

Ben Juarez

ALM 116 - PS5

① $n = 10$, $1:30 - 5:30$ pm, $\Delta t = 30$ min, T_1, \dots, T_n independent exponential

② Find $E[W_i]$

$$W_i = \begin{cases} 0 & \text{if } T_i \leq 30 \\ T_i - 30 & \text{if } T_i > 30 \end{cases} \Rightarrow E[W_i] = P(T_i \leq 30) \cdot 0 + P(T_i > 30) \cdot E[W_i | T_i > 30]$$
$$= P(T_i > 30) \cdot E[W_i | T_i > 30]$$
$$= \int_{30}^{\infty} (T-30) \frac{1}{30} \cdot e^{-\frac{T}{30}} dT$$
$$= 30e^{-1}$$
$$E[W_i] = \frac{30}{e}$$

③ See attached MATLAB code

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② Let $\{N_t, t \geq 0\}$ be a Poisson process w/ intensity λ

① Find prob. of no events in $(t_1, t_2]$

$$\Rightarrow P(N_{t_2} - N_{t_1} = 0) = e^{-\lambda(t_2 - t_1)} \cdot \frac{(\lambda(t_2 - t_1))^0}{0!} = e^{-\lambda(t_2 - t_1)}$$

② Find probability that there is exactly 1 event in each of given intervals

intervals $(0, 1], (1, 2], \dots, (n-1, n]$

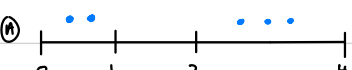
$$\Rightarrow \Delta t = 1 \text{ for each increment} \Rightarrow P(N_{t_1} - N_{t_0} = 1) = P(N_{t_2} - N_{t_1} = 1) = \dots = P(N_{t_n} - N_{t_{n-1}} = 1)$$

* independent stationary increments

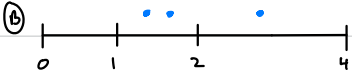
$$\Rightarrow P(N_{t_n} - N_{t_{n-1}} = 1) \text{ for each } n \text{ is equivalent to } (P(N_1 - N_0 = 1))^n$$

$$\Rightarrow (P(N_1 - N_0 = 1))^n = (e^{-\lambda} \frac{\lambda^1}{1!})^n = (\lambda e^{-\lambda})^n$$

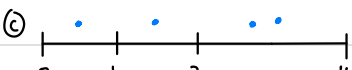
③ Find probability that there are two events in $(0, 2]$ and three in $(1, 4]$

①  $\Rightarrow P(N_1 - N_0 = 2) \cdot P(N_2 - N_1 = 0) \cdot P(N_4 - N_2 = 3) = (e^{-\lambda} \frac{\lambda^2}{2!}) (e^{-\lambda}) (e^{-\lambda} \frac{(\lambda^3)}{6})$

$$= e^{-4\lambda} \left(\frac{\lambda^5}{2 \cdot 6} \right)$$

②  $\Rightarrow P(N_1 - N_0 = 0) \cdot P(N_2 - N_1 = 2) \cdot P(N_4 - N_2 = 1) = (e^{-\lambda}) (e^{-\lambda} \frac{\lambda^2}{2!}) (e^{-\lambda} \lambda)$

$$= e^{-4\lambda} (\lambda^3)$$

③  $\Rightarrow P(N_1 - N_0 = 1) \cdot P(N_2 - N_1 = 1) \cdot P(N_4 - N_2 = 2) = (e^{-\lambda} \lambda) (e^{-\lambda} \lambda) (e^{-\lambda} \frac{(\lambda^2)}{2})$

$$= e^{-4\lambda} (2\lambda^4)$$

$$\Rightarrow P(N_2 - N_0 = 2, N_4 - N_2 = 3) = P_A + P_B + P_C$$

$$= e^{-4\lambda} \left(\frac{\lambda^5}{2} \right) + e^{-4\lambda} (\lambda^3) + e^{-4\lambda} (2\lambda^4)$$

$$= \lambda^5 e^{-4\lambda} \left(\frac{1}{2} \lambda^2 + 1 + 2\lambda \right)$$

③ • let $\{N_t, t \geq 0\}$ be Poisson process w/ rate λ w/ $T_1 \sim \text{Exp}(\lambda)$

Find cond. distribution of T_1 given $t_0 \uparrow N_t = 1$

$$\Rightarrow f_{T_1|N_t=1}(t_1|1) = \frac{f_{T_1, N_t}(t_1, 1)}{P(N_t=1)} = \frac{P(T_1 \leq t_1, N_t=1)}{P(N_t=1)} \quad \text{for } t_1 \leq t$$

$$\begin{aligned} \Rightarrow \frac{P(T_1 \leq t_1, N_t=1)}{P(N_t=1)} &= \frac{P(T_1 \leq t_1, N_t=1)}{(\lambda t) e^{-\lambda t}} \\ &= \frac{P(N_t - N_0 = 1, N_t - N_{t_1} = 0)}{(\lambda t) e^{-\lambda t}} \\ &= \frac{P(N_{t_1} - N_0 = 1) \cdot P(N_t - N_{t_1} = 0)}{(\lambda t) e^{-\lambda t}} \\ &= \frac{(\lambda t_1) e^{-\lambda t_1} \cdot e^{-\lambda(t-t_1)}}{(\lambda t) e^{-\lambda t}} \\ &= \frac{(\lambda t_1) e^{-\lambda t_1} \cdot e^{-\lambda t} \cdot e^{\lambda t_1}}{(\lambda t) e^{-\lambda t}} \end{aligned}$$

$$\Rightarrow f_{T_1|N_t=1}(t_1|1) = \frac{t_1}{t}$$

\Rightarrow Uniform distribution $U(\alpha=0, \beta=t]$

④ a) Find probability that student's wait time is 0

$$P(N_{\tau+t} - N_t = 0) = e^{-\lambda \tau} \cdot \frac{(\lambda \tau)^0}{0!} = e^{-\lambda \tau}$$

b) Find expected waiting time

• Let W denote waiting time

• Let car i be next car

• Let t_i denote time until $\Rightarrow W = \begin{cases} 0 & \text{if } t_i > \tau \\ t_i + W & \text{if } t_i \leq \tau \end{cases}$

car i passes

$$\Rightarrow E[W] = P(t_i > \tau) \cdot 0 + P(t_i \leq \tau) \cdot (E[t_i | t_i \leq \tau] + E[W])$$

$$E[W] = P(t_i \leq \tau) \cdot (E[t_i | t_i \leq \tau] + E[W])$$

$$E[W] = P(t_i \leq \tau) \cdot E[t_i | t_i \leq \tau] + P(t_i \leq \tau) \cdot E[W]$$

$$E[W] - P(t_i \leq \tau) \cdot E[W] = P(t_i \leq \tau) \cdot E[t_i | t_i \leq \tau]$$

$$E[W] = \frac{P(t_i \leq \tau) \cdot E[t_i | t_i \leq \tau]}{1 - P(t_i \leq \tau)}$$

$$= \frac{P(t_i \leq \tau) E[t_i | t_i \leq \tau]}{1 - (1 - P(t_i > \tau))}$$

$$= \frac{\int_0^\tau t \lambda e^{-\lambda t} dt}{e^{-\lambda \tau}}$$

$$= \frac{\frac{1}{\lambda} (1 - (\lambda \tau + 1) e^{-\lambda \tau})}{e^{-\lambda \tau}}$$

$$= \frac{\frac{1}{\lambda} (1 - \lambda \tau e^{-\lambda \tau} - e^{-\lambda \tau})}{e^{-\lambda \tau}}$$

$$\Rightarrow E[W] = \frac{1}{\lambda} (e^{-\lambda \tau} - \lambda \tau - 1)$$

* $P(t_i \leq \tau) = 1 - P(t_i > \tau) \approx$ at least one car comes before next τ seconds

$$* P(t_i > \tau) = e^{-\lambda \tau}$$

⑤ a) Find probability of winning under τ -strategy

• stop after 1st customer after $\tau \leq \tau^*$ (assume $\tau^* > \frac{1}{\lambda}$)

$$\Rightarrow P(N_{\tau^*} - N_{\tau} = 1) = \lambda(\tau^* - \tau) \cdot e^{-\lambda(\tau^* - \tau)}$$

⑥ Find optimal value τ_{opt} of τ that maximizes probability of winning

\Rightarrow Let p_w denote $P(N_{\tau^*} - N_{\tau} = 1)$

$$\begin{aligned} \Rightarrow \frac{dp_w}{d\tau} &= \frac{d}{d\tau} (\lambda(\tau^* - \tau) \cdot e^{-\lambda(\tau^* - \tau)}) = 0 = \lambda(\tau^* - \tau) \cdot \lambda e^{-\lambda(\tau^* - \tau)} - \lambda e^{-\lambda(\tau^* - \tau)} \\ &= \lambda e^{-\lambda(\tau^* - \tau)} (\lambda(\tau^* - \tau) - 1) \\ &= \lambda(\tau^* - \tau) - 1 \end{aligned}$$

$$1 = \lambda\tau^* - \lambda\tau \quad \Rightarrow \quad \tau_{opt} = \tau^* - \frac{1}{\lambda} \quad * \tau^* > \frac{1}{\lambda}$$

⑦ What is probability of winning under optimal strategy?

$$\Rightarrow P(N_{\tau^*} - N_{\tau} = 1 \mid \tau = \tau_{opt}) = \lambda(\tau^* - (\tau^* - \frac{1}{\lambda})) \cdot e^{-\lambda(\tau^* - (\tau^* - \frac{1}{\lambda}))} = e^{-1} = \frac{1}{e}$$

⑧ What is expected # of croissants under optimal strategy?

$$\begin{aligned} \Rightarrow E[C] &= E[N_{\tau_{opt}} - N_0] + P(N_{\tau^*} - N_{\tau_{opt}} \geq 1) \\ &= \lambda\tau_{opt} + (1 - P(N_{\tau^*} - N_{\tau_{opt}} = 1)) \\ &= \lambda\tau_{opt} + (1 - P(N_{\tau^*} - N_{\tau_{opt}} = 0)) \\ &= \lambda\tau_{opt} + (1 - e^{-\lambda(\tau^* - \tau_{opt})}) \\ &= \lambda(\tau^* - \frac{1}{\lambda}) + (1 - e^{-\lambda(\tau^* - \tau^* + \frac{1}{\lambda})}) \\ &= \lambda\tau^* - 1 + 1 - e^{-1} \end{aligned}$$

$$\Rightarrow E[C] = \lambda\tau^* - \frac{1}{e} \quad \text{croissants}$$

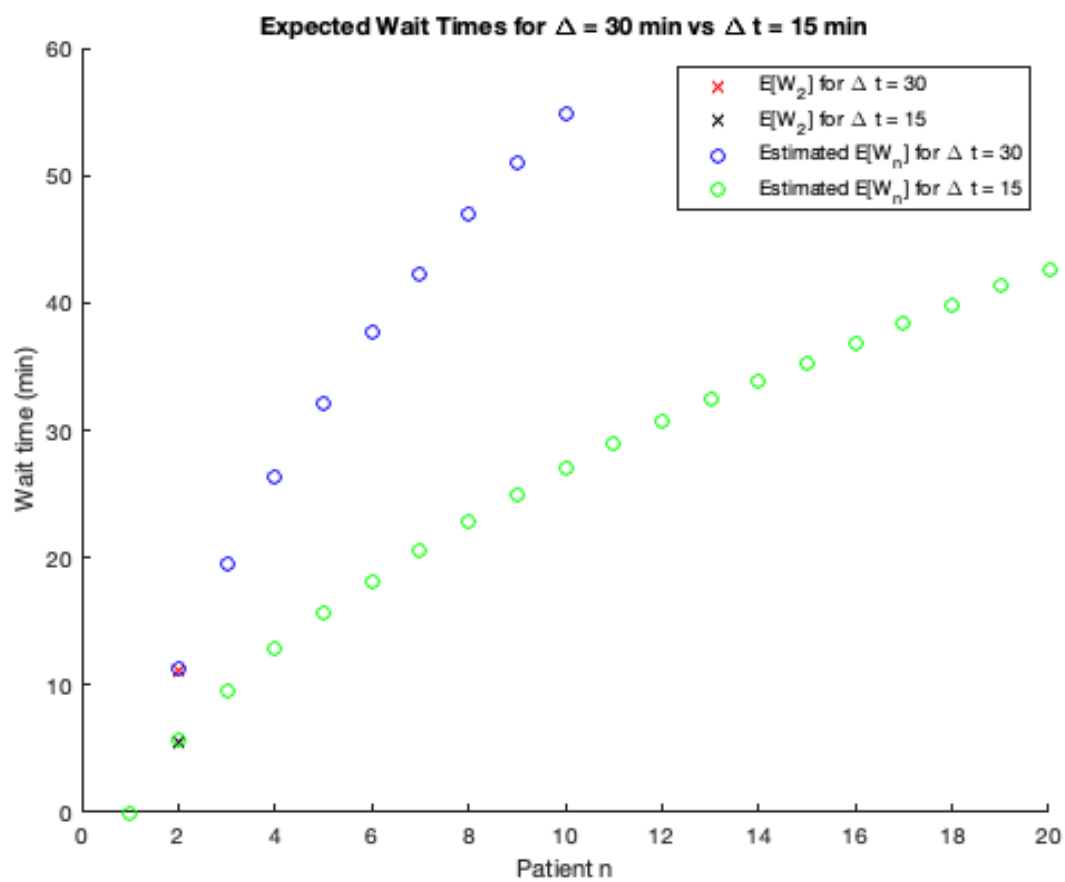
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% Ben Juarez - PS4Q1bc

% Part b
delta_t1 = 30;
delta_t2 = 15;
n1 = 10;
n2 = 20;
N=10^4;
E_W1 = zeros(0,n1);
E_W2 = zeros(0,n2);
for i = 1:N
    W_n1 = zeros(1,n1);
    W_n2 = zeros(1,n2);
    for j = 1:n1
        if j == 1
            W_n1(j) = 0;
        else
            W_n1(j) = max(0, W_n1(j-1) + exprnd(delta_t1) - delta_t1);
        end
    end
    for j = 1:n2
        if j == 1
            W_n2(j) = 0;
        else
            W_n2(j) = max(0, W_n2(j-1) + exprnd(delta_t2) - delta_t2);
        end
    end
    E_W1 = [E_W1; W_n1];
    E_W2 = [E_W2; W_n2];
end
E_W1 = mean(E_W1);
E_W2 = mean(E_W2);

hold on
scatter(2, delta_t1 / exp(1), "X", "red")
scatter(2, delta_t2 / exp(1), "X", "black")
scatter(1:1:n1, E_W1, "o", "blue")
scatter(1:1:n2, E_W2, "o", "green")
legend("E[W_2] for \Delta t = 30", ...
    "E[W_2] for \Delta t = 15", ...
    "Estimated E[W_n] for \Delta t = 30", ...
    "Estimated E[W_n] for \Delta t = 15")
title("Expected Wait Times for \Delta = 30 min vs \Delta t = 15 min")
xlabel("Patient n")
ylabel("Wait time (min)")
hold off
% Part c
% We can see that the wait time is estimated to be lower for the 20th
% patient with delta t = 15 min (approx 43-44 min) compared to the 10th
% patient with delta t = 30 min (approx 53-54 min). It is better to be
% be the 20th patient with delta t = 15 min!

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