

# Ec/ACM/CS 112 - PS6

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## Question 1

### preliminaries

```
rm(list=ls())
set.seed(123)
library(LaplacesDemon)
library(coda)
```

### Step 2: Program the Gibbs sampler

```
GibbsSampling = function(paramInit, y1, y2, y3, y4, y5, y6, nSamples) {
  n1 = length(y1)
  n2 = length(y2)
  n3 = length(y3)
  n4 = length(y4)
  n5 = length(y5)
  n6 = length(y6)
  nTot = n1 + n2 + n3 + n4 + n5 + n6
  nDim = length(paramInit)
  paramSamples = array(rep(NA, nSamples * nDim), dim = c(nSamples, nDim))
  theta_hat = function(mu, sigma, tau, n, y) {
    numerator = mu/tau^2 + (n*mean(y))/sigma^2
    denom = (1/tau^2) + (n/sigma^2)
    return(numerator/denom)
  }
  sd_theta = function(sigma, tau, n) {
    denom = (1/tau^2) + (n/sigma^2)
    return(sqrt(1/denom))
  }
  paramLast = paramInit
  for (t in 1:nSamples) {
    theta_1_Inner = rnorm(1, theta_hat(mu = paramLast[7], sigma = paramLast[8],
                                       tau = paramLast[9], n = n1, y = y1),
                        sd_theta(sigma = paramLast[8], tau = paramLast[9], n = n1))
    theta_2_Inner = rnorm(1, theta_hat(mu = paramLast[7], sigma = paramLast[8],
                                       tau = paramLast[9], n = n2, y = y2),
                        sd_theta(sigma = paramLast[8], tau = paramLast[9], n = n2))
    theta_3_Inner = rnorm(1, theta_hat(mu = paramLast[7], sigma = paramLast[8],
                                       tau = paramLast[9], n = n3, y = y3),
                        sd_theta(sigma = paramLast[8], tau = paramLast[9], n = n3))
    theta_4_Inner = rnorm(1, theta_hat(mu = paramLast[7], sigma = paramLast[8],
```

```

        tau = paramLast[9], n = n4, y = y4),
        sd_theta(sigma = paramLast[8], tau = paramLast[9], n = n4))
theta_5_Inner = rnorm(1, theta_hat(mu = paramLast[7], sigma = paramLast[8],
        tau = paramLast[9], n = n5, y = y5),
        sd_theta(sigma = paramLast[8], tau = paramLast[9], n = n5))
theta_6_Inner = rnorm(1, theta_hat(mu = paramLast[7], sigma = paramLast[8],
        tau = paramLast[9], n = n6, y = y6),
        sd_theta(sigma = paramLast[8], tau = paramLast[9], n = n6))
theta_Inner = c(theta_1_Inner, theta_2_Inner, theta_3_Inner, theta_4_Inner,
        theta_5_Inner, theta_6_Inner)
mu_Inner = rnorm(1, mean(theta_Inner), paramLast[9] / sqrt(nGroups))
sampVar_1 = sum((y1 - theta_1_Inner)^2)
sampVar_2 = sum((y2 - theta_2_Inner)^2)
sampVar_3 = sum((y3 - theta_3_Inner)^2)
sampVar_4 = sum((y4 - theta_4_Inner)^2)
sampVar_5 = sum((y5 - theta_5_Inner)^2)
sampVar_6 = sum((y6 - theta_6_Inner)^2)
sigma_Hat_2 = (sampVar_1 + sampVar_2 + sampVar_3 + sampVar_4
        + sampVar_5 + sampVar_6) / nTot
sigma_2_Inner = rinvcchisq(1, nTot, sigma_Hat_2)
tau_Hat_2 = sum((theta_Inner - mu_Inner)^2) / (nGroups - 1)
tau_2_Inner = rinvcchisq(1, nGroups - 1, tau_Hat_2)
paramSamples[t,] = c(theta_1_Inner, theta_2_Inner, theta_3_Inner, theta_4_Inner,
        theta_5_Inner, theta_6_Inner, mu_Inner,
        sqrt(sigma_2_Inner), sqrt(tau_2_Inner))
paramLast = paramSamples[t,]
}
return(paramSamples)
}

```

### step 3: Sampling and diagnostics

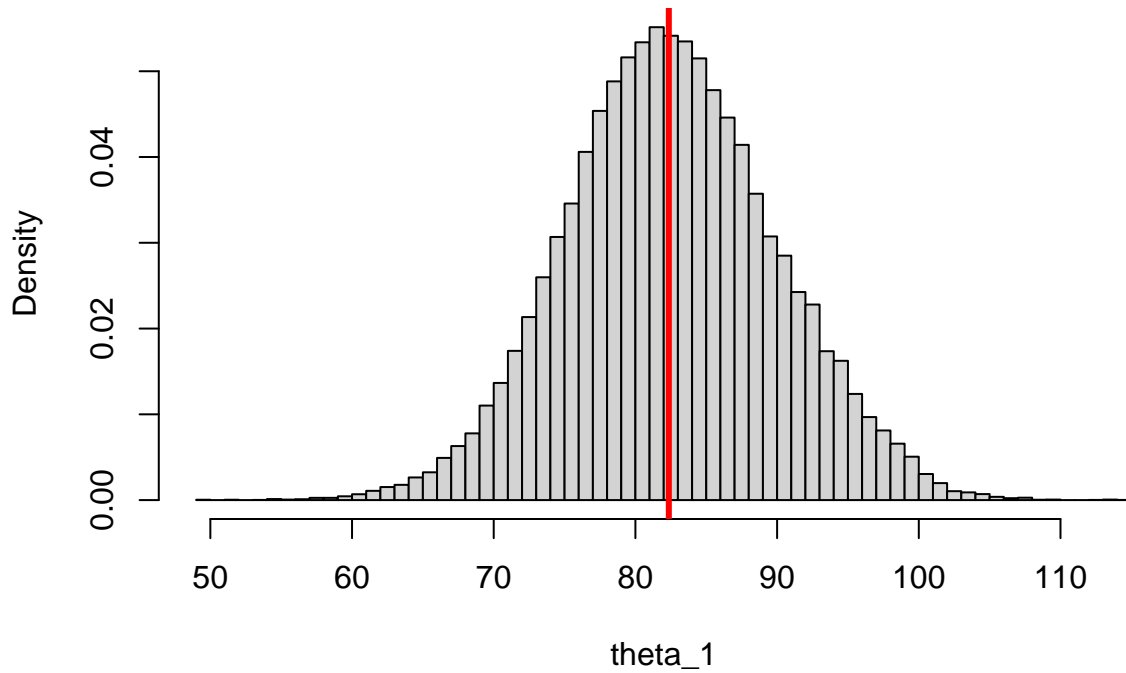
```

nGroups = 6
y1 = c(83,92,92,46)
y2 = c(117,109,114,104)
y3 = c(101,93,92,86)
y4 = c(105,119,116,102)
y5 = c(79,97,103,79)
y6 = c(57,92,104,77)
nSamples = 100000
paramInit = function() {
  theta_1 = sample(y1, 1)
  theta_2 = sample(y2, 1)
  theta_3 = sample(y3, 1)
  theta_4 = sample(y4, 1)
  theta_5 = sample(y5, 1)
  theta_6 = sample(y6, 1)
  mu = mean(c(y1, y2, y3, y4, y5, y6))
  sigma = runif(1, 0.1, 5)
  tau = runif(1, 0.1, 5)
  return(c(theta_1, theta_2, theta_3, theta_4, theta_5, theta_6, mu, sigma, tau))
}
postSamples = GibbsSampling(paramInit(), y1, y2, y3, y4, y5, y6, nSamples)

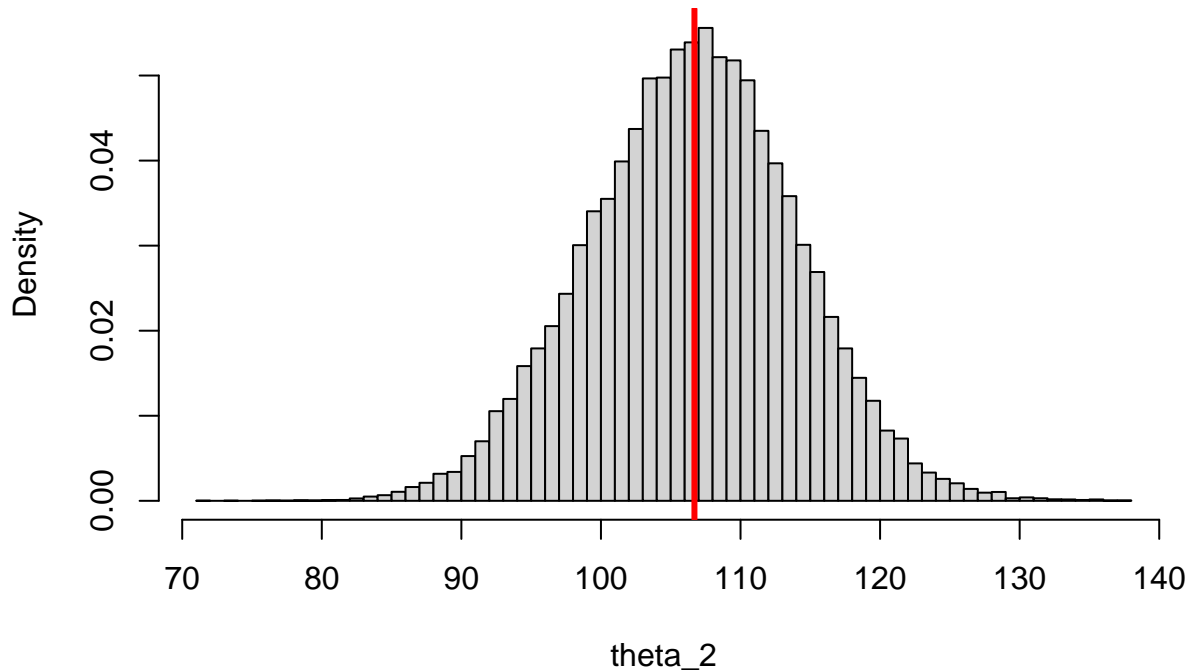
```

```
postSamples = postSamples[(nSamples/2 + 1):nSamples,]
```

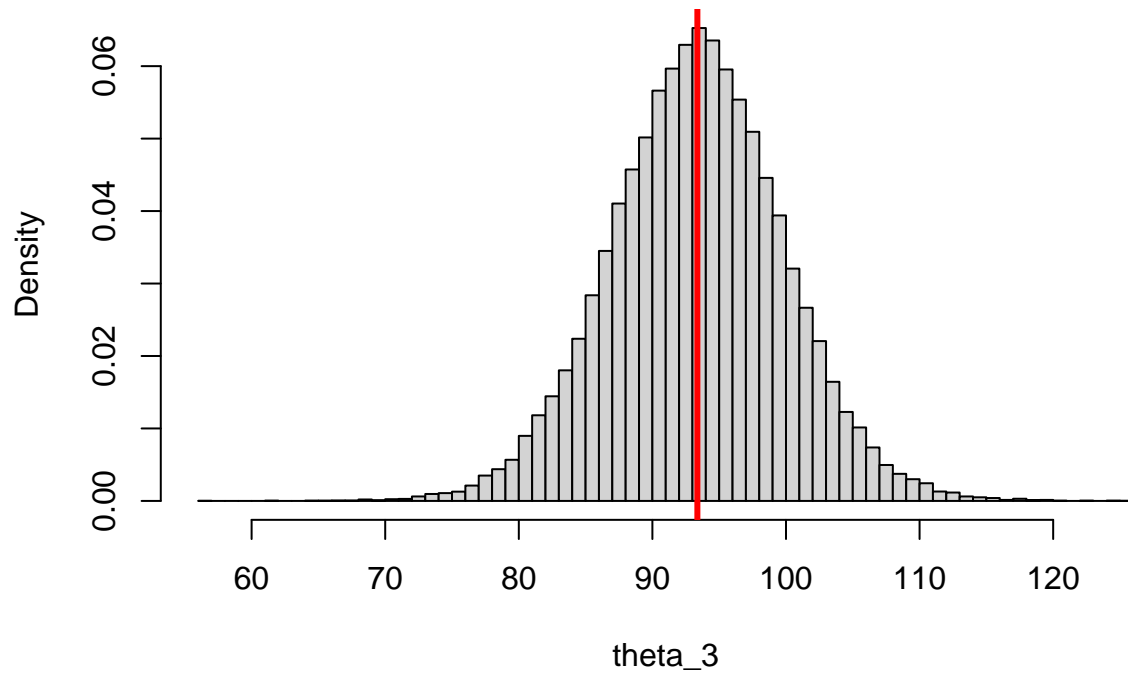
```
hist(postSamples[,1], xlab="theta_1", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,1]), col="red", lwd=3)
```



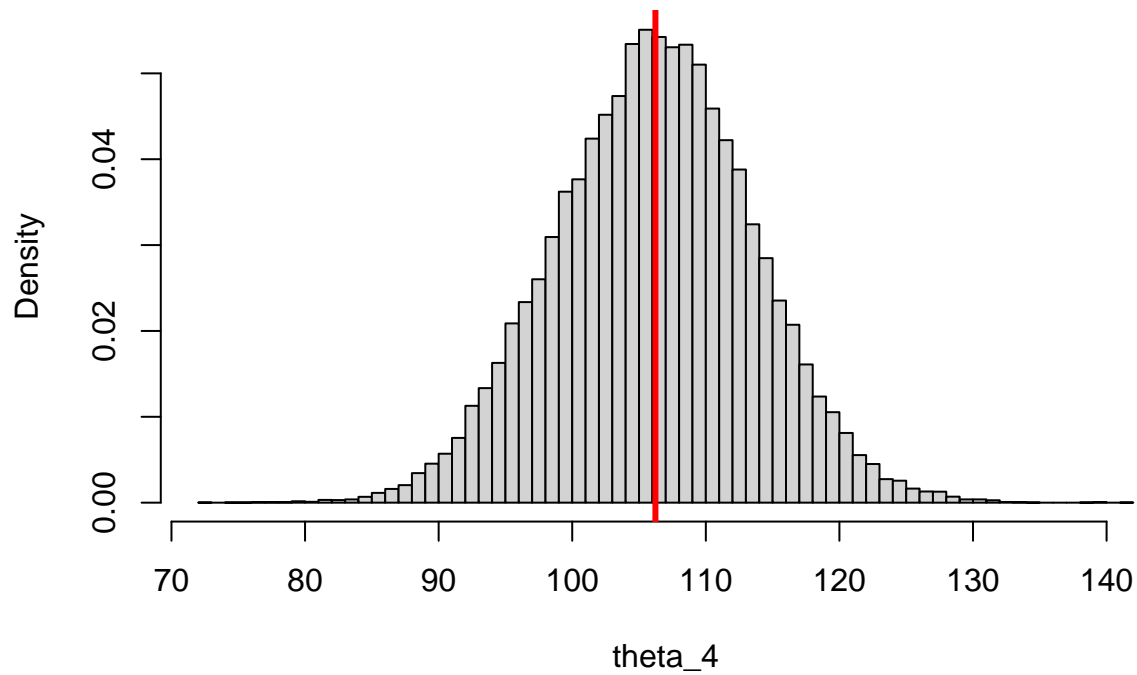
```
hist(postSamples[,2], xlab="theta_2", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,2]), col="red", lwd=3)
```



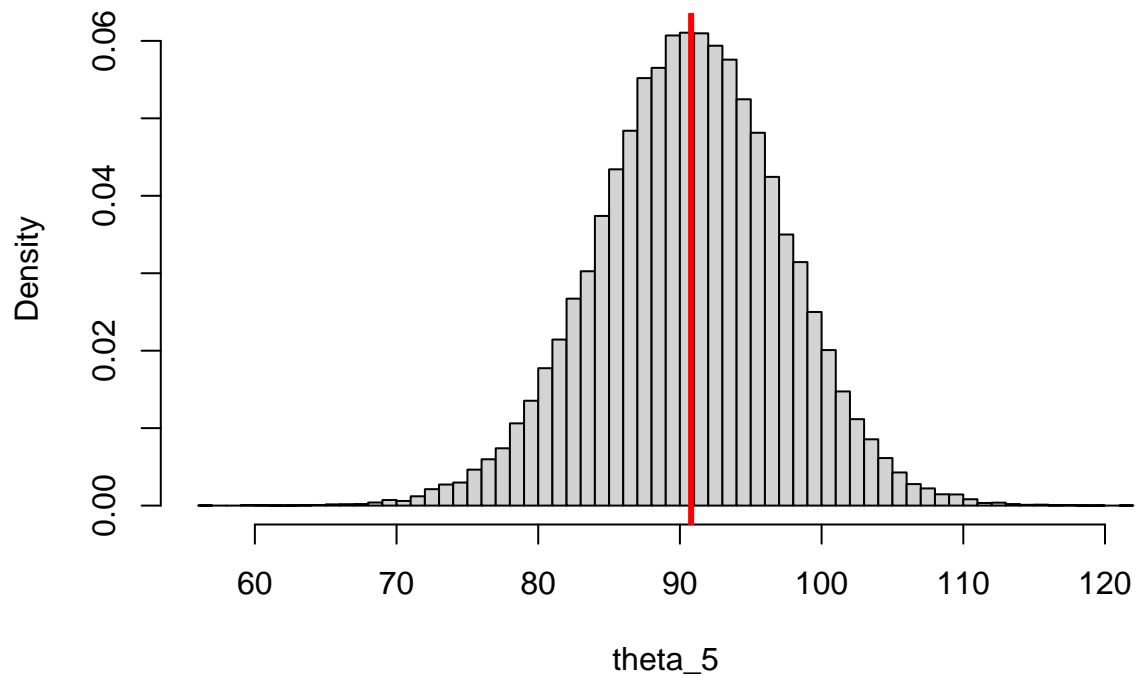
```
hist(postSamples[,3], xlab="theta_3", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,3]), col="red", lwd=3)
```



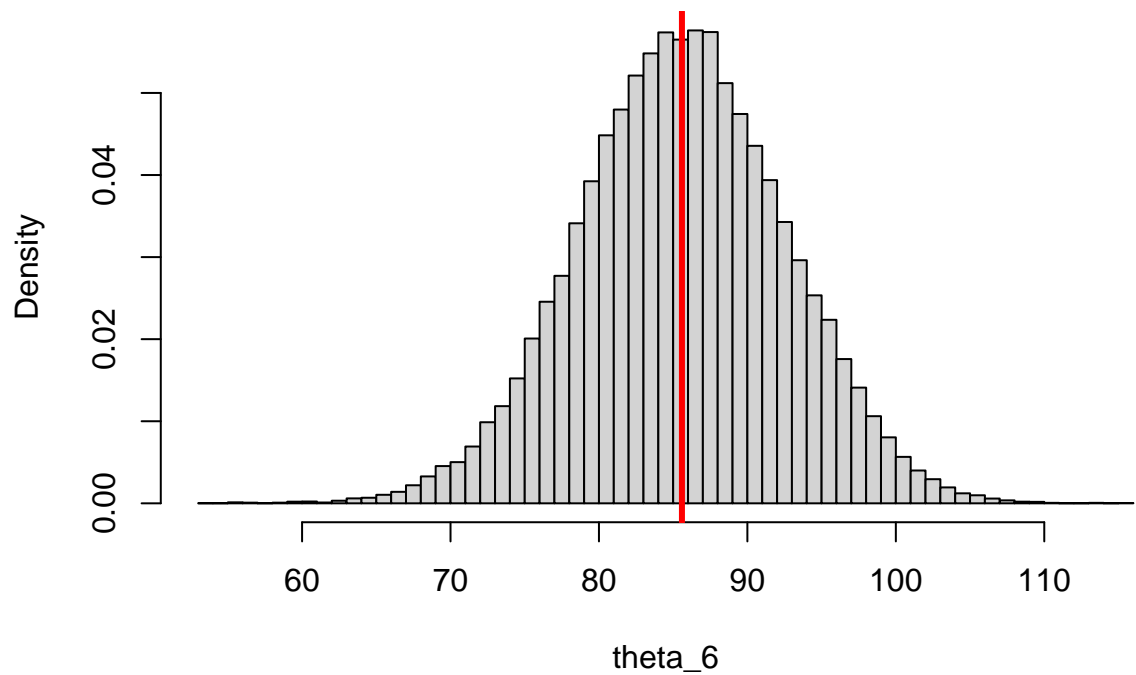
```
hist(postSamples[,4], xlab="theta_4", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,4]), col="red", lwd=3)
```



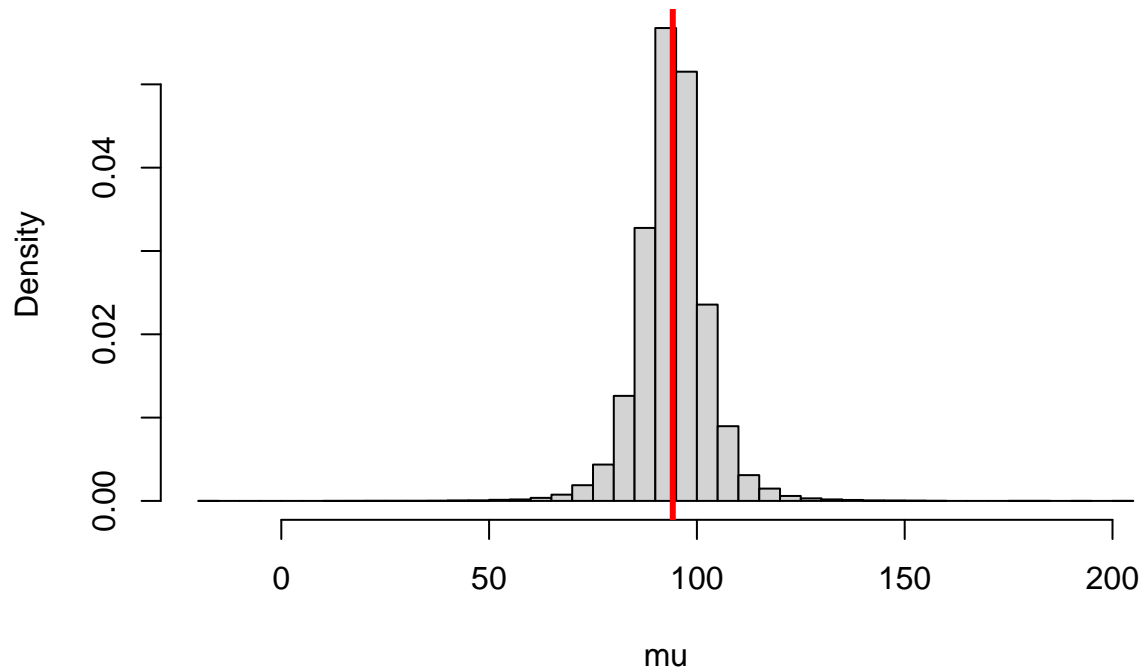
```
hist(postSamples[,5], xlab="theta_5", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,5]), col="red", lwd=3)
```



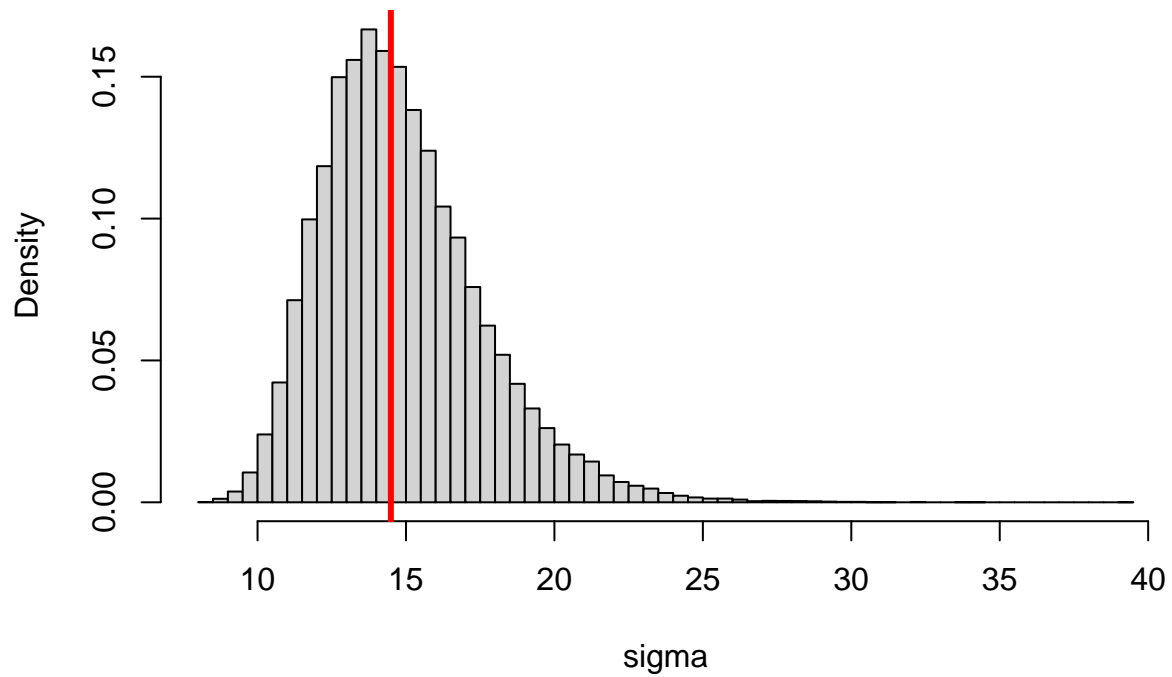
```
hist(postSamples[,6], xlab="theta_6", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,6]), col="red", lwd=3)
```



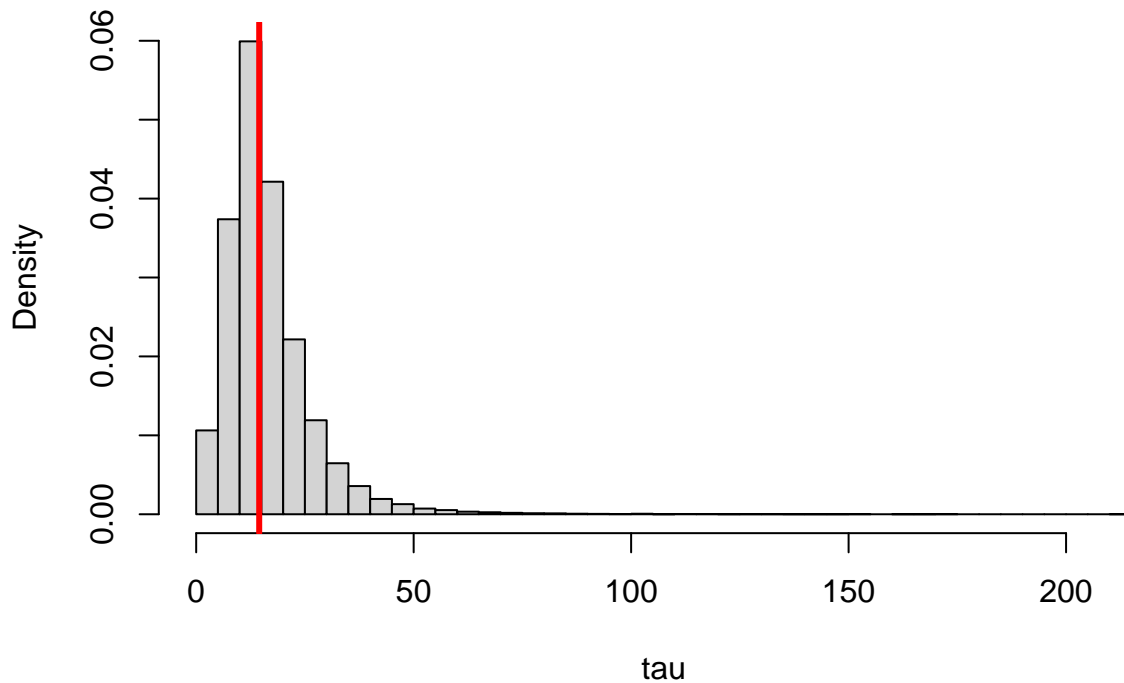
```
hist(postSamples[,7], xlab="mu", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,7]), col="red", lwd=3)
```



```
hist(postSamples[,8], xlab="sigma", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,8]), col="red", lwd=3)
```



```
hist(postSamples[,9], xlab="tau", breaks=50, prob=TRUE, main="")
abline(v = median(postSamples[,8]), col="red", lwd=3)
```



Looks like Markov chain is sampling properly. Theta parameters have converged quite well to normal distributions.

#### step 4: marginal posterior summary statistics

```
library(knitr)
data = matrix(c(quantile(postSamples[,1], 0.01),
                  quantile(postSamples[,1], 0.05),
                  mean(postSamples[,1]), median(postSamples[,1]),
                  quantile(postSamples[,1], 0.95),
                  quantile(postSamples[,1], 0.99),
                  quantile(postSamples[,2], 0.01),
                  quantile(postSamples[,2], 0.05),
                  mean(postSamples[,2]), median(postSamples[,2]),
                  quantile(postSamples[,2], 0.95),
                  quantile(postSamples[,2], 0.99),
                  quantile(postSamples[,3], 0.01),
                  quantile(postSamples[,3], 0.05),
                  mean(postSamples[,3]), median(postSamples[,3]),
                  quantile(postSamples[,3], 0.95),
                  quantile(postSamples[,3], 0.99),
                  quantile(postSamples[,4], 0.01),
                  quantile(postSamples[,4], 0.05),
                  mean(postSamples[,4]), median(postSamples[,4]),
                  quantile(postSamples[,4], 0.95),
                  quantile(postSamples[,4], 0.99),
                  quantile(postSamples[,5], 0.01),
                  quantile(postSamples[,5], 0.05),
                  mean(postSamples[,5]), median(postSamples[,5]),
                  quantile(postSamples[,5], 0.95),
                  quantile(postSamples[,5], 0.99),
                  quantile(postSamples[,6], 0.01),
```

```

quantile(postSamples[,6], 0.05),
mean(postSamples[,6]),median(postSamples[,6]),
quantile(postSamples[,6], 0.95),
quantile(postSamples[,6], 0.99),
quantile(postSamples[,7], 0.01),
quantile(postSamples[,7], 0.05),
mean(postSamples[,7]),median(postSamples[,7]),
quantile(postSamples[,7], 0.95),
quantile(postSamples[,7], 0.99),
quantile(postSamples[,8], 0.01),
quantile(postSamples[,8], 0.05),
mean(postSamples[,8]),median(postSamples[,8]),
quantile(postSamples[,8], 0.95),
quantile(postSamples[,8], 0.99),
quantile(postSamples[,9], 0.01),
quantile(postSamples[,9], 0.05),
mean(postSamples[,9]),median(postSamples[,9]),
quantile(postSamples[,9], 0.95),
quantile(postSamples[,9], 0.99)), ncol = 6, byrow = TRUE)
colnames(data) <- c('1%-quantile', '5%-quantile', 'mean', 'median', '95%-quantile',
'99%-quantile')
rownames(data) <- c('theta_1', 'theta_2', 'theta_3', 'theta_4', 'theta_5', 'theta_6',
'mu', 'sigma', 'tau')
t = as.table(data)
kable(t)

```

	1%-quantile	5%-quantile	mean	median	95%-quantile	99%-quantile
theta_1	65.287430	70.613362	82.52412	82.35844	95.03721	99.83041
theta_2	89.087264	94.134249	106.56796	106.69873	118.62369	123.86688
theta_3	77.835659	82.617471	93.34225	93.36741	103.91076	109.09174
theta_4	88.866137	93.790834	106.12940	106.22806	118.18398	123.51455
theta_5	74.428433	79.737991	90.71289	90.78935	101.28891	106.00111
theta_6	68.920311	74.137801	85.58360	85.59126	97.06357	101.56061
mu	71.443483	81.359559	94.17198	94.19404	107.17510	117.44062
sigma	10.118133	11.151167	14.86056	14.49052	19.85487	22.89024
tau	0.886097	4.816062	16.41065	14.30932	34.42551	53.97245

step 5: compute  $P(\text{theta}_5 < \text{theta}_1)$

```

print(paste("P(theta_5 < theta_1) =",
sum(postSamples[,5] < postSamples[,1])/
length(postSamples[,5] < postSamples[,1])))

```

```
## [1] "P(theta_5 < theta_1) = 0.18354"
```

## Question 2

prelimns

```

set.seed(123)
library(mvtnorm)

```



```
##
## Attaching package: 'mvtnorm'

## The following objects are masked from 'package:LaplacesDemon':
##
##      dmvtn, rmvtn

library(LaplacesDemon)
library(coda)
```

## step 1: function to simulate data

```
simData = function(numBooks, numSubjects, mu, tau_sq, sigma_sq) {
  meanEffects = rnorm(n = numBooks, mean = mu, sd = tau_sq)
  scores = rmvnorm(n = numSubjects, mean = meanEffects, sigma = diag(numBooks)*sigma_sq)
  return(list(scores, meanEffects))
}
```

## step 2: function to estimate posteriors using Gibbs sampling

```
GibbsSampling2 = function(paramInit, y, nSamples) {
  n = c()
  l = ncol(y)
  for (i in 1:l) {
    n[i] = length(y[,i])
  }
  nTot = sum(n)
  nDim = length(paramInit)
  paramSamples = array(rep(NA, nSamples * nDim), dim = c(nSamples, nDim))

  theta_hat = function(mu, sigma, tau, n, y) {
    numerator = mu/tau^2 + (n*mean(y))/sigma^2
    denom = (1/tau^2) + (n/sigma^2)
    return(numerator/denom)
  }
  sd_theta = function(sigma, tau, n) {
    denom = (1/tau^2) + (n/sigma^2)
    return(sqrt(1/denom))
  }
  paramLast = paramInit
  for (t in 1:nSamples) {
    theta_Inner = c()
    for (m in 1:l) {
      theta_Inner[m] = rnorm(1, theta_hat(mu = paramLast[l+1],
                                          sigma = paramLast[l+2],
                                          tau = paramLast[l+3],
                                          n = n[m],
                                          y = y[,m]),
                           sd_theta(sigma = paramLast[l+2],
                                     tau = paramLast[l+3],
                                     n = n[m]))
    }
    mu_Inner = rnorm(1, mean(theta_Inner), paramLast[l+3]/sqrt(nGroups))
    sampVar = c()
```

```

    for (m in 1:l) {
      sampVar[m] = sum((y[,m]-theta_Inner[m])^2)
    }
    sigma_Hat_2 = sum(sampVar)/nTot
    sigma_2_Inner = rinvchisq(1, nTot, sigma_Hat_2)
    tau_Hat_2 = sum((theta_Inner - mu_Inner)^2) / (nGroups - 1)
    tau_2_Inner = rinvchisq(1, nGroups - 1, tau_Hat_2)
    paramSamples[t,] = c(theta_Inner, mu_Inner, sqrt(sigma_2_Inner), sqrt(tau_2_Inner))
    paramLast = paramSamples[t,]
  }
  return(paramSamples)
}

pSamples = function(data) {
  nGroups = ncol(data)
  nSamples = 10000
  paramInit = function() {
    theta = c()
    for (i in 1:nGroups) {
      theta[i] = sample(data[,i],1)
    }
    mu = mean(data)
    sigma = runif(1,0.1,5)
    tau = runif(1,0.1,5)
    return(c(theta, mu, sigma, tau))
  }
  postSamples = GibbsSampling2(paramInit(), data, nSamples)
  postSamples = postSamples[(nSamples/2 + 1):nSamples,]
  return(postSamples)
}

```

### step 3: function to compute expected square errors

```

ESE = function(trueParams, postSamples) {
  true_means = trueParams[1:(length(trueParams)-3)]
  mu = trueParams[length(trueParams)-2]
  sigma_sq = trueParams[length(trueParams)-1]
  tau_sq = trueParams[length(trueParams)]

  ncol = ncol(postSamples)

  sample_mu = sample(postSamples[, ncol-2], 10000, replace=TRUE)
  sample_sigma_sq = sample(postSamples[, ncol-1], 10000, replace=TRUE)
  sample_tau_sq = sample(postSamples[, ncol], 10000, replace=TRUE)
  ese_mu = sum((sample_mu - mu)^2)/10000
  ese_tau_sq = sum((sample_tau_sq - tau_sq)^2)/10000
  ese_sigma_sq = sum((sample_sigma_sq - sigma_sq)^2)/10000

  avg_ese_theta = c()
  for (i in 1:length(true_means)) {
    theta = sample(postSamples[,i], 10000, replace=TRUE)
    avg_ese_theta[i] = sum((theta - true_means[i])^2)/10000
  }
}

```

```

    return(list(ese_mu, ese_tau_sq, ese_sigma_sq, avg_ese_theta))
}

```

#### step 4: run simulations and plot results

```

J = c(10, 20, 40, 80)
numDatasets = 10
mu = -10
tau_sq = 100
sigma_sq = 25

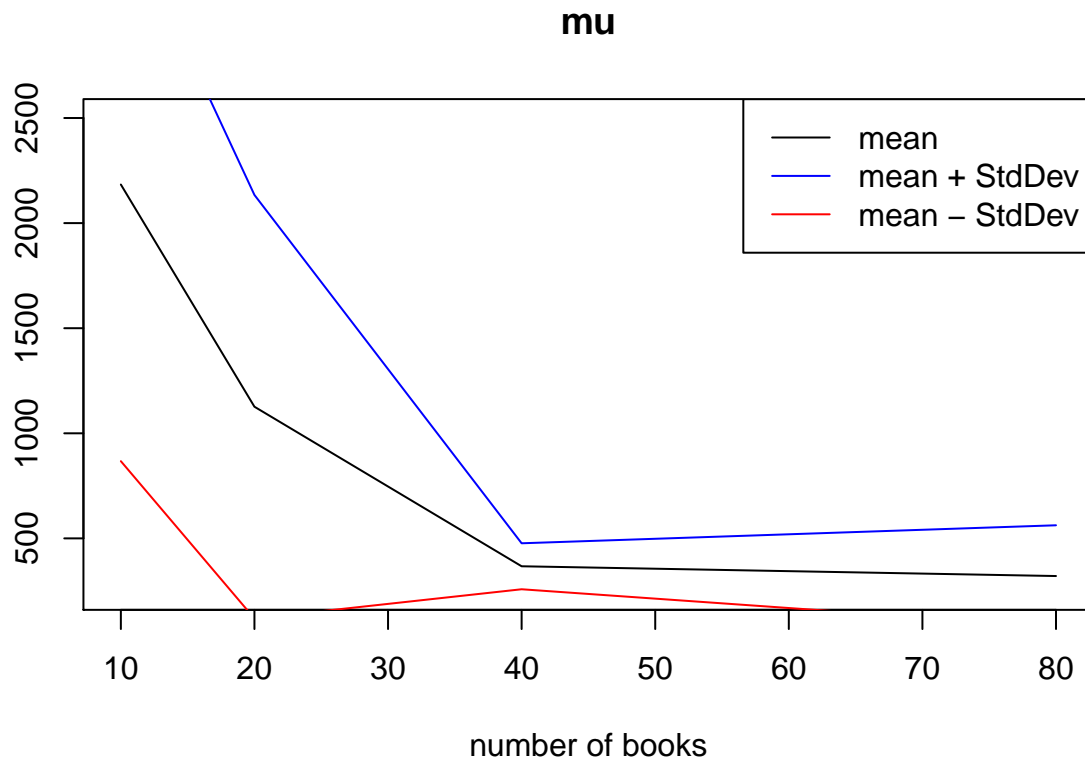
e1 = array(rep(-1, 40), dim = c(10,4))
e2 = array(rep(-1, 40), dim = c(10,4))
e3 = array(rep(-1, 40), dim = c(10,4))
e4 = array(rep(-1, 40), dim = c(10,4))

for (j in 1:length(J)) {
  for (n in 1:numDatasets) {
    sim_data = simData(numBooks = J[j], numSubjects = 960/J[j], mu = mu, tau_sq = tau_sq,
                      sigma_sq = sigma_sq)
    data = sim_data[[1]]
    nGroups = ncol(data)
    postSamples = pSamples(data = data)
    errors = ESE(trueParams = c(sim_data[[2]], mu, sigma_sq, tau_sq),
                postSamples = postSamples)
    ese_mu = errors[[1]]
    ese_tau_sq = errors[[2]]
    ese_sigma_sq = errors[[3]]
    avg_ese_theta = errors[[4]]

    if (j == 1) {
      e1[n, ] = c(ese_mu, ese_tau_sq, ese_sigma_sq, mean(avg_ese_theta))
    } else if (j == 2) {
      e2[n, ] = c(ese_mu, ese_tau_sq, ese_sigma_sq, mean(avg_ese_theta))
    } else if (j == 3) {
      e3[n, ] = c(ese_mu, ese_tau_sq, ese_sigma_sq, mean(avg_ese_theta))
    } else if (j == 4) {
      e4[n, ] = c(ese_mu, ese_tau_sq, ese_sigma_sq, mean(avg_ese_theta))
    }
  }
}

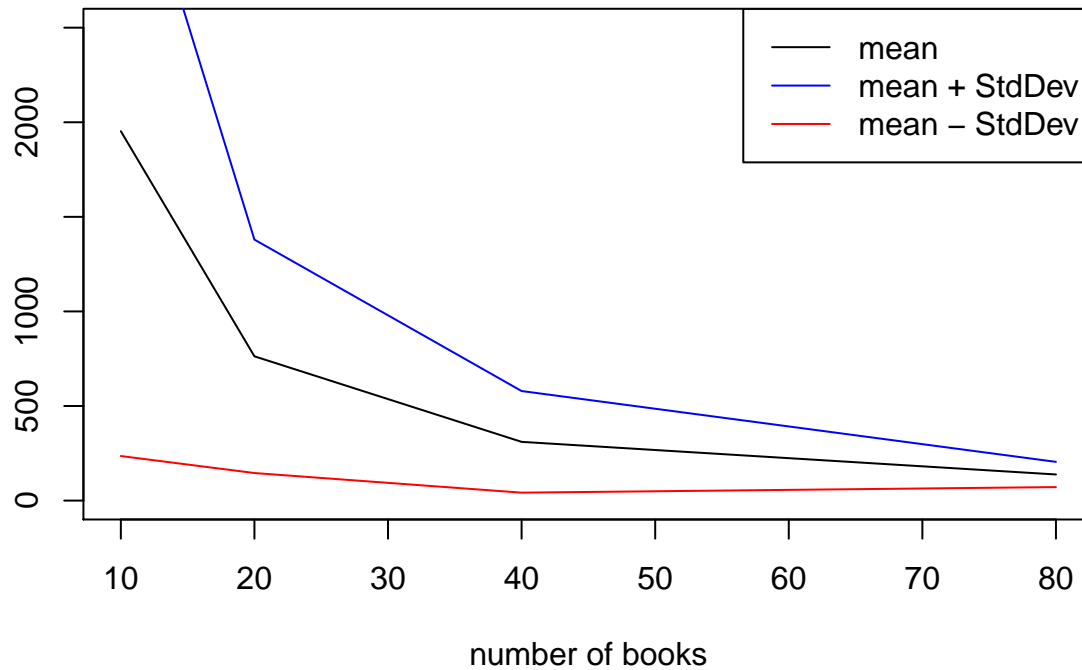
plot(J, c(mean(e1[,1]), mean(e2[,1]), mean(e3[,1]), mean(e4[,1])), type = "l",
      col = "black", ylab = "", xlab = "number of books", main = "mu", ylim = c(250, 2500))
lines(J, c(mean(e1[,1])-sd(e1[,1]),mean(e2[,1])-sd(e2[,1]),mean(e3[,1])-sd(e3[,1]),
          mean(e4[,1])-sd(e4[,1])), col = "red")
lines(J, c(mean(e1[,1])+sd(e1[,1]),mean(e2[,1])+sd(e2[,1]),mean(e3[,1])+sd(e3[,1]),
          mean(e4[,1])+sd(e4[,1])), col = "blue")
legend("topright", c("mean", "mean + StdDev", "mean - StdDev"), lty = c(1,1,1),
      col = c("black", "blue", "red"))

```



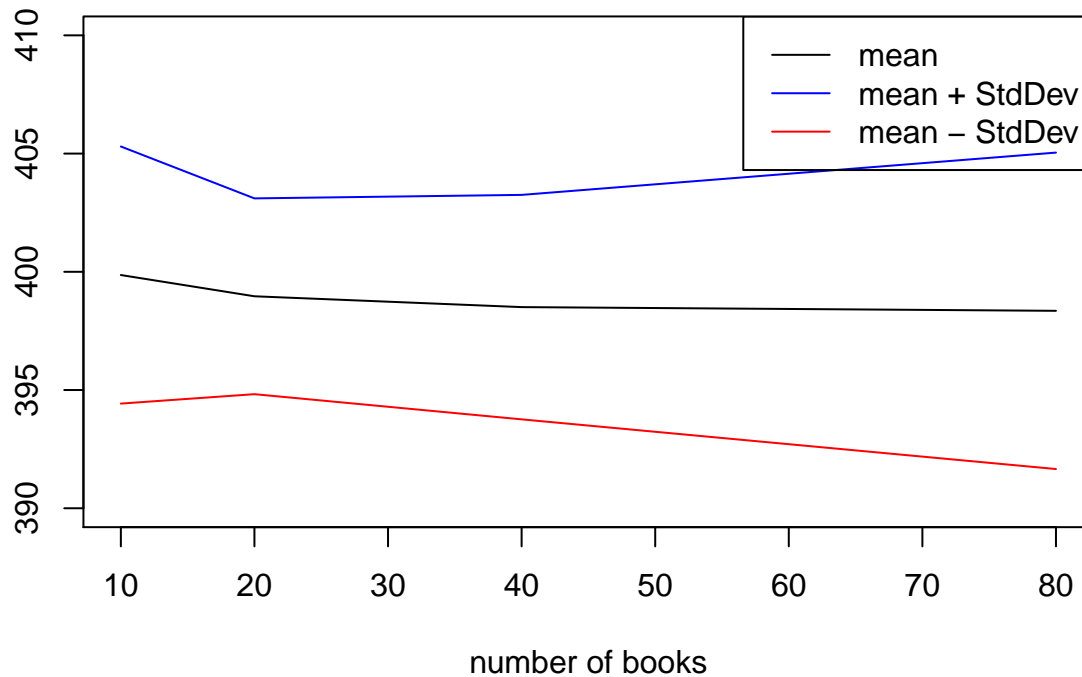
```
plot(J, c(mean(e1[,2]), mean(e2[,2]), mean(e3[,2]), mean(e4[,2])), type = "l",
      col = "black", ylab = "", xlab = "number of books", main = "tau_sq",
      ylim = c(0, 2500))
lines(J, c(mean(e1[,2])-sd(e1[,2]), mean(e2[,2])-sd(e2[,2]), mean(e3[,2])-sd(e3[,2]),
           mean(e4[,2])-sd(e4[,2])), col = "red")
lines(J, c(mean(e1[,2])+sd(e1[,2]), mean(e2[,2])+sd(e2[,2]), mean(e3[,2])+sd(e3[,2]),
           mean(e4[,2])+sd(e4[,2])), col = "blue")
legend("topright", c("mean", "mean + StdDev", "mean - StdDev"), lty = c(1,1,1),
      col = c("black", "blue", "red"))
```

## tau\_sq

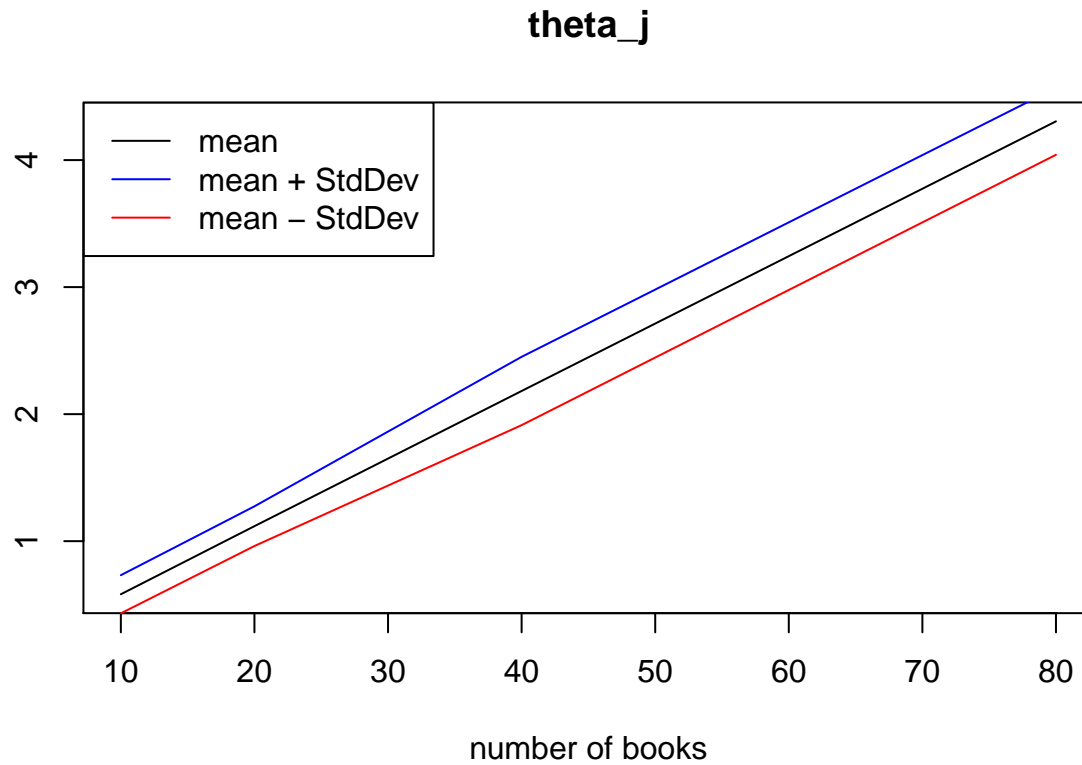


```
plot(J, c(mean(e1[,3]), mean(e2[,3]), mean(e3[,3]), mean(e4[,3])), type = "l",
      col = "black", ylab = "", xlab = "number of books", main = "sigma_sq",
      ylim = c(390, 410))
lines(J, c(mean(e1[,3])-sd(e1[,3]),mean(e2[,3])-sd(e2[,3]),mean(e3[,3])-sd(e3[,3]),
           mean(e4[,3])-sd(e4[,3])), col = "red")
lines(J, c(mean(e1[,3])+sd(e1[,3]),mean(e2[,3])+sd(e2[,3]),mean(e3[,3])+sd(e3[,3]),
           mean(e4[,3])+sd(e4[,3])), col = "blue")
legend("topright", c("mean", "mean + StdDev", "mean - StdDev"), lty = c(1,1,1),
      col = c("black", "blue", "red"))
```

## sigma\_sq



```
plot(J, c(mean(e1[,4]), mean(e2[,4]), mean(e3[,4]), mean(e4[,4])), type = "l",
      col = "black", ylab = "", xlab = "number of books", main = "theta_j")
lines(J, c(mean(e1[,4])-sd(e1[,4]), mean(e2[,4])-sd(e2[,4]), mean(e3[,4])-sd(e3[,4]),
           mean(e4[,4])-sd(e4[,4])), col = "red")
lines(J, c(mean(e1[,4])+sd(e1[,4]), mean(e2[,4])+sd(e2[,4]), mean(e3[,4])+sd(e3[,4]),
           mean(e4[,4])+sd(e4[,4])), col = "blue")
legend("topleft", c("mean", "mean + StdDev", "mean - StdDev"), lty = c(1,1,1),
      col = c("black", "blue", "red"))
```



**step 5.**

Looking at the previous step, we can see that as the number of books increases (and thus fewer subjects per book), the mean expected mean square errors of  $\mu$  and  $\tau$  squared squared decreases while it increases for the  $\theta_j$ 's. However, we see that these values remain somewhat constant for  $\sigma^2$ . Looks like having more books is better since these mean values  $\mu$  and  $\tau$  are decreasing. Would choose 80 books.