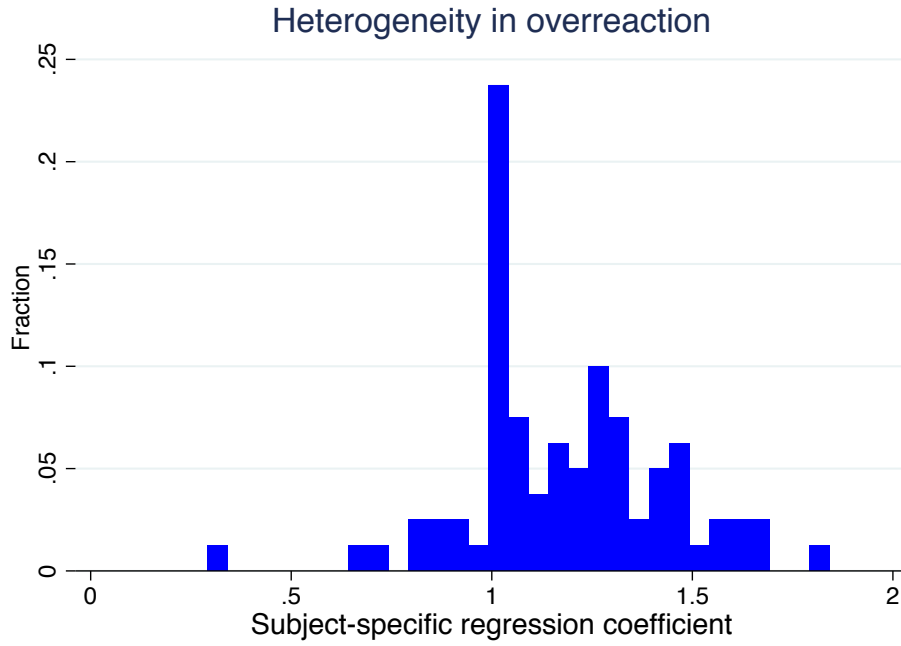


Figure 4: Subject-level distribution of regression coefficients of last signal in treatment *Main* (N=80). To estimate these coefficients, we run regressions akin to column (1) in Table 2 except that in each regression the sample is restricted to only one subject. Moreover, to adequately identify overreaction in the presence of cued recall, the sample is restricted to tasks in which a subject observed at least one first-period signal. A rational subject would have a coefficient of one.

5 Exogenous Memory Manipulations

5.1 The Role of Imperfect Memory

To provide causal evidence for the role of imperfect memory in expectation overreaction, we manipulate the extent to which memory can actually play in the role in the experiment. Conceptually, treatment *Reminder* is designed to set $r = 1$. To this effect, we reminded participants of their first-period belief immediately before they received the second-period signal.

Given the explicit focus on memory in this section, the analysis considers both (i) the financially incentivized second-period beliefs and (ii) the unincentivized recall data. As specified in the pre-analysis plan, we analyze the recall data by computing the difference between recall of positive and recall of negative news and multiplying this difference by 10 so that the variable has the same scale as the beliefs data. This summary statistic of a subject's recall is highly correlated with actual second-period beliefs ($\rho = 0.95$), suggesting that the recall data are meaningful.⁶

Table 4 presents the results of the treatment comparison between *Main* and *Reminder*.

⁶Table 10 in Appendix C replicates the results of Tables 2 for this summary statistic of recall, as specified in the pre-analysis plan.

Table 3: Treatment *Main*: Heterogeneity analysis

	Dependent variable: 2nd period belief					
	(1)	(2)	(3)	(4)	(5)	(6)
2nd period signal	1.26*** (0.08)	1.25*** (0.09)	1.23*** (0.04)	1.21*** (0.04)	1.14*** (0.04)	1.14*** (0.04)
2nd period signal × Raven score	-0.030** (0.01)	-0.028* (0.01)				
2nd period signal × Memory for non-cued signals			-0.23*** (0.06)	-0.20*** (0.06)		
2nd period signal × Response time					-0.37 (0.29)	-0.35 (0.24)
Belief in 1st period	0.75*** (0.03)		0.75*** (0.03)		0.75*** (0.03)	
Session FE	No	Yes	No	Yes	No	Yes
1st period signal history FE	No	Yes	No	Yes	No	Yes
Company FE	No	Yes	No	Yes	No	Yes
Order FE	No	Yes	No	Yes	No	Yes
Observations	800	800	800	800	800	800
Adjusted R^2	0.80	0.80	0.80	0.80	0.80	0.80

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The table suppresses the coefficients of Raven score (columns (1)–(2)), memory for non-cued signals (columns (3)–(4)), and response time (columns (5)–(6)). Response times are measured in minutes. The sample includes all observations from treatment *Main* where subjects observed a second-period signal. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

As specified in the pe-registration, we again analyze our data by means of OLS regressions in which we relate subjects' second-period beliefs (or recall) to the value of the second-period signal, except that now we also interact the second-period signal with a treatment dummy. The table also includes a treatment dummy, which is suppressed for brevity. Our prediction, spelled out in Sections 2 and 3.3, is that the value of the second-period signal should matter more in treatment *Main* than in *Reminder*.

The results provide supporting evidence for this prediction. The interaction term is quantitatively large and statistically significant at all conventional levels. This is true when we consider the beliefs data as well as when we directly look at reported recall, compare columns (4)–(6). In *Main*, subjects respond 12–14% more to the value of the last signal than subjects in *Reminder*. Again, this pattern is a specific prediction of our framework, but not of an account of recency effects.

Moreover, as we can see from the regression coefficient of the second-period signal

Table 4: Treatments *Main* vs. *Reminder*

	Dependent variable:					
	2nd period belief			Δ Recall [Pos. – Neg.]		
	(1)	(2)	(3)	(4)	(5)	(6)
2nd period signal	0.99*** (0.01)	0.98*** (0.01)	0.98*** (0.01)	0.95*** (0.02)	0.95*** (0.02)	0.94*** (0.02)
2nd period signal \times 1 if <i>Main</i> , 0 if <i>Reminder</i>	0.12*** (0.03)	0.13*** (0.03)	0.14*** (0.03)	0.11** (0.04)	0.12*** (0.04)	0.13*** (0.04)
Belief in 1st period	0.84*** (0.02)			0.83*** (0.02)		
Stock price in 1st period		0.84*** (0.02)			0.83*** (0.02)	
Treatment FE	Yes	Yes	Yes	Yes	Yes	Yes
Session FE	No	No	Yes	No	No	Yes
1st period signal history FE	No	No	Yes	No	No	Yes
Company FE	No	No	Yes	No	No	Yes
Order FE	No	No	Yes	No	No	Yes
Subject FE	No	No	Yes	No	No	Yes
Observations	1300	1300	1300	1300	1300	1300
Adjusted R^2	0.86	0.86	0.86	0.82	0.83	0.83

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The sample includes all observations from treatments *Main* and *Reminder* where subjects observed a second-period signal. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

(which captures the coefficient in treatment *Reminder* only), in treatment *Reminder* we find no overreaction: the OLS coefficient is 0.94–0.99, if anything even less than one. These results provide direct causal evidence that imperfect memory is necessary in order for overreaction to arise in our setup, as predicted by our key equation (5). Figure 5 visualizes this result and the difference between *Main* and *Reminder*. Again, we see that there is no overreaction in treatment *Reminder*, which is reassuring evidence that our experimental setup is not misconstrued by subjects: in the absence of memory constraints, the second-period signal is incorporated in a rational fashion.

5.2 The Role of Associative Memory

To provide direct causal evidence for the role of associative recall, we proceed by directly manipulating the strength of associativeness of memory. According to equation (5), if there is no associative recall ($a = 0$) there should be no overreaction.

As a direct test of this hypothesis, we compare treatments *Main* and *NoCue*. Table 5 presents the results and follows a similar logic as Table 4. Again, as specified in the pre-analysis plan, we link participants' second-period beliefs to the second-period signal,

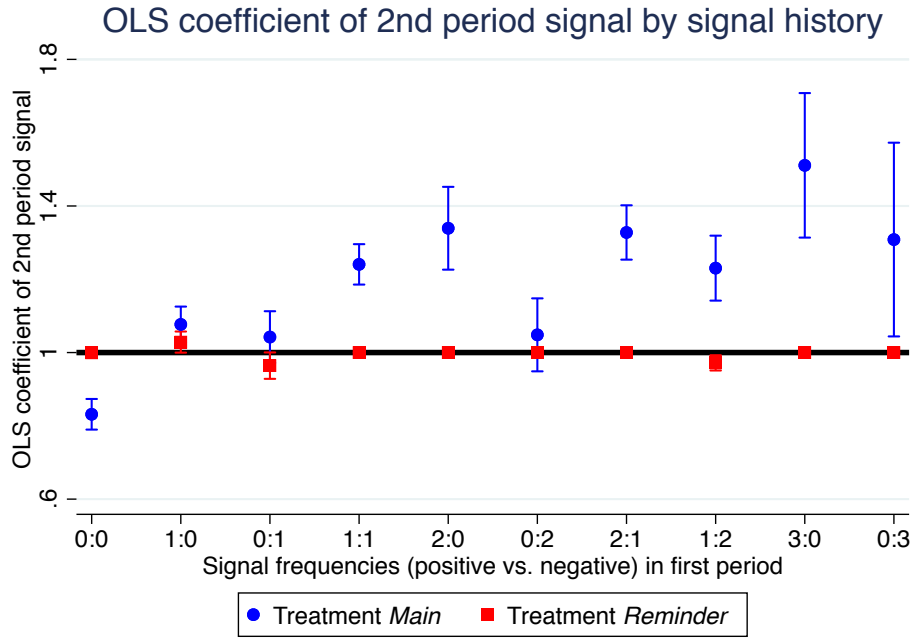


Figure 5: OLS coefficient in a regression of second-period beliefs on the last signal, separately for each set of signal frequencies in the first part. The regressions control for a subject’s first-period belief. Standard error bars are computed based on clustering at the subject level.

interacted with a treatment dummy. Again the table also includes a treatment dummy, which is suppressed for brevity.

The regression results document that, as predicted, the interaction term is positive and statistically significant. On average, subjects in *Main* react 21–30% more to the value of the second-period signal than subjects in *NoCue*. This is true for both the beliefs and the direct recall data. These results provide direct causal evidence for the role of associative recall in generating overreaction. Figure 6 visualizes the difference in regression coefficients between these two treatments.

6 Over- vs. Underreaction

Next, we turn to investigating predictions 6 and 7 in Section 3.3, which conceptually correspond to exogenously varying the parameter ρ in the simple model. For this purpose, as discussed in Section 3, we implemented treatments *Underreaction* and *Underreaction reminder*. Here, second-period signals were communicated with those contexts that belonged to the respective opposite signal in the first period. Thus, a signal now cues the recollection of the opposite past signals.

Table 6 presents the results. Columns (1) and (2) show that, within treatment *Underreaction*, the coefficient of the second-period signal is 0.74–0.76, substantially smaller

Table 5: Treatments *Main* vs. *NoCue*

	<i>Dependent variable:</i>					
	2nd period belief			Δ Recall [Pos. – Neg.]		
	(1)	(2)	(3)	(4)	(5)	(6)
2nd period signal	0.89*** (0.04)	0.88*** (0.04)	0.88*** (0.04)	0.76*** (0.05)	0.75*** (0.05)	0.77*** (0.06)
2nd period signal \times 1 if <i>Main</i> , 0 if <i>NoCue</i>	0.21*** (0.04)	0.22*** (0.04)	0.22*** (0.05)	0.29*** (0.06)	0.29*** (0.06)	0.30*** (0.07)
Belief in 1st period	0.62*** (0.03)			0.60*** (0.03)		
Stock price in 1st period		0.62*** (0.03)			0.61*** (0.03)	
Treatment FE	Yes	Yes	Yes	Yes	Yes	Yes
Session FE	No	No	Yes	No	No	Yes
1st period signal history FE	No	No	Yes	No	No	Yes
Company FE	No	No	Yes	No	No	Yes
Order FE	No	No	Yes	No	No	Yes
Subject FE	No	No	Yes	No	No	Yes
Observations	1600	1600	1600	1600	1600	1600
Adjusted R^2	0.68	0.68	0.67	0.62	0.63	0.63

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The sample includes all observations from treatments *Main* and *No Cue* where subjects observed a second-period signal. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

than one, indicating meaningful underreaction to the second-period signal. Columns (3) and (4) leverage exogenous variation in signal histories to document that underreaction strongly increases in the number of first-period signals that were communicated in the same context as the second-period signal, see the negative and statistically significant interaction term.⁷

These results suggest that underreaction is generated by asymmetric recall. To further corroborate this claim, Figure 9 in Appendix B analyzes the self-reported recall patterns in treatment *Underreaction* as a function of the signal history, akin to Figure 3 in Section 4. Here, we see that, in contrast to treatment *Main*, subjects' recall is much more precise for those first-period signals that *differ* in value from the second-period signal than for those signals that take on the same realization as the second-period signal. Again, this pattern is expected because those first-period signals that take on a different value as the second-period signal get cued by the second-period context.

Finally, columns (5) and (6) compare treatments *Underreaction* and *Underreaction re-*

⁷Table 12 in Appendix C shows that very similar results hold when we consider the recall data rather than subjects' beliefs.

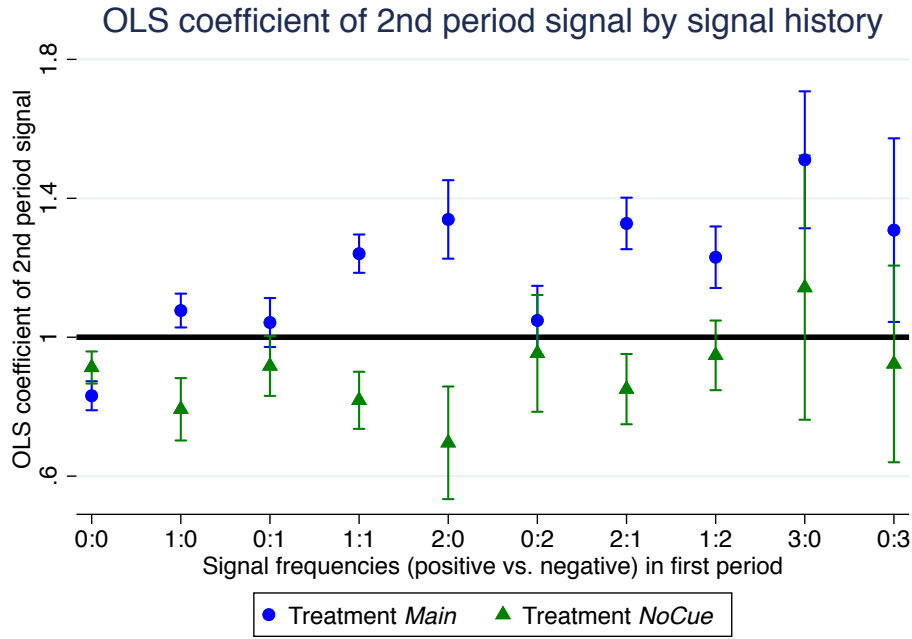


Figure 6: OLS coefficient in a regression of second-period beliefs on the last signal, separately for each set of signal frequencies in the first part. The regressions control for a subject’s first-period belief. Standard error bars are computed based on clustering at the subject level.

minder. As in Section 5.1, the coefficient of interest is now the interaction term between the second-period signal and a treatment dummy. The dummy is negative and statistically highly significant. The magnitude suggests that underreaction is 25–28% stronger in *Underreaction*. In contrast, as we can infer from the coefficient of the second-period signal, there is no underreaction in treatment *Underreaction reminder*, with a coefficient of 1.01, statistically indistinguishable from one. This pattern is again predicted because in our framework associative asymmetric recall cannot occur if memory is exogenously set to have no imperfections ($r = 1$).

Figure 7 visualizes the results in both treatments, separately for each signal history. The figure confirms that underreaction is present at all signal histories except for when there are no first-period signals, as predicted by our formal framework. In summary, the results show that associative memory can generate predictable patterns of both over- or underreaction, depending on which contexts a piece of news is associated with.

In summary, Sections 4, 5, and 6 have subjected equation (5) in Section 2 to a systematic experimental test. Based on initial evidence for overreaction in expectations (Table 2), we have exogenously manipulated the number of first-period signals that are identical to the second-period signal; we exogenously varied both the relevance of imperfect memory (Table 4) and the strength of associative recall (Table 5); and we exogenously varied the nature of the correlation between news and contexts (Table 6). Throughout, the results are in line with the predictions of equation (5).

Table 6: Treatments *Underreaction* and *Underreaction reminder*

	Dependent variable: 2nd period belief					
	Treatments:				+ Reminder	
	<i>Underreaction</i>					
	(1)	(2)	(3)	(4)	(5)	(6)
2nd period signal	0.76*** (0.04)	0.74*** (0.04)	0.99*** (0.04)	0.95*** (0.05)	1.01*** (0.02)	1.01*** (0.02)
Belief in 1st period	0.65*** (0.04)		0.50*** (0.05)		0.77*** (0.03)	
2nd period signal × # 1st period signals in same context			-0.31*** (0.06)	-0.28*** (0.07)		
2nd period signal × 1 if <i>Underreaction</i> , 0 if <i>Reminder</i>					-0.25*** (0.04)	-0.28*** (0.05)
Treatment FE	No	No	No	No	Yes	Yes
Session FE	No	Yes	No	Yes	No	Yes
Signal history FE	No	Yes	No	Yes	No	Yes
Company FE	No	Yes	No	Yes	No	Yes
Order FE	No	Yes	No	Yes	No	Yes
Subject FE	No	Yes	No	Yes	No	Yes
Observations	800	800	800	800	1300	1300
Adjusted R^2	0.67	0.68	0.68	0.70	0.79	0.79

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The table suppresses the coefficients of the number of identical first-period signals (columns (3)–(4)) and of a binary indicator for whether the signals are mixed (columns (5)–(6)). The sample includes all observations from treatments *Underreaction* and *Underreaction reminder* where subjects observed a second-period signal. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

7 Robustness: Extended Time Lag

7.1 Experimental Design

All treatments reported up to this point relied on a design in which the time lag between the first and second period was 15 minutes. As an extension and robustness check, we now investigate whether the effects of associative memory on expectation formation also prevail under a slightly longer time lag. For this purpose, we conducted treatments *Time lag* and *Time lag reminder*. These treatments followed the same procedure as treatments *Main* and *Reminder*, except that the time lag between the first and second period of the experiment was three days. When subjects signed up for the experiment, they registered for two separate sessions that were conducted at the same time of day, on a Tuesday and Friday.

On the first day, subjects completed the first period of the experiment, using the same experimental instructions and control questions as in the baseline treatments reported

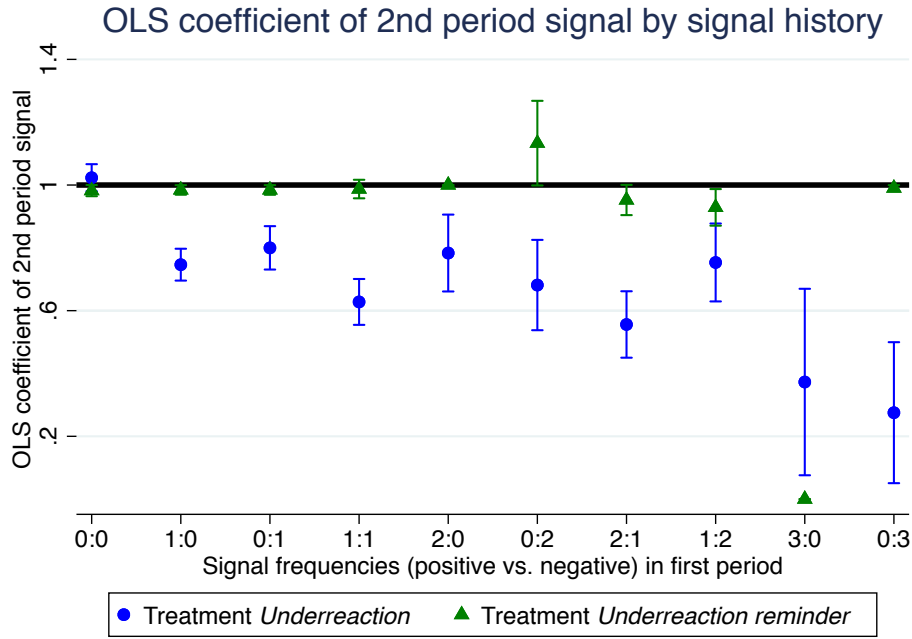


Figure 7: OLS coefficient in a regression of second-period beliefs on the last signal, separately for each set of signal frequencies in the first part. The regressions control for a subject's first-period belief. Standard error bars are computed based on clustering at the subject level.

above (with obvious minor modifications regarding the timing of the second period). Thus, throughout it was obvious to participants that the information in the first period would be relevant for the second period three days later. After the first period, participants completed the real effort task, the Raven matrices test as well as the demographic questionnaire.⁸ On the second day, participants re-read the original instructions and completed the same set of control questions again. Then, they completed the second period of the experiment.

These two treatments were also pre-registered in the original pre-registration described in Section 3.5. 80 subjects participated in treatment *Time lag* and 50 in treatment *Time lag reminder*. The treatments were randomized within experimental sessions and implemented in the BonnEconLab of the University of Bonn. The sessions for the first period lasted up to 60 minutes and those for the second period up to 45 minutes. Average earnings were 25 euros, including a 10 euros show-up fee. Participants were only paid in case they returned for the second part. In our experiments, attrition was negligible: 95% of subjects returned for the second period three days after the initial sessions.

⁸Thus, subjects could not take notes right after the first period.

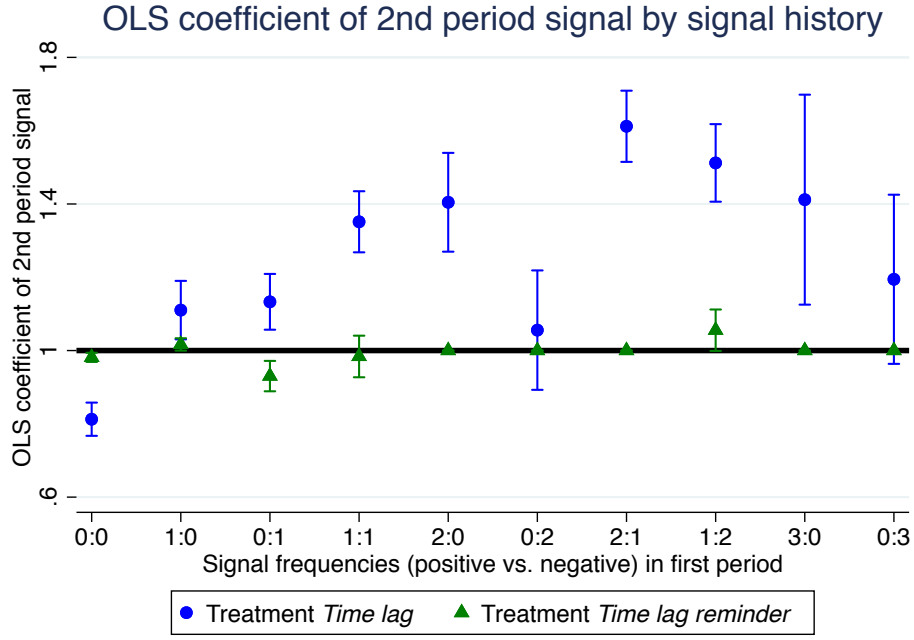


Figure 8: OLS coefficient in a regression of second-period beliefs on the last signal, separately for each set of signal frequencies in the first part. The regressions control for a subject's first-period belief. Standard error bars are computed based on clustering at the subject level.

7.2 Results

Figure 8 summarizes the results, which are very similar to those in treatments *Main* and *Reminder*: we see (i) overreaction; (ii) stronger overreaction when more first-period signals get cued by the second-period signal; and (iii) stronger overreaction relative to a treatment with a reminder. Tables 13 and 14 in Appendix C present corresponding regression analyses.

8 Estimating the Model

All analyses reported up to this point are motivated and structured through the formal framework laid out in Section 2. To supplement these reduced-form analyses, we now explicitly estimate this model, in particular its key memory parameters.

Specifically, we estimate the parameters $\hat{\beta}$, \hat{r} , and \hat{a} by minimizing the sum of squared residuals for the non-linear regression equation (compare equation (4) above):

$$b = 100 + \beta n_{k+1} + r \sum_{x=1}^k n_x + (1-r)a \sum_{x=1}^z n_x + \epsilon \quad (9)$$

Table 7 summarizes the estimates across treatments, where our interest is in the parameters that govern the strength of memory \hat{r} as well as of associative recall \hat{a} .

Table 7: Estimates of model parameters across treatments

Treatment	Memory imperfection ($1 - \hat{\tau}$)	Associative recall \hat{a}
<i>Main</i>	0.41*** (0.03)	0.79*** (0.08)
<i>Reminder</i>	0.01 (0.01)	-1.59 (4.75)
<i>NoCue</i>	0.51*** (0.04)	0.01 (0.12)
<i>Underreaction</i>	0.68*** (0.03)	0.35*** (0.06)
<i>Underreaction reminder</i>	0.05*** (0.01)	0.73** (0.33)
<i>Time lag</i>	0.51*** (0.03)	0.61*** (0.07)
<i>Time lag reminder</i>	0.02 (0.02)	0.31 (1.06)

Notes. Estimates of equations (9), standard errors (clustered at subject level) reported in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

The quantitative estimates resonate with the results documented above and provide interesting cross-treatment comparisons. In treatment *Main*, we estimate a substantial role for associative recall. The estimates imply that participants recall non-cued signals with probability 59% and cued ones with probability 91%. In treatment *Reminder*, we confirm that imperfect memory entirely disappears (by construction of the treatment), so that associative recall cannot be measured with reasonable precision (compare the large standard error). Similar patterns prevail in *Underreaction reminder* and *Time lag reminder*. Analogously, we see that in treatment *NoCue*, associative recall collapses to zero, again by construction of the treatment.

Finally, note that estimated memory imperfection is larger in *Underreaction* and *Time lag* than in *Baseline*. This is intuitive since (i) in *Underreaction* participants had to deal with the additional difficulty that the mapping between news and contexts changed between the first and second period and (ii) in *Time lag*, subjects had to memorize news for three days rather than 15–30 minutes.

9 Discussion

This paper has provided a theoretically-structured experimental analysis of the roles of imperfect and associative memory for belief formation. The notion of associative recall has recently received increased attention from economic theorists, yet direct experimen-

tal evidence on the importance of cued recall in structured economic decision environments is limited.

In our experiments, (i) participants' beliefs strongly overreact to the latest piece of news; (ii) the presence and magnitude of such overreaction depend on the precise signal history in predictable ways; (iii) exogenously manipulating the degree of memory imperfection provides causal evidence that without imperfect memory, overreaction does not occur; (iv) exogenously manipulating the strength of associative recall provides causal evidence that in our context associative memory is necessary in order for overreaction to arise; (v) associative memory generates predictable over- or underreaction in expectations, depending on the precise mapping that links types of news with certain contexts; (vi) the effect of associative memory on expectation overreaction holds both over a rather short (15 minutes) and a somewhat longer (three days) time horizon; and (vii) a direct estimation of our simple model suggests that associative memory plays a quantitatively large role in generating observed beliefs.

In combination, we view our experiments as clean and theoretically-founded experimental investigation of the relevance of associative recall for belief formation. We believe that by offering a new experimental paradigm in which these types of effects can be studied, our paper opens up the possibility for further experimental research in an agenda on memory imperfections and belief formation. To take just one example, the simple model that structures our analysis rests on the assumption that people are fully naïve about their tendency to selectively recall signals that get cued by the current context. While our results corroborate this assumption at least partly (since otherwise we would not have identified overreaction in its different variations), there is scope for more direct measurements of people's sophistication or naïveté about their own recall technology.

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ONLINE APPENDIX

A Additional Derivations

A.1 Partial Naïveté

The two forms of naivete would be captured in slightly different ways:

- Suppose the DM to some extent (captured by naivete parameter α , $\alpha = 0$ captures full naivete, $\alpha = 1$ captures full sophistication) fails to realize that he sometimes forgets. When he does realize, however, then he correctly (in a Bayesian sense) infers the realization of the information based on his memory parameters r and a . Specifically, $Pr(n_{nr} = n_{k+1} | norecall) = \frac{1-a}{2-a}$. The DM's forecast would thus be given by $f = v + n_{k+1} + \sum_{x=1}^k m_x n_x + \alpha \sum_{x=1}^k (1 - m_x) \frac{-a}{2-a} n_{k+1}$
- Suppose the DM instead fully realizes that he sometimes forgets, but is naive in the way he infers what he forgot. This form of naivete could be captured by DM's belief \hat{a} about memory parameter a , $\hat{a} \leq a$ ($\hat{a} = 0$ captures full naivete, $\hat{a} = 1$ captures full sophistication). The DM would do inference as outlined above, but would use \hat{a} instead of a to infer the value of signals he does not recall. The DM's forecast would thus be given by $f = v + n_{k+1} + \sum_{x=1}^k m_x n_x + \sum_{x=1}^k (1 - m_x) \frac{-\hat{a}}{2-\hat{a}} n_{k+1}$.

A.2 Derivation of OLS Estimator

We formally derive the relationship between equation (5) and the OLS estimator $\hat{\beta}_1$ in equation (8). Keeping the notation that b is the belief and n the news, then with N observations (subject-tasks) the OLS estimator is given by

$$E[\hat{\beta}] = E \left[\frac{\sum n_i b_i - 1/N \sum n_i \sum b_i}{\sum n_i^2 - 1/N (\sum n_i^2)} \right] \quad (10)$$

This expectation of a ratio can be approximated by the ratio of the expectations (also, the expectation of a ratio equals the ratio of probability limits). Denote $c = v + r \sum n_x$, which is not a function of the last signal. Substitute in for the forecast. Observing that

$E[n_i] = 0$, we get

$$E[\hat{\beta}] = \frac{\sum E[n_i b_i] - 1/N \sum E[n_i] \sum E[b_i]}{\sum E[n_i^2] - 1/N E(\sum n_i^2)} \quad (11)$$

$$= \frac{\sum E[n_i [n_i (1 + z_i (1 - r) \rho a) + c]]}{\sum E[n_i^2]} \quad (12)$$

$$= 1 + (1 - r) \rho a \frac{\sum E[n_i^2 z_i]}{\sum E[n_i^2]} \quad (13)$$

$$= 1 + (1 - r) \rho a \frac{\sum E[n_i^2] E[z_i]}{\sum E[n_i^2]} \quad (14)$$

$$= 1 + (1 - r) \rho a \frac{\bar{z} \sum E[n_i^2]}{\sum E[n_i^2]} \quad (15)$$

$$= 1 + (1 - r) \rho a \bar{z} \quad (16)$$

Because z_i and n_i are independent.

B Additional Figures

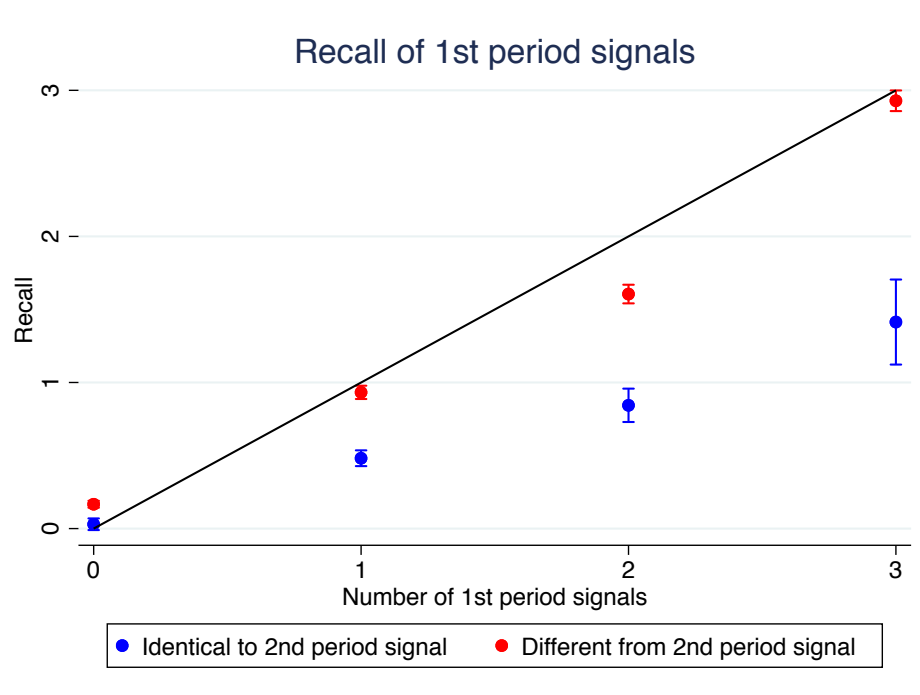


Figure 9: Recall of first-period signals in Treatment *Underreaction*, depending on whether the second-period signal was identical to or different from the first-period signals. We construct the recall variables as follows. In the case of recall of signals that are different from the second-period signal, we use the reported recall quantity. In the case of recall of signals that are identical to the second-period signal, we use the reported recall minus one. That is, we make the arguably very plausible assumption that subjects always remember the value of the second-period signal that they just saw a few seconds ago.

C Additional Tables

Table 8: Beliefs in the first period

	<i>Dependent variable: 1st period belief</i>					
	<i>Main</i>		<i>Reminder</i>		<i>No Cue</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
Stock price in 1st period	0.98*** (0.01)	0.97*** (0.01)	1.00*** (0.00)	1.00*** (0.00)	0.99*** (0.01)	1.00*** (0.01)
Session FE	No	Yes	No	Yes	No	Yes
Company FE	No	Yes	No	Yes	No	Yes
Order FE	No	Yes	No	Yes	No	Yes
Subject FE	No	Yes	No	Yes	No	Yes
Observations	800	800	500	500	800	800
Adjusted R^2	0.96	0.97	1.00	1.00	0.99	0.99

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 9: Beliefs in the second period in case of no signal in second period

	<i>Dependent variable: 2nd period belief</i>					
	<i>Main</i>		<i>Reminder</i>		<i>No Cue</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
Belief in 1st period	0.56*** (0.09)	1.00*** (0.00)	1*** (0.00)	1.00*** (0.00)	0.51*** (0.08)	1.37* (0.81)
Session FE	No	Yes	No	Yes	No	Yes
1st period signal history FE	No	Yes	No	Yes	No	Yes
Company FE	No	Yes	No	Yes	No	Yes
Order FE	No	Yes	No	Yes	No	Yes
Subject FE	No	Yes	No	Yes	No	Yes
Observations	160	160	100	100	160	160
Adjusted R^2	0.39	1.00	1.00	1.00	0.36	0.34

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 10: Treatment *Main*: Recall data

	<i>Dependent variable:</i> Δ Recall [Pos. – Neg.]						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
2nd period signal	1.05*** (0.04)	1.06*** (0.03)	1.07*** (0.04)	0.81*** (0.05)	0.83*** (0.05)	0.79*** (0.05)	0.83*** (0.05)
Belief in 1st period	0.74*** (0.03)					0.57*** (0.05)	
Stock price in 1st period		0.74*** (0.03)					
Value of cued 1st period signals				0.92*** (0.03)	0.90*** (0.03)		
Value of non-cued 1st period signals				0.58*** (0.05)	0.58*** (0.05)		
2nd period signal \times # 1st period signals in same context						0.36*** (0.05)	0.31*** (0.05)
Session FE	No	No	Yes	No	Yes	No	Yes
1st period signal history FE	No	No	Yes	No	No	No	Yes
Company FE	No	No	Yes	No	Yes	No	Yes
Order FE	No	No	Yes	No	Yes	No	Yes
Subject FE	No	No	Yes	No	Yes	No	Yes
Observations	800	800	800	800	800	800	800
Adjusted R^2	0.76	0.77	0.77	0.78	0.78	0.78	0.78

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 11: Treatment *Main*: Heterogeneity analysis (recall data)

	<i>Dependent variable:</i> Δ Recall [Pos. – Neg.]					
	(1)	(2)	(3)	(4)	(5)	(6)
2nd period signal	1.29*** (0.13)	1.27*** (0.13)	1.21*** (0.06)	1.19*** (0.05)	1.15*** (0.05)	1.15*** (0.05)
2nd period signal \times Raven score	-0.045* (0.02)	-0.042* (0.02)				
2nd period signal \times Memory for non-cued signals			-0.30*** (0.08)	-0.26*** (0.08)		
2nd period signal \times Response time recall					-0.45** (0.19)	-0.50*** (0.19)
Belief in 1st period	0.74*** (0.03)		0.74*** (0.03)		0.74*** (0.03)	
Session FE	No	Yes	No	Yes	No	Yes
1st period signal history FE	No	Yes	No	Yes	No	Yes
Company FE	No	Yes	No	Yes	No	Yes
Order FE	No	Yes	No	Yes	No	Yes
Observations	800	800	800	800	800	800
Adjusted R^2	0.76	0.77	0.76	0.77	0.76	0.77

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The table suppresses the coefficients of Raven score (columns (1) and (4)), memory for non-cued signals (columns (2) and (5)), and response time (columns (3) and (6)). Response times are measured in minutes. The sample includes all observations from treatment *Main* where subjects observed a second-period signal. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 12: Treatments *Underreaction* and *Reminder underreaction*: Recall data

	<i>Dependent variable:</i> 2nd period belief					
	Treatments:					
	<i>Underreaction</i>				+ <i>Reminder</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
2nd period signal	0.62*** (0.05)	0.60*** (0.05)	0.86*** (0.05)	0.81*** (0.06)	0.89*** (0.03)	0.91*** (0.04)
Belief in 1st period	0.66*** (0.04)		0.49*** (0.05)		0.76*** (0.03)	
2nd period signal × # 1st period signals in same context			-0.33*** (0.06)	-0.28*** (0.06)		
2nd period signal × 1 if <i>Underreaction</i> , 0 if <i>Reminder</i>					-0.28*** (0.06)	-0.32*** (0.06)
Treatment FE	No	No	No	No	Yes	Yes
Session FE	No	Yes	No	Yes	No	Yes
Signal history FE	No	Yes	No	Yes	No	Yes
Company FE	No	Yes	No	Yes	No	Yes
Order FE	No	Yes	No	Yes	No	Yes
Subject FE	No	Yes	No	Yes	No	Yes
Observations	800	800	800	800	1300	1300
Adjusted R^2	0.59	0.61	0.61	0.63	0.71	0.72

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The table suppresses the coefficients of the number of identical first-period signals (columns (3)–(4)) and of a binary indicator for whether the signals are mixed (columns (5)–(6)). The sample includes all observations from treatments *Time lag* and *Reminder time lag* where subjects observed a second-period signal. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 13: Treatments *Time lag* and *Time lag reminder*

	<i>Dependent variable: 2nd period belief</i>					
	Treatments:					
	<i>Time lag</i>				<i>+ Reminder time lag</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
2nd period signal	1.17*** (0.04)	1.16*** (0.04)	0.85*** (0.05)	0.84*** (0.06)	1.02*** (0.02)	1.03*** (0.02)
Belief in 1st period	0.52*** (0.04)		0.33*** (0.05)		0.71*** (0.03)	
2nd period signal × # 1st period signals in same context			0.44*** (0.05)	0.45*** (0.07)		
2nd period signal × 1 if <i>Time lag</i> , 0 if <i>Reminder t. l.</i>					0.18*** (0.04)	0.15*** (0.05)
Treatment FE	No	No	No	No	Yes	Yes
Session FE	No	Yes	No	Yes	No	Yes
Signal history FE	No	Yes	No	Yes	No	Yes
Company FE	No	Yes	No	Yes	No	Yes
Order FE	No	Yes	No	Yes	No	Yes
Subject FE	No	Yes	No	Yes	No	Yes
Observations	800	800	800	800	1300	1300
Adjusted R^2	0.68	0.67	0.70	0.70	0.77	0.77

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The table suppresses the coefficients of the number of identical first-period signals (columns (3)–(4)) and of a binary indicator for whether the signals are mixed (columns (5)–(6)). The sample includes all observations from treatments *Time lag* and *Time lag reminder* where subjects observed a second-period signal. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table 14: Treatments *Time lag* and *Reminder time lag*: Recall data

	<i>Dependent variable:</i> Δ Recall [Pos. – Neg.]					
	Treatments:					
	<i>Time lag</i>			<i>+ Reminder time lag</i>		
	(1)	(2)	(3)	(4)	(5)	(6)
2nd period signal	0.97*** (0.05)	0.96*** (0.06)	0.66*** (0.05)	0.66*** (0.06)	0.97*** (0.03)	0.98*** (0.03)
Belief in 1st period	0.52*** (0.04)		0.33*** (0.05)		0.70*** (0.03)	
2nd period signal \times # 1st period signals in same context			0.43*** (0.06)	0.43*** (0.07)		
2nd period signal \times 1 if <i>Time lag</i> , 0 if <i>Reminder t. l.</i>					0.022 (0.06)	-0.0011 (0.06)
Treatment FE	No	No	No	No	Yes	Yes
Session FE	No	Yes	No	Yes	No	Yes
Signal history FE	No	Yes	No	Yes	No	Yes
Company FE	No	Yes	No	Yes	No	Yes
Order FE	No	Yes	No	Yes	No	Yes
Subject FE	No	Yes	No	Yes	No	Yes
Observations	800	800	800	800	1300	1300
Adjusted R^2	0.57	0.57	0.60	0.60	0.70	0.70

Notes. OLS estimates, robust standard errors (in parentheses) are clustered at the subject level. The table suppresses the coefficients of the number of identical first-period signals (columns (3)–(4)) and of a binary indicator for whether the signals are mixed (columns (5)–(6)). The sample includes all observations from treatments *Time lag* and *Reminder time lag* where subjects observed a second-period signal. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

D Experimental Instructions

to be added

E Experimental Materials

E.1 Screenshots of signal, story and image presentation



Company N tries to advertise its products through commercials with German celebrities, like, for instance, Boris Becker, Helene Fischer or Til Schweiger. Recently, a new advertisement campaign with a celebrity worked extremely well.

The value of the company increased by 10 points.

Figure 10: Example screenshot of how a piece of positive news for Company N is communicated to subjects. The signal is displayed in the last line of the text. A story and an image accompany the signal.



The head of sales of Company K is a choleric. Every once in a while, he engages in temper tantrums during which he yells at customers of Company K and insults them. These customers hence take their business elsewhere. Just now, another temper tantrum occurred.

The value of the companied decreased by 10 points.

Figure 11: Example screenshot of how a piece of negative news for Company K is communicated to subjects. The signal is displayed in the last line of the text. A story and an image accompany the signal.