

ACM 118: Stochastic processes and regression

Instructor: Houman Owhadi

TA: Pau Battle Franck

Use Piazza for all Q&A and discussion

ACM 118 should not be your 1st course
on prob. theory

Focus: Gaussian Processes & Kernel Methods

Objective: Learn to reason, model and
organize computational knowledge
with GPs & kernel methods

Grading: 4 pb sets, 25% each

Interactive course

- Questions are welcome
- Pace can be adjusted
- Digressions are okay

Characteristic functions

X : rand. vect. on \mathbb{R}^d

ϕ_X : characteristic function of X

$$\begin{array}{ccc} \mathbb{R}^d & \xrightarrow{\quad} & \mathbb{C} \\ \lambda & \mapsto & \mathbb{E}[e^{i\lambda^T X}] \end{array}$$

Thm $\phi_X(\lambda) = \phi_Y(\lambda) \quad \forall \lambda \in \mathbb{R}^d$

$$\iff X \stackrel{\text{distribution}}{=} Y$$

$$\iff \mathbb{E}[f(X)] = \mathbb{E}[f(Y)]$$

$$\forall f: \mathbb{R}^d \rightarrow \mathbb{R}$$

bounded & measurable
continuous

. X_1, \dots, X_m are indep. iff $\forall \lambda_1, \dots, \lambda_m$
 $\mathbb{E}[e^{i \sum_{j=1}^m \lambda_j^T X_j}] = \prod_{j=1}^m \mathbb{E}[e^{i \lambda_j^T X_j}]$

. ($d=1$) If $\mathbb{E}[|X|^k] < \infty$ then

$$\phi_X^{(k)}(0) = i^k \mathbb{E}[X^k]$$

If $\phi_x^{(k)}(0)$ well defined then

$$\mathbb{E}[|X|^k] < \infty$$

$$\begin{aligned} X_n &\xrightarrow{\text{clst}} X \Leftrightarrow \mathbb{E}[f(X_n)] \xrightarrow[n \rightarrow \infty]{\longrightarrow} \mathbb{E}[f(X)] \\ &\forall f \in C_b(\mathbb{R}^d) \\ \Leftrightarrow \phi_{X_n}(\lambda) &\xrightarrow[n \rightarrow \infty]{\longrightarrow} \phi_X(\lambda) \\ \forall \lambda \in \mathbb{R}^d \end{aligned}$$

$$\phi_X(\lambda) := \mathbb{E}[e^{i\lambda^T X}]$$

Thm 1. $\phi_X(0) = 1$ & $|\phi_X(\lambda)| \leq 1$

2. $\phi_X(\lambda)$ is uniformly continuous $|\phi_X(\lambda)| \leq \mathbb{E}[|e^{i\lambda^T X}|] = 1$

3. $\forall \lambda_1, \dots, \lambda_m \in \mathbb{R}^d, \forall z_1, \dots, z_m \in \mathbb{C}$

$$\sum_{j,k=1}^m \phi_X(\lambda_j - \lambda_k) z_j \bar{z}_k \geq 0$$

$$\mathbb{E}\left[\left(\sum_j z_j e^{i\lambda_j^T X}\right)^2\right] \geq 0$$

4. Bochner's theorem

$\phi: \mathbb{R}^d \rightarrow \mathbb{C}$ is the characteristic function of some rand. vect. X iff it satisfies 1, 2 & 3

Thm X_n : sequ. of \mathbb{R}^d valued rand. var.

If $\lim_{n \rightarrow \infty} \phi_{X_n}(\lambda)$ exists $\forall \lambda \in \mathbb{R}^d$

and $\phi(\lambda) := \lim_{n \rightarrow \infty} \phi_{X_n}(\lambda)$ is continuous at $\lambda = 0$, then ϕ is the characteristic function of some r.v. X & $X_n \xrightarrow[n \rightarrow \infty]{\text{distrib}} X$

$$\underline{\text{Ex}} \quad \mathbb{E}[e^{i\lambda X}]$$

$$\phi_x(\lambda)$$

$$\phi_x(0) = 1$$

$$\lim_{\lambda \rightarrow 0} \mathbb{E}[e^{i\lambda X}] = \mathbb{E}\left[\lim_{\lambda \rightarrow 0} e^{i\lambda X}\right] = 1$$

$$\nabla_{\lambda} \phi_X(\lambda) = E[e^{i\lambda X} iX]$$

$$\nabla_{\lambda} \phi_X(0) = i E[X]$$

$$\begin{aligned}\phi'_X(\lambda) &= \lim_{\varepsilon \rightarrow 0} \frac{\phi_X(\lambda + \varepsilon) - \phi_X(\lambda)}{\varepsilon} \\ &= \lim_{\varepsilon \rightarrow 0} E \left[\frac{e^{i(\lambda+\varepsilon)X} - e^{i\lambda X}}{\varepsilon} \right] \\ &= \lim_{\varepsilon \rightarrow 0} E \left[e^{i\lambda X} \frac{e^{i\varepsilon X} - 1}{\varepsilon} \right]\end{aligned}$$

dominated convergence

Gaussian random variables

$$d=1$$

Def X is a Gaussian rand. var. with mean 0 and variance 1
 $X \sim \mathcal{N}(0, 1)$

if $f_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$
 ↴ pdf of X

$$\text{Thm } E[e^{i\lambda X}] = e^{-\lambda^2/2} \quad \forall \lambda \in \mathbb{R}$$

$$f(z) = E[e^{zX}] = e^{z^2/2} \quad \forall z \in \mathbb{R}$$

↓ is analytic on \mathbb{C}

$$\Rightarrow E[e^{zX}] = e^{z^2/2} \quad \forall z \in \mathbb{C}$$

f analytic $\Leftrightarrow f(z) = \sum_{n=0}^{\infty} a_n z^n \quad \forall z \in \mathbb{R}$

$$\sum_{n=0}^{\infty} |a_n| |z|^n < \infty \quad \forall z \in \mathbb{R}$$

$$\tilde{f}(z) = \sum_{n=0}^{\infty} a_n z^n \quad \forall z \in \mathbb{C}$$

$$\sum_{n=0}^{\infty} |a_n| |z|^n < \infty \quad \forall z \in \mathbb{C}$$

$$\tilde{f}(z) = f(z) \quad \forall z \in \mathbb{R}$$

$$e^z = \sum_{n=0}^{\infty} \frac{z^n}{n!}$$

$$f(z) = e^{z^2/2} = \sum_{n=0}^{\infty} \frac{z^{2n}}{2^n n!}$$

$$\begin{aligned} E[e^{i\lambda X}] &= 1 + i\lambda E[X] + \dots + \frac{(i\lambda)^n}{n!} E[X^n] \\ &\quad + O((\lambda)^{n+1}) \\ &= e^{-\lambda^2/2} \\ X &\sim \mathcal{N}(0, 1) \end{aligned}$$

$$E[X] = 0$$

$$E[X^2] = 1$$

$$E[X^{2n+1}] = 0$$

$$E[X^{2n}] = \frac{(2n)!}{2^n n!}$$

Def Y is Gaussian with mean m and variance σ^2 , $Y \sim \mathcal{N}(m, \sigma^2)$

if one of the equivalent properties is satisfied

$$(i) Y = \sigma X + m \quad \text{where } X \sim \mathcal{N}(0, 1)$$

$$(ii) f_Y(y) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{(y-m)^2}{2\sigma^2}\right)$$

$$(iii) \quad \mathbb{E}[e^{i\lambda Y}] = e^{im\lambda - \sigma^2\lambda^2/2}$$

$$\mathbb{E}[e^{i\lambda(\omega X_m)}] \stackrel{\text{def}}{=} e^{i\lambda m} e^{-\frac{(\lambda\sigma)^2}{2}}$$

Ex $X \sim \mathcal{N}(\mu_1, \sigma_1^2)$
 $Y \sim \mathcal{N}(\mu_2, \sigma_2^2)$

$X \& Y$ indep

$$X+Y \stackrel{?}{\sim} \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

$$\begin{aligned} \mathbb{E}[e^{i\lambda(X+Y)}] &= \mathbb{E}[e^{i\lambda X}] \mathbb{E}[e^{i\lambda Y}] \\ &\stackrel{X,Y \text{ indep}}{=} e^{i\lambda\mu_1 - \sigma_1^2 \frac{\lambda^2}{2}} e^{i\lambda\mu_2 - \sigma_2^2 \frac{\lambda^2}{2}} \\ &= e^{i\lambda(\mu_1 + \mu_2) - (\sigma_1^2 + \sigma_2^2) \frac{\lambda^2}{2}} \end{aligned}$$

The Gaussian distribution is a stable distribution

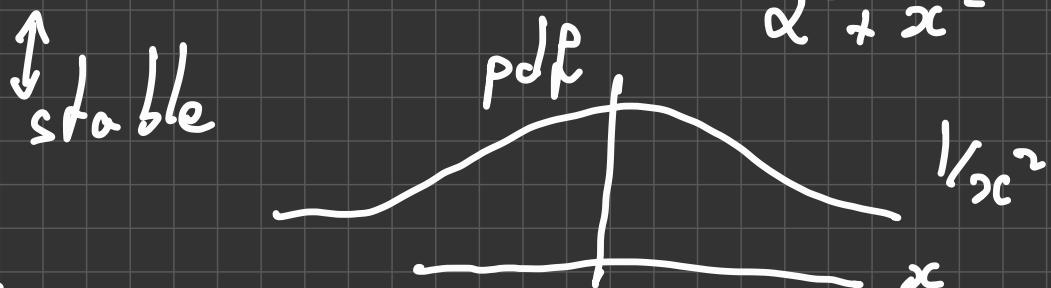
Def distrib. (X) is stable iff $\forall a, b > 0$

X_1, X_2 indep copies of X

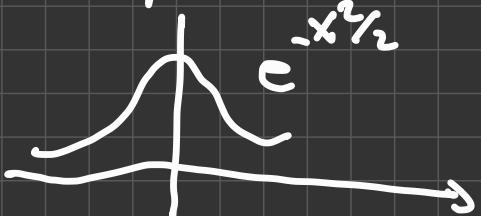
$$(*) \quad aX_1 + bX_2 \sim cX + d \quad \text{for } c > 0 \\ d$$

The distr. of X is said to be strictly stable if $(*)$ holds with $d = 0$

Rk X is Cauchy $\text{pdf}(X) \sim \frac{1}{\alpha^2 + x^2}$



pdf Gaussian



X heavy-tailed \Leftrightarrow

$$\lim_{x \rightarrow \infty} e^{tx} P[X > x] = \infty \quad t > 0$$

Gaussian r.v. are the only stable r.v. that does not have a heavy tail

$X \sim \text{Cauchy}(\alpha)$ $E[X]$ not defined

X_1, \dots, X_n iid Cauchy (α)

$\frac{X_1 + \dots + X_n}{n} \sim \text{Cauchy}(\alpha)$

E_x $X \sim \mathcal{N}(0, 1)$

$Y \sim \mathcal{N}(0, 1)$

Is $X+Y$ Gaussian ?

$X = -Y$

$X+Y=0 \sim \mathcal{N}(0, 0)$

$Y = \varepsilon X$ $\varepsilon = \begin{cases} +1 & \text{with prob } 1/2 \\ -1 & \end{cases}$
 ε indep. from X

$X+Y = (1+\varepsilon)X$

$\text{Var}(X+Y) > 0$

$P[X+Y=0] = 1/2$

$\Rightarrow X + Y$ is not a Gaussian r.v.

Rk $f_Y(y) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(y-\mu)^2}{2\sigma^2}\right)$

maximize the entropy $- \int f(x) \ln f(x) dx$

s.t. $\int_{\mathbb{R}} x f(x) dx = \mu$

$\int_{\mathbb{R}} (x - \mu)^2 f(x) dx = \sigma^2$

Gaussian vectors in \mathbb{R}^d

Def If X is an \mathbb{R}^d -valued rand. var.
st. $E[|X|^2] < \infty$ then

$\text{Cov}(X)$, the covariance matrix of X
is the $d \times d$ matrix defined by

$$\text{Cov}[X]_{i,j} = E[(x_i - E[x_i])(x_j - E[x_j])]$$

$$\text{Cov}(X) = E[(X - E[X])(X - E[X])^T]$$

$$\underset{\substack{v \in \mathbb{R}^d \\ \|}}{E_x} \text{Var}[v^T X] = v^T \text{Cov}(X) v$$

$$E[(v^T(X - E[X]))^2]$$

$$E[v^T(X - E[X])(X - E[X])^T v]$$

Gaussian rand. vectors

X : a \mathbb{R}^d valued rand. vect.

X is a Gaussian vector with mean μ and covariance matrix K

$$X \sim \mathcal{N}(\mu, K)$$

if

$$f_X(x) = \frac{1}{(2\pi)^{\frac{d}{2}} \sqrt{\det K}}$$

$$e^{-\frac{1}{2}(x-\mu)^T K^{-1} (x-\mu)}$$

$$\mu \in \mathbb{R}^d$$

K : symmetric positive definite $d \times d$ matrix

$$k_{i,j} = k_{j,i}$$

$$v^T K v \geq \alpha v^T v \text{ for some } \alpha > 0$$

$$v^T K v > 0 \quad \forall v \neq 0$$

Properties $E[X] = \mu$

$$\text{Cov}[X] = K$$

$\lambda \in \mathbb{R}^d$

$$\begin{aligned}\mathbb{E}[e^{i\lambda^T X}] &= e^{-\frac{1}{2}\lambda^T K \lambda + i\lambda^T \mu} \\ &= e^{-\frac{1}{2}\lambda^T \text{cov}(X)\lambda + i\lambda^T E[X]}\end{aligned}$$

Ex $X \sim \mathcal{N}(\mu, K)$

$v^T X \sim ?$

$v \in \mathbb{R}^d$

$\lambda = t v$

$$\begin{aligned}\phi(t) &= \mathbb{E}[e^{itv^T X}] = \mathbb{E}[e^{i\lambda^T X}] \\ &= \exp\left(-\frac{t^2}{2}v^T K v + itv^T \mu\right)\end{aligned}$$

characteristic function of

$v^T X \sim \mathcal{N}(v^T \mu, v^T K v)$

Ex $X = (X_1, X_2, X_3) \quad X \sim \mathcal{N}(0, K)$

$$K = \begin{pmatrix} 30 & 0 & 2 \\ 0 & 20 & 1 \\ 2 & 1 & 20 \end{pmatrix}$$

$$\underbrace{X_2 + X_3}_{v^T X} \sim \mathcal{N}(0, \underbrace{V^T K V}_{42})$$

$$v = \begin{pmatrix} 0 \\ 1 \\ 1 \end{pmatrix}$$

Ex $X: \mathbb{R}^d$ valued rand. vect.

s.t. $v^T X$ is Gaussian $\forall v \in \mathbb{R}^d$

Is X a Gaussian vector?

$$\text{Sol } \mathbb{E}[e^{i v^T X}] = e^{-\frac{1}{2} \text{Var}[v^T X] + i \mathbb{E}[v^T X]}$$

$v^T X$ Gaussian

Y is Gaussian \Rightarrow

$$\mathbb{E}[e^{i Y}] = e^{-\frac{1}{2} \text{Var}[Y] + i \mathbb{E}[Y]}$$

$$= \exp\left(-\frac{1}{2} v^T \text{Cov}(X) v + i v^T \mathbb{E}[X]\right)$$

↓
characteristic function of a Gaussian
vector

$\Rightarrow X$ is Gaussian

Def We say that X is a Gaussian vector in \mathbb{R}^d with mean μ and covariance matrix K (symm. positive) if

$$\forall v \in \mathbb{R}^d, \quad v^T X \sim \mathcal{N}(v^T \mu, v^T K v)$$

Ex $X \sim \mathcal{N}(\mu, K)$ X is \mathbb{R}^d valued

M : $d \times d$ matrix

$$MX \sim ? \quad \mathcal{N}(M\mu, MKM^T)$$

$$v^T (MX) = (M^T v)^T X \quad \uparrow$$

$$\sim \mathcal{N}(v^T M\mu, v^T M K M^T v)$$

Ex Z_1, \dots, Z_d iid $\mathcal{N}(0, 1)$

$$Z = (Z_1, \dots, Z_d)$$

$M: d \times d$ matrix

$$Mz \sim \mathcal{N}(0, MM^T)$$

Ex $X \sim \mathcal{N}(\mu, K)$

How to sample from X ?

Assume that you can only sample from

$$Z_c (Z_1, \dots, Z_d) \quad Z_i \text{ iid } \mathcal{N}(0, 1)$$

Sol find $L: d \times d$ | $K = LL^T$

$$X = X - \mu + \mu$$

$$X - \mu \sim L Z$$

$$X \sim \mu + LZ$$

$$LZ \sim \mathcal{N}(0, LL^T)$$

K symm.
positive \Rightarrow $\exists L$ lower triangular/
 $K = LL^T$

$K = LL^T$ Cholesky decomposition of K
factorization

Cost $\sim \mathcal{O}(d^3)$ \rightarrow naïve

$\sim \mathcal{O}(d \text{ polylog}(\frac{d}{\epsilon}))$ \rightarrow fast
Cholesky fact.

E_x $X \sim \mathcal{N}(0, K)$

λ_i, v_i : eigenpairs of K

$$K v_i = \lambda_i v_i$$

v_1, \dots, v_d forms an orthonormal basis of \mathbb{R}^d

$$X = \sum_{j=1}^d \langle v_j, X \rangle v_j$$

$$\langle v_j, X \rangle = v_j^T X$$

K invertible $\Rightarrow \lambda_j > 0$

$$X = \sum_{j=1}^d \sqrt{\lambda_j} Z_j v_j$$

$$Z_j = \frac{\langle v_j, X \rangle}{\sqrt{\lambda_j}}$$

$$Z = (Z_1, \dots, Z_d)$$

$$Z \sim ?$$

$$\mathbb{E}[e^{i \langle t, Z \rangle}] = \mathbb{E}[e^{i \langle v, X \rangle}] = e^{-\sqrt{t} K v}$$

$$t = (t_1, \dots, t_d) \quad v = \sum_{j=1}^d \frac{t_j e_j}{\sqrt{\lambda_j}} \quad e_1, \dots, e_d \text{ orthon. basis of } \mathbb{R}^d$$

$$= \exp\left(-\frac{1}{2} \sum_{j=1}^d t_j^2\right)$$

$$v^T K v = \sum_{j=1}^d t_j^2$$

$$\mathbb{E}[e^{i \langle t, Z \rangle}] = e^{-\frac{1}{2} \sum_{j=1}^d t_j^2}$$

$$Z \sim \mathcal{N}(0, I_d)$$

CLT in \mathbb{R}^d

Ihm X_n : sequ. of \mathbb{R}^d valued r.v.
i.i.d.

$$\mathbb{E}[X_i] = \mu \quad \text{Cor}(X_i) = K$$

$$S_n := X_1 + \dots + X_n$$

$$\frac{S_n - n\mu}{\sqrt{n}} \xrightarrow[n \rightarrow \infty]{\text{distrib.}} \mathcal{N}(0, K)$$

Rk $K = LL^T$ $X_i \sim \mathcal{N}(0, K)$ $\underbrace{\quad}_{(L^T)^{-1} X_i \sim \mathcal{N}(0, L^{-1} K (L^T)^{-1})}$ $\underbrace{\quad}_{\text{Id}}$

$$(L^T)^{-1} \frac{S_n - n\mu}{\sqrt{n}} \xrightarrow[n \rightarrow \infty]{\text{distr.}} \mathcal{N}(0, \text{Id})$$

Proof Assume $\mu = 0$ X_j indep.

$$\mathbb{E}\left[e^{i\lambda^T \frac{S_n}{\sqrt{n}}}\right] = \prod_{j=1}^n \mathbb{E}\left[e^{i\lambda^T X_j / \sqrt{n}}\right]$$

$$= \left(\mathbb{E}[e^{i\lambda^T X_i/\sqrt{n}}] \right)^n$$

↗
\$X_i\$ are i.i.d. dist.

Taylor expansion

$$\mathbb{E}[e^{iv^T X}] = 1 + i \mathbb{E}[v^T X] - \frac{1}{2} \mathbb{E}[|v^T X|^2] + o(|v|^2)$$

$$= 1 - \frac{1}{2} v^T \underbrace{\mathbb{E}[XX^T]v}_{\text{Cov}(X)} + o(|v|^2)$$

↗
 $\mathbb{E}[X] = 0$

$$= \left(1 - \frac{\lambda^T K \lambda}{2n} + o\left(\frac{|\lambda|^2}{n}\right) \right)^n$$

→
 $n \rightarrow \infty$

$$\exp\left(-\frac{\lambda^T K \lambda}{2}\right)$$

charact. funct. of $N(0, K)$

conv. of charact. funct. \Rightarrow conv. in distr.

Ex Proof of the Chi-Square test

An experiment with d possible outcomes

outcome j occurs with prob. p_j

n repetitions of experiment are done

$$X_{kj} = \begin{cases} 1 & \text{outcome of } k\text{th experiment is } j \\ 0 & \text{else} \end{cases}$$

$$\vec{X}_i = (X_{i1}, \dots, X_{id})$$

$$\vec{X}_n = (X_{n1}, \dots, X_{nd})$$

X_i are iid vectors of \mathbb{R}^d

N_j : # of times outcome j appears in
 n repetition of exper.

$$N_j = X_{1j} + X_{2j} + \dots + X_{nj}$$

$$\chi^2(n) = \sum_{j=1}^d \frac{(N_j - np_j)^2}{np_j}$$

Thm (Karl Pearson)

$$D^2_{(n)} \xrightarrow[n \rightarrow \infty]{\text{distr.}} \chi^2_{(d-1)}$$
$$\sim Z_1^2 + \dots + Z_{d-1}^2$$
$$Z_i \text{ iid } \mathcal{N}(0, 1)$$

Proof

$$Y_i = \left(\frac{X_{i1} - p_1}{\sqrt{p_1}}, \dots, \frac{X_{id} - p_d}{\sqrt{p_d}} \right)$$

$$Y_i \text{ iid } E[Y_i] = 0$$

$$\left(\frac{N_1 - np_1}{\sqrt{np_1}}, \dots, \frac{N_d - np_d}{\sqrt{np_d}} \right) = \frac{Y_1 + \dots + Y_n}{\sqrt{n}}$$

$$D^2_{(n)} = \|v\|^2$$

$$\frac{Y_1 + \dots + Y_n}{\sqrt{n}} \xrightarrow[n \rightarrow \infty]{\text{CLT}} \mathcal{N}(0, C)$$

$$C_{ij} = E[Y_{1,i} Y_{1,j}]$$

$$= E\left[\frac{X_{1,i} - p_i}{\sqrt{p_i}} \frac{X_{1,j} - p_j}{\sqrt{p_j}} \right]$$

$$= \sqrt{p_i p_j} + \frac{E[X_{1,i} X_{1,j}]}{\sqrt{p_i p_j}}$$

$$- \frac{E[X_{1,i}]}{\sqrt{p_i p_j}} p_j - \frac{E[X_{1,j}]}{\sqrt{p_i p_j}} p_i$$

$$= \sqrt{p_i p_j} + \delta_{ij} - \sqrt{p_i p_j} - \sqrt{p_i p_i}$$

$$= \delta_{ij} - \sqrt{p_i p_j}$$

$$\sqrt{P} = (\sqrt{p_1}, \dots, \sqrt{p_d})$$

$$C = I - \sqrt{P} \sqrt{P}^\top$$

$$C \sqrt{P} = 0 \quad \lambda_1, \dots, \lambda_d \text{ eig of } C$$

$$C v = v \quad \text{for } v \perp \sqrt{P} \Rightarrow \lambda_1 = \dots = \lambda_{d-1} = 1$$

$$\lambda_d = 0$$

$$D^2(n) \xrightarrow[n \rightarrow \infty]{\text{distr.}} \|Y\|^2$$

$$Y \sim \mathcal{N}(0, C)$$

previous ex

$$Y = \sum_{j=1}^{d-1} z_j v_j$$

\downarrow

eigenvectors of C

iid $\mathcal{N}(0, 1)$

$$\Rightarrow \|Y\|^2 \sim \sum_{j=1}^{d-1} z_j^2 \sim \chi^2(d-1)$$

Application

6 faced die

100 rolls : $30 \rightarrow 6$ $22 \rightarrow 2$

$12 \rightarrow 5$ $8 \rightarrow 1$

$10 \rightarrow 4$

$18 \rightarrow 3$

Is the die fair or not?

$$X_i = (X_{i1}, \dots, X_{i6})$$

$$X_{ij} = \begin{cases} 1 & \text{if outcome roll } i \text{ is } j \\ 0 & \text{otherwise} \end{cases}$$

$$N_j = X_{1j} + X_{2j} + \dots + X_{nj}$$

$$D^2(n) = \sum_{j=1}^d \frac{(N_j - n p_j)^2}{n p_j}$$

$$p_j = \frac{1}{6}$$

$$\left[P\left[\underbrace{Z_1^2 + \dots + Z_5^2}_{X^2(5)} \geq D^2(n) \right] \right]$$

$$= 0.000001 \\ \rightarrow \text{not fair}$$

$$= 0.7 \rightarrow \text{fair}$$

Gaussian spaces and Gaussian condition.

$(\Omega, \mathcal{U}, \text{IP})$: probability space

Ω : sample space

\mathcal{U} : σ -algebra (set of subsets of Ω stable under countable $\cap, \cup, {}^c$)

IP: $\mathcal{U} \rightarrow [0, 1]$

prob. measure

$L^2(\Omega, \mathcal{U}, \text{IP})$: set of r.v. $X: \Omega \rightarrow \mathbb{R}$

\mathcal{U} -measurable

$$\mathbb{E}[X^2] < \infty$$

Def A (centered) Gaussian space is
a closed subspace of $L^2(\Omega, \mathcal{U}, \text{IP})$
whose elements are centered Gaussian r.v.

Ex X_1, \dots, X_d iid $\mathcal{N}(0, 1)$

$$\begin{aligned} H &= \text{span} \{X_1, \dots, X_d\} \\ &= \left\{ \sum_{i=1}^d \lambda_i X_i \mid \lambda \in \mathbb{R}^d \right\} \end{aligned}$$

H : Gaussian space

Ex X_1, X_2 Gaussian r.v

↓?

Is $\text{Span} \{X_1, X_2\}$ a Gaussian space ?

No $X_1 + X_2$ may not be Gaussian

Ex (X_1, \dots, X_d) is Gaussian vector

↓?

$\text{Span} \{X_1, \dots, X_d\}$ is a Gaussian space

YES

Ex $(X_1, X_2, X_3) \sim \mathcal{N}(0, K)$

$$K = \begin{pmatrix} 3 & 0 & 1 \\ 0 & 2 & 1 \\ 1 & 1 & 2 \end{pmatrix}$$

Q1 $X_2 + X_3 \sim \mathcal{N}(0, \sqrt{K_{22}})$

$$v = \begin{pmatrix} 0 \\ 1 \\ 1 \end{pmatrix}$$

b1 Are $X_2 + X_3$ & X_1 indep?

$$\text{Cov}(X_1, X_2 + X_3) = w^T K v = 0 + 1 = 1 \neq 0$$
$$\begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix} \quad \begin{pmatrix} 0 \\ 1 \\ 1 \end{pmatrix}$$

NO

$\text{Cov}(Y, Z) \neq 0 \Rightarrow Y, Z$ not indep

Is it true that

$\text{Cov}(Y, Z) = 0$ $\xrightarrow[NO]{?} Y, Z \text{ indep.}$

Y, Z are both Gaussian

i. $Y \sim \mathcal{N}(0, 1)$

$$Z \sim \mathcal{E} Y \quad \mathcal{E} = \begin{bmatrix} 1 & \text{prob. } V_2 \\ -1 & V_2 \end{bmatrix}$$

$$\text{Cov}(Y, Z) = 0$$

but not indep. $|Z| = |Y|$

$Z, Y \text{ indep.} \xrightarrow[YES]{?} |Z| \text{ indep. } |Y|$

$Y, Z \text{ indep.} \quad YES$
 $Y, Z \text{ Gaussian} \Rightarrow Y, Z \text{ jointly Gaussian}$

$$(v, w) \quad \begin{pmatrix} Y \\ Z \end{pmatrix} = \underbrace{v^T Y}_{\text{Gaussian}} + \underbrace{w^T Z}_{\text{Gaussian}} \quad \text{indep.} \quad \xrightarrow{\quad} \text{Gaussian}$$

Is it true that

YES

$$\text{Cov}(Y, Z) = 0$$

(Y, Z) is Gaussian

? $\Rightarrow Y, Z$ indep

$$E[e^{i\lambda Y + i\nu Z}] \stackrel{?}{=} E[e^{i\lambda Y}] E[e^{i\nu Z}]$$

$$e^{i(\lambda, \nu) \left(E[Y] - \frac{1}{2} (\lambda, \nu) \text{Cov}(Y, Z) \right)}$$

$$e^{i\lambda E[Y] + i\nu E[Z] - \frac{1}{2} \lambda^2 \text{Var}[Y]}$$

$$- \frac{1}{2} \nu^2 \text{Var}(Z)$$

$$- \lambda \nu \text{Cov}(Y, Z)$$

"0"

E_x $(X_1, X_2, X_3) \sim \mathcal{N}(0, K)$

$$K = \begin{pmatrix} 1 & 0 & 1 \\ 0 & 2 & 1 \\ 1 & 1 & 2 \end{pmatrix}$$

cl Are $X_2 + X_3$ and $X_2 - X_3$ indep?

$(X_2 + X_3, X_2 - X_3)$ is Gaussian

$\text{Cov}(X_2 + X_3, X_2 - X_3)$

$$\begin{pmatrix} 0 & 1 & 1 \end{pmatrix} K \begin{pmatrix} 0 \\ 1 \\ -1 \end{pmatrix} = 2 - 2 = 0$$

YES

$X = (X_1, \dots, X_d)$ Gaussian vector

X_1, \dots, X_d indep $\Rightarrow \text{Cov}(X)$
diagonal

$$\leftarrow ? \cdot Y^{-1}$$

because characteristic function is a product

$$E[e^{i \sum_{j=1}^d \lambda_j X_j}] = \prod_{j=1}^d E[e^{i \lambda_j X_j}]$$

Thm Let H be a Gaussian space

and let $(H_i, i \in I)$ be a family of
closed sub-vector spaces of H

Then the subspaces H_i for $i \in I$ are

orthogonal in L^2 ($\Leftrightarrow \forall X_i \in H_i, X_j \in H_j$
 $E[X_i X_j] = 0$)

i if the \mathcal{G} -algebra $\mathcal{G}(H_i)$ for $i \in I$
 are indep. ($\mathcal{G}(H_i)$ is the \mathcal{G} -algebra
 generated by rand. var.
 elements of H_i)

$\mathcal{G}(H_i)$ indep $\mathcal{G}(H_j)$

$\forall X_{i,1}, \dots, X_{i,m} \in H_i$

$X_{j,1}, \dots, X_{j,n} \in H_j$

$\forall f, g \in C_b$

$f(X_{i,1}, \dots, X_{i,m})$ indep $g(X_{j,1}, \dots, X_{j,n})$

$H_i \perp\!\!\!\perp^2 H_j \Leftrightarrow$ All r.v. in H_i

$\forall i \neq j$ indep from all r.v. in H_j

$H_i \subset H$

H Gaussian

easy

$\forall i \neq j$



charact. funct.

$$\underline{E}_x \quad (X_1, X_2, X_3) \sim \mathcal{N}(0, K)$$

$$K = \begin{pmatrix} 3 & 0 & 1 \\ 0 & 2 & 1 \\ 1 & 1 & 2 \end{pmatrix}$$

d $E[X_1 | X_2, X_3]$?

Reminder $E[X_1 | X_2, X_3] = f(X_2, X_3)$

where f minimizes

$$E[(X_1 - f(X_2, X_3))^2]$$

$$X_1 = X_1 - aX_2 - bX_3 + aX_2 + bX_3$$

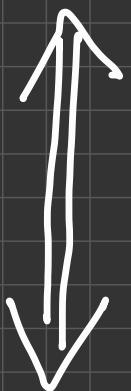
$$E[X_1 | X_2, X_3] = E[X_1 - aX_2 - bX_3 | X_2, X_3]$$

$$+ E[aX_2 + bX_3 | X_2, X_3]$$

$\underbrace{aX_2 + bX_3}_{\text{red}}$

If $X_1 - aX_2 - bX_3$ $\sim \mathbb{E}[X_1 - aX_2 - bX_3 | X_2, X_3]$
 indep from $X_2, X_3 \Rightarrow$

$$\mathbb{E}[X_1 - aX_2 - bX_3] \stackrel{\parallel}{=} 0$$



$$\text{Cov}(X_1 - aX_2 - bX_3, X_2) = 0$$

$$\text{Cor}(X_1 - aX_2 - bX_3, X_3) = 0$$

$$\begin{cases} -2a - b = 0 \\ 1 - a - 2b = 0 \end{cases} \quad \begin{array}{l} \xrightarrow{\text{Solve}} \\ \begin{aligned} a &= -1/3 \\ b &= 2/3 \end{aligned} \end{array}$$

$$\begin{aligned} \mathbb{E}[X_1 | X_2, X_3] &= -\frac{X_2}{3} + \frac{2}{3}X_3 \\ &= aX_2 + bX_3 \quad \text{where} \end{aligned}$$

a, b minimize

$$\mathbb{E}[|X_1 - aX_2 - bX_3|^2]$$

et $\mathcal{L}aw(X_1 | X_2, X_3)$?

$$X = (X_1, X_2, X_3) \sim \mathcal{N}(0, K)$$

$$K = \begin{pmatrix} 3 & 0 & 1 \\ 0 & 2 & 1 \\ 1 & 1 & 2 \end{pmatrix}$$

$$a = -\frac{1}{3}, \quad b = \frac{2}{3}$$

$$X_1 = \underbrace{E[X_1 | X_2, X_3]}_{aX_2 + bX_3} + \underbrace{X_1 - E[X_1 | X_2, X_3]}_{X_1 - aX_2 - bX_3}$$

behaves like a cst if $X_2 \neq X_3$ known

2 indep from X_2, X_3

$\mathcal{N}(0, \sigma^2)$

$$\sigma^2 = \text{Var}(X_1 - aX_2 - bX_3) = \frac{\sigma^2}{9}$$

because $X_1 - aX_2 - bX_3$

indep from X_2, X_3

$$\mathcal{L}aw(X_1 - aX_2 - bX_3 | X_2, X_3) = \mathcal{L}aw(X_1 - aX_2 - bX_3)$$

$$\text{Given } (X_1 | X_2, X_3) \sim \mathcal{N}\left(-\frac{X_2}{3} + \frac{2}{3}X_3, \frac{\sigma^2}{9}\right)$$

$$\underline{Ex} \quad \underline{X} = (X_1, X_2, X_3) \sim \mathcal{N}(m, K)$$

\downarrow
not centered

$$K = \begin{pmatrix} 3 & 0 & 1 \\ 0 & 2 & 1 \\ 1 & 1 & 2 \end{pmatrix}$$

$$E[X_1 | X_2, X_3] = ?$$

$$X_1 = X_1 - m_1 + m_1$$

$$E[X_1 | X_2, X_3] = m_1 + E[X_{1-m_1} | X_2, X_3]$$

$$= m_1 + E[X_{1-m_1} | \underbrace{X_2 - m_2}_{Y_1}, \underbrace{X_3 - m_3}_{Y_2}]$$

$$| Y \sim \mathcal{N}(0, K)$$

$$\downarrow \\ = E[X_1] - \frac{1}{3} (X_2 - E[X_2]) \\ + \frac{2}{3} (X_3 - E[X_3])$$

Thm H: Gaussian space

V : closed sub-vector space of H

Pr : orthogonal projection on V

$\text{Pr}: H \rightarrow V$

$X \rightarrow \text{Pr}(X) = \underset{z \in V}{\arg \min} [E[X-z]^2]$

Let $X \in H$, then

(i) $E[X|G(V)] = \text{Pr}(X)$

(ii) $\hat{v}^2 := E[(X - \hat{v})^2]$ then

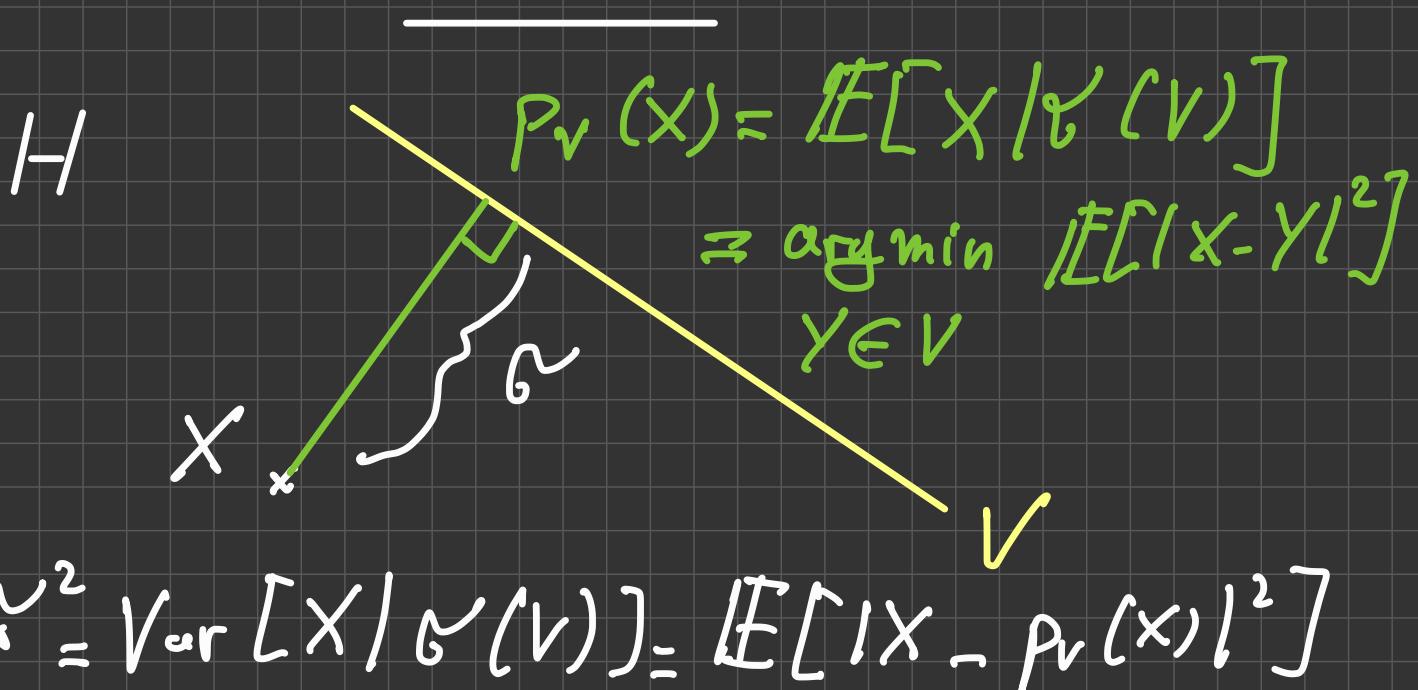
$$X|_{\mathcal{G}(V)} \sim \mathcal{N}(\hat{v}, \hat{v}^2)$$

$\forall P \in \mathcal{B}(\mathbb{R})$

$$P[X \in P | \mathcal{G}(V)] = \int_{\mathcal{G}(V)} dy e^{-\frac{(y - \hat{v})^2}{2\hat{v}^2}}$$

Rk important

in general $E[X|\mathcal{G}(V)] = P_{L^2(\Omega, \mathcal{G}(V), P)}(x)$
 here $= \hat{v}$



$$= \min_{Y \in V} \mathbb{E}[|X - Y|^2]$$

Proof

$$X = (X - \text{Pr}_V(X)) + \text{Pr}_V(X)$$

$$X - \text{Pr}_V(X) \perp \text{Pr}_V(X)$$

$\Rightarrow X - \text{Pr}_V(X)$ indep. from $\text{Pr}_V(X)$

$$\mathbb{E}[X | \mathcal{V}(V)] = \mathbb{E}[\text{Pr}_V(X) | \mathcal{V}(V)] = \text{Pr}_V(X)$$

$$+ \mathbb{E}[X - \text{Pr}_V(X) | \mathcal{V}(V)]$$

$\mathbb{E}[X - \text{Pr}_V(X)] = 0$

$$\mathbb{E}[X | \mathcal{V}(V)] = \text{Pr}_V(X)$$

$$X | \mathcal{V}(V) = (X - \text{Pr}_V(X)) | \mathcal{V}(V) + \text{Pr}_V(X) | \mathcal{V}(V)$$

$X - \text{Pr}_V(X) \sim \mathcal{N}(0, \sigma^2)$ cst

$$X|\sigma(v) \sim \mathcal{N}(\Pr(x), \sigma^2)$$

Cor $X \sim \mathcal{N}(0, K)$ Gaussian vector
on \mathbb{R}^d $d \geq 2$

K : non-degenerate

$\phi: m \times d$ matrix
rank m $0 < m < d$

$$E[X | \phi X = y] = K \phi^\top (\phi K \phi^\top)^{-1} y$$

$$X | \phi X = y \sim \mathcal{N}(E[X | \phi X = y], C)$$

$$C = K - K \phi^\top (\phi K \phi^\top)^{-1} \phi K$$

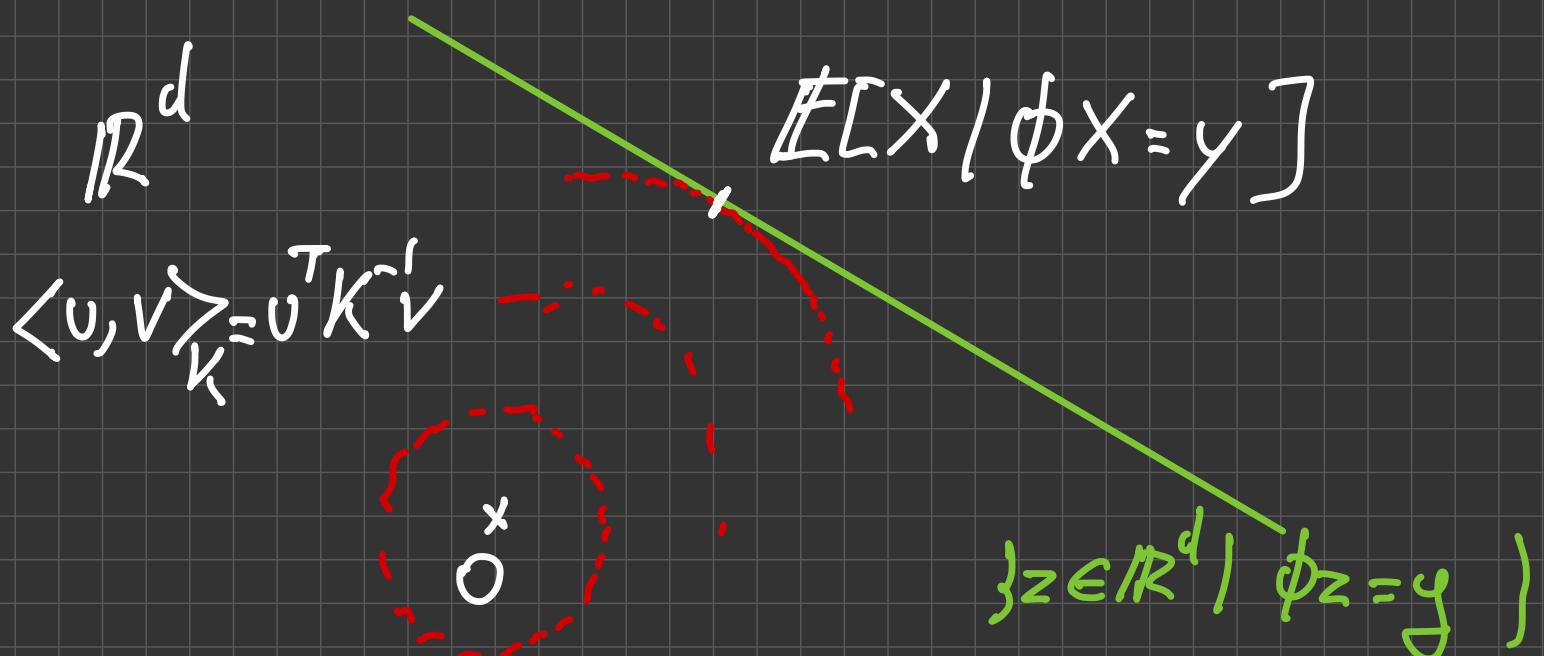
$$E[X | \phi X = y] = \arg \min_{\text{over } z \in \mathbb{R}^d} z^\top K^{-1} z$$

(optimal recovery) s.t. $\phi z = y$

For $v \in \mathbb{R}^d$

$$v^T C v = \min_{w \in \mathbb{R}^m} (v - \phi^T w)^T K (v - \phi^T w)$$

Geometric interpretation



$$\text{pdf}(x) \propto \exp\left(-\frac{z^T K^{-1} z}{2}\right)$$

$$C = \text{Cov}(X - E[X | \phi X])$$

$$V^T C V = V^T K^{\phi^T} V = \min_{w \in \mathbb{R}^m} (v - \phi^T w)^T K (v - \phi^T w)$$

\Downarrow

Schur complement of matrix K
operator K

$$\text{w.r.t. } \{\phi^T w, w \in \mathbb{R}^m\} \subset \mathbb{R}^d$$

(Shorted operator)

Proof $X | \phi X$

$$v^T X | \phi X \quad \forall v \in \mathbb{R}^d$$

$$\text{Prop} \quad v^T X | \phi X \sim \mathcal{N}(E[v^T X | \phi X], v^2)$$

$$E[v^T X | \phi X] = w^T \phi X$$

$$w = \arg \min_w E[|v^T X - w^T \phi X|^2]$$

$$\begin{aligned}
 F(w) &= E[|v^T X - w^T \phi X|^2] \\
 &= E[v^T X X^T v + w^T \phi X X^T \phi^T w \\
 &\quad - 2 w^T \phi X X^T v] \\
 &= v^T K v + w^T \phi^T K \phi w - 2 w^T \phi^T K v
 \end{aligned}$$

$$E[X X^T] = K$$

$$\begin{aligned}
 \nabla_w F(w) = 0 \iff \phi^T K \phi w = \phi^T K v \\
 w = (\phi^T K \phi)^{-1} \phi^T K v
 \end{aligned}$$

$$\begin{aligned}
 \Rightarrow E[v^T X | \phi X] &= w^T \phi X \\
 &= v^T K \phi^T (\phi^T K \phi)^{-1} \phi X \\
 \Leftarrow E[X | \phi X] &= K \phi^T (\phi^T K \phi)^{-1} \phi X
 \end{aligned}$$

representer formula

$$G^2 = \left[E\left[\| \nu^\top X - E[X | \phi X] \|^2 \right] \right]$$

$$C = \text{Cov}(X - E[X | \phi X])$$

$$= E[(X - E[X | \phi X])(X - E[X | \phi X])^T]$$

$$= E[XX^T] + E[E[X | \phi X] E[X | \phi X]^T]$$

$$- E[X E[X | \phi X]^T] - E[E[X | \phi X] X^T]$$

$$K = E[XX^T]$$

$$E[X | \phi X] = K \phi^T (\phi K \phi^T)^{-1} \phi X$$

$$= K + K \phi^T (\phi K \phi^T)^{-1} \phi K \phi^T (\phi K \phi^T)^{-1} \phi K$$

~~$$- K \phi^T (\phi K \phi^T)^{-1} \phi K - K \phi^T (\phi K \phi^T)^{-1} \phi K$$~~

$$= K - K \phi^T (\phi K \phi^T)^{-1} \phi K$$

$v^T X | \phi X$ Gaussian $\forall v$

\Downarrow
 $X | \phi X$ Gaussian

$$\sim \mathcal{N}(E[X|\phi X], C)$$

Now let's prove

$$(1) E[X|\phi X = y] = \arg \left\{ \begin{array}{l} \min \\ \text{over } z \in \mathbb{R}^d \\ \text{s.t. } \phi z = y \end{array} \right. \begin{array}{l} z^T K^{-1} z \\ \end{array}$$

(optimal recovery)

For $v \in \mathbb{R}^d$

$$(2) v^T C v = \min_{w \in \mathbb{R}^m} (v - \phi^T w)^T K (v - \phi^T w)$$

We have

$$E[X|\phi X = y] = K \phi^T (\phi^T K \phi^T)^{-1} y$$

$$\text{Write } V := K \phi^\top (\phi K \phi^\top)^{-1} y$$

$$\phi V = \phi K \phi^\top (\phi K \phi^\top)^{-1} y = y$$

$\Rightarrow V$ satisfies const. of (1)

$$\text{Let } w \in \mathbb{R}^d / \phi w = y$$

$$W = V + w - v$$

$$w^\top K^{-1} w = v^\top K^{-1} v + (w-v)^\top K^{-1} (w-v) \\ + 2 v^\top K^{-1} (w-v)$$

$$v^\top K^{-1} (w-v) = y^\top (\phi K \phi^\top)^{-1} \phi K \cancel{K^{-1}} (w-v) \\ = 0 \quad \text{because } \phi (w-v) = 0$$

$$\Rightarrow w^\top K^{-1} w = v^\top K^{-1} v + (w-v)^\top K^{-1} (w-v)$$

$$\Rightarrow V = \arg \min_{z \in \mathcal{Z}} z^\top K^{-1} z \\ \text{s.t. } \phi z = y$$

Now let us prove (2)

$$E[V^T X | \phi X] = w^T \phi X$$

$$w = \arg \min_{\bar{w}} E[|V^T X - \bar{w}^T \phi X|^2]$$

$$\text{G}^2 = V^T C V = E[|V^T X - w^T \phi X|^2]$$

$$\text{G}^2 = \min_{\bar{w}} E[|V^T X - \bar{w}^T \phi X|^2]$$

$$= \min_{\bar{w}} E[(V - \bar{\phi} \bar{w})^T (X^T) (V - \bar{\phi} \bar{w})]$$

$$= \min_{\bar{w}} (V - \bar{\phi} \bar{w})^T K (V - \bar{\phi} \bar{w})$$

$$V^T C V = \min_{w} (V - \bar{\phi} w)^T K (V - \bar{\phi} w)$$

Conditional covariance and precision matrix

Y, Z finite dim rand. vecs.

$\text{Cov}(Y)$: cov matrix of Y

$$(\text{Cov}(Y))_{ij} = \text{Cor}(Y_i, Y_j)$$

$\text{Cov}(Y|Z)$: cond cov of Y given Z

$$(\text{Cov}(Y|Z))_{ij} = \text{Cor}(Y_i, Y_j | Z)$$

$$\text{Cov}(Y) = E[(Y - E[Y])(Y - E[Y])^T]$$

$$\text{Cov}(Y, Z) = E[(Y - E[Y])(Z - E[Z])^T]$$

If $\text{Cov}(Y)$ invertible

$\text{Cov}(Y)^{-1}$: precision matrix of Y

X : r.v. on \mathbb{R}^n

$n \geq 3$ $X = (X_1, \dots, X_n) \sim \mathcal{N}(m, \Theta)$

$X_1, X_2 | X_3, \dots, X_n$

$\text{Cor}(X_1, X_2 | X_3, \dots, X_n)$

as a function $\Theta_{1,2}'$

X r.v. on \mathbb{R}^n

M $m \times n$ matrix

$\text{Cov}(MX) = M \text{Cov}(X) M^\top$

$n \geq 3$ $X = (X_1, \dots, X_n) \sim \mathcal{N}(m, \Theta)$

$Y = (X_1, X_2)$ $Z = (X_3, \dots, X_n)$

$$m = (m_y, m_z)$$

$$\Theta = \begin{pmatrix} \Theta_{yy} & \Theta_{yz} \\ \Theta_{zy} & \Theta_{zz} \end{pmatrix}$$

$$\Theta_{yy} = \text{Cov}(y) \quad \Theta_{yz} = \text{Cov}(y, z)$$

$$Y \sim \mathcal{N}(m_y, \Theta_{yy}) \quad Z \sim \mathcal{N}(m_z, \Theta_{zz})$$

$$\underline{\text{Prop}} \quad E[Y|Z] = m_y + \Theta_{yz} \Theta_{zz}^{-1} (z - m_z)$$

$$\text{Cov}(Y|Z) = \Theta_{yy} - \Theta_{yz} \Theta_{zz}^{-1} \Theta_{zy}$$

proof

$$Z = \phi X \quad \phi : (n-z) \times n$$

previous prop \rightarrow know
 $X | \phi X$

$$E[Y|z] = m_y + E[Y - m_y | z - m_z]$$

|| (n-2) red.

$$Y = Y - m_y + m_y$$

$\beta(z - m_z)$

$\beta: 2 \times (n-2)$ matrix

$$E[Y|z] = m_y + \beta(z - m_z)$$

$$\text{Cov}(Y - m_y - \beta(z - m_z), z - m_z) = 0$$

||

$$E[(Y - m_y - \beta(z - m_z))(z - m_z)^\top]$$

C

$$E[Y|z] = m_y + E[Y - m_y - \beta(z - m_z) | z - m_z]$$

$$+ E[\beta(z - m_z) | z - m_z]$$

$$Y = Y - m_Y - \beta (Z - m_Z) \\ + \beta (Z - m_Z) + m_Y$$

$$\mathbb{E}[Y|Z] = \mathbb{E}[Y - m_Y - \beta(Z - m_Z)|Z - m_Z] \\ + \beta(Z - m_Z) + m_Y$$

$$Y - m_Z - \beta (Z - m_Z) \perp Z - m_Z$$

↓

$$Y - m_Z - \beta (Z - m_Z) \text{ indep from } [Z - m_Z]$$

↓

$$\mathbb{E}[(Y - m_Y - \beta(Z - m_Z)) | Z - m_Z]$$

||

$$\mathbb{E}[Y - m_Y - \beta(Z - m_Z)] = 0$$

$$\mathbb{E}[Y|Z] = m_Y + \beta(z - m_Z)$$

$$\text{Cov}(Y - m_Y - \beta(z - m_Z), z - m_Z) = 0$$

||

$$E[(Y - m_Y - \beta(z - m_Z))(z - m_Z)^\top] = 0$$

||

$$\Theta_{YZ} - \beta \Theta_{ZZ} = 0$$

$$\beta = \Theta_{YZ} \Theta_{ZZ}^{-1}$$

$$\mathbb{E}[Y|Z] = m_Y + \Theta_{YZ} \Theta_{ZZ}^{-1} (z - m_Z)$$

$$\text{Cov}(Y|Z) = \text{Cov}(Y - m_Y - \beta(z - m_Z))$$

↙ why ?

$$Y|Z = \underbrace{Y - m_Y - \beta(z - m_Z)}_{\text{indep from } Z} + \underbrace{m_Y + \beta(z - m_Z)}_{\text{cst given } Z}$$

Q Why $Y - m_Y - \beta(Z - m_Z)$

and $(Z - m_Z)$
are jointly Gaussian

Because $\forall v \in \mathbb{R}^2, \forall w \in \mathbb{R}^{n-2}$

$$v^T (Y - m_Y - \beta(Z - m_Z)) + w^T (Z - m_Z) = \lambda^T (X - m_X)$$
$$\lambda \in \mathbb{R}^n$$

$$\text{Cov}(Y|Z) = \text{Cov}(Y - m_Y - \beta(Z - m_Z))$$

$$= E[(Y - m_Y - \beta(Z - m_Z))(Y - m_Y - \beta(Z - m_Z))^T]$$

$$= \underbrace{\Theta_{YY} - \Theta_{YZ} \Theta_{ZZ}^{-1} \Theta_{ZY}}_{\beta = \Theta_{YZ} \Theta_{ZZ}^{-1}}$$

$$\Theta = \begin{pmatrix} \Theta_{YY} & \Theta_{YZ} \\ \Theta_{ZY} & \Theta_{ZZ} \end{pmatrix}$$

Schur complement of
block ZZ of matrix Θ

$$\text{Cov}(y|z) = \Theta_{yy} - \Theta_{yz} \Theta_{zz}^{-1} \Theta_{zy}$$

$$\begin{aligned} \text{Var}[P^T Y | z] &= P^T \text{Cov}(Y|z) P \\ &= \text{Var}[P^T Y - P^T \beta(z - m_z)] \\ &= \inf_{v \in \mathbb{R}^{n-2}} \text{Var}[P^T Y - v^T (z - m_z)] \\ &= \inf_{v \in \mathbb{R}^{n-2}} (P, -v) \Theta \begin{pmatrix} P \\ -v \end{pmatrix} \\ &= \inf_{v \in \mathbb{R}^{n-2}} (P, v) \Theta \begin{pmatrix} P \\ v \end{pmatrix} \end{aligned}$$

variational formulation

of the Schur complement

Shorted Operator (if $\dim(z) = \infty$)

$$\text{Var}[P^T Y | z] = \text{Var}[P^T Y - P^T E[Y|z]]$$

$$\rho^T Y = \underbrace{\rho^T Y - \rho^T E[Y|Z]}_{\text{indep from } Z} + \underbrace{\rho^T E[Y|Z]}_{\text{Cst given } Z}$$

$$\text{Var}[\rho^T Y|Z] \approx \text{Var}[\rho^T Y - \rho^T E[Y|Z])$$

$$\begin{aligned} &= \text{Var}[\rho^T Y - \rho^T \beta(z - m_z)] \\ &= \inf_{v \in \mathbb{R}^{n-2}} \text{Var}[\rho^T Y - v^T (z - m_z)] \end{aligned}$$

$$E[Y|Z] = m_y + \beta(z - m_z)$$

$$\text{Var}[\rho^T Y - \rho^T \beta(z - m_z)]$$

$$\approx \text{Var}[\rho^T Y - \rho^T E[Y - m_y|Z]]$$



$$= \text{Var} [\rho^T (Y - m_Y) - \rho^T E[(Y - m_Y) | Z]]$$

$$= \text{Var} [\rho^T (Y - m_Y) - E[\rho^T (Y - m_Y) | Z - m_Z]]$$

$$E[\rho^T (Y - m_Y) | Z - m_Z]$$

is \perp proj. of $\rho^T (Y - m_Y)$ onto

$$\text{Span}_{\cup} \left\{ U^T (Z - m_Z) \right\}$$

$$= \min_U \text{Var} [\rho^T (Y - m_Y) - U^T (Z - m_Z)]$$

Prop

$$\text{Cov}(Y_1, Y_2 | Z) = \frac{-\Theta_{1,2}^{-1}}{\Theta_{1,1}^{-1} \Theta_{2,2}^{-1} - (\Theta_{1,2}^{-1})^2}$$

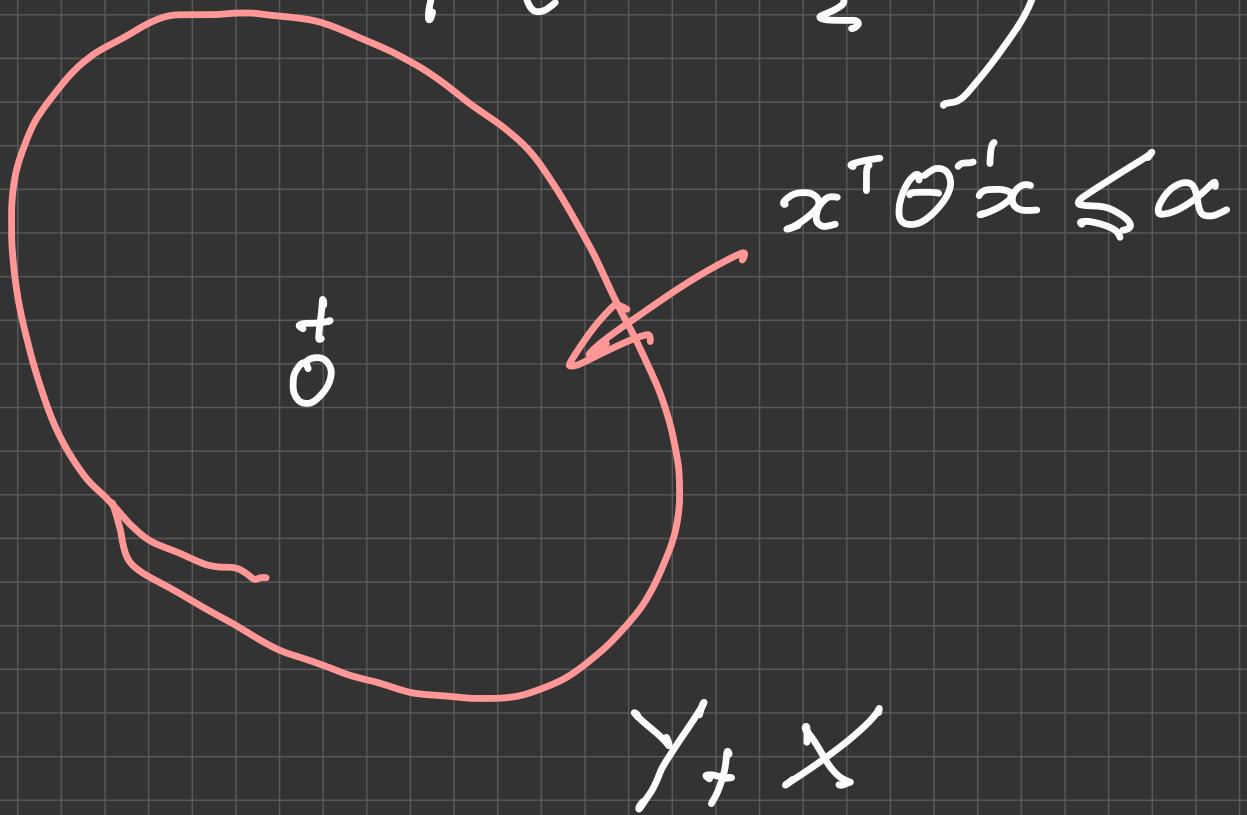
$$\text{Var}[Y_1 | Z] = \frac{\Theta_{2,2}^{-1}}{\Theta_{1,1}^{-1} \Theta_{2,2}^{-1} - (\Theta_{1,2}^{-1})^2}$$

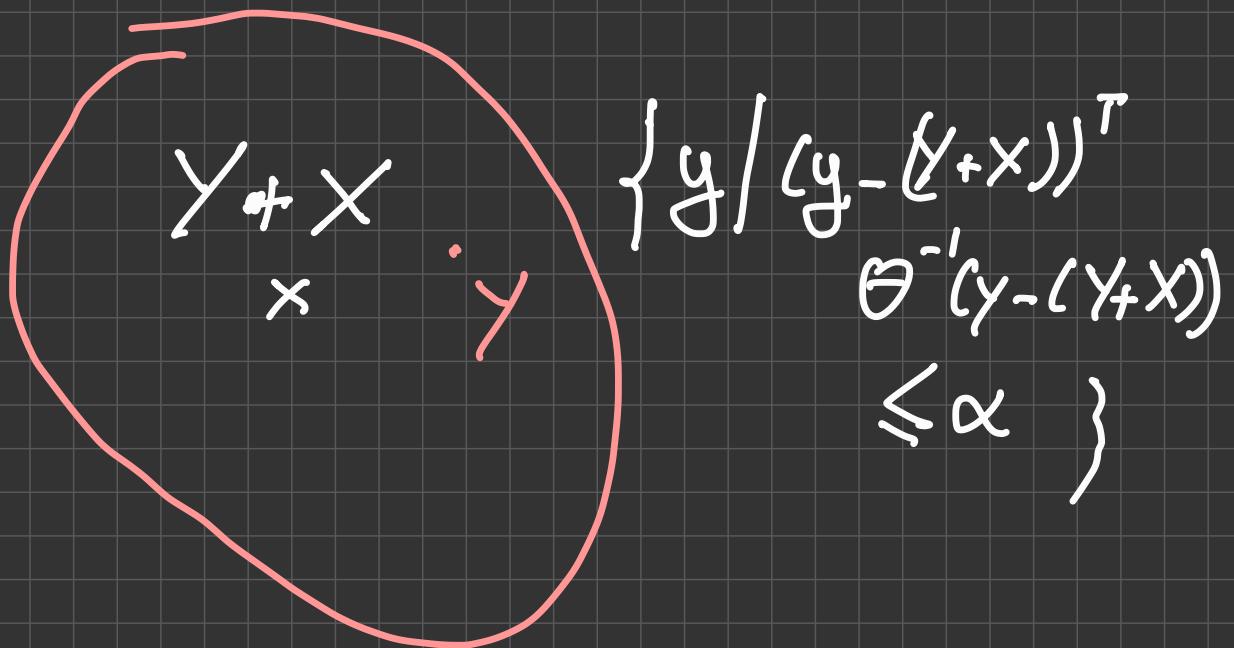
$$\text{Cor}[Y_1, Y_2 | Z] := \frac{\text{Cov}(Y_1, Y_2 | Z)}{\sqrt{\text{var}(Y_1 | Z)} \sqrt{\text{var}(Y_2 | Z)}}$$

$$= -\frac{\Theta_{1,2}^{-1}}{(\Theta_{1,1}^{-1})^{1/2} (\Theta_{2,2}^{-1})^{1/2}}$$

$$X \sim \mathcal{N}(m, \Theta)$$

$$P(x) \propto \exp\left(-\frac{x^\top \Theta^{-1} x}{2}\right)$$





Proof

$$\Theta_{zz}^{-1} = \begin{pmatrix} (\Theta_{yy} - \Theta_{yz} \Theta_{zz}^{-1} \Theta_{zy})^{-1} & 0 \\ -\Theta_{zz}^{-1} \Theta_{zy} (\Theta_{yy} - \Theta_{yz} \Theta_{zz}^{-1} \Theta_{zy})^{-1} & 0 \end{pmatrix}$$

$$(\Theta_{zz}^{-1})_{2,2} = (\Theta_{zz} - \Theta_{zy} \Theta_{yy}^{-1} \Theta_{yz})^{-1}$$

$$(\Theta_{yz}^{-1})_{2,2} = -\Theta_{yy}^{-1} \Theta_{yz} (\Theta_{zz} - \Theta_{zy} \Theta_{yy}^{-1} \Theta_{yz})^{-1}$$

$$\text{Cor}(Y|Z) = \Theta_{yy} - \Theta_{yz} \Theta_{zz}^{-1} \Theta_{zy}$$

$$\left(\text{Cov}(Y|Z) \right)^{-1} = \begin{pmatrix} \Theta_{1,1}^{-1} & \Theta_{1,2}^{-1} \\ \Theta_{2,1}^{-1} & \Theta_{2,2}^{-1} \end{pmatrix}$$



$$\Theta =$$

$$= \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

$$\Theta^{-1} =$$

$$= \begin{pmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{pmatrix}$$

Gaussian Processes

T : arbitrary set

Def A family of rand. var $\{X_t, t \in T\}$
is a (centered) Gaussian Process (GP)

if $\forall n \in \mathbb{N}, \forall t_1, \dots, t_n \in T$

$\forall \lambda_1, \dots, \lambda_n \in \mathbb{R}$

$$\lambda_1 X_{t_1} + \dots + \lambda_n X_{t_n}$$

is a (centered) Gaussian rand. var

Ex $T = \{1, 2, \dots, n\}$

$X = (X_1, \dots, X_n)$ is a GP

if X is a Gaussian
vector

Def A stochastic process X is a collection
 $\{X_t, t \in T\}$ of rand. var on a prob space

$(\Omega, \mathcal{U}, \mathbb{P})$

Def For a stochastic process X
we define $\mathcal{G}(X)$ as the σ -algebra
generated by events

$$(X_{t_1}, \dots, X_{t_k}) \in A \quad A \in \mathcal{B}(\mathbb{R}^k)$$

Thm The probability distribution of X
is uniquely determined by its finite
dimensional distributions

$$N_{t_1, \dots, t_k}(A) = \mathbb{P}[(X_{t_1}, \dots, X_{t_k}) \in A] \quad A \in \mathcal{B}(\mathbb{R}^k)$$

Ex $(X_t)_{t \in T}$

$(X_{t_1}, \dots, X_{t_k})$ is a Gaussian vector

because $\sum_i \lambda_i X_{t_i}$ is Gaussian $\forall \lambda \in \mathbb{R}^k$

$$(X_{t_1}, \dots, X_{t_K}) \sim \mathcal{N}\left(\dots \mathbb{E}[X_t] \dots\right),$$

$$\left(\begin{array}{c} \text{Cor}(X_{t_i}, X_{t_j}) \end{array} \right)$$

$$\mathbb{E}[X_t], \text{Cor}(X_t, X_s) \quad \forall t, s$$

Def Let $(X_t)_{t \in T}$ be a GP

The covariance function of X is
the function

$$\Gamma: T \times T \rightarrow \mathbb{R}$$

$$\Gamma(s, t) := \text{Cor}(X_s, X_t)$$

$$\left(= \mathbb{E}[X_s X_t] \right. \\ \left. \text{if } X \text{ is centered} \right)$$

Thm The prob. law of X is uniquely

determined if

$$m(t) = E[X_t]$$

$$\Gamma(s, t) = \text{Cor}(X_s, X_t)$$

$X \sim GP(m, \Gamma) \Leftrightarrow X$ is a GP with
mean m , cov. fct. Γ

$$\uparrow \\ \sim \mathcal{N}(m, \Gamma)$$

Ex $\Gamma: T \times T \rightarrow \mathbb{R}$

$$s, t \quad \text{Cor}(X_s, X_t)$$

$$\Gamma(s, t) = \Gamma(t, s)$$

$\forall k, t_1, \dots, t_k \in T$

$\Theta_{i,j} = \Gamma(t_i, t_j)$ must be positive

$$\Leftrightarrow \forall c \in \mathbb{R}^k, \sum_{i,j} c_i c_j \Gamma(t_i, t_j) \geq 0$$

Def We say that a function Γ on $T \times T$ is symmetric and of positive type if

$$\Gamma(s, t) = \Gamma(t, s) \quad \forall s, t$$

$$\forall k, t_1, \dots, t_k, c \in \mathbb{R}^k$$

$$\sum_{i,j} c_i c_j \Gamma(t_i, t_j) \geq 0$$

Ihm Let $m: T \rightarrow \mathbb{R}$

$$\Gamma: T \times T \rightarrow \mathbb{R}$$

symmetric positive

then \exists a GP $X \sim GP(m, \Gamma)$

proof Kolmogorov existence theorem

$$\underline{Ex} \quad T = \mathbb{R}_+ \quad \Gamma(s, t) \leq \min(s, t) \quad \Rightarrow \quad X \sim GP(0, \Gamma)$$

↓
Brownian Motion

$$T = [0, 1]^2 \quad \Gamma(s, t) = \min(s_1, t_1) \min(s_2, t_2)$$

$\rightarrow X \sim GP(0, \Gamma)$
 \rightarrow Brownian Sheet

$$T = \mathcal{L}^2(E, \mathcal{E}, \nu)$$

$$\left\{ f: E \rightarrow \mathbb{R} \mid \int_E f^2 d\nu < \infty \right\}$$

$$\Gamma(f, g) = \langle f, g \rangle_{\mathcal{L}^2}$$

$$= \int_E f(x) g(x) \nu(dx)$$

Gaussian measure
Gaussian field

$$\underline{Ex} \quad \xi_1 \sim GP(0, K_1) \quad \xi_1, \xi_2$$

indep

$$\xi_2 \sim GP(0, K_2)$$

$$\xi_1 + \xi_2$$

ξ_1 and ξ_2 indep

$$(\xi_1(t_1), \dots, \xi_1(t_k))$$

$$\text{indep from } (\xi_2(s_1), \dots, \xi_2(s_{k'}))$$

$$\xi_1 + \xi_2 \sim GP(0, K_1 + K_2)$$

$$\underline{Ex} \quad \xi_1 \mid \xi_1 + \xi_2 \quad ?$$

Kolmogorov existence theorem

The prob. dist of stochastic process X on T
 is uniquely determined by its finite dim
 distributions

$$* \quad N_{t_1, \dots, t_k}(A) = \mathbb{P}[X_{t_1}, \dots, X_{t_k} \in A] \quad \forall A \in \mathcal{B}(\mathbb{R}^k)$$

$\forall k \in \mathbb{N}, \forall t_1, \dots, t_k \in T$

$\xrightarrow{(*)}$ two consistency properties

$$(1) \quad \forall k \in \mathbb{N}, t_1, \dots, t_k \in T, A_1 \times \dots \times A_k \in \mathcal{B}(\mathbb{R}^k)$$

and π a permutation of $\{1, \dots, k\}$

$$N_{t_1, \dots, t_k}(A_1 \times \dots \times A_k) = N_{t_{\pi_1}, \dots, t_{\pi_k}}(A_{\pi_1} \times \dots \times A_{\pi_k})$$

$$(2) \quad \forall k \in \mathbb{N}, t_1, \dots, t_k \in T, A_1 \times \dots \times A_{k-1} \in \mathcal{B}(\mathbb{R}^{k-1})$$

$$N_{t_1, \dots, t_{k-1}}(A_1 \times \dots \times A_{k-1}) = N_{t_1, \dots, t_k}(A_1 \times \dots \times A_{k-1} \times \mathbb{R})$$

$$A_1 \times A_2 = \{ (x_1, x_2) \mid x_1 \in A_1, x_2 \in A_2 \}$$

$$A_1 \times A_2 \times \dots \times A_{k-1} \times \mathbb{R}$$

$$= \left\{ (x_1, \dots, x_k) \in \mathbb{R}^k \mid \begin{array}{l} x_1 \in A_1 \\ \vdots \\ x_{k-1} \in A_{k-1} \\ x_k \in \mathbb{R} \end{array} \right\}$$

Thm If N_{t_1, \dots, t_k} is a system of distributions satisfying (1) & (2) then \exists some prob. space $(\Omega, \mathcal{U}, \mathbb{P})$, a stochastic process $(X_t)_{t \in T}$ having N_{t_1, \dots, t_k} as its finite dimensional distributions

$$\mu: (t_1, \dots, t_k) \in T^k \xrightarrow{k \geq l} \mu_{t_1, \dots, t_k} \in \mathcal{S}(\mathbb{R}^k)$$

Brownian Motion as a Gaussian Process

Def For $T = \mathbb{R}_+$ and $\Gamma(s, t) = \min(s, t)$
the centered GP($0, \Gamma$) noted
 $(B_t, t \in \mathbb{R}_+)$ is called a Brownian Motion

E_x Assume that X is a BM

$$\cdot X_0 = 0 \quad (\sim \mathcal{N}(0, 0))$$

$$\cdot X_t \sim \mathcal{N}(0, t)$$

$$X_t - X_s \sim \mathcal{N}(0, t-s)$$

$$\mathbb{E}[(X_t - X_s)^2] \quad t > s$$
$$\min(t, t) + \min(s, s) - 2 \min(t, s)$$

$$(\mathbb{E}[X_t^2] + \mathbb{E}[X_s^2] - 2 \mathbb{E}[X_t X_s])$$

$$X \sim GP(0, \Gamma) \quad \Gamma(t, s) = \min(s, t)$$

$$\Gamma(t, t) + \Gamma(s, s) - 2 \Gamma(t, s)$$

$$\text{Cov}(X_t - X_s, X_r) = \text{Cov}(X_t, X_r) - \text{Cov}(X_s, X_r)$$

$t > s > r$

$$= \min(t, r) - \min(s, r)$$

$t > s > r_1, \dots, r_k \Rightarrow X_t - X_s \text{ indep from } X_{r_1}, \dots, X_{r_k}$

$\Rightarrow X_t - X_s \text{ and } X_r \text{ are indep.}$

Prop $(X_t, t \in \mathbb{R}^+)$ is a B.M



(i) $X_0 = 0$ a.s.

(ii) $\forall 0 \leq s \leq t$

$X_t - X_s \sim \mathcal{N}(0, t-s)$

indep from $\mathcal{G}(X_r, r \leq s)$

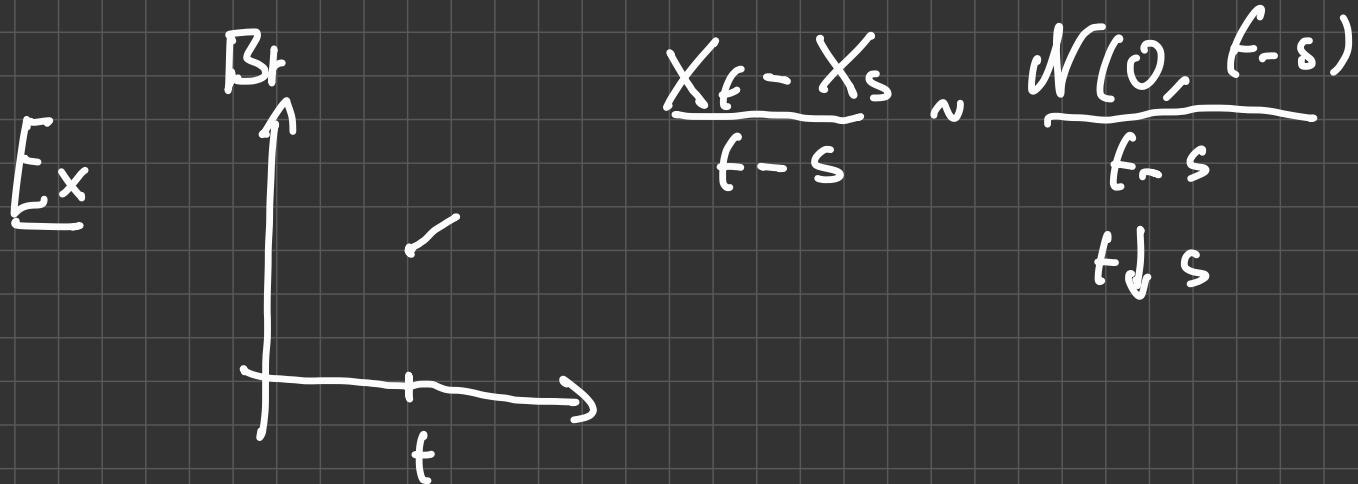
Proof

(i) $\Rightarrow X_t$ is a GP because

$\forall k, t_1, \dots, t_k, \lambda_1, \dots, \lambda_k$

$$\sum_i \lambda_i X_{t_i} = c_1 X_{t_1} + \sum_{i \geq 2} c_i (X_{t_{i+1}} - X_{t_i})$$

~~~~~ Gaussian  
~~~~~ Gaussian  
~~~~~ indep Gaussian



Cor  $(X_t, t \in \mathbb{R}_+)$  is a BM.

$$X_0 = 0 \text{ a.s.}$$

and  $\forall t_0=0 < t_1 < \dots < t_n$

the r.v.  $X_{t_i} - X_{t_{i-1}}$  are indep

and  $X_{t_i} - X_{t_{i-1}} \sim \mathcal{N}(0, t_i - t_{i-1})$

$$f(y_1, \dots, y_n) = \frac{1}{(2\pi)^{n/2}} \frac{1}{(t_1(t_2-t_1)) \dots (t_n(t_{n+1}-t_n))^{y_1}}$$

$$x \exp \left( - \sum_{i=1}^n \frac{(y_i - y_{i-1})^2}{2(t_i - t_{i-1})} \right)$$

Ex Let  $B$  be a BM.

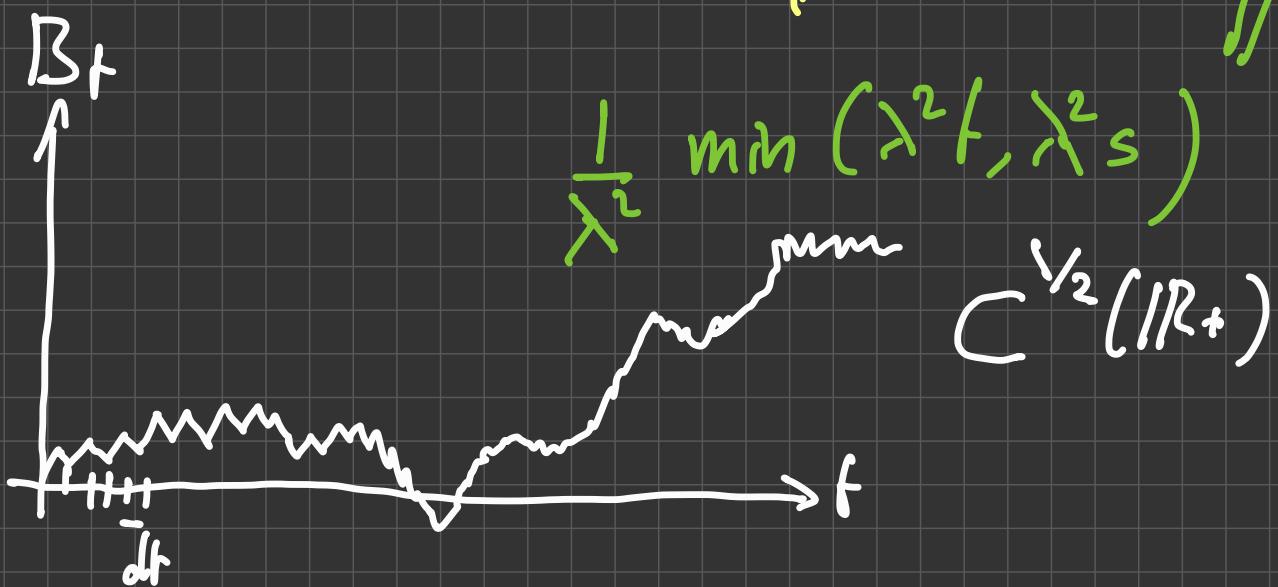
(i)  $-B$  is a BM

(ii)  $\forall \lambda > 0$ ,  $B_t^\lambda = \frac{1}{\lambda} B_{\lambda^2 t}$

is a BM.

$$\text{Cov}(B_t^\lambda, B_s^\lambda) = \frac{1}{\lambda^2} \text{Cov}(B_{\lambda^2 t}, B_{\lambda^2 s})$$

$$= \frac{1}{\lambda^2} \min(\lambda^2 t, \lambda^2 s)$$



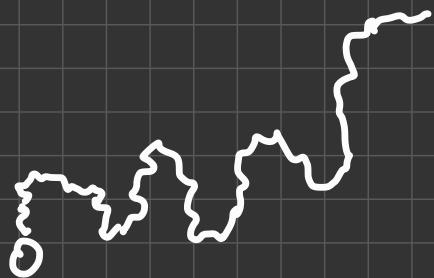
$$B_{t+dt} - B_t = dB_t \sim \mathcal{N}(0, dt)$$

$$\mathbb{E}[B_t^2] = t$$

$$(\mathbb{E}[B_t^2])^{1/2} \sim \sqrt{t}$$

Fick's law  
diffusive process

Brown:

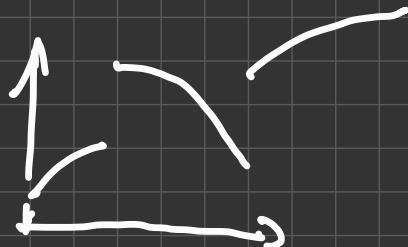


$$B_t = (B_t^1, B_t^2)$$

↳ 2 dim. Brownian Motion

Wiener process

Ex  $f: \mathbb{R}_+ \rightarrow \mathbb{R}$



$$P(s, f) = P(s) P(f)$$

$$X \sim GP(0, R) \quad Z \sim \mathcal{N}(0, I)$$

$$X_f = f(f) \quad Z$$

$$\mathbb{E}[X_f X_s] = f(f) f(s) \quad \mathbb{E}[Z^2]$$

$$= f(f) f(s) = P(s, f)$$



$X(t)$  : Stochastic Process

$\sigma(X_r | 0 \leq r \leq s)$

history of  $X$  up to time  $s$

$X$  is a Markov Process  $\Leftrightarrow t > s > 0$

$\forall B \subset \mathbb{R}^n$ , open

$$P[X_t \in B | \sigma(X_r | 0 \leq r \leq s)] = P[X_t \in B | X_s]$$

A B.M. is a Markov Process

$$t > s \quad B_t = \underbrace{B_t - B_s}_{\text{indep from } \mathcal{G}(B_r, r \leq s)} + B_s$$

$$\mathbb{P}[B_t \in A \mid \mathcal{G}(B_r, r \leq s)]$$

$$= \mathbb{P}[B_t \in A \mid B_s]$$

↑  
Markov Prop

Wiener process

$$W_t = (W_t^1, \dots, W_t^n)$$

$W_t^i$ : iid 1d BM.

$W$  is a Markov Process  
d-dim B.M.

# Gaussian Process Regression

$$(\xi(x))_{x \in \mathcal{X}} \sim GP(0, K)$$

$$E[\xi(x)] = 0 \quad \forall x \in \mathcal{X}$$

$$E[\xi(x) \xi(x')] = K(x, x') \quad \forall x, x' \in \mathcal{X}$$

$\mathcal{X}$  arbitrary

$$\begin{aligned} \xi: \mathcal{X} &\longrightarrow \text{Gaussian space} \\ x &\longrightarrow \xi(x) \sim \mathcal{N}(0, K(x, x)) \end{aligned}$$

$$\text{Let } X := (X_1, \dots, X_N) \in \mathcal{X}^N$$

$$\text{Write } \xi(X) := (\xi(X_1), \dots, \xi(X_N))$$

$$Ex \quad \xi(X) \sim \mathcal{N}(0, K(X, X))$$

$K(X, X)$ :  $N \times N$  matrix with entries  $K(X_i, X_j)$

Def Call  $K$  non-degenerate if

$K(X, X)$  is invertible  $\forall N, X_1, \dots, X_N /$   
 $X_i \neq X_j$   
for  $i \neq j$

Ex  $X = \mathbb{R}^d \Leftrightarrow K(x, x') = \exp\left(-\frac{\|x - x'\|^2}{2\sigma^2}\right)$   
non-degenerate

Assume  $X_i \neq X_j$  for  $i \neq j$

$K$  non-degenerate

Thm Let  $Y \in \mathbb{R}^N$ . Then  $\forall x \in X$

$$(i) E[\xi(x) | \xi(X) = Y] = m(x)$$

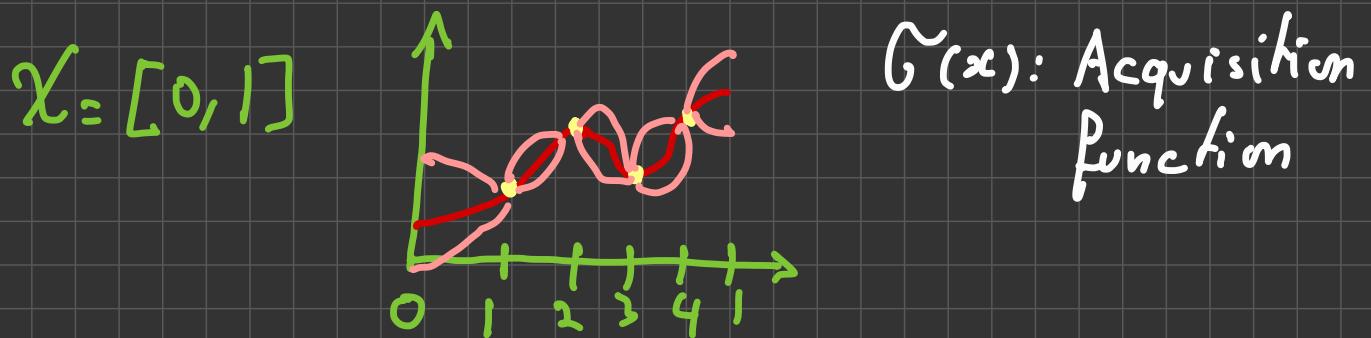
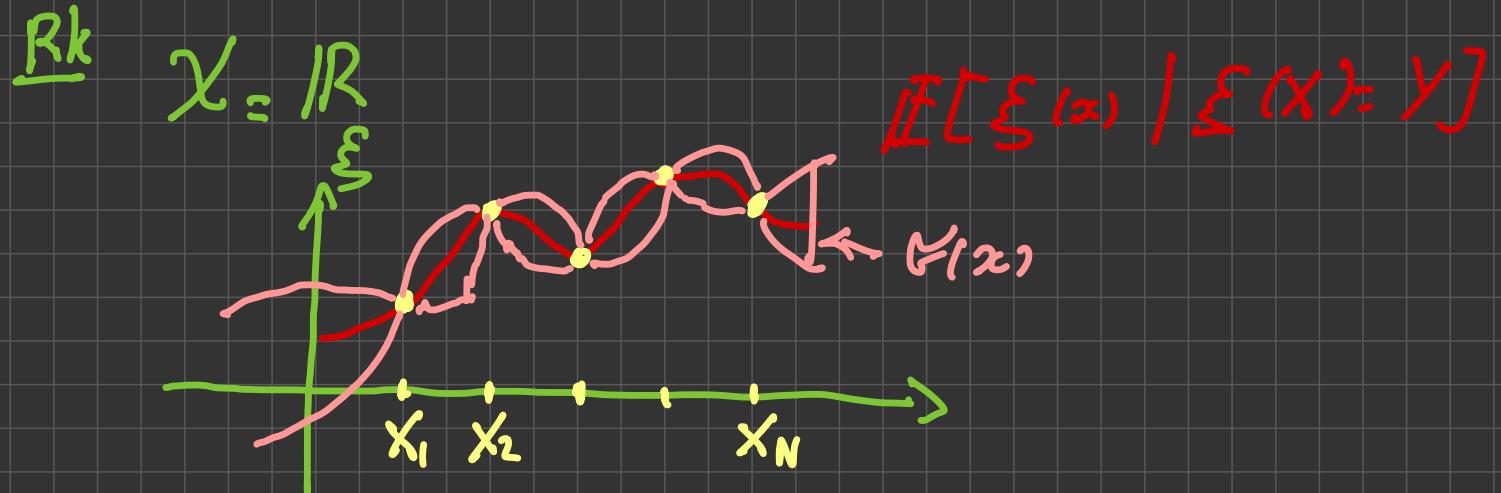
$$= K(x, X) K(X, X)^{-1} Y$$

$K(x, X)$ :  $N$ -vector with entries  
 $K(x, X_i)$

$$(2) \text{Var}[\xi(x) | \xi(X) = y] = G^2(x)$$

$$= K(x, x) - K(x, X) K(X, X)^{-1} K(X, x)$$

↓ does not depend on  $y$



$$(3) \xi(x) \mid \xi(X) = Y \sim \mathcal{N}(m(x), G^2(x))$$

$$(\xi(x))_{x \in X} \mid \xi(X) = Y \sim GP(m(x), G)$$

$$G(x, x') = K(x, x') - \frac{K(x, X) K(X, x')}{K(X, X')}$$

Proof

$$(ξ(x), ξ(x_1), \dots, ξ(x_N))$$

Gaussian vector on  $\mathbb{R}^{N+1}$

Get (1) + (2) from previous thm  
on cond. Gaussian vectors

$$Z = (ξ(x), ξ(x_1), \dots, ξ(x_N))$$

Z is a centered Gaussian vector

$$Z_1 | Z_2, \dots, Z_{N+1}$$

Proof of (3)

$$x_1, \dots, x_m \in \mathcal{X}^m$$

$(\xi(x_1), \dots, \xi(x_m), \xi(X_1), \dots, \xi(X_N))$   
 Gaussian vector on  $\mathbb{R}^{m+N}$

$(\xi(x_1), \dots, \xi(x_m)) \mid \xi(X)$  Gaussian  
 vector on  $\mathbb{R}^m$

$\Rightarrow$  linear combination of its entries  
 are Gaussian

$\Rightarrow (\xi(x))_{x \in X} \mid \xi(X)$  is a GP

we know its mean

$$\begin{aligned} \text{Cov}(\xi(x_1), \xi(x_2) \mid \xi(X)) \\ = K(x_1, x_2) - K(x_1, X) K(X, X)^{-1} K(X, x_2) \end{aligned}$$

Rk

$$E[\xi(x) \mid \xi(X) = Y] = m(x)$$

$$= K(x, X) K(X, X)^{-1} Y$$

$$= \sum_{i=1}^N Y_i \Psi_i(x)$$

$$\Psi_i(x) = K(x, X) K(X, X)^{-1} \cdot$$

$$= \sum_j K(x, X_j) (K(X, X))^{-1}_{j,i}$$

$\Psi_i$ : Optimal Recovery splines

$$\Psi_i(x) = E[\xi(x) \mid \xi(X_j) = \delta_{ij}, V_j]$$

elementary gambler / bet gambler

$$\underline{Ex} \quad f(t) = \sin(t) e^{t^2 - \sin(t+1)}$$

Approximate

$$I = \int_0^1 f(t) dt$$

$$t_k = \frac{k}{N} \quad k \in \{0, 1, \dots, N\}$$

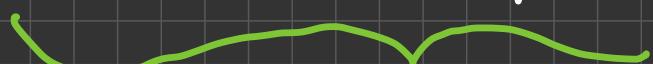
given  $f(t_k)$  try to approximate  $I$

$I_{\text{true}}$   $f$  deterministic  $\Leftrightarrow \beta_t: 1d BM$

$$\beta \sim GP(0, K)$$

$$K(t, t') = \min(t, t')$$

$$I \approx \mathbb{E} \left[ \int_0^1 \beta_t dt \mid \beta_{t_i} = f(t_i) \forall i \right]$$

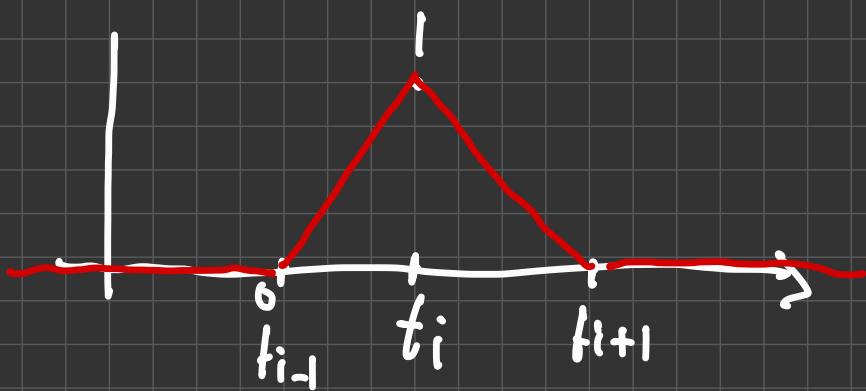
$$= \int_0^1 \mathbb{E} [\beta_t \mid \beta_{t_i} = f(t_i) \forall i] dt$$


$$\sum_i f(t_i) \Psi_i(t) \quad \stackrel{\min(t, t_j)}{=} \quad$$

$$\Psi_i(t) = \sum_j K(t, t_j) \Theta_{j,i}$$

$$\Theta_{i,j} = \min(t_i, t_j) = K(t_i, t_j)$$

$$\Psi_i(t) = \mathbb{E}[B_t \mid \begin{array}{l} B_{t_i} = 1 \\ B_{t_j} = 0 \quad \forall j \neq i \end{array}]$$



$$I \approx \int_0^1 \sum_{i=0}^N f(t_i) \Psi_i(t)$$

piecewise linear interpolation of  $f$   
 rediscover trapezoidal quadrature rule

$$\int_{[0,1]^2} f(t) dt$$

$$f \leftrightarrow X_t \sim GP(0, \Gamma)$$

$$\Gamma(s, t) = \min(s_1, t_1) \min(s_2, t_2)$$

$$\underline{\int_0^t} f(t) dt$$

$$f \leftrightarrow B_t$$

$$f \leftrightarrow X_t = \int_0^t B_s ds \sim GP(0, \Gamma)$$

$$\Gamma(s, t) = E[X_t, X_s]$$

$$= E\left[\int_s^t B_r dr \int_0^t B_{r'} c(r') dr'\right]$$

$$= E\left[\int_{\substack{0 \leq r \leq s \\ 0 \leq r' \leq t}} B_r B_{r'} c(r') dr dr'\right]$$

$$= \int_{\substack{0 \leq r \leq s \\ 0 \leq r' \leq t}} \min(r, r') dr dr'$$

$$I = \int_0^t p(t) dt = \int_0^t \underbrace{\mathbb{E}[X_t] X_{t_i}}_{\text{cubic spline interpolant of } p} dt$$

cubic spline interpolant of  $p$

$$p \leftrightarrow \int_0^t \int_s^t B_s ds \Leftrightarrow \text{spline of order 5}$$

$$\underline{Ex} - B_t: I \subset BM$$

$GP(0, K)$

$$\mathbb{E}[B_t B_{t'}] = K(t, t') = \min(t, t')$$

$$= \mathbb{E}\left[\int_{\begin{array}{c} 0 \leq r \leq s \\ 0 \leq r' \leq t \end{array}} B_r B_{r'} dr dr'\right]$$

$$B_r B_{r'} \leq |B_r B_{r'}|$$

$$\mathbb{E}\left[\int_{\begin{array}{c} 0 \leq r \leq s \\ 0 \leq r' \leq t \end{array}} |B_r B_{r'}| dr dr'\right] < \infty$$

$$\int_{\substack{0 \leq r \leq s \\ 0 \leq r' \leq t}} \mathbb{E}[\langle B_r, B_{r'} \rangle] dr dr' \stackrel{?}{<} \infty$$

"

$$\leq \underbrace{\mathbb{E}[B_r^2]^{\frac{1}{2}}}_{\sqrt{r}} \underbrace{\mathbb{E}[B_{r'}^2]^{\frac{1}{2}}}_{\sqrt{r'}}$$

↓

$$\leq \int_{\substack{0 \leq r \leq s \\ 0 \leq r' \leq t}} \sqrt{r} \sqrt{r'} dr dr' < \infty$$

$$\underline{\text{Ex}} \quad I = \int_{[0,1]^{50}} f(x) dx$$

$$f: \mathbb{R}^{50} \rightarrow \mathbb{R}$$

Approximate  $I$

$$\underline{\text{Sol}} \quad X_1, \dots, X_N \stackrel{\text{iid}}{\sim} \text{Uniform}([0,1]^{50})$$

$$f \leftarrow \xi_N \mathcal{N}(0, K)$$

$$I \approx E\left[\int_{[0,1]^{50}} \xi(x) dx \mid \xi(X) = f(X)\right]$$

$\rightarrow$  quadrature rule in claim 50

$$I \approx \int_{[0,1]^{50}} \sum_{i=1}^N f(x_i) \Psi_i(x)$$

$$\Psi_i(x) = K(x, X) K(X, X)_{:, i}^{-1}$$

$$\approx \sum_{i=1}^N f(x_i) c_i$$

$$c_i = \int_{[0,1]^{50}} \Psi_i(x) dx$$

$\downarrow$  quadrature coefficients

can be pre-computed

Bayesian Quadrature Rule

$$\underline{Ex} \quad \xi_1 \sim GP(0, K_1)$$

$$\xi_2 \sim GP(0, K_2)$$

$\xi_1, \xi_2$  indep

$$\xi_1 + \xi_2 \sim GP(0, K)$$

$$K = K_1 + K_2$$

What is

$$\xi_1 | \xi_1 + \xi_2 ?$$

$$\underline{Ex} \quad X \in (X_1, \dots, X_n) \in \mathcal{X}^N$$

$$\xi_1(X) / (\xi_1 + \xi_2)(X) = Y$$

$$\underline{Sol} \quad (\xi_1(X), \xi_1(X) + \xi_2(X))$$

$$\sim \mathcal{N}(0, C)$$

$$C = \begin{pmatrix} K_1(X, X) & K_1(X, \tilde{X}) \\ K_1(\tilde{X}, X) & K(\tilde{X}, \tilde{X}) \end{pmatrix}$$

$$\xi_1(x) \Big| \xi_1(x) + \xi_2(x) = y \\ \sim \mathcal{N}(m, C')$$

$$m = K_1(X, X) K(X, \tilde{X})^{-1} y$$

$$C' = K_1(X, X) - K_1(X, X) K(X, X)^{-1} \\ K_1(\tilde{X}, X)$$

$$Ex \quad \xi_1 \Big| \xi_1 + \xi_2 \sim \begin{array}{c} \xrightarrow{\text{+ + + + + + +}} \\ x_1 \dots x_N \end{array} \rightarrow X$$

$$\xi_1 \Big| \xi_1 + \xi_2 \sim GP(m, C)$$

$$m(x) = K_1(x, \cdot) K(\cdot, \cdot)^{-1} (\xi_1 + \xi_2)$$

$$C = K_1 - K_1 K^{-1} K_1$$

↓  
operator

What is this

Need to define covariance operators

Later  
↔  
Gaussian fields

# Gaussian Measure

Wiener integral

$B_t: \text{Id } BM$

$f \in C([0, 1])$

What is  $\int_0^1 f(s) dB_s$  ?

$$t_i = \frac{i}{n}$$

$$\int_0^1 f(s) ds \approx \sum_i f(t_i) (t_{i+1} - t_i)$$

take limit  $n \rightarrow \infty$

$$\int_0^1 f(s) dB_s \approx \sum_{i=0}^{n-1} f(t_i) (B_{t_{i+1}} - B_{t_i})$$

$I_n$        $n \rightarrow \infty$  ?

$$I_n \sim \mathcal{N}(0, \sum_{i=0}^{n-1} f^2(t_i) (t_{i+1} - t_i))$$

$$\begin{aligned} & \mathbb{E}\left[\left(\sum_{i=0}^{n-1} f(t_i) (B_{t_{i+1}} - B_{t_i})\right)^2\right] \\ &= \sum_{i,j} f(t_i) f(t_j) \underbrace{\mathbb{E}[(B_{t_{i+1}} - B_{t_i})(B_{t_{j+1}} - B_{t_j})]}_{\begin{array}{l} = 0 \quad i \neq j \\ = t_{i+1} - t_i \quad i = j \end{array}} \end{aligned}$$

$$I_n \sim \mathcal{N}(0, \sum_{i=0}^{n-1} f^2(t_i) (t_{i+1} - t_i))$$

$$\downarrow n \rightarrow \infty \qquad \qquad \qquad \downarrow n \rightarrow \infty$$

$$I \sim \mathcal{N}(0, \int_0^1 f^2(t) dt)$$

$$\int_0^1 f(s) dB_s \sim \mathcal{N}(0, \int_0^1 f^2(t) dt)$$

$$\mathbb{L}^2(X, \mu) = \left\{ f: X \rightarrow \mathbb{R} \mid \int f^2(x) \mu(dx) < \infty \right\}$$

$$G: \mathbb{L}^2(X, \mu) \xrightarrow{\text{Gaussian space}} \mathcal{H}$$

$$f \xrightarrow{} G(f) \sim \mathcal{N}(0, \|f\|_{\mathbb{L}^2(X, \mu)}^2)$$

$$g \in C([0, 1])$$

What is

$$\left( \int_0^1 f(s) dB_s, \int_0^1 g(s) dB_s \right) ?$$

$$\sim \mathcal{N}(0, C)$$

$$\approx \left( \sum_{i=0}^{n-1} f(t_i) (B_{t_{i+1}} - B_{t_i}), \sum_{i=0}^{n-1} g(t_i) (B_{t_{i+1}} - B_{t_i}) \right) \sim \mathcal{N}(0, C_n)$$

$$C_n = \begin{pmatrix} \sum_{i=0}^{n-1} P^2(f_i) (f_{i+1} - f_i) \\ \vdots \\ \sum_{i=0}^{n-1} P(f_i) g(f_i) (f_{i+1} - f_i) \quad \sum_{i=0}^{n-1} g^2(f_i) (f_{i+1} - f_i) \end{pmatrix}$$

$\downarrow n \rightarrow \infty$

$$C = \begin{pmatrix} \int_0^1 P(f) df \\ \int_0^1 P(f) g(f) df \\ \int_0^1 g^2(f) df \end{pmatrix}$$

Ex Applications of the Wiener Integral

Ex  $B_t$  1d BM

$$\int_0^t f(s) dB_s \sim \mathcal{N}(0, \int_0^t f^2(s) ds)$$

$$f(s) = \begin{cases} s & s \leq t \\ 0 & s > t \end{cases}$$

$$\int_0^t s dB_s \sim \mathcal{N}\left(0, \int_0^t s^2 ds\right)$$

$$\underline{\text{Ex}} \quad (B_t, \int_0^t \cos(s) dB_s) \sim \mathcal{N}(0, C)$$

$$\int_0^t dB_s$$

$$C = \begin{pmatrix} t & \int_0^t \cos(s) ds \\ \int_0^t \cos(s) ds & \int_0^t \cos^2(s) ds \end{pmatrix}$$

Ex Estimating the response function of  
of a neuron

$x(t) \rightarrow$  neuron  $\rightarrow y(t)$

$$y(t) = \int_0^t K(t-s) dx(s) + Z_t$$

$s \rightarrow K(s)$ : linear response kernel

$t \rightarrow Z_t$ : stochastic process whose dist.  
is unknown and indep. of  $t \rightarrow x(t)$

Sol  $y_t(\beta, z) = \int_0^t K(t-s) d\beta_s + Z_t$

$$f \in L^2([0,t], dx)$$

$$\begin{aligned} E[y_t(\beta, z) \int_0^t f(s) d\beta_s] \\ = \int_0^t f(s) K(t-s) ds \end{aligned}$$

$$\mathbb{E} \left[ y_t(\beta, z) \int_0^t f(s) dB_s \right]$$

$$\begin{aligned}
&= \mathbb{E} \left[ \int_0^t K(t-s) dB_s \int_0^t f(s) dB_s \right] \\
&\quad + \underbrace{\mathbb{E} \left[ Z_t \int_0^t f(s) dB_s \right]}_{\text{II}} \\
&\stackrel{\downarrow}{=} \int_0^t K(t-s) f(s) ds
\end{aligned}$$

$$\mathbb{E} \left[ \int_0^t K(t-s) dB_s \int_0^t f(s) dB_s \right]$$

Generalization

$(E, \mathcal{E})$ : measurable space

$\mu$ :  $\mathcal{G}$ -finite measure  $(E, \mathcal{E})$



$\exists (A_i)_{1 \leq i \leq \infty}$  s.t.  $E = \bigcup_i A_i$  and  
 $\forall i, \mu(A_i) < \infty$

Ex  $E = \mathbb{R}^d$

$\mathcal{E} = \mathcal{B}(\mathbb{R}^d)$  (Borel subsets of  $\mathbb{R}^d$ )

$\mu(dx) = c_x$  Lebesgue measure

Def

A Gaussian measure of intensity  $\mu$

is an isometry (linear map preserving the inner product) from  $\mathbb{L}^2(E, \mathcal{E}, \mu)$

to a Gaussian space

$G: \mathbb{L}^2(E, \mathcal{E}, \mu) \rightarrow \mathcal{F}$   
 $f \quad \mapsto \quad Gf$

$\forall \lambda \in \mathbb{R}, f, g \in \mathbb{L}^2(E, \mathcal{E}, \mu)$

$$G(f+g) = G(f) + G(g)$$

$$E[G(f)G(g)] = \langle f, g \rangle_{\mathcal{L}^2(E, \mathcal{E}, \mu)}$$

$$\mathcal{L}^2(E, \mathcal{E}, \mu) := \left\{ f: E \rightarrow \mathbb{R} \mid \int_E f^2(x) \mu(dx) < \infty \right\}$$

$$\exists f \in \mathcal{L}^2(E, \mathcal{E}, \mu)$$

$$G(f) \sim \mathcal{N}(0, E[G(f)^2])$$

$$\|f\|_{\mathcal{L}^2(E, \mathcal{E}, \mu)}^2$$

$$A \in \mathcal{E}, \mu(A) < \infty \quad f = \mathbf{1}_A$$

$$G(A) \approx G(\mathbf{1}_A) \sim \mathcal{N}(0, \mu(A))$$

$$\exists E = [0, 1]$$

$$\mathcal{E} = \mathcal{B}([0, 1])$$

$$\mu(dx) = dx$$

$$\mu([a, b]) = b - a \quad \|f\|_{L^2(E, \mathcal{E}, \mu)}^2$$

$$f: [0, 1] \rightarrow \mathbb{R} \mid \int_0^1 f^2(x) dx < \infty$$

$$G_p(f) \sim \mathcal{N}(0, \int_0^1 f^2(x) dx)$$

$$\nu(dx) = \cos^2(\pi x) dx$$

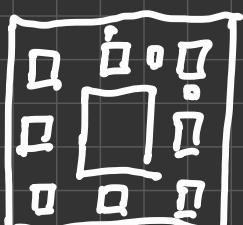
$$\nu([a, b]) = \int_a^b \cos^2(\pi x) dx$$

$$f: [0, 1] \rightarrow \mathbb{R} \mid \int_0^1 f^2(x) \cos^2(\pi x) dx < \infty$$

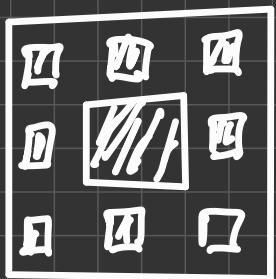
$$G_\nu(f) = \mathcal{N}(0, \underbrace{\int_0^1 f^2(x) \cos^2(\pi x) dx}_{\|f\|_{L^2(E, \mathcal{E}, \nu)}^2})$$

$$\|f\|_{L^2(E, \mathcal{E}, \nu)}^2$$

$$E_x \quad E = [0, 1]^2$$



$\mu$ : measure on Sierpinski carpet



$$f: [0, 1]^2 \rightarrow \mathbb{R}$$

$$G_p(f) = \mathcal{N}(0, \int_S f(x)^2 \mu(dx))$$

$E_x(E, \mathcal{E}, \mu)$

$$A_1, \dots, A_n \in \Sigma \quad \begin{array}{l} \mu(A_i) < \infty \\ A_i \cap A_j = \emptyset \text{ for } i \neq j \end{array}$$

$$(G(A_1), \dots, G(A_n)) \sim \mathcal{N}(0, C)$$

"

$G(\eta_{A_i})$

$$C_{i,j} = 0 \quad i \neq j$$

$$C_{i,i} = \mu(A_i)$$

$G(A_i) \in \mathcal{H}$  (Gaussian space)

$$\mathbb{E}[G(A_i) G(A_j)] = \int \mathbf{1}_{A_i} \mathbf{1}_{A_j} \mu(dx) = 0$$

Existence and construction of Gaussian Measures

Thm If  $(E, \mathcal{E})$  is a measurable space and  $\mu$  is a  $\sigma$ -finite measure on  $(E, \mathcal{E})$  such that  $L^2(E, \mathcal{E}, \mu)$  is separable (i.e. it contains a dense countable subset, e.g.  $E = \mathbb{R}_+$ ,  $\mu(dx) = dx$ ), then there exists a Gaussian measure of intensity  $\mu$  on  $(E, \mathcal{E})$

Proof  $L^2(E, \mathcal{E}, \mu)$  is separable

$\Rightarrow \exists \gamma_n$ : orthonormal basis of  $L^2(E, \mathcal{E}, \mu)$

$$\varphi_n \in \mathbb{L}^2(E, \mathcal{E}, \nu)$$

$$\langle \varphi_n, \varphi_m \rangle_{\mathbb{L}^2} = \delta_{n,m} = \begin{cases} 1 & n=m \\ 0 & n \neq m \end{cases}$$

$$\forall f \in \mathbb{L}^2(E, \mathcal{E}, \nu)$$

$$f = \sum_{n=1}^{\infty} \langle f, \varphi_n \rangle_{\mathbb{L}^2} \varphi_n$$

$\xi_n$ : iid  $\mathcal{N}(0, 1)$  r.v

Define

$$G(f) := \sum_{n=1}^{\infty} \langle f, \varphi_n \rangle_{\mathbb{L}^2} \xi_n$$

$G(f)$  is Gaussian

$$\mathbb{E}[G(f)] = 0$$

$$\forall f, g \in \mathbb{L}^2(E, \mathcal{E}, \nu)$$

$$g = \sum_{n=1}^{\infty} \langle g, \varphi_n \rangle_{\mathbb{L}^2} \varphi_n$$

$$\mathbb{E}[G(f)G(g)] =$$

$$\mathbb{E}\left[\sum_{n=1}^{\infty} \langle f, \varphi_n \rangle \xi_n \sum_{m=1}^{\infty} \langle g, \varphi_m \rangle \xi_m\right]$$

$$= \sum_{n,m=1}^{\infty} \langle f, \varphi_n \rangle \langle g, \varphi_m \rangle \underbrace{E[\xi_n \xi_m]}_{\delta_{n,m}}$$

$$= \sum_s \langle f, \varphi_n \rangle \langle g, \varphi_n \rangle = \langle f, g \rangle_{L^2}$$


---

$$\|f\|_{L^2}^2 = \sum_n \langle f, \varphi_n \rangle^2$$

$$E\left[ \sum_{n=1}^{\infty} \langle f, \varphi_n \rangle \xi_n \sum_{m=1}^{\infty} \langle g, \varphi_m \rangle \xi_m \right]$$

||

$$E\left[ \sum_{n,m=1}^{\infty} \langle f, \varphi_n \rangle \langle g, \varphi_m \rangle \xi_n \xi_m \right]$$

$$E\left[ \sum_{n,m=1}^{\infty} \langle f, \varphi_n \rangle \langle g, \varphi_m \rangle |\xi_n| |\xi_m| \right] < \infty$$


---

$$\text{Rk } f = \sum_{n=1}^{\infty} \langle f, \varphi_n \rangle_{L^2} \varphi_n$$

$\xi_n \sim N(0, 1)$   
iid

$$G(f) = \sum_{n=1}^{\infty} \langle f, \varphi_n \rangle_{L^2} \xi_n$$

Define (formally)

$$\xi = \sum_{n=1}^{\infty} \xi_n \varphi_n$$

$\xi$  is not a random element of  $L^2(E, \mathcal{E}, \mu)$

$$\|\xi\|_{L^2}^2 = \sum_{n=1}^{\infty} \xi_n^2 \text{ is a.s. infinite}$$

However we formally have

$$\begin{aligned} \langle \xi, f \rangle_{L^2(E, \mathcal{E}, \mu)} &= \sum_{n=1}^{\infty} \xi_n \langle \varphi_n, f \rangle_{L^2} \\ &= G(f) \end{aligned}$$

$\forall m, \xi^m := \sum_{n=1}^m \xi_n \varphi_n$  is a random element  
of  $L^2(E, \mathcal{E}, \mu)$

$$G(f) = \lim_{m \rightarrow \infty} \langle \xi^m, f \rangle_{L^2(E, \mathcal{E}, \mu)}$$

$G$ : is not a prob. measure on  $L^2(E, \mathcal{E}, \mu)$   
but a cylinder measure (limit of  
prob. measures)

In the sense that the cylinder measure  
associated with  $\xi$  can be defined as a  
weak limit of Gaussian measures on  
 $L^2(E, \mathcal{E}, \mu)$

---

Gaussian vectors  $\times$  r.e. of  $\mathbb{R}^d$   
 $(L^2(E, \mathcal{E}, \mu))$

$$G(f) \sim \mathcal{N}_{\text{GO}}(0, \|f\|_{L^2}^2)$$

$T \in L^2(E, \mathcal{E}, \mu)$

$$\mathbb{M}(f, g) = \langle f, g \rangle_{L^2}$$

$G(f)$ : GP with mean zero and cov. function  $\mathbb{M}$

$$\underline{E} \quad E = [0, 1] \quad \mu(dx) = dx$$

$$\mathbb{L}^2(E, \mathcal{E}, \mu) = \mathbb{L}^2([0, 1], \mathcal{B}([0, 1]), dx)$$

$$Y_n : \left\{ 1, \sqrt{2} \sin(2\pi n t), \sqrt{2} \cos(2\pi n t) \right\}$$

$$\xi = \xi_0 + \sum_{n=1}^m \sqrt{2} \sin(2\pi n t) \xi_{n,s} \sim \mathcal{N}(0, 1)$$

$$+ \sum_{n=1}^m \sqrt{2} \cos(2\pi n t) \xi_{n,c} \sim \mathcal{N}(0, 1)$$

$\xi^m$ : rand. el. of  $\mathbb{L}^2([0, 1], dx)$

$\xi^\infty$ : is not a random element of  $\mathbb{L}^2([0, 1], dx)$

$$\|\xi^\infty\|_{L^2} = \xi_0 + \sum_{n=1}^\infty \xi_{n,s}^2 + \xi_{n,c}^2 = \infty$$

$$\|\xi^\infty\|_{H^{-1}} \leq 1 + \sum_{n=1}^\infty \frac{\xi_{n,s}^2}{n^2} + \frac{\xi_{n,c}^2}{n^2} < \infty$$

$\hookrightarrow$  Sobolev spaces

$\xi^\infty$  rand. element of  $H^{-1}([0, 1], dx)$

$$H^1([0,1], dx) =$$

$$= \left\{ f: [0,1] \rightarrow \mathbb{R} \mid \int_0^1 |\nabla f|^2(x) dx < \infty \right\}$$

$$H^{-1}([0,1], dx)$$

$$= \left\{ f: [0,1] \rightarrow \mathbb{R} \mid \sup_g \frac{\left( \int_0^1 f g \right)^2}{\int_0^1 |\nabla g|^2 dx} < \infty \right\}$$

$$\xi^m = \sum_{n=1}^m \varphi_n \xi_n \in L^2(E, \mathcal{E}, \mu)$$

$$V = \left\{ f: \sum_n \frac{\langle f, \varphi_n \rangle^2}{n^2} < \infty \right\}$$

$$L^2(E, \mathcal{E}, \mu) \subset V$$

$$\xi^m \rightarrow \xi^\infty \in V$$

$$P[\xi^\infty \in L^2(E, \mathcal{E}, \mu)] = 0$$

$$P[\xi^\infty \in V] = 1$$

$$\underline{E} \subset E = [0, 1] \quad \mu(dx) = dx$$

$$\mathbb{L}^2(E, \mathcal{E}, \mu) = \mathbb{L}^2([0, 1], \mathcal{B}([0, 1]), dx)$$

$$\varphi_n: \left\{ 1, \underbrace{\sqrt{2} \sin(2\pi n t)}_{\varphi_{n,s}}, \underbrace{\sqrt{2} \cos(2\pi n t)}_{\varphi_{n,c}} \right\}$$

$$f = \langle f, \varphi_0 \rangle \varphi_0 + \sum_n \langle f, \varphi_{n,s} \rangle \varphi_{n,s} + \langle f, \varphi_{n,c} \rangle \varphi_{n,c}$$

$$G(f) = \langle f, \varphi_0 \rangle \xi_0 + \sum_n \langle f, \varphi_{n,s} \rangle \xi_{n,s} + \langle f, \varphi_{n,c} \rangle \xi_{n,c}$$

$\mathcal{N}(0, 1)$

$$G(f) \sim \mathcal{N}(0, \int_0^1 f^2(x) dx)$$

Haar Functions

$$\{\varphi_{j,n} \mid j=1, \dots, 2^{n-1}, n=1, 2, \dots\}$$

$\varphi_0(f) \equiv 1$  and for  $k \leq 2j - 1$

$$\varphi_{j,n} = \begin{cases} 2^{\frac{(n-1)/2}{2^n}} & \frac{k-1}{2^n} \leq f < \frac{k}{2^n} \\ -2^{\frac{(n-1)/2}{2^n}} & \frac{k}{2^n} \leq f < \frac{k+1}{2^n} \\ 0 & \text{else.} \end{cases}$$



$$f = \langle f, \varphi_0 \rangle \varphi_0 + \sum_{j \in \mathbb{N}} \langle f, \varphi_{j,n} \rangle \varphi_{j,n}$$

$$G(f) = \langle f, \varphi_0 \rangle \zeta_0 + \sum_{j \in \mathbb{N}} \langle f, \varphi_{j,n} \rangle \zeta_{j,n}$$

$\overset{2}{\mathcal{N}(\omega_1, I)}$        $\overset{2}{\mathcal{N}(\omega_1, I)}$   
ii d

$$G(f) \sim \mathcal{N}(0, \int_0^1 f^2(x) dx)$$

# Brownian Motion from a Gaussian Measure

$$(E, \mathcal{E}, \nu) = (\mathbb{R}_+, \mathcal{B}(\mathbb{R}_+), dx)$$

$(\varphi_n)_{n \in \mathbb{N}}$  : orthonormal basis of  
 $L^2(\mathbb{R}_+, \mathcal{B}(\mathbb{R}_+), dx)$   
 (Fourier, Haar, Wavelets)

$$G: L^2(\mathbb{R}_+, \mathcal{B}(\mathbb{R}_+), dx) \longrightarrow H$$

$$f = \sum_n \langle f, \varphi_n \rangle \varphi_n \longrightarrow G(f)$$

$$\sum_n \langle f, \varphi_n \rangle \xi_n$$

$\xi_n$  iid  
 $N(0, 1)$

Thm  $B_t = G(1_{[0, t]})$  is a 1d B.M.

Proof  $(B_t, t \in \mathbb{R}_+)$ : centered GP

because  $\forall \lambda_1, \dots, \lambda_n \in \mathbb{R}$

$\sum_i \lambda_i B_{t_i} = G\left[\sum_i \lambda_i \mathbf{1}_{[0, t_i]}\right]$  is Gaussian

$$\Gamma(s, t) = E[B_s B_t]$$

$$\begin{aligned} &= \int_{\mathbb{R}_+} \mathbf{1}_{[0, s]} \mathbf{1}_{[0, t]} dx \\ &= \min(s, t) \end{aligned}$$

$\Rightarrow B_t$  is a 1d BM.

Wiener integral with respect to a Brownian Motion

Def If  $B$  is a BM.

and  $G$  the associated Gaussian Measure

$$(B_t = G(\mathbf{1}_{[0, t]}))$$

We write

$$\int_0^\infty f(s) dB_s := G(f)$$

for  $\int_0^\infty f^2(x) dx < \infty$

$$\int_0^t f(s) dB_s := G(f 1_{[0,t]})$$

The mapping  $f \mapsto \int_0^\infty f(s) dB_s$  is called  
the Wiener integral w.r. to the BM. B.

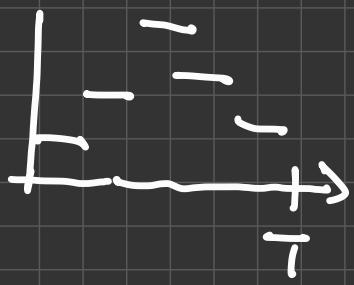
Rk If  $v < v'$

$$\begin{aligned} \int_v^{v'} dB_s &= G([v, v']) \\ &= B_{v'} - B_v \end{aligned}$$

Def A function  $f \in L^2(0, T)$  ( $\int_0^T f^2(x) dx < \infty$ )

is called a step function if  $\exists$  a partition  $0 = t_0 < t_1 < \dots < t_m = T$  s.t.

$$f(t) = p_k \quad \text{for } t_k \leq t < t_{k+1} \quad (k=0, \dots, m-1)$$



Prop If  $f \in L^2(0, T)$  is a step function  
then

$$\int_0^T f(s) dB_s = \sum_{k=0}^{m-1} f_k (B_{t_{k+1}} - B_{t_k})$$

Lemma If  $f \in L^2(0, T)$  then  $\exists$  a sequence  
of step functions  $f_n \in L^2(0, T)$

$$\int_0^T |f - f_n|^2(s) ds \xrightarrow[n \rightarrow \infty]{} 0$$

Moreover by isometry

$$\int_0^T f_n(s) dB_s \xrightarrow[n \rightarrow \infty]{L^2(\Omega)} \int_0^T f(s) dB_s$$

Lemma  $\forall a, b \in \mathbb{R}$ ,  $f, g \in L^2(0, T)$

$$\int_0^T (af + bg) dB_s = a \int_0^T f(s) dB_s + b \int_0^T g(s) dB_s$$

$$\mathbb{E} \left[ \int_0^T f(s) dB_s \right] = 0$$

$$\mathbb{E} \left[ \int_0^T f(s) dB_s \int_0^T g(s) dB_s \right] = \int_0^T f(s) g(s) ds$$

$$\int_0^T f(s) dB_s \sim \mathcal{N}(0, \int_0^T f^2(s) ds)$$

Wiener integral

(Itô integral: ( $f$  stochastic process)  
beyond ACM 118)

Stationary processes.

Def A stochastic process  $(X_t)_{t \in \mathbb{R}}$  is stationary if

$$\text{dist}((X_t)_{t \in \mathbb{R}}) = \text{dist}((X_{t+s})_{t \in \mathbb{R}}) \quad \forall s \in \mathbb{R}$$

$X_t$ : iid  $\mathcal{N}(0, 1)$

$$R(t, s) = \begin{cases} 1 & \text{if } t=s \\ 0 & \cdot \quad t \neq s \end{cases}$$

$(X_t)_{t \in \mathbb{R}}$ : GP(0, R)

$$R(s, t) = R(t-s)$$

X stationary

$$R(s, t) = \exp\left(-\frac{(t-s)^2}{2\sigma^2}\right)$$

$B_t^1, B_t^2$ : incdp lcl BN

$$B_t = \begin{cases} B_t^1 & t \geq 0 \\ B_{-t}^2 & t < 0 \end{cases}$$

$\varphi$ : continuous function on  $\mathbb{R}$  /

$$\varphi(t) \equiv 0 \text{ for } t \notin [-1, 1]$$

$$X_t = \int_{t-1}^{t+1} \varphi(s-t) dB_s$$

.  $X_t$  is a GP

$$\mathbb{E}[X_t] = 0$$

$$P(s, t) = \mathbb{E}[X_t X_s]$$

$$= \int_{\mathbb{R}} \varphi(v-t) \varphi(v-s) dv$$

$$= \int_{\mathbb{R}} \varphi(x) \varphi(x+t-s) dx$$

$$= \Gamma(\epsilon - s)$$

cor. fact invariant under time shifts

## Gaussian vectors in Euclidean Spaces

$E$ : Euclidean space of dimension  $d$   
( $E$  is isomorphic to  $\mathbb{R}^d$ )

$\langle u, v \rangle$ . or  $\langle u, v \rangle_E$ : scalar product of  $E$

Rk  $E$  endowed with scalar product  $\langle u, v \rangle$   
is also known as a finite-dimensional  
Hilbert space

Ex.  $E = \mathbb{R}^d$   $\langle u, v \rangle = u^\top v$   $u, v \in \mathbb{R}^d$

.  $E = \mathbb{R}^{n \times n}$   $\langle A, B \rangle = \sum_{i,j=1}^n A_{i,j} B_{i,j}$

. Let  $\varphi_1, \dots, \varphi_d$  be  $d$  linearly independent elements of

$$L^2((0, 1), dx) = \left\{ f: [0, 1] \rightarrow \mathbb{R} \mid \int_0^1 f^2(x) dx < \infty \right\}$$

$$E = \text{Span} \left\{ \varphi_1, \dots, \varphi_d \right\}$$

$$\langle \psi, \varphi \rangle = \int_0^1 \psi(x) \varphi(x) dx \quad \forall \psi \in E$$

Def  $X: \Omega \rightarrow E$  is a Gaussian vector  
if

$\langle v, X \rangle$  is a Gaussian r.v.  $\forall v \in E$

$$E_x \quad E = \text{Span} \left\{ \varphi_1, \dots, \varphi_d \right\} = \left\{ \varphi = \sum_{i=1}^d c_i \varphi_i \mid c \in \mathbb{R}^d \right\}$$

$$\langle \psi, \varphi \rangle = \int_0^1 \varphi(x) \psi(x) dx$$

$$X = \sum_{i=1}^d X_i \varphi_i \quad \text{random element of } E$$

$X_1, \dots, X_d$ : iid  $\mathcal{N}(0, 1)$

$\Rightarrow X$  is a Gaussian vector on  $E$

$$\Psi \in E$$
$$\langle \Psi, X \rangle = \sum_{i=1}^d X_i \underbrace{\int_0^1 \Psi(x) \varphi_i(x) dx}_{\text{is Gaussian}}$$

$\Downarrow$   
 $X$  is a Gaussian vector on  $E$

Only need  $(X_1, \dots, X_d)$  to be a Gaussian vector of  $\mathbb{R}^d$

—

Rk  $X$  Gaussian vector on  $E$

$e_1, \dots, e_d$ , basis of  $E$

$(\langle e_1, X \rangle, \dots, \langle e_d, X \rangle)$

$$\langle u, X \rangle \quad \forall u \in E$$

$$\langle u, X \rangle \sim \mathcal{N}(\langle u, m_x \rangle, \frac{?}{E})$$

Def A mapping  $B: E \times E \rightarrow \mathbb{R}$   
is called a symmetric bilinear form on  $E$

if  $\forall \lambda, u, v, w \in \mathbb{R} \times E \times E \times E$

$$B(\lambda u + w, v) = \lambda B(u, v) + B(w, v)$$

$$\& B(u, v) = B(v, u)$$

Ex  $E = \mathbb{R}^d$   $B(u, v) = u^T v$

$$B(u, v) = u^T A v$$

$A: d \times d$  symmetric matrix

$E_x$   $E_x = \text{Span} \{ \varphi_1, \dots, \varphi_d \}$

$\varphi_i : [0, 1] \rightarrow \mathbb{R} \quad \int_0^1 \varphi_i^2(x) dx < \infty$

$B(\varphi, \psi) = \int_0^1 \varphi(x) \psi(x) dx$

(\*)  $B(\varphi, \psi) = \int_0^1 \int_0^1 \varphi(x) G(x, y) \psi(y) dx dy$

$G$ : symmetric function of  $x$  &  $y$

$$(G(x, y) = G(y, x))$$

$E_x$  In previous ex all bilinear sym forms

are of the form \*

$$\Psi = \sum_{i=1}^d v_i \varphi_i \quad \varphi = \sum_{i=1}^d v_i \varphi_i$$

$$B(\varphi, \psi) = \sum_{i,j=1}^d v_i v_j B(\varphi_i, \varphi_j)$$
$$A_{i,j} = B(\varphi_i, \varphi_j)$$

A sym. matrix

$\phi_1, \dots, \phi_d$ : basis of  $E$  that is biorthogonal to  $\varphi_1, \dots, \varphi_d$

$$\langle \phi_i, \varphi_j \rangle = \delta_{i,j} \quad \forall i, j$$

$$B(\varphi_i, \varphi_j) = \int \int \int \varphi_i(x) \phi_i(x) B(\varphi_i, \varphi_j) \phi_j(y) \varphi_j(y) dx dy$$

$\Rightarrow \forall \psi, \varphi \in E$

$$B(\varphi, \psi) = \int \int \int \varphi(x) G(x, y) \psi(y) dx dy$$

$$G(x, y) = \sum_{i,j=1}^d \phi_i(x) B(\varphi_i, \varphi_j) \phi_j(y)$$

Def  $Q: E \rightarrow \mathbb{R}$  is a positive quadratic form on  $E$ , if  $\exists$  a symmetric bilinear form  $B$  on  $E$  /  $Q[u] = B[u, u]$

$$\& Q[u] \geq 0 \quad \forall u \in E$$

Rk Can you recover  $B$  from  $Q$ ? YES

$$B[u, v] = \frac{1}{2} [Q[u+v] - Q[u] - Q[v]]$$

$$\underline{\text{Ex}} \quad E = \mathbb{R}^d, \quad Q[u] = u^T u$$

$$Q[u] = u^T A u$$

$A$ : symmetric positive  
definite matrix

$$\underline{\text{Ex}} \quad E = \text{Span} \{ \varphi_1, \dots, \varphi_d \} \quad \varphi_i : [0, 1] \rightarrow \mathbb{R}$$

$$\int_0^1 \varphi_i(x) dx < \infty$$

$$Q[\varphi] = \int_0^1 \varphi_{(x)}^2 dx$$

$$Q[\varphi] = \int_0^1 \int_0^1 \varphi(x) \sigma(x, y) \varphi(y) dx dy$$

$G$ : symmetric /  $\int_0^1 \int_0^1 \varphi(x) \sigma(x, y) \varphi(y) dx dy \geq 0$

$\downarrow$   
positive kernel

If  $\nabla \varphi_i \in L^2([0, 1], dx)$

$$Q[\varphi] = \int_0^1 |\nabla \varphi|^2(x) dx$$

Prop If  $X$  is a Gaussian vector on  $E$ , then

$\exists m_X \in E$  and a positive quadratic form  
 $Q_X$  on  $E$  /  $\forall v \in E$

$$\mathbb{E}[\langle v, X \rangle] = \langle v, m_X \rangle$$

$$\text{Var}[\langle v, X \rangle] = Q_X[v]$$

$$E[\exp(i\langle v, X \rangle)] = \exp\left(i\langle v, m_X \rangle - \frac{1}{2} Q_X[v]\right)$$

$$\langle v, X \rangle \sim \mathcal{N}(\langle v, m_X \rangle, Q_X[v])$$

We say

$$X \sim \mathcal{N}(m_X, Q_X)$$

Proof  $(e_1, \dots, e_d)$  orthonormal basis of  $E$

$$X = \sum_{i=1}^d X_i e_i$$

$X_i = \langle e_i, X \rangle$  are Gaussian

$$m_X = \sum_{i=1}^d E[X_i] e_i$$

$$v = \sum_{i=1}^d v_i e_i$$

$$Q_X[v] = \text{Var}[\langle v, X \rangle]$$

$$= \sum_{i,j} v_i v_j \operatorname{Cor}(X_i, X_j)$$

$$\langle v, X \rangle \sim \mathcal{N}(\langle v, m_X \rangle, Q_X[v])$$

$$m_X^e = \begin{pmatrix} \mathbb{E}[\langle e_1, X \rangle] \\ \vdots \\ \mathbb{E}[\langle e_d, X \rangle] \end{pmatrix}$$

$C^e$ :  $d \times d$  matrix

$$C_{i,j}^e = \operatorname{Cov}(\langle e_i, X \rangle, \langle e_j, X \rangle)$$

$X$  is defined indep from basis

$m_X^e, C_X^e$ : depen on basis

$(f_1, \dots, f_d)$ : another orthonormal basis of  $E$

$$f_i = \sum_{j=1}^d A_{i,j} e_j$$

$A$ :  $d \times d$  orthonormal matrix ( $A^T A = I_d$ )

$$X^f = \begin{pmatrix} \langle f_1, X \rangle \\ \vdots \\ \langle f_d, X \rangle \end{pmatrix}$$

$C^f$ :  $d \times d$  matrix

$$C_{i,j}^f = \text{Cov} [\langle f_i, X \rangle, \langle f_j, X \rangle]$$

$$m_x^f = A m_x^e$$

$$C^f = A C^e A^T$$

$P_{\text{rep}}$   $(e_1, \dots, e_d)$  orthonormal basis of  $E$

$$X = \sum_{i=1}^d X_i e_i$$

$$X \sim \mathcal{N}(m_X, Q_X)$$

Then the r.v.  $X_1, \dots, X_d$  are indep.

if  $(\text{Cov}(X_i, X_j))_{i,j}$  is diagonal

$\Leftrightarrow Q_X$  is diagonal in basis  $e_1, \dots, e_d$

$$Q_x[v] = B[v, v]$$

$$B[e_i, e_j] = \lambda_i \delta_{i,j}$$

Def  $\gamma$  is a symmetric positive endomorphism on  $E$  iff  $\gamma$  is a linear function mapping  $E \rightarrow E$  s.t.  $\forall u, v \in E$

$$\langle u, \gamma[v] \rangle = \langle \gamma[u], v \rangle$$

$$\langle u, \gamma[u] \rangle \geq 0$$

Rk  $Q_x$  is a quadratic form iff  
 $\exists$  a symmetric positive endomorphism associated to  $Q_x$  s.t.  $\forall u \in E$

$$Q_x[v] = \langle u, \gamma_x[u] \rangle$$

The matrix of  $\gamma_x$  in the basis  $(e_1, \dots, e_d)$   
is  $(\text{cov}(X_i, X_j))_{i,j}$   $X_i = \langle x, e_i \rangle$

$$\langle v, \gamma_x(v) \rangle = \frac{1}{4} [Q_x(v+v) - Q_x(v-v)]$$

Thm If holds true that

1. If  $\gamma$  is a symmetric positive endomorphism on  $E$ , then there exists a centered Gaussian vector  $X$  s.t.  $\gamma_X = \gamma$

2. Let  $X$  be a centered Gaussian vector   
 $(\varepsilon_1, \dots, \varepsilon_r)$  a basis of  $E$  diagonalizing

$\gamma_X$

$$\gamma_X \varepsilon_j = \lambda_j \varepsilon_j \quad \lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_r > 0 = \lambda_{r+1} \dots = \lambda_d = 0$$

Then

$$X = \sum_{j=1}^r Y_j \varepsilon_j$$

$Y_j$  indep  $\mathcal{N}(0, \lambda_j)$  r.v.

topological support of  $X$  is span  $\{\varepsilon_1, \dots, \varepsilon_r\}$

Thm Let  $X \sim \mathcal{N}(m_x, Q_x)$

$v_1, \dots, v_m \in E$

$X | \langle v_1, X \rangle, \dots, \langle v_m, X \rangle \sim \mathcal{N}(m_x^\perp, Q_x^\perp)$

$$m_x^\perp = \mathbb{E}[X | \langle v_1, X \rangle, \dots, \langle v_m, X \rangle]$$

$$= m_x + \sum_{i,j=1}^m \langle v_i, X - m_x \rangle \Theta_{i,j}^{-1} \nabla_X v_j$$
$$\Theta_{i,j} = \langle v_i, \nabla_X v_j \rangle$$

$$Q_x^\perp[u] = \min_{v \in \text{Span}\{v_1, \dots, v_m\}} Q_x[v - u]$$

$$Q_x^\perp[u] = \langle u, \nabla_X u \rangle$$

$$- \sum_{i,j} \langle u, \nabla_X v_i \rangle \Theta_{i,j}^{-1} \langle u, \nabla_X v_j \rangle$$

$$= \langle u, \nabla_X^\perp u \rangle$$

$$\mathcal{D}_x^\perp = \mathcal{D}_x - \sum_{i,j} \Theta_i j^{-1} \mathcal{D}_x v_i \otimes \mathcal{D}_x v_j$$

$$\mathcal{D}_x v_i \otimes \mathcal{D}_x v_j : E \rightarrow E$$

$$v \mapsto \mathcal{D}_x v_i \langle v, \mathcal{D}_x v_j \rangle$$

Proof  $v \in E$

$$\text{Find } \langle v, x \rangle / \langle v_1, x \rangle, \dots, \langle v_m, x \rangle$$

$$\langle v, x \rangle, \langle v_1, x \rangle, \dots, \langle v_m, x \rangle$$

Gaussian vector of  $\mathbb{R}^{m+1}$

# Gaussian measures on a Hilbert space

$H$ : separable Hilbert space

$\langle \cdot, \cdot \rangle_H$ : inner product on  $H$

We say a linear operator

$$T: H \rightarrow H$$

is symmetric and positive w.r. to  $\langle \cdot, \cdot \rangle_H$

$$\text{if } \begin{cases} \langle \varphi, T\phi \rangle_H = \langle \phi, T\varphi \rangle_H & \forall \varphi, \phi \in H \\ \langle \phi, T\phi \rangle_H \geq 0 & \forall \phi \in H \end{cases}$$

Trace of  $T$

$$\text{tr}[T] = \sum_{i=1}^{\infty} \langle T e_i, e_i \rangle_H \quad (\text{incep from } (e_i)_{i \in \mathbb{N}})$$

$(e_i)_{i \in \mathbb{N}^*}$ : any orthonormal basis of  $H$

Def  $T$  is trace class  $\Leftrightarrow \text{tr}[T] < \infty$

## Gaussian measure

Def  $\mathcal{P}(X)$ : set of Borel prob measures  
on a topological space  $X$ , and, when  
 $X$  is a Banach space, we let  
 $\mathcal{P}_2(X)$  denote those with finite  
second moments

$$(E[\|X\|^2] < \infty)$$

Def A Borel prob. measure  $\mu \in \mathcal{P}(H)$   
on the Hilbert space  $H$  is said to be  
a Gaussian measure if the random vector  
 $X$  of  $H$  defined by  $\mu$  is such that  
 $\langle \phi, X \rangle_H$  is a Gaussian rand. var  
 $\forall \phi \in H$

Thm If  $\nu \in \mathcal{S}(H)$  is a Gaussian measure  
then  $\nu \in \mathcal{S}_2(H)$ , i.e.

$$E_{X \sim \nu} [\|X\|^2] < \infty$$

Furthermore,  $\exists m \in H$  and a  
linear symmetric positive trace class operator

$$T: H \rightarrow H$$

s.t. the random vector  $X$  of  $H$  defined  
by  $\nu$  is s.t.

$$\langle \phi, X \rangle_H \sim \mathcal{N}(\langle \phi, m \rangle_H, \langle \phi, T\phi \rangle_H) \quad \forall \phi \in H$$

In particular

$$E_{X \sim \nu} [\langle \phi_1, X_m \rangle_H \langle \phi_2, X_m \rangle_H] = \langle \phi_1, T\phi_2 \rangle_H \quad \forall \phi_1, \phi_2 \in H$$

We say that  $\nu \in \mathcal{S}(H)$  is a Gaussian measure with mean  $m$  & covariance operator  $T$

$$\times_n \mathcal{N}(m, T)$$

Thm Given  $m \in H$  and a linear symmetric positive trace class operator  $T: H \rightarrow H$   
 $\exists$  a Gaussian measure  $\nu \in \mathcal{S}(H)$  with mean  $m$  and covariance operator  $T$

Proof Let  $e_1, e_2, \dots$  be an orthonormal basis of  $H$

diagonalizing  $T$

$$Te_i = \lambda_i e_i$$

$T$  positive, trace class  $\Rightarrow \lambda_i \geq 0$

$$\sum_i \lambda_i < \infty$$

$Z_1, Z_2, \dots$  iid  $\mathcal{N}(0, 1)$  r.v.

$$X := m + \sum_i \sqrt{\lambda_i} Z_i e_i$$

$$X \sim \mathcal{N}(m, T)$$

$T$  trace class  $\Rightarrow X$  is a random element of  $H$

## Gaussian fields on Hilbert Space

$H$ : separable Hilbert space

$\langle \cdot, \cdot \rangle_H$ : inner product

$\|\cdot\|_H$ : norm

Def.) A centered Gaussian field  $\xi$  on  $H$  is an isometry mapping  $H$  to a centered Gaussian space

$\xi: H \rightarrow$  Gaussian Space (centered)  
 $\phi \mapsto \xi(\phi) \sim \mathcal{N}(0, \|\phi\|_H^2)$

$$\text{Cov}(\xi(\phi), \xi(\varphi)) = \langle \phi, \varphi \rangle_H$$

Thm Let  $\phi_1, \phi_2, \dots$  be an orthonormal basis of  $H$   
 $Z_1, Z_2, \dots$  iid  $\mathcal{N}(0, 1)$

Then  $\xi: H \rightarrow \text{Gaussian space}$

$$\begin{aligned} \phi &\mapsto \xi(\phi) \\ \sum_{i=1}^{\infty} c_i \phi_i &\mapsto \sum_{i=1}^{\infty} c_i Z_i \end{aligned}$$

is a Gaussian field in the sense of Def \*

Proof  $E[\xi(\phi)^2] = \sum_i c_i^2 = \|\phi\|_H^2$

Rk Formally  $\xi = \sum_i Z_i \phi_i$

$$\phi = \sum_i c_i \phi_i$$

$$\xi(\phi) = \langle \xi, \phi \rangle_H$$

If  $\dim(\mathcal{H}) = \infty$  then  $\xi$  is not a rand. element of  $\mathcal{H}$

$$\|\xi\|_{\mathcal{H}}^2 = \sum_i z_i^2 = \infty \text{ a.s.}$$

However

$$V = \left\{ \sum_i c_i \phi_i \mid c_i \in \mathbb{R} \text{ and } \sum_i \frac{|c_i|^2}{i^2} < \infty \right\}$$

$$\|\phi\|_V^2 = \sum_i \frac{|c_i|^2}{i^2} \quad \text{for } \phi = \sum_i c_i \phi_i$$

then  $\xi$  is a.s. an element of  $V$  ( $V \supset \mathcal{H}$ )

$$\|\xi\|_V^2 = \sum_i \frac{z_i^2}{i^2} < \infty \text{ a.s.}$$

$\xi$ : Gaussian measure on  $V$

Ex  $\mathcal{H} = L^2(\Omega)$   
 $\Omega \subset \mathbb{R}^d$

$$\|\varphi\|_{H^2}^2 = \int_{\Omega} \varphi^2(x) dx$$

$\xi$  is not a random element of  $L^2(\Omega)$

but of  $H^{-s}(\Omega)$  for  $s > \frac{d}{2}$

$\phi_i$ : eigenfunctions of  $(-\Delta)$  on  $H_0^1(\Omega)$

$$-\Delta \phi_i = \lambda_i \phi_i$$

Weyl's estimate:  $\lambda_i \propto i^{2/d}$

$$H^{-s}(\Omega) := \left\{ \sum_k c_k \phi_k \mid \sum_k \frac{c_k^2}{\lambda_k^s} < \infty \right\}$$

$$\Rightarrow \|\xi\|_{H^{-s}(\Omega)} < \infty \text{ a.s.}$$

$\xi$  is a Gaussian measure on  $H^{-s}(\Omega)$

Gaussian measures and fields on Banach spaces

Gaussian fields:

$\mathcal{B}$ : Banach space (separable)

$\mathcal{B}^*$ : dual space (space of linear functionals  
on  $\mathcal{B}$ )

$[\cdot, \cdot]$ : dual pairing between  $\mathcal{B}^*$  and  $\mathcal{B}$

$\phi \in \mathcal{B}^*, v \in \mathcal{B}, [\phi, v]$

Let  $T: \mathcal{B}^* \rightarrow \mathcal{B}$

be a symmetric, positive, not necessarily injective, linear operator

.  $[\varphi, T\phi] = [\phi, T\varphi] \quad \forall \varphi, \phi \in \mathcal{B}^*$

.  $[\phi, T\phi] > 0 \quad \forall \phi \in \mathcal{B}^*$

Def  $\xi$  is a Gaussian field on  $B$   
with mean  $v \in B$  and covariance operator  $T$

$$\xi \sim N(v, T)$$

if  $\xi: B^* \rightarrow \text{Gaussian space}$  s.t.

$$\xi(\phi) := [\phi, \xi] \sim N([\phi, v], [\phi, T\phi]) \quad \forall \phi \in B^*$$

Ex  $B = \mathbb{R}^N$

$$\|x\|^2 = x^T A x$$

$A: N \times N$  positive definite symmetric

$$[x, y] = x^T y$$

$B^* = \mathbb{R}^N$

$$\|x\|_*^2 = \sup_{y \in B} \left( \frac{[x, y]}{\|y\|} \right)^2 = x^T A^{-1} x$$

$E$   $\times$   $D$  bounded domain of  $\mathbb{R}^d$

$$B = \{f: D \rightarrow \mathbb{R} \mid \int_D f^2(x) dx < \infty\}$$

$$\|f\|^2 = \int_D f^2(x) dx \quad [f, g] = \int_D f g$$

$$B^* = B \quad \|\cdot\|_* = \|\cdot\|$$

$E$   $B = H_0'(D)$

$$\|f\|^2 = \int_D |\nabla f|^2(x) dx$$

$$[f, g] = \int_D f g$$

$B$ : closure  $C_0^\infty(D)$  w.r.t.  $\|\cdot\|$

$$B^* = H^{-1}(D)$$

$$\|\phi\|_*^2 = \int_D \phi(x) G(xy) \phi(y) dx dy$$

G: Green's function of  $(-\Delta)$

s.l. of  $\begin{cases} -\Delta u = \phi & \text{on } D \\ u=0 & \text{at } \partial D \end{cases}$

is  $u(x) = \int_D G(x,y) \phi(y) dy$

$B$   $B = H_0^1(0,1)$

$$= \left\{ f: (0,1) \rightarrow \mathbb{R} \mid \int_0^1 |\nabla f|^2(x) dx < \infty \right. \\ \left. f(0) = f(1) = 0 \right\}$$

$$[\phi, v] = \int_0^1 \phi(x) v(x) dx$$

$$B^* = H^{-1}(0,1)$$

$$[\varphi, v] = \int_0^1 \nabla \varphi \nabla v(x) dx$$

$$B^* = H_0^1(\Omega)$$

$$\phi \in H^{-1}(\Omega) \quad v \mapsto \int_0^1 \phi(x) v(x) dx$$

Linear functional on  $B$

$$p \in H_0^1(\Omega) \quad v \mapsto \int_0^1 \nabla p(x) \nabla v(x) dx$$

Linear functional on  $B$

$H^{-1}(\Omega)$ : Sobolev space

$$\{ \phi \mid \left| \int_0^1 \phi(x) v(x) dx \right| \leq C \| v \|_{H_0^1} \}$$

$$H^{-s}(\Omega) = \{ \phi \mid \left| \int_0^1 \phi(x) v(x) dx \right| \leq C \| v \|_{H_0^s(\Omega)} \}$$

$\mathcal{S} \sim \mathcal{N}(v, T)$

Let  $\phi_1, \dots, \phi_m \in \mathcal{B}^*$

Write  $[\phi, \zeta] = ([\phi_1, \zeta], \dots, [\phi_m, \zeta])$

Ex  $[\phi, \zeta] \sim \mathcal{N}([\phi_v], [\phi_T \phi])$

$[\phi_T \phi]$ :  $m \times m$  matrix

$$[\phi_T \phi]_{i,j} = [\phi_i, T \phi_j]$$

Assume  $[\phi_T \phi]$  invertible

Thm  $E[\zeta | [\phi, \zeta]]$

$$= v + \sum_{i,j=1}^m T \phi_i \Theta_{i,j}^{-1} [\phi_j, \zeta - v]$$
$$\Theta_{i,j} = [\phi_i, T \phi_j]$$

$$\xi[\phi, \zeta] = \mathcal{N}\left(\mathbb{E}[\zeta | [\phi, \zeta]], \phi^\perp(\tau)\right)$$

Def The short  $\phi^\perp(\tau) : \mathcal{B}^* \rightarrow \mathcal{B}$

of the symmetric positive operator  $T : \mathcal{B}^* \rightarrow \mathcal{B}$   
to the annihilator  $\phi^\perp = \{v \in \mathcal{B} \mid [\phi_i, v] = 0 \forall i\}$

is the symmetric positive operator defined

by

$$[\psi, \phi^\perp(\tau)\psi] = \inf_{\phi \in \text{Span}\{\phi_1, \dots, \phi_m\}} [(\psi - \phi), T(\psi - \phi)]$$

Proof  $\psi \in \mathcal{B}^*$

$$[\psi, \zeta] / [\phi^1, \zeta], \dots, [\phi^m, \zeta]$$

$$\psi \in \text{Span}(\phi^1, \dots, \phi^m)$$

$$[\psi, \xi] \sim \mathcal{N}(\mathbb{E}[\psi, \xi] / [\phi^1, \xi], \dots, [\phi^m, \xi]),$$

" "

$$[\psi, \phi^+(T) \psi]$$

$$\inf_{\phi \in \text{Span}\{\phi_1, \dots, \phi_m\}} [\psi - \phi, T(\psi - \phi)] = 0$$

$([\psi, \xi], [\phi^1, \xi], \dots, [\phi^m, \xi])$

$\rightarrow$  Gaussian vector of  $\mathbb{R}^{m+1}$

## Canonical Gaussian Field

$(\mathcal{B}, \|\cdot\|)$ : Separable Banach Space

$$\|v\|^2 = [Q^{-1}v, v]$$

$Q: \mathcal{B}^* \rightarrow \mathcal{B}$

continuous, symmetric, positive  
and invertible

$$[Q^{-1}v, v] = \|v\|^2$$

$$[\phi, Q\phi] = \|\phi\|_*^2 = \sup_{v \in \mathcal{B}} \frac{([\phi, v])^2}{\|v\|^2}$$

$$[Q^{-1}v, v] = \langle v, v \rangle$$

$$[\phi, Q\varphi] = \langle \phi, \varphi \rangle_*$$

$$Q: \mathcal{B}^* \rightarrow \mathcal{B}$$

$$[\phi, Q\varphi] = \frac{1}{4} \left[ [\phi + \varphi, Q(\phi + \varphi)] - [(\phi - \varphi), Q(\phi - \varphi)] \right]$$

Canonical (  $\overline{\text{Gaussian Field}}$  )

$$\xi \sim \mathcal{N}(0, Q)$$

$\xi$ : Linear isometry

$$(\mathcal{B}^*, \|\cdot\|_*) \rightarrow \text{Gaussian space}$$

$$[\phi, \xi] \sim \mathcal{N}(0, \|\phi\|_*^2)$$

$$\mathbb{E}[[\mathcal{Y}\xi][\phi, \xi]] = \langle \mathcal{Y}\phi \rangle_*$$

$$[\mathcal{Y}, Q \phi]$$

Ex  $\mathcal{B} = \mathbb{R}^N$

$$\|\mathbf{x}\|^2 = \mathbf{x}^\top A \mathbf{x}$$

$A$ :  $N \times N$  positive definite matrix

$$[\mathbf{x}, \mathbf{y}] = \mathbf{x}^\top \mathbf{y}$$

$$\mathcal{B}^* = \mathbb{R}^N$$

$$\|x\|_F^2 = \sup_{y \in \mathcal{B}} \left( \frac{[x, y]}{\|y\|} \right)^2 = x^\top A^{-1} x$$

$$Q = A^{-1}$$

$$\xi \sim \mathcal{N}(0, Q)$$

$$\phi \in \mathcal{B}^* = \mathbb{R}^N$$

$$[\phi, \xi] = \phi^\top \xi \sim \mathcal{N}(0, \phi^\top A^{-1} \phi)$$

$\xi$ : Gaussian vector  
mean 0 and cor matrix  $A^{-1}$

$$f: p d f(\xi)$$

$$f(x) \propto \exp\left(-\frac{x^\top A x}{2}\right)$$

$$\propto \exp\left(-\frac{\|x\|^2}{2}\right)$$

$\phi_1, \phi_2, \dots, \phi_m \in \mathcal{B}^*$  (linear indep.)

$$[\phi, v] = ([\phi_1, v], \dots, [\phi_m, v])$$
$$v \in \mathcal{B}$$

Prop

$$\xi | [\phi, \xi] \sim \mathcal{N}(E[\xi | [\phi, \xi]], \phi^\perp(Q))$$

$$E[\xi | [\phi, \xi]] = \sum_{i,j} [\phi_i, \xi] \Theta_{i,j}^{-1} Q \phi_j$$
$$\Theta_{i,j} = [\phi_i, Q \phi_j]$$

$$\phi^\perp = \{ v \in \mathcal{B} \mid [\phi, v] = 0 \}$$

Short  $\phi^\perp(Q) : \mathcal{B}^* \rightarrow \mathcal{B}$

$$[\varphi, \phi^\perp(Q)\varphi] = \inf_{\phi \in \text{Span}\{\phi_1, \dots, \phi_m\}} [(\varphi - \phi), Q(\varphi - \phi)]$$

$$= \inf_{\phi \in \text{Span}\{\phi_1, \dots, \phi_m\}} \|\varphi - \phi\|_*^2$$

Prop  $E[\varepsilon | [\phi, \varepsilon] = y]$   $y = \begin{pmatrix} y_1 \\ \vdots \\ y_m \end{pmatrix}$

is the minimizer of

$$\begin{cases} \min \|v\| \\ \text{s.t. } v \in \mathcal{B} \text{ & } [\phi, v] = y \end{cases}$$

$$\phi^\perp(Q) = Q - \sum_{i,j} \theta_{i,j}^{-1} Q \phi_i \otimes Q \phi_j$$

for  $\varphi \in \mathcal{B}^*$

$$(Q \phi_i \otimes Q \phi_j)(\varphi) = [\varphi, Q \phi_j] Q \phi_i$$

$$Q \phi_i \otimes Q \phi_j : \mathcal{B}^* \rightarrow \mathcal{B}$$

$$\varphi \rightarrow [\varphi, Q \phi_i] Q \phi_i$$

$\mathcal{E}$   $\mathcal{B} =$  Space of functions on  $\mathbb{R}$

$\mathcal{B}^*$  = Space of generalized functions  
(distributions) on  $\mathbb{R}$

For  $\phi \in \mathcal{B}^*$

$$\int_{\mathbb{R}} \phi(x) v(x) dx$$

Assume that  $\mathcal{B} \subset C^1(\mathbb{R})$

Then  $[\phi, v] = \int_{\mathbb{R}} \phi(x) v(x) dx$

$$\forall x \in \mathbb{R}, \delta_x \in \mathcal{B}^*$$

$$\delta_x \circ \delta_x \in \mathcal{B}^*$$

$$[\delta_x, v] = v(x)$$

$$[\delta_x \circ \delta_x, v] = \delta_x v(x)$$

For  $\phi \in \mathcal{B}^*$

$$[\phi, Q\phi] = \int_{\mathbb{R}^2} \phi(x) K(x, y) \phi(y) dx dy$$

$$K: \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}$$

$$\delta_x \in \mathcal{B}^*$$

$$K(x, y) = [\delta_x, Q\delta_y]$$

$$\xi \sim \mathcal{N}(0, Q) \Leftrightarrow \xi \sim \mathcal{N}(0, K)$$

↓  
cor operator      ↓  
                        cor function

—

$$X_1, \dots, X_N \in \mathbb{R}$$

What is

$$v(x) = E[\xi(x) \mid \xi(X_1) = Y_1, \dots, \xi(X_N) = Y_N \\ \partial_x \xi(X_1) = Z_1, \dots, \partial_x \xi(X_N) = Z_N]$$



$$v(x) = E[\phi(x, \xi) \mid [\phi_1, \xi] = Y_1, \dots, [\phi_N, \xi] = Y_N \\ [\phi_{N+1}, \xi] = Z_1, \dots, [\phi_{2N}, \xi] = Z_N]$$

$$\phi_1 = \delta_{X_1}, \dots, \phi_N = \delta_{X_N}$$

$$\phi_{N+1} = \delta_{X_1} \circ \partial_x, \dots, \phi_{2N} = \delta_{X_N} \circ \partial_x$$

$$v(x) = K(x, \phi) K(\phi, \phi)^{-1} \begin{pmatrix} Y \\ Z \end{pmatrix}$$

$K(x, \phi)$ :  $2N$  vec for with entries

$$K(x, \phi_i) = \int K(x, y) \phi_i(y) dy$$

$K(\phi, \phi) : 2N \times 2N$  ma trix  
with en frien

$$K(\phi_i, \phi_j) := \int \phi_i(x) K(x, y) \phi_j(y) dx dy$$

$$\begin{pmatrix} y \\ z \end{pmatrix} = \begin{pmatrix} y_1 \\ \vdots \\ y_N \\ z_1 \\ \vdots \\ z_N \end{pmatrix}$$

Gptima (Recovery)

$$\xi \sim \mathcal{N}(0, Q)$$

$$\phi_1, \dots, \phi_m \in \mathcal{B}^*$$

Player I

$$v \in \mathcal{B}$$

Player II

$$\text{see } [\phi, v] = ([\phi_1, v], \dots, [\phi_m, v])$$

choose  $\Psi: \mathbb{R}^m \rightarrow \mathcal{B}$

max

min

$$\mathcal{E}(v, \Psi) = \frac{\|v - \Psi([\phi, v])\|}{\|v\|}$$

$$\|v\|^2 = [Q^{-1}v, v]$$

$$\sup_{v \in \mathcal{B}} \inf_{\Psi} \mathcal{E}(v, \Psi)$$

$$m < \dim(\mathcal{B})$$

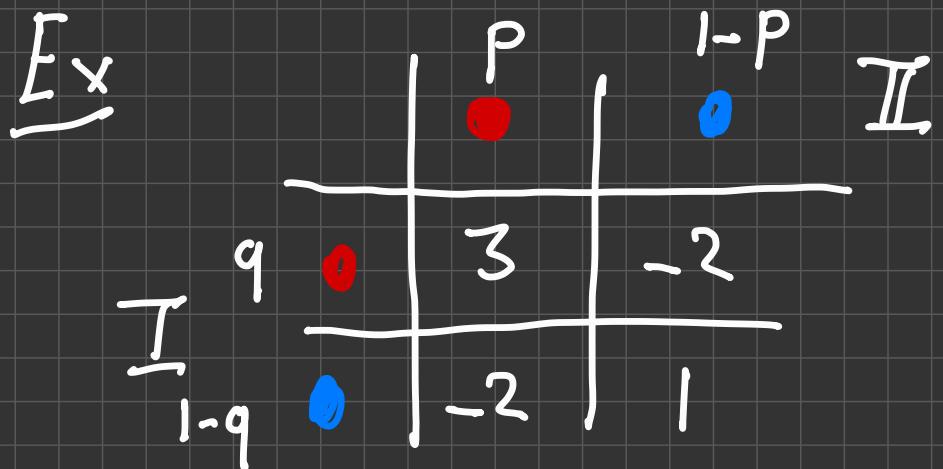
$$\inf_{\Psi} \sup_{v \in \mathcal{B}} \mathcal{E}(v, \Psi) \neq 0$$

Thm  $E[\xi | [\phi, \xi] = [\phi, v]]$

is the minimizer of

$$\min_{v \text{ s.t. } [f, v] = [\phi, v]} \max_{v \in \mathcal{B}} \frac{\|v - v\|}{\|v\|^2}$$

$$\Psi[[\phi, v]] = \sum_{i,j} [\phi_i, v] \Theta_{i,j}^{-1} Q \phi_j$$
$$\Theta_{i,j} = [\phi_i, Q \phi_j]$$



Gain of Player II  
Loss of Player I

Average Gain of Player II

$$3qp + |(1-p)(1-q) - 2(1-q)p - 2(1-p)q| \\ = 1 + p(8q - 3) - 3q$$

$$q = \frac{3}{8} \quad \text{Av. Gain Player II} = 1 + px^0 - \frac{q}{8} \\ = -\frac{1}{8}$$

The lifted game

$\mathcal{P}_2(\mathcal{B})$ : set of prob. measures on  $\mathcal{B}$   
with finite second moments

I

$\nu \in \mathcal{P}_2(\mathcal{B})$

II

$\psi: \mathbb{R}^m \rightarrow \mathbb{R}$

$$\mathcal{E}(\nu, \psi) = \frac{\mathbb{E}_{v \sim \nu} [\|v - \psi[\phi_v]\|^2]}{\mathbb{E}_{v \sim \nu} [ \|v\|^2]}$$

Thm If  $\dim(\text{span}\{\phi_1, \dots, \phi_m\}) < \dim(B)$



$$\inf_{\psi} \sup_{\nu \in \mathcal{P}_2(B)} \mathcal{E}(\nu, \psi) = \sup_{\nu} \inf_{\psi} \mathcal{E}(\nu, \psi) =$$

Optimal strategy for Player I

$$\xi - E[\xi | \phi, \xi] \sim \mathcal{N}(0, \Phi^\perp(Q))$$

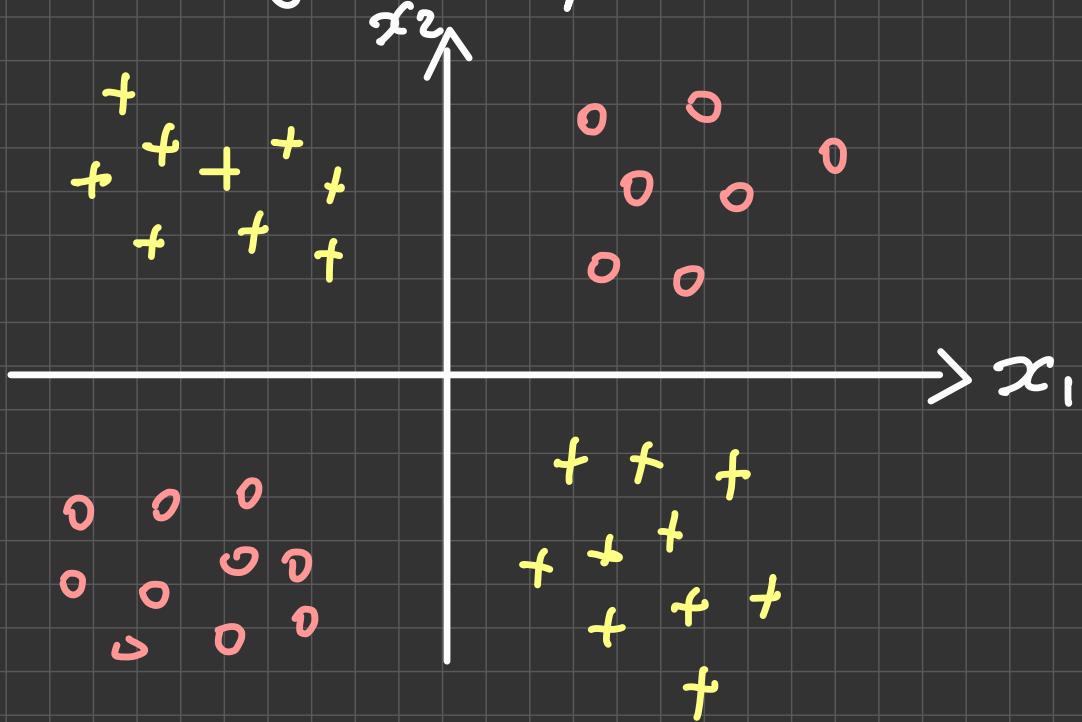
$$\xi \sim \mathcal{N}(0, Q)$$

Optimal strategy of Player II is

$$\psi([\phi, v]) = E[\xi | \phi, \xi = [\phi, v]]$$

# Kernel Methods

Motivating example : XOR



$N$  cluster points  $(x_i, y_i)_{1 \leq i \leq N}$

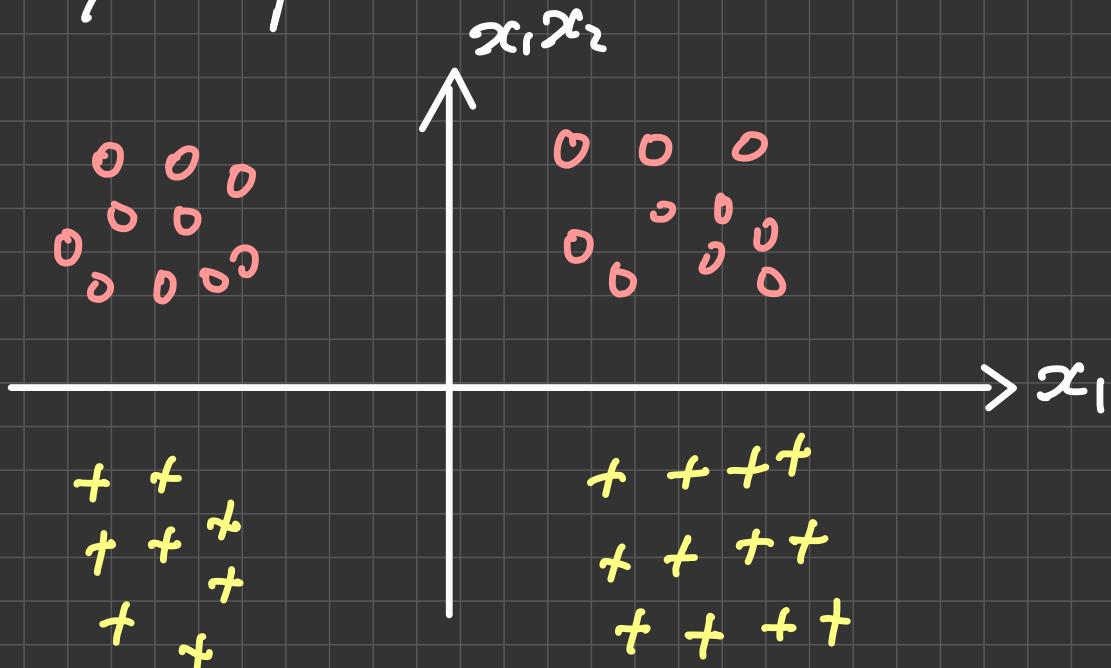
$$x_i \in \mathbb{R}^2, y_i \in \{0, +\}$$

are not linearly separable

idea Consider instead cluster points  
 $(\phi(x_i), y_i)_{1 \leq i \leq N}$

$$\text{for } \mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \quad \phi(\mathbf{x}) = \begin{pmatrix} x_1 \\ x_1 x_2 \end{pmatrix}$$

These lifted  $N$  cluster points are linearly separable



Construction of kernels

Def (Inner product)

Let  $H$  be a vector space over  $\mathbb{R}$   
 A function  $\langle \cdot, \cdot \rangle_H : H \times H \rightarrow \mathbb{R}$

is said to be an inner product on  $H$  if

$$1. \langle \alpha_1 f_1 + \alpha_2 f_2, g \rangle_H = \alpha_1 \langle f_1, g \rangle_H + \alpha_2 \langle f_2, g \rangle_H$$

$$2. \langle f, g \rangle_H = \langle g, f \rangle_H$$

$$3. \langle f-f, f-f \rangle_H = 0$$
$$\langle f-f, f-f \rangle_H = 0 \Rightarrow f = 0$$

norm

$$\|f\|_H = \sqrt{\langle f, f \rangle_H}$$

Hilbert space:  $(H, \langle \cdot, \cdot \rangle_H)$  s.t.

$H$  is complete (contains limit of Cauchy sequences)

Def  $\mathcal{X}$ : non-empty set

A function  $k: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$

is called a kernel if  $\exists$  a  $\mathbb{R}$ -Hilbert space

and a map  $\phi: \mathcal{X} \rightarrow H$  s.t.

$\forall x, x' \in \mathcal{X}$

$$k(x, x') = \langle \phi(x), \phi(x') \rangle_H$$

Rk  $\phi$  called Feature map

$H$ : Feature space

Rk No conditions on  $\mathcal{X}$

ex  $\mathcal{X} = \{ \text{text documents} \}_d$

$H = \mathbb{R}^d$

$$\phi(x) = (\phi_1(x), \dots, \phi_d(x))$$

$\phi_i(x)$ : # of times word  $i$  appears in doc.  $x$

$\{l, \dots, c\} = \{cat, dog, kennel, fury, \dots\}$

Ex Given  $k$  is  $\phi$  unique?

$$k(x, x') = x \cdot x'$$

$$\mathcal{X} = \mathbb{R} \quad \phi(x) = x \quad \mathcal{H} = \mathbb{R}$$

$$\phi = \begin{pmatrix} x \\ 0 \end{pmatrix} \quad \mathcal{H} = \mathbb{R}^2$$

$$\phi = \begin{pmatrix} x/\sqrt{2} \\ x/\sqrt{2} \end{pmatrix} \quad \mathcal{H} = \mathbb{R}^2$$

$$\begin{aligned} k(x, x') &= \phi(x)^T \phi(x') \\ &= \Phi_{(x)}^T A^T A \Phi_{(x')} \quad \text{if } A^T A = I \end{aligned}$$

$$\mathcal{H} = \overbrace{\text{Span} \left\{ x \rightarrow k(x, y) \right\}}$$

Lemma (Sums of kernels are kernels)

Given  $\alpha > 0$  and  $k, k_1$  and  $k_2$  all kernels on  $X$

$\alpha k$  and  $k_1 + k_2$  are kernels on  $X$

Proof  $k \leftrightarrow \phi, H$

$\alpha k \leftrightarrow \sqrt{\alpha} \phi, H$

$k_1 \leftrightarrow \phi_1, H_1$       |  $\Rightarrow$   
 $k_2 \leftrightarrow \phi_2, H_2$

$k_1 + k_2 \leftrightarrow \phi = (\phi_1, \phi_2)$        $H = H_1 \times H_2$

$(k_1 + k_2)(x, x') = \langle \phi(x), \phi(x') \rangle_H$

$$= \langle \phi_1(x), \phi_1(x') \rangle_{H_1} \\ + \langle \phi_2(x), \phi_2(x') \rangle_{H_2}$$

Rk Difference of kernels may not be kernels

Lemma  $\mathcal{X}, \tilde{\mathcal{X}}$ : sets  
 $A: \mathcal{X} \rightarrow \tilde{\mathcal{X}}$

Let  $k$  be a kernel on  $\tilde{\mathcal{X}}$

Then  $K(x, x') = k(A(x), A(x'))$   
 is a kernel on  $\mathcal{X}$

Proof  $k \leftrightarrow \phi, H$   
 $K \leftrightarrow \phi \circ A, H$

Warping  
kernels

$$K(x, x') = \langle \psi \circ A(x), \psi \circ A(x') \rangle_H$$

Lemma Product of kernels are kernels

Let  $k_1$  be a kernel on  $\mathcal{X}_1$

$$k_2 \quad \text{on} \quad \mathcal{X}_2$$

Then

(i)  $K((x_1, x_2), (x'_1, x'_2)) = k_1(x_1, x'_1) k_2(x_2, x'_2)$   
 is a kernel on  $\mathcal{X}_1 \times \mathcal{X}_2$

(ii) If  $\mathcal{X}_1 = \mathcal{X}_2 = \mathcal{X}$

$\Gamma(x, x') = k_1(x, x') k_2(x, x')$   
 is a kernel on  $\mathcal{X}$

proof Only need to prove (ii)  
since (ii)  $\Rightarrow$  (i)

(ii)  $\Rightarrow$  (i) If (ii) is true then

$(x_1, x_2, x'_1, x'_2) \rightarrow k_1(x_1, x'_1)$  kernel  
on  $\mathcal{X} = \mathcal{X}_1 \times \mathcal{X}_1$

$(x_1, x_2, x'_1, x'_2) \rightarrow k_2(x_2, x'_2)$  kernel  
on  $\mathcal{X} = \mathcal{X}_1 \times \mathcal{X}_2$

$(x_1, x_2, x'_1, x'_2) \rightarrow k_1(x_1, x'_1) k_2(x_2, x'_2)$   
kernel on  $\mathcal{X} = \mathcal{X}_1 \times \mathcal{X}_2$

proof of (ii)

(ii) If  $\mathcal{X}_1 = \mathcal{X}_2 = \mathcal{X}$

$\Gamma(x, x') = k_1(x, x') k_2(x, x')$   
is a kernel on  $\mathcal{X}$

Need the following

Thm  $k: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$

$$x, x' \rightarrow k(x, x')$$

is a kernel on  $\mathcal{X}$  if

$\forall x_1, \dots, x_m \in \mathcal{X}$

$$\Theta = (\Theta_{i,j}) \quad \Theta_{i,j} = k(x_i, x_j)$$

is symmetric & positive

—————

Let  $x_1, \dots, x_m \in \mathcal{X}$

$$\Theta'_{i,j} = k_1(x_i, x_j) \quad \Theta^2_{i,j} = k_2(x_i, x_j)$$

$$\Theta' = L L^T \quad \left( \begin{pmatrix} \text{Lower} \\ \text{Upper} \end{pmatrix} = \begin{pmatrix} L \\ U \end{pmatrix} \right) = \begin{pmatrix} \text{Cholesky factor} \end{pmatrix}$$

$\rightarrow$  Cholesky factorization of  $\Theta'$

$$c \in \mathbb{R}^m$$

$$\begin{aligned} \sum_{i,j} c_i \cap(x_i, x_j) c_j &= \sum_{i,j} c_i \theta_{i,j}^1 \theta_{i,j}^2 c_j \\ &= \sum_{i,j,k} c_i c_j L_{ik} L_{jk} \theta_{i,j}^2 \\ &= \sum_k (\Sigma^k)^T \theta^2 \Sigma^k \geq 0 \end{aligned}$$

$$Z_k = \begin{pmatrix} c_1 L_{1k} \\ \vdots \\ c_m L_{mk} \end{pmatrix}$$

Lemma (Polynomial kernels)

Let  $x, x' \in \mathbb{R}^d$ ,  $d \geq 1$

$m \geq 1$   $c \geq 0$  ( $c \in \mathbb{R}$ )

$m \in \mathbb{N}$

Then  $k(x, x') = (\langle x, x' \rangle + c)^m$

is a valid kernel

proof

$$\phi(x) = \begin{pmatrix} x \\ \sqrt{c} \end{pmatrix} \in \mathbb{R}^{c+1} = H$$

$$\phi(x) \cdot \phi(x') = x^T x' + c$$

$\Downarrow$  stability under product  
is a kernel

$(\langle x, x' \rangle + c)^m$  is a kernel

Def The space  $\ell_p$  of  $p$ -summable sequences is defined as all sequences  $(a_i)_{i \geq 1}$  for which

$$\sum_{i=1}^{\infty} |a_i|^p < \infty$$

$\ell_2$ : Hilbert space  $\langle a, b \rangle = \sum_i a_i b_i$

Lemma Given a non-empty set  $X$  and a sequence of functions  $(\phi_i(x))_{i \geq 1}$

in  $\ell_2$  where  $\phi_i : \mathcal{X} \rightarrow \mathbb{R}$  is the  $i$ th coordinate of the feature map  $\phi(\mathbf{x})$

Then

$$K(\mathbf{x}, \mathbf{x}') = \sum_{i=1}^{\infty} \phi_i(\mathbf{x}) \phi_i(\mathbf{x}')$$

is a well defined kernel on  $\mathcal{X}$

proof

$$\begin{aligned} |K(\mathbf{x}, \mathbf{x}')| &= \left| \sum_i \phi_i(\mathbf{x}) \phi_i(\mathbf{x}') \right| \\ &\leq \|\phi(\mathbf{x})\|_{\ell_2} \|\phi(\mathbf{x}')\|_{\ell_2} \end{aligned}$$

C.S.

$K$  is well defined  $\forall \mathbf{x}, \mathbf{x}' \in \mathcal{X}$

Prop (Taylor series kernel)

Assume we can define the Taylor series

$$f(z) = \sum_{n=0}^{\infty} a_n z^n, \quad |z| < r, \quad z \in \mathbb{R}$$

for  $r \in (0, \infty]$ , with  $a_n > 0$   $\forall n$

Define  $\mathcal{X} = \left\{ x \in \mathbb{R}^d \mid \|x\| < \sqrt{r} \right\}$

Then

$$K(x, x') = \sum_{n=0}^{\infty} a_n \langle x, x' \rangle^n \quad (= f(\langle x, x' \rangle))$$

is a valid kernel on  $\mathcal{X}$

proof

$$K(x, x') = \sum_{n=0}^{\infty} a_n \langle x, x' \rangle^n$$

converges

CS inequ  $|\langle x, x' \rangle| \leq \|x\| \|x'\| < r$

Ex  $K(x, x') = \underbrace{\exp(\langle x, x' \rangle)}$  is a kernel  
on  $\mathbb{R}^d$

$$f(z) = e^z = \sum_{n=0}^{\infty} \frac{z^n}{n!}$$

↓ exponential kernel

$$\underline{\text{Ex}} \quad G(x, x') = \exp\left(-\frac{\|x - x'\|^2}{2\sigma^2}\right)$$

$\downarrow$  is a kernel on  $\mathbb{R}^d$

Gaussian kernel

Normalized kernels

Prop Let  $K$  be a kernel on  $\mathcal{X}$

Let

$$K'(x, x') = \begin{cases} 0 & \text{if } K(x, x) = 0 \\ & \text{or } K(x', x') = 0 \\ \frac{K(x, x')}{\sqrt{K(x, x)} \sqrt{K(x', x')}} & \text{otherwise.} \end{cases}$$

Then  $K'$  is a kernel on  $\mathcal{X}$

proof Let  $x_1, \dots, x_m \in \mathcal{X}, c \in \mathbb{R}^m$

$$K \leftrightarrow \phi, H$$

$$\begin{aligned}
& \sum_{i,j} c_i c_j K'(x_i, x_j) \\
&= \sum_{i,j} c_i c_j \frac{k(x_i, x_j)}{\sqrt{k(x_i, x_i) k(x_j, x_j)}} \\
&= \sum_{i,j} c_i c_j \underbrace{\langle \phi(x_i), \phi(x_j) \rangle_H}_{\| \phi(x_i) \|_H \| \phi(x_j) \|_H} \\
&= \left\| \sum_i c_i \frac{\phi(x_i)}{\| \phi(x_i) \|_H} \right\|_H^2 \geq 0
\end{aligned}$$

$$K' \Leftrightarrow \phi'(x) = \frac{\phi(x)}{\| \phi(x) \|_H}, H$$

$$\text{Ex } G(x, x') = \exp\left(-\frac{\|x - x'\|^2}{2\sigma^2}\right)$$

$\downarrow$  is a kernel on  $\mathbb{R}^d$

Gaussian kernel

Sol

$$K(x, x') = \exp\left(\frac{\langle x, x' \rangle}{\sigma^2}\right)$$

is a kernel on  $\mathbb{R}^n$

$$G(x, x') = \frac{K(x, x')}{\sqrt{K(x, x) K(x', x')}}$$

$$= \frac{e^{x \cdot x' / \sigma^2}}{e^{\|x\|^2 / \sigma^2} e^{\|x'\|^2 / \sigma^2}}$$

$$= \exp\left(-\frac{1}{2\sigma^2} \|x - x'\|^2\right)$$

is a kernel

Def A symmetric function

$$k: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$$

is of positive type if  $\sum_{i,j} c_i c_j k(x_i, x_j) \geq 0$ ,  $c \in \mathbb{R}^m$

$$x_1, \dots, x_m \in \mathcal{X}$$

$$\sum_{i,j} c_i c_j k(x_i, x_j) \geq 0$$

( $k$  is also called PDS, positive definite symmetric)

The function  $k$  is strictly positive definite (or non-degenerate) if

$$\sum_{i,j} c_i c_j k(x_i, x_j) = 0 \quad \begin{cases} c = 0 \\ \& x_i \neq x_j \text{ for } i \neq j \end{cases}$$

Lemma

Let  $H$  be any Hilbert space,  $X$

a non-empty set and  $\phi: X \rightarrow H$

Then  $k(x, x') = \langle \phi(x), \phi(x') \rangle_H$   
is a symmetric function of positive type  
(PDS)

Proof  $\sum_{i,j} c_i c_j k(x_i, x_j) = \sum_{i,j} c_i c_j \langle \phi(x_i), \phi(x_j) \rangle_H$

$$= \left\| \sum_i c_i \phi(x_i) \right\|_H^2 \geq 0$$

Lemma

Let  $k: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$  be PDS

Then  $\forall x, x'$

$$(k(x, x'))^2 \leq k(x, x) k(x', x')$$

Proof

$$\text{Let } \Theta = \begin{pmatrix} k(x, x) & k(x, x') \\ k(x', x) & k(x', x') \end{pmatrix}$$

$$\Theta \succcurlyeq 0 \Rightarrow \det(\Theta) \geq 0$$

"

$$k(x, x) k(x', x') - (k(x, x'))^2 \geq 0$$

# Thm Reproducing Kernel Hilbert Space (RKHS)

Let  $k: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$  be a PDS function.

Then there exists a Hilbert space  $\mathcal{H}$  and a mapping  $\phi: \mathcal{X} \rightarrow \mathcal{H}$  s.t.

$$(1) \quad k(x_1, x') = \langle \phi(x_1), \phi(x') \rangle_{\mathcal{H}}$$
$$\forall x_1, x' \in \mathcal{X}$$

Furthermore  $\mathcal{H}$  can be taken to be a space of functions  $\mathcal{X} \rightarrow \mathbb{R}$  with the following property known as the reproducing property

$$(2) \quad h(x) = \langle h, k(x, \cdot) \rangle_{\mathcal{H}}$$

$\mathcal{H}$  is called the reproducing kernel Hilbert Space (RKHS) associated to  $k$

Rk (1) & (2)  $\Rightarrow (\mathcal{H}, \langle \cdot, \cdot \rangle_{\mathcal{H}})$  is unique

Proof

proof of (1). Based on GPs

$k$  PDS  $\Rightarrow k$  is a covariance function

$\Rightarrow \exists$  a GP  $\xi \sim \mathcal{N}(0, k)$

i.e.  $k(x_1, x') = E[\xi(x), \xi(x')]$

$\Rightarrow \mathcal{H}$  = Gaussian Space

containing all  $(\xi(x))_{x \in X}$

$$\phi(x) = \xi(x)$$

$$\langle \cdot, \cdot \rangle_H = \text{L_2 inner product in } H$$

Q Is  $H$  the RKHS associated with  $k$ ?

→ no!

This  $H$  satisfies (1) but not (2)

$$H \neq \{X \rightarrow \mathbb{R}\}$$

$$= \{X \rightarrow \text{Gaussian space}\}$$

Proof (1) & (2)

$\forall x \in X$ , define  $\phi(x): X \rightarrow \mathbb{R}$

$$\forall x' \in X, \quad \phi(x)(x') = k(x, x')$$

$$\phi(x)(\cdot) = k(x, \cdot)$$

$$H_0 = \left\{ \sum_{i \in \mathcal{I}} c_i \phi(x_i) \mid c_i \in \mathbb{R}, x_i \in \mathcal{X} \right\}$$

$\text{Car}(\mathcal{I}) < \infty$

$$\langle \cdot, \cdot \rangle : H_0 \times H_0 \rightarrow \mathbb{R}$$

$$(f, g) \mapsto \langle f, g \rangle$$

||

$$\begin{aligned} \sum_{i \in \mathcal{I}} a_i b_i k(x_i, x_j) \\ \sum_{j \in \mathcal{J}} " b_j f(x_j) \\ \sum_{i \in \mathcal{I}} " a_i g(x_i) \end{aligned}$$

$\langle \cdot, \cdot \rangle$ : is symmetric, bilinear

$\langle f, g \rangle$ : does not depend on particular representation of  $f$  &  $g$

$$k \text{ PDS} \Rightarrow \langle f, f \rangle = \sum_{i,j \in I} \alpha_i \alpha_j k(x_i, x_j) \geq 0$$

$$\forall f = \sum_{i \in I} \alpha_i \phi(x_i) \in H_0$$

$$\sum_{i,j=1}^m c_i c_j \langle f_i, f_j \rangle = \left\langle \sum_{i=1}^m c_i f_i, \sum_{i=1}^m c_i f_i \right\rangle \geq 0$$

$\forall f_1, \dots, f_m \in H_0$   
 $c_1, \dots, c_m \in \mathbb{R}$

$\langle \cdot, \cdot \rangle$  is PDS in  $H_0$

Prev. lemma  $\Rightarrow$

$$\langle f, \phi(x) \rangle^2 \leq \langle f, f \rangle \langle \phi(x), \phi(x) \rangle$$

$\forall f \in H_0, x \in X$

By def of  $\langle \cdot, \cdot \rangle$

$$Af = \sum_{i \in \mathcal{X}} a_i \phi(x_i) \in H_0$$

(\*)  $f(x) = \sum_{i \in \mathcal{X}} a_i k(x_i, x) = \langle f, \phi(x) \rangle$

$$|f(x)|^2 \leq \langle f, f \rangle k(x, x)$$

$\Rightarrow \langle \cdot, \cdot \rangle$ : defines an inner product  
on  $H_0$  which thereby becomes a  
pre-Hilbert space

$\Rightarrow \langle \cdot, \cdot \rangle$ : defines an inner product  
on  $H_0$  which thereby becomes a  
pre-Hilbert space

$H_0$  can be closed (completed)  
to form a Hilbert space in which  
 $H_0$  is dense

By Cauchy - Schwartz inequ.

$\forall x \in \mathcal{X}$

$$f \mapsto \langle f, \phi(x) \rangle$$

is Lipschitz

and therefore continuous

$\ell_2$  is dense in  $H \Rightarrow$  reproducing prop \*  
holds in  $H$

Def A Hilbert space  $H$  of  
functions is a RKHS if the  
evaluation functionals

$f \mapsto f(x)$  are bounded, i.e.

$\forall x \in \mathcal{X}, \exists C > 0$

$$|f(x)| \leq C \|f\|_H, \forall f \in H$$

Thm (Riesz representation)

If  $A: \mathcal{H} \rightarrow \mathbb{R}$  is a bounded linear operator in a Hilbert space  $\mathcal{H}$ , there exists some  $g_A \in \mathcal{H}$  s.t.

$$A[f] = \langle f, g_A \rangle_{\mathcal{H}} \quad \forall f \in \mathcal{H}$$

Prop (Reproducing prop)

If  $\mathcal{H}$  is a RKHS then  $\forall x \in X$  there exists a function  $k_x \in \mathcal{H}$  s.t.

$$f(x) = \langle k_x, f \rangle_{\mathcal{H}} \quad \forall f \in \mathcal{H}$$

Rk  $k(x, y) = \langle k_x, k_y \rangle_{\mathcal{H}}$

$$= \langle k_y, k_x \rangle = k(y, x)$$

kernel associated with RKHS  $H$

$$\sum_{i,j} c_i c_j k(x_i, x_j) = \left\| \sum_i c_i k_{x_i} \right\|_H^2 \geq 0$$

Thm For every (positive definite symmetric function) kernel on  $\mathcal{X} \times \mathcal{X}$  there exists a unique RKHS with  $k$  as its reproducing kernel.

Conversely the reproducing kernel of a RKHS is unique  
(and PDS)

$$k(x, y) = \langle k_x, k_y \rangle_H$$

Ex RKHS associated with GP  
 Let  $\xi \sim \mathcal{N}(0, A^{-1})$   $A^{-1}_{(i,j)}$

Gaussian vector on  $\mathbb{R}^N$

$A$ :  $N \times N$  PDS matrix

Here :  $\chi = \{1, 2, \dots, N\}$

RKHS :  $\mathbb{R}^N$  endowed with norm

$$\|v\|^2 = v^T A v \quad \langle u, v \rangle = u^T A v$$

$H : \{\chi_1, \dots, \chi_N\} \rightarrow \mathbb{R} \iff \mathbb{R}^N$   
 $i \qquad \qquad \qquad v_i \qquad \qquad \qquad v \in \mathbb{R}^N$

Feature Map ( $E[\xi_i \xi_j] = k(i, j)$ )

$\phi(i) : \chi \rightarrow \mathbb{R} \rightarrow k(i, \cdot)$   
 $j \qquad \qquad \qquad \phi(i)(j) = k(i, j) = A^{-1}_{i,j}$

$$\langle \phi(i), \phi(j) \rangle = A_{i,i}^{-1} A_{\cdot,j} A_{\cdot,i}^{-1} = A_{i,j}^{-1} = k(i,j)$$

Reproducing property

$$v: \{1, \dots, N\} \rightarrow \mathbb{R}$$

$$\langle v, \phi(i) \rangle = v(i)$$

$$\approx v^T A^{-1} A_{\cdot,i}$$

Ex Let  $\xi \sim \mathcal{N}(0, \mathcal{L}^{-1})$  be a GP on  $\Omega \subset \mathbb{R}^d$

$\mathcal{Q}: H_0^s(\Omega) \longrightarrow H^{-s}(\Omega)$  enclosed

with norm  $\|v\|^2 = \int_{\Omega} v^T \mathcal{L} v$  for  $v \in H_0^s(\Omega)$

$s > \frac{d}{2}$ ,  $\mathcal{L}$ : symm. pos. definite (elliptic)

$$\int_U \nabla \mathcal{L} V = \int_U \nabla \mathcal{L} \circ$$

$$\int_U \nabla \mathcal{L} V \geq 0$$

$$\int_U \nabla \mathcal{L} V = 0 \Rightarrow V = 0$$

$v \in H_0^s(\Omega) : v: \Omega \rightarrow \mathbb{R}$   
 $x \mapsto v(x)$

$$\int_U \nabla \mathcal{L} V = \int_{\Omega} v(x) \mathcal{L} V(x) dx$$

kernel: Green's function of  $\mathcal{L}$

$$G(x, y) = \int_{\Omega} \delta_x \mathcal{L}^{-1} \delta_y$$

$$\phi(x) = G(x, \cdot)$$

$$\langle \phi(x), \phi(y) \rangle = \int_{\Omega} \phi(x)(x') \mathcal{L} \phi(y)(x')$$

$$\langle v, v \rangle = \int_{\Omega} v \mathcal{L} v$$

$$\phi(x) : \Omega \rightarrow \mathbb{R}$$

$$x' \mapsto G(x, x')$$

$$\langle \phi(x), \phi(y) \rangle$$

$$G(x, y) = \int_{\Omega} G(x, x') \underbrace{\delta(x' - y)}_{\delta y} dx'$$

$$= \int_{\Omega} G(x, x') \delta(x' - y) dx = f(x)$$

$$v(x) = \int v \delta G(\cdot, x) = \langle v, \phi(x) \rangle$$

—

$H$ : the reproducing kernel Hilbert space associated to  $k$

Any Hilbert space  $H$  s.t.

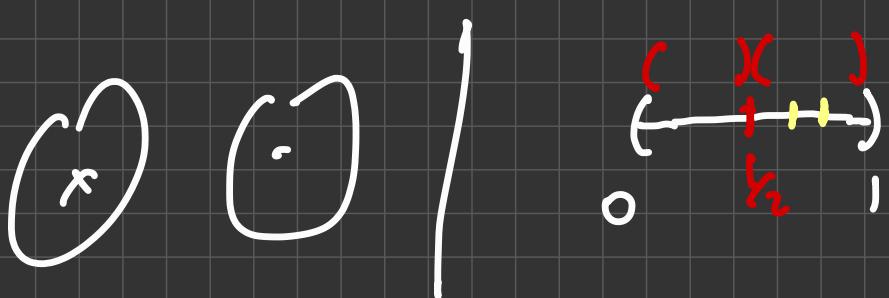
$$\exists \phi: X \rightarrow H$$

with  $k(x, x') = \langle \phi(x), \phi(x') \rangle$   
 $\forall x, x'$

is called a feature space associated  
and  $\phi$  is called a feature map.

## Mercer's thm

Hausdorff space: topological space  
in which each pair of distinct points  
can be separated by disjoint open sets



Borel measure: Any measure defined  
on  $\sigma$ -algebra of Borel sets

Finite measure: measure that always takes finite values

Thm Let  $\mathcal{X}$  be a compact Hausdorff space and  $\nu$  a finite Borel measure with support  $\mathcal{X}$

Suppose  $k$  is a continuous SPD kernel on  $\mathcal{X}$ , and define the integral operator

$$T_k: L_2(\mathcal{X}, \nu) \rightarrow L_2(\mathcal{X}, \nu) \text{ by}$$

$$(T_k f)(\cdot) = \int_{\mathcal{X}} k(\cdot, x) f(x) \nu(dx)$$

which is positive definite, i.e.  $\forall f \in L_2(\mathcal{X}, \nu)$

$$\int_{\mathcal{X}^2} k(u, v) f(u) f(v) \nu(du) \nu(dv) \geq 0$$

(Mercer's condition)

Then there is an orthonormal basis  $\{\psi_i\}$  of  $L_2(X, \mu)$  consisting of eigenvectors of  $T_k$  s.t. the corresponding sequence of eigenvalues  $\{\lambda_i\}$  are non-negative.

The eigenfunctions corresponding to non-zero eigenvalues can be taken as continuous functions on  $X$  and  $k(u, v)$  has the representation

$$T_k \psi_i = \lambda_i \psi_i$$

$$k(u, v) = \sum_{i=1}^{\infty} \lambda_i \psi_i(u) \psi_i(v)$$

where the convergence is absolute and uniform

Rk

$$k(u, v) = \left\langle (\sqrt{\lambda_1} \psi_1(u), \sqrt{\lambda_2} \psi_2(u), \dots), (\sqrt{\lambda_1} \psi_1(v), \sqrt{\lambda_2} \psi_2(v), \dots) \right\rangle_{L_2}$$

$$\phi(x) = (\sqrt{\lambda_1} \psi_1(x), \sqrt{\lambda_2} \psi_2(x), \dots)$$

$\{t = \rho_{x_i}\} (a_1, \dots, \dots) \}$

$$\langle a, b \rangle = \sum_{i=1}^{\infty} a_i b_i$$

## Bochner's theorem

Thm  $k(x, x') = \Psi(x - x')$  is a kernel  
on  $\mathbb{R}^d$  iff  $\exists$  a finite non-negative  
Borel measure  $\Lambda$  on  $\mathbb{R}^d$

$$\Psi(x - x') = \int_{\mathbb{R}^d} e^{i\omega^T(x - x')} \Lambda(d\omega)$$

Rk One may normalize  $k$ . s.t.  $\Psi(0) = 1$   
 $\Rightarrow \Lambda$  is a prob. measure and  
 $k$  corresponds to its characteristic function

$$\Rightarrow \Psi(x - x') = \mathbb{E}_{\omega \sim \Lambda} [e^{i\omega^T(x - x')}]$$

$\omega_j$  iii ~  $\wedge$

$$\Psi(x - x') \approx \frac{1}{N} \sum_{j=1}^N e^{i \omega_j^T (x - x')}$$

$$\approx \left( \frac{e^{i \omega_1^T x}}{\sqrt{N}}, \frac{e^{i \omega_2^T x}}{\sqrt{N}}, \dots, \frac{e^{i \omega_N^T x}}{\sqrt{N}} \right)$$

$\cdot \underbrace{\quad}_{\left( \frac{e^{i \omega_1^T x'}}{\sqrt{N}}, \dots, \frac{e^{i \omega_N^T x'}}{\sqrt{N}} \right)}$

Rk  $k(x, x') = \Psi(x - x')$

$$(X_1, Y_1), \dots, (X_n, Y_M)$$

$$k(X, X)^{-1} \rightarrow M$$

$$\phi(x) = (\phi_1(x), \dots, \phi_N(x))$$

$$f(x) = \sum_{i=1}^N c_i \phi_i(x)$$

$M >> N$

$$\min_{\Phi} \sum_{i=1}^M \|f(\mathbf{x}_i) - y_i\|^2 + \lambda \|\Phi\|_K^2$$

↗

$$K_f(x, x') = \sum_{i=1}^N \phi_i(x) \phi_i(x')$$

inverting

$$\|\Phi\|_K^2 = \sum_{i=1}^N c_i^2$$

MxM  
matrix

$$\min_c \sum_{i=1}^M \left| \sum_{j=1}^N c_j \phi_j(x_i) - y_i \right|^2 + \lambda \sum_{j=1}^N c_j^2$$

↗ ihrer Erbringung a NxN matrix

$$\Psi(x-x') = \int_{\mathbb{R}^d} e^{i\omega(x-x')} c(\Lambda(\omega))$$

$$\underline{\text{Ex}} \quad k(x, x') = \exp\left(-\frac{\|x - x'\|^2}{2\sigma^2}\right)$$

$$\uparrow \downarrow$$

$$K(\omega) = \left(\frac{2\pi}{\zeta^2}\right)^{-d/2} \exp\left(-\zeta^2 \frac{\|\omega\|_2^2}{2}\right)$$

$$k(x, x') = \exp\left(-\frac{\|x - x'\|_1}{\sigma}\right)$$

$$\uparrow \downarrow$$

$$K(\omega) = \prod_{i=1}^d \frac{1}{1 + \omega_i^2}$$

## Radial kernels

Def The kernel  $k(x, x')$  is said to be radial if  $k(x, x') = \Psi(\|x - x'\|)$

for some (positive definite)  $\Psi: [0, \infty) \rightarrow \mathbb{R}$

Thm ( Schoenberg's thm.)

A continuous function  $\Psi: [0, \infty) \rightarrow \mathbb{R}$

is positive definite and radial

(i.e. it defines a kernel  $k(x, x') = \Psi(\|x-x'\|)$   
on  $\mathbb{R}^d$  for all  $d$ ) if it is of the form

$$\Psi(r) = \int_0^\infty \exp(-r^2 f^2) \nu(df)$$

where  $\nu$  is a finite non-negative  
Borel measure on  $[0, \infty)$

$\Rightarrow$  radial kernels are mixtures of  
Gaussians

Rk Thm,  $k: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$

s.t.  $\forall x_1, \dots, x_N, \forall N$

$$\Theta_{i,j} = k(x_i, x_j)$$

$\Theta$  is symm. positive

$$\Rightarrow k(x, x') = \langle \phi(x), \phi(x') \rangle_H$$

—

## Ridge regression

data  $(X_1, Y_1), \dots, (X_n, Y_n)$   
 $X_i \in \mathbb{R}^d \quad Y_i \in \mathbb{R}$

Ridge regression Linear regression with  $L^2$  penalty

$X_{n+1} \rightarrow$  predict  $Y_{n+1}$

idea:  $a^\top X_{n+1} \quad a \in \mathbb{R}^d$

$$\min_{a \in \mathbb{R}^d} \sum_{i=1}^n |Y_i - a^\top X_i|^2 + \lambda a^\top a$$

Solution

$$X_i \rightarrow \begin{pmatrix} X_{i,1} & X_{i,2} & \dots & X_{i,n} \\ \vdots & & & \vdots \\ X_{i,1} & X_{i,2} & \dots & X_{i,n} \end{pmatrix} \in \mathbb{R}^{d \times n}$$

$n$

$$X_{i,j} := (X_i)_j \quad (X_{i,j})_i = (X_j)_i$$

$$Y_i = \begin{pmatrix} Y_{i,1} \\ \vdots \\ Y_{i,n} \end{pmatrix}$$

$$\begin{aligned} \text{Loss: } & \|Y - X^T a\|_2^2 + \lambda \|a\|_2^2 \\ &= Y^T Y - 2 Y^T X^T a + a^T X X^T a + \lambda a^T a \\ &= Y^T Y - 2 Y^T X^T a + a^T (X X^T + \lambda I) a \end{aligned}$$

minimized for

$$a = a^* = (X X^T + \lambda I)^{-1} X^T Y$$

which is the classical regularized least square solution

# Kernel ridge regression

data  $(X_1, Y_1), \dots, (X_n, Y_n)$

$$X_i \in \mathcal{X}$$

$$Y_i \in \mathbb{R}$$

$k(x, x')$ : kernel  $\mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$

$\phi$ : feature map

$H$ : feature space (RKHS)

$$k(x, x') = \langle \phi(x), \phi(x') \rangle_H$$

Idea:  $X_{n+1} \xrightarrow{\text{predict}} Y_{n+1}$

predict  $Y_{n+1}$  with  $f(X_{n+1})$

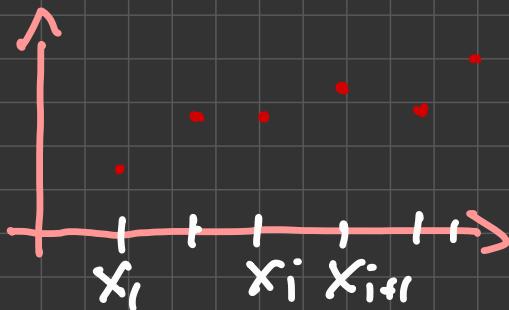
$$f \in \mathcal{H}$$

$$(1) \min_{f \in \mathcal{H}} \sum_{i=1}^n |Y_i - f(X_i)|^2 + \lambda \|f\|_H^2$$

using  $\langle f, \phi(X_i) \rangle = f(\phi(X_i))$

$$(2) \min_{f \in H} \sum_{i=1}^n |Y_i - \langle f, \phi(x_i) \rangle_H|^2 + \lambda \|f\|_H^2$$

Q: Why  $\exists \infty \# \text{ of } f / \sum_i |Y_i - f(x_i)|^2 = 0$



Thm The minimizer of (1)  $\Leftrightarrow$  (2) is

$$f = \sum_{i=1}^n c_i \phi(x_i)$$

Proof  $f = \underbrace{\sum_{i=1}^n c_i \phi(x_i)}_{f_\phi} + f_\perp$

$$f_\phi \in \text{span} \{ \phi(x_1), \dots, \phi(x_n) \}$$

$f_\perp$ : orthog. complement in  $H$

$\mathcal{L}$ : loss function

$$\mathcal{L}(f) = \sum_{i=1}^n |y_i - \langle f_\phi + f_\perp, \phi(x_i) \rangle|^2 + \lambda \|f_\phi + f_\perp\|_H^2$$

$$\langle f_\perp, \phi(x_i) \rangle = 0$$

$$= \sum_{i=1}^n |y_i - \underbrace{\langle f_\phi, \phi(x_i) \rangle}_\text{does not depend on } f_\perp|^2 + \lambda (\|f_\phi\|_H^2 + \|f_\perp\|_H^2)$$

does not depend on  $f_\perp$

minimize d for  $f_\perp = 0$

$$\Rightarrow \mathcal{L}(f) \text{ minimized for } f \in \text{Span} \{ \phi(x_1), \dots, \phi(x_n) \}$$

$k(X, X)$ : Gram matrix ( $n \times n$ )

$$(k(X, X))_{i,j} = k(x_i, x_j)$$

$$\text{For } f = \sum_{i=1}^n c_i \phi(x_i)$$

$$\mathcal{L}(f) = \|Y - k(X, X)C\|_2^2 + \lambda C^T k(X, X)C$$

$$\langle \phi(x_i), \phi(x_j) \rangle_H = k(x_i, x_j)$$

Minimizer in  $C$  is

$$C^* = (k(X, X) + \lambda I)^{-1} Y$$

for  $x \in \mathcal{X}$ ,  $k(X, x)$ : n dim vector  
with entries

$$k(x_i, x)$$

$$\begin{aligned} f_{(x)}^* &= Y^T (k(X, X) + \lambda I)^{-1} k(X, x) \\ &= k(x, X) (k(X, X) + \lambda I)^{-1} Y \end{aligned}$$

$$f^*(x) = \sum_{i=1}^n c_i^* \underbrace{\phi(x_i)}_{k(x_i, x) = k(x, x_i)}$$

Let  $\xi$  be the centered GP with covariance function  $k$

$$\xi \sim \mathcal{N}(0, k)$$

$\xi$  defined by quadratic norm  $\|\cdot\|_H$

$$E[\xi(x) \xi(x')] = k(x, x')$$

( $H$  self-dual, using  $\langle \cdot, \cdot \rangle_H$  to define duality)

$$\text{for } \phi \in H, \quad \langle \phi, \xi \rangle \sim \mathcal{N}(0, \|\phi\|_H^2)$$

$$\begin{aligned} \xi: H &\rightarrow \text{Gaussian space} \\ \phi &\mapsto \xi(\phi) \end{aligned}$$

Thm For

$$f^*_{\lambda} = \underset{f \in H}{\operatorname{argmin}} \sum_{i=1}^n |y_i - f(x_i)|^2 + \lambda \|f\|_H^2$$

we have

$$f^*_{\lambda}(x) = E[\xi(x) \mid \xi(x_i) = y_i + z_i] \quad \forall i = 1, \dots, n$$

where  $z_i$  are i.i.d.  $\sim \mathcal{N}(0, \lambda)$

If  $k(x, x)$  invertible

as  $\lambda \downarrow 0$ ,  $f_\lambda^* \rightarrow f_0^*$  (pointwise in  $\| \cdot \|_H$  norm)

$f_0^*$  is the minimizer of

$$\begin{cases} \min & \|f\|_H \\ \text{s.t.} & f(x_i) = y_i \end{cases}$$

and  $f_0^*(x) = E[\xi(x) \mid \xi(X_i) = y_i]_{i=1, \dots, n}$

proof Consider Gaussian vector

$$(\underbrace{\xi(X_1) - z_1, \dots, \xi(X_n) - z_n}_{V}, \xi(x))$$

$$\text{Cov}(V, V) = \text{Cor}(V) = k(X, X) + \lambda I$$

$$\text{Cor}(V, \xi(x)) = k(X, x)$$

$$\begin{aligned} E[\xi(x) \mid V] &= V^T \text{Cov}(V)^{-1} \text{cor}(V, \xi(x)) \\ &= \text{Cor}(\xi(x), V) \text{Cov}(V)^{-1} V \end{aligned}$$

$$\text{Cov}(\xi(x) - V^T \text{Cov}(V)^{-1} \text{Cov}(V, \xi(x)), V) = 0$$

Q: Unpack convergence

$$f_\lambda^*(x) = k(x, X) (k(X, X) + \lambda I)^{-1} y$$

$$f_0^*(x) = k(x, X) k(X, X)^{-1} y$$

$$\begin{cases} \min \|f\|_H \\ \text{s.t. } f(x_i) = y_i \end{cases}$$

$$f = f_\phi + f_\perp \quad f_\phi \in \text{span } \{\phi(x_1), \dots, \phi(x_n)\}$$

$f_\perp \in \text{orth. comp. in } H$

$$\begin{cases} \min \|f\|_H^2 \\ \text{s.t. } f(x_i) = y_i \end{cases} \Leftrightarrow \begin{cases} \min \|f_\phi\|_H^2 + \|f_\perp\|_H^2 \\ \text{s.t. } \langle f_\phi, \phi(x_i) \rangle = y_i \end{cases}$$

$$\langle f_\phi, \phi(x_i) \rangle$$

minimized for

$$f_\perp = 0$$

$$f = \sum_{i=1}^n c_i \phi(x_i)$$

$$\Leftrightarrow \begin{cases} \min & c^\top k(X, X) \\ \text{s.t.} & c^\top k(X, X) = Y \end{cases}$$

$$\Leftrightarrow c^* = k(X, X)^{-1} Y$$

$$f_\lambda^*(x) = k(x, X) (k(X, X) + \lambda I)^{-1} Y$$

$$f_0^*(x) = k(x, X) k(X, X)^{-1} Y$$

$$\begin{aligned} f_\lambda^*(x) &\rightarrow f_0^*(x) & \forall x \in \mathcal{X} \\ f_\lambda^* &\xrightarrow{\mathcal{H}} f_0^* & (\text{in RKHS norm} \\ && \| \cdot \|_{\mathcal{H}}) \end{aligned}$$

$$\| f_\lambda^* - f_0^* \|_{\mathcal{H}}^2 = \left\| \sum_{i=1}^n (c_{\lambda,i}^* - c_{0,i}^*) \phi(x_i) \right\|_{\mathcal{H}}^2$$

$$c_\lambda^* = (k(X, X) + \lambda I)^{-1} Y$$

$$C_0 = k(X, X)^{-1} Y$$

$$\lim_{\lambda \rightarrow 0} C_\lambda^* = C_0$$

$$A(\lambda) = k(X, X) + \lambda I$$

$$\frac{d}{d\lambda} A(\lambda)^{-1} \Big|_{\lambda=0} = -A(0)^{-1} \frac{d}{d\lambda} A(0) A(0)^{-1}$$

## Representer theorem

Thm Let  $k$  be a kernel on  $\mathcal{X}$   
and let  $H$  be its associated RKHS

For  $x_1, \dots, x_n \in \mathcal{X}$  consider the opt. pb

$$(*) \min_{f \in H} D(f(x_1), \dots, f(x_n)) + P(\|f\|_H^2)$$

where  $P$  is non decreasing and

$D$  depends on  $f$  only through  $f(x_1), \dots, f(x_n)$

If  $(*)$  has a minimizer then if

has a minimizer of the form

$$f(x) = \sum_{i=1}^n c_i k(x, x_i) \quad \text{where } c_i \in \mathbb{R}$$

Furthermore if  $P$  is strictly increasing  
then every minimizer of  $(*)$  has this form.

Proof Write

$$\mathcal{J}(f) = D(f(x_1), \dots, f(x_n)) + P(\|f\|_H^2)$$

Consider  $S := \text{span}\{k(\cdot, x_1), \dots, k(\cdot, x_n)\}$

$$S \subset H$$

$S$  finite dim  $\Rightarrow S$  closed

projection thm  $\Rightarrow H = S \oplus S^\perp$

orth comp. in  $H$

$\Rightarrow \forall f \in H$ ,  $f$  can be written

$$f = f_{\parallel} + f_{\perp} \quad \text{where } \begin{array}{l} f_{\parallel} \in S_{\perp} \\ f_{\perp} \in S^{\perp} \end{array}$$

$$\langle f_{\perp}, k(\cdot, x_i) \rangle_H \approx 0 \quad \forall i$$

reproducing property  $\Rightarrow$

$$\begin{aligned} f(x_i) &= \langle f, k(\cdot, x_i) \rangle_H \\ &= \langle f_{\parallel}, k(\cdot, x_i) \rangle_H + \langle f_{\perp}, k(\cdot, x_i) \rangle_H \\ &= f_{\parallel}(x_i) \end{aligned}$$



$$\begin{aligned} J(f) &= D(f(x_1), f(x_2), \dots, f(x_n)) + P(\|f\|_H^2) \\ &= D(f_{\parallel}(x_1), \dots, f_{\parallel}(x_n)) + P(\|f_{\parallel}\|_H^2) \end{aligned}$$

$$\geq D(f_{\parallel}(x_1), \dots, f_{\parallel}(x_n)) + P(\|f_{\parallel}\|_H^2)$$

$$= J(f_{\parallel}) \quad \|f\|_H^2 = \|f_{\parallel}\|_H^2 + \|f_{\perp}\|_H^2$$

$\Rightarrow$  If  $f$  is a minimizer of  $J(f)$

so is  $f_{\parallel}$

Second statement  $\Leftarrow$   $P$  is strictly  $\nearrow$

$$J(f) \underset{\text{if } f \notin S}{\overset{\Downarrow}{>}} J(f_{\parallel})$$

Alternate proof

$$\min_f D(f(x_1), \dots, f(x_n)) + P(\|f\|_H^2)$$

$$\min_{z \in \mathbb{R}^n} D(z_1, \dots, z_n) + P(\|f\|_H^2)$$

$$f \in \mathcal{H}$$

$$\text{s.t. } f(x_i) = z_i$$

$$\min_{z \in \mathbb{R}^n} \min_{f \in \mathcal{H}} D(z_1, \dots, z_n) + P(\|f\|_H^2)$$

$$\text{s.t. } f(x_i) = z_i$$

# Kernel PCA

$M$  centered observations

$x_k : k = 1, \dots, M, x_k \in \mathbb{R}^N$

$$\sum_{k=1}^M x_k = 0$$

PCA diagonalizes the cov. matrix

$$C = \frac{1}{M} \sum_{i=1}^M x_i x_i^\top$$

→ solve eigenvalue prob

$$* C v = \lambda v$$

for eigenvalues  $\lambda > 0$  and  $v \in \mathbb{R}^N \setminus \{0\}$

$N=2$



$(N \gg M)$

$$Cv = \lambda v$$

$$\Downarrow$$
$$Cv = \sum_{i=1}^M (x_i^T v) x_i = \lambda v$$

$$\lambda > 0 \Rightarrow v \in \text{Span}\{x_1, \dots, x_M\}$$

$$\lambda = 0 \Rightarrow v \in \text{Span}\{x_1, \dots, x_M\} \oplus V_\perp$$

$$\Rightarrow \text{All } v \in \text{Span}\{x_1, \dots, x_M\}$$

$$Cv = \lambda v \quad C(v+w) = Cv = \lambda v$$
$$w \in V_\perp \quad \Rightarrow \quad \lambda = 0 \text{ you can assume } w=0$$

$$V_\perp = \{w \in \mathbb{R}^N, w^T x_i = 0 \quad \forall i\}$$

Generalization of PCA in feature space  
(Kernel PCA)

$H$ : Hilbert space (feature space)

$\phi: \mathbb{R}^N \rightarrow H$  (feature map)

Kernel PCA  $\rightarrow$  do PCA on  $\phi(x_i)$   $i=1, \dots, M$

Assume data is contained in  $H$

$$\sum_{i=1}^M \phi(x_i) = 0$$

$$\bar{C}: H \rightarrow H$$

Cov matrix in  $H$

$$\bar{C} = \frac{1}{M} \sum_{i=1}^M \phi(x_i) \phi(x_i)^T$$

$\phi(x_i) \phi(x_i)^T$ : linear operator

$$H \rightarrow H$$

$$f \mapsto \phi(x_i) \langle \phi(x_i), f \rangle$$

We want eigenvectors of  $\bar{C}$

$$(v, \lambda) \quad \lambda \geq 0$$

$$\bar{C}v = \lambda v$$

By same argument as for PCA (in  $\mathbb{R}^N$ )  
 $v \in \text{Span } \{\phi(x_1), \dots, \phi(x_M)\}$

$$\bar{v} = \lambda v$$

$\uparrow\downarrow$

$$(1) \quad \lambda \phi(x_k)^T v = \phi(x_k)^T \bar{v} \quad \forall k=1, \dots, M$$

$$\rightarrow \exists \alpha_i \quad (i=1, \dots, M) \text{ s.t.}$$

$$(2) \quad v = \sum_{i=1}^M \alpha_i \phi(x_i)$$

$$(1) + (2) \Rightarrow$$

$$(3) \quad \lambda \sum_{i=1}^M \alpha_i \phi(x_k)^T \phi(x_i) =$$

$$\frac{1}{M} \sum_{i=1}^M \alpha_i \phi(x_k)^T \sum_{j=1}^M \phi(x_j) (\phi(x_j)^T \phi(x_i))$$

$$k=1, \dots, M$$

Define  $M \times M$  matrix  $K$  by

$$K_{i,j} = \phi(x_i)^T \phi(x_j) = k(x_i, x_j)$$

↑  $\phi$   
kernel

$$(3) \Leftrightarrow M \lambda K \alpha = K^T \alpha$$

$$\Leftrightarrow M \lambda \alpha = K \alpha$$

→ only need to diagonalize  $K$

$0 \leq \lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_M$  : eigenvalues

$\alpha^1 \quad \dots \quad \alpha^M$  eigenvectors

$$\lambda_i \alpha^i = K \alpha^i$$

$\lambda_p$ : first non zero eigenvalue

Normalize  $\alpha^p, \dots, \alpha^M$  by

$$(4) \quad \left\langle v^k, v^k \right\rangle_H = 1 \quad \forall k = p, \dots, M$$

$$\begin{aligned} (4) \Leftrightarrow 1 &= \sum_{i,j=1}^n \alpha_i^k \alpha_j^k \phi(x_i)^T \phi(x_j) \\ &= \sum_{i,j=1}^n \alpha_i^k \alpha_j^k K_{i,j} \\ &= \alpha^{k,T} K \alpha^k \\ &= \lambda_k \alpha^{k,T} \alpha^k \end{aligned}$$

$$\Leftrightarrow \alpha^{k,T} \alpha^k = 1/\lambda_k$$

For principal component extraction

need to compute projections on eigenvectors

$$v^k \in H, k = 1, \dots, n$$

Let  $x$  be a test point with an image  $\phi(x) \in H$

$$\begin{aligned} v^{k,T} \phi(x) &= \sum_{i=1}^M \alpha_i^k \phi(x_i)^T \phi(x) \\ &= \sum_{i=1}^M \alpha_i^k k(x_i, x) \end{aligned}$$

## KPCA

1. Compute matrix

$$K_{i,j} = k(x_i, x_j) \quad (M \times M \text{ matrix})$$

2. Compute eigenvectors

$$\lambda_j \alpha^j = K \alpha^j$$

$$\text{normalize} \Leftrightarrow \alpha^{j,T} \alpha^j \lambda_j = 1$$

3. Compute projections of a test point onto eigenvectors

$$\langle v^j, \phi(x) \rangle = \sum_{i=1}^M \alpha_i^j k(x_i, x)$$

Centering in feature space

drop assumption  $\sum_i \phi(x_i) = 0$

Use

$$\tilde{\phi}(x_i) = \phi(x_i) - \frac{1}{M} \sum_{i=1}^M \phi(x_i)$$

instead (they are centered in H)

prev. analysis holds with cor. matrix

$$\tilde{K}_{i,j} = \tilde{\phi}(x_i) \tilde{\phi}(x_j)^T$$

eig p<sup>l</sup>  $\rightarrow \tilde{\lambda} \tilde{\alpha} = \tilde{k} \tilde{\alpha}$

$\tilde{\alpha}$   $\rightarrow$  expansion coeff. of an eigenvector  
 $\tilde{v} = \sum_{i=1}^M \tilde{\alpha}_i \tilde{\phi}(x_i)$

express  $\tilde{K}$  in terms of  $K$

11:  $M \times M$  matrix,  $\Pi_{i,j} = 1 - K_{i,j}$

$$\begin{aligned}\tilde{K}_{i,j} &= \tilde{\phi}(x_i) \tilde{\phi}(x_j)^T \\ &= \left( \phi(x_i) - \frac{1}{M} \sum_{m=1}^M \phi(x_m) \right)^T \left( \phi(x_j) - \frac{1}{M} \sum_{n=1}^M \phi(x_n) \right) \\ &= K - \frac{1}{M} \Pi K - \frac{1}{M} K \Pi + \frac{1}{M^2} \Pi K \Pi\end{aligned}$$

Normalize the  $\tilde{v}^k$  in  $H$

$$\overbrace{\tilde{v}^k}^{\lambda_k} (\tilde{\alpha}^{k,T} \tilde{\alpha}^k) = 1$$

Feature extraction

$f \in \mathbb{R}^N$ : test point

$$\begin{aligned}\tilde{v}^{k,T} \tilde{\phi}(f) &= \sum_{i=1}^M \tilde{\alpha}_i^k \tilde{\phi}(x_i)^T \tilde{\phi}(f) \\ &= \sum_{i=1}^M \tilde{\alpha}_i^k \tilde{k}(x_i, f)\end{aligned}$$

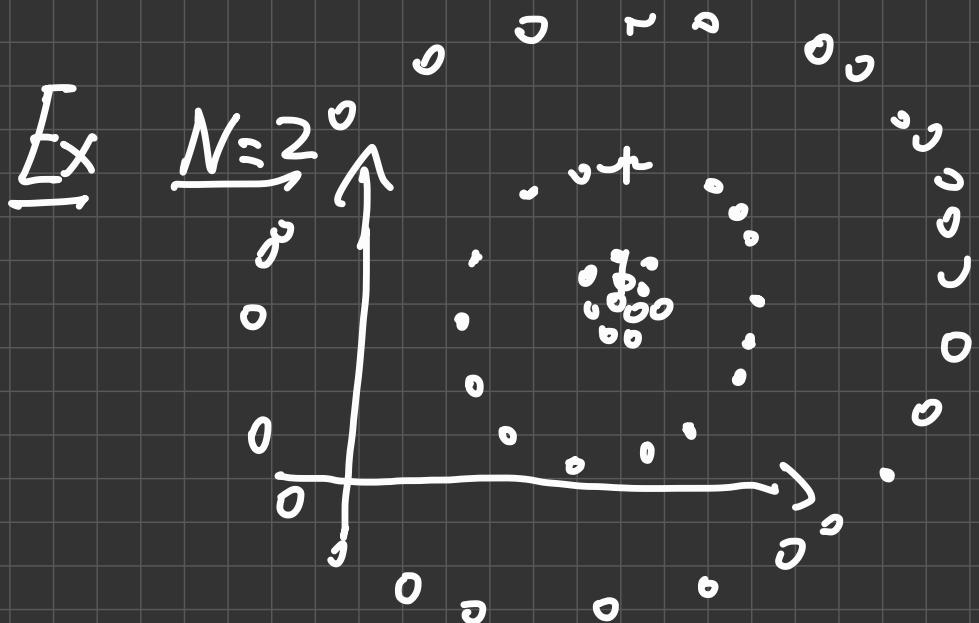
L. test points:  $f_1, \dots, f_L$

$$K_{i,j}^{\text{test}} = \phi(f_i)^T \phi(x_j) = k(f_i, x_j)$$

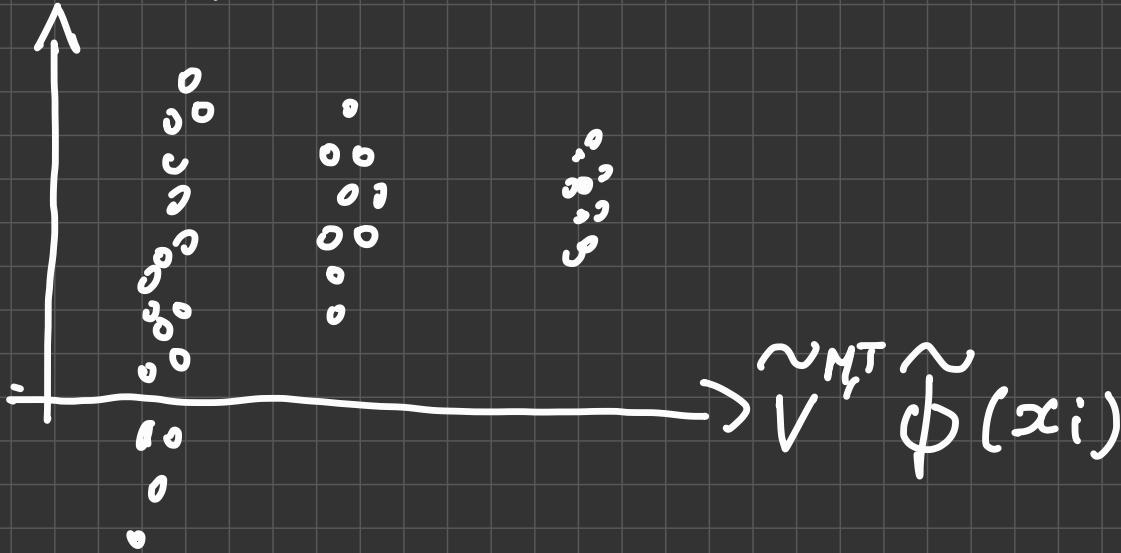
$$\tilde{K}_{i,j}^{\text{test}} = \left( \phi(f_i) - \frac{1}{M} \sum_{m=1}^M \phi(x_m) \right)^T \left( \phi(x_j) - \frac{1}{M} \sum_{n=1}^M \phi(x_n) \right)$$

$$\tilde{K}^{\text{test}} = K^{\text{test}} - \frac{1}{M} \mathbf{1} \mathbf{1}^T - \frac{1}{M} K^{\text{test}} \mathbf{1} \mathbf{1}^T + \frac{\mathbf{1} K \mathbf{1}^T}{M^2}$$

$\mathbf{1} \mathbf{1}^T$ :  $L \times M$  matrix  $\mathbf{1}_{i,j} = 1$



$$\tilde{V}^{M-1, \top} \tilde{\phi}(x_i)$$



Ex Image denoising

$x_i$ : images ( $MN \times T$ )

$x^*$ : noisy image

$$P_d \tilde{\phi}(x^*) = (\tilde{V}^{M-d+1})^\top \tilde{\phi}(x^*) \tilde{V}^{M-d+1} + \dots + (\tilde{V}^M)^\top \tilde{\phi}(x^*) \tilde{V}^M$$

projection of  $\tilde{\phi}(x^*)$  onto

span eigenvectors associated with  
d largest eigenvalues

$x^*$ : noisy image

$y^*$ : denoised image

$$y^* = \underset{y \in X}{\operatorname{arg\,min}} \quad \| \tilde{\phi}(y) - P_{\mathcal{D}} \tilde{\phi}(x^*) \|_H$$

Kernel linear discriminant analysis

Fisher Linear discriminant analysis (LDA)

LDA

classification prob

C classes

Each class has  $N_i$  m-dim samples

$i=1, \dots, C$

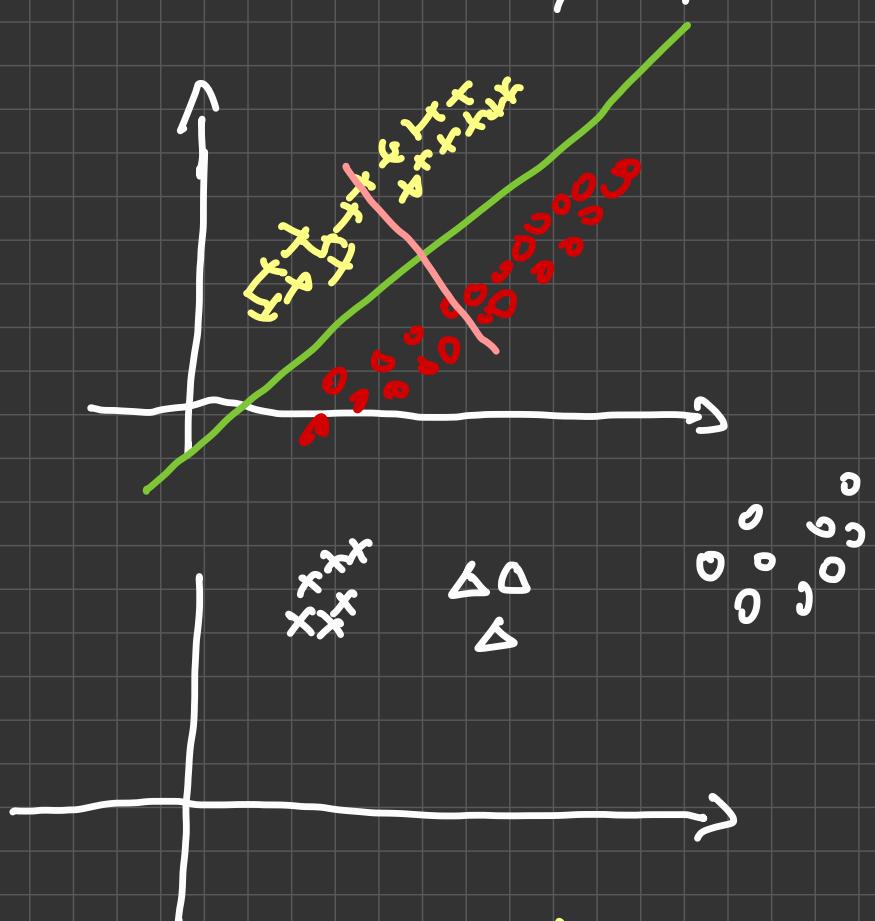
For each class i

samples:  $\{x_1^i, \dots, x_{N_i}^i\}$

$$X = \begin{bmatrix} x_1^1 & \dots & x_{N_1}^1 \\ \vdots & & \vdots \\ x_1^m & \dots & x_{N_m}^m \\ \vdots & & \vdots \\ x_1^c & \dots & x_{N_c}^c \end{bmatrix}$$

$N_1 + \dots + N_c = N$

→ seek to separate classes by projecting  
the  $x_i^j$  onto hyperplane of dim  $C-1$



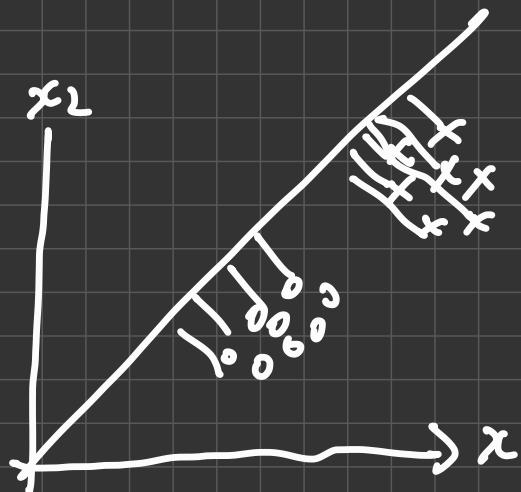
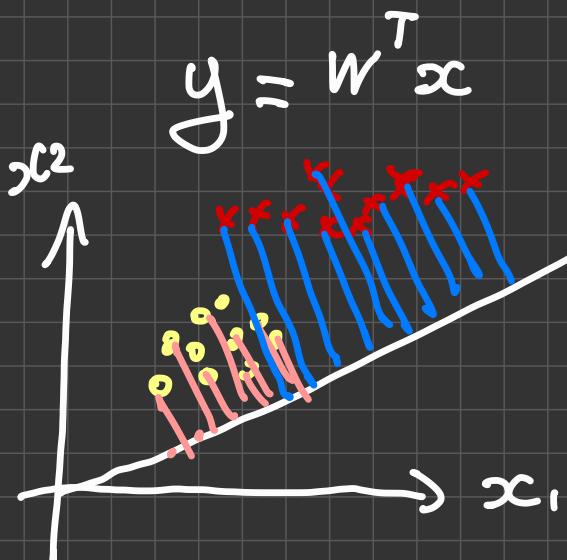
LDA with 2 classes

$$N_1 + N_2 = N \quad m \text{ class samples } \{x_1, \dots, x_N\}$$

$N_1$ : class 1

$N_2$ : 2

seek obtain  $y$  by projecting  $x_i$  onto a line  
( $C-1$  dim space)



Need a good measure of separation

$w_i$ : set of points of class 1

$$\bar{n}_i = \frac{1}{N_i} \sum_{x \in w_i} x$$

$$\tilde{n}_i = w^T \bar{n}_i = \frac{1}{N_i} \sum_{x \in w_i} w^T x$$

idea  $J(w) = |\tilde{n}_1 - \tilde{n}_2|$

$$= |w^\top (\mu_1 - \mu_2)|$$



scatter of class i

$$\tilde{s}_i^2 = \sum_{x \in \omega_i} (w^\top x - \tilde{\mu}_i)^2$$

measure of variability within class i  
of feature projection

$$\tilde{s}_1^2 + \tilde{s}_2^2$$

Quantity to be maximized

$$\mathcal{J}(w) = \frac{|\tilde{\mu}_1 - \tilde{\mu}_2|^2}{\tilde{s}_1^2 + \tilde{s}_2^2}$$

Need to find

$$w^* = \arg \max_w \mathcal{J}(w)$$

Scatter matrix

$$S_i = \sum_{x \in w_i} (x - \mu_i) (x - \mu_i)^T$$

$$S = S_1 + S_2$$

$$\tilde{S}_i^2 = w^T S_i w$$

$$\tilde{S}_1^2 + \tilde{S}_2^2 = w^T S w$$

Between class scatter matrix

$$S_B = (\mu_1 - \mu_2) (\mu_1 - \mu_2)^T$$

$$w^T S_B w = |\tilde{\mu}_1 - \tilde{\mu}_2|^2$$

$$\mathcal{J}(w) = \frac{|\tilde{\mu}_1 - \tilde{\mu}_2|^2}{\tilde{S}_1 + \tilde{S}_2} = \frac{w^T S_B w}{w^T S w}$$

$$\frac{d\mathcal{J}}{dw} = 0 \Leftrightarrow (w^T S_w) S_B w - (w^T S_B w) S_w = 0$$

$$\Leftrightarrow S_B w - \mathcal{J}(w) S_w = 0$$



$$S_B w = \lambda S_w$$

$$\lambda = J(w)$$

↓  
generalized eigenvalue problem

So |

$$w^* = \underset{w}{\operatorname{argmax}} J(w) = \underset{w}{\operatorname{argmax}} \frac{w^T S_B w}{w^T S_w}$$

↓

$$= S^{-1}(\mu_1 - \mu_2)$$

Fisher linear discriminant

LDA with  $C$  classes

$C$  classes

seek  $(C-1)$  projection vectors  $w_i$

$$W = \begin{bmatrix} w_1 & \dots & w_{C-1} \end{bmatrix} \Bigg|_m$$

$\underbrace{\hspace{10em}}$   
 $C-1$

$C-1$  projections:  $y_1, \dots, y_{C-1}$

$$y_i = w_i^T x$$



$$\mathbf{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_{C-1} \end{pmatrix}$$

$$\mathbf{x} = \begin{pmatrix} x_1 \\ \vdots \\ x_m \end{pmatrix}$$

$$x \in \mathbb{R}^m$$
  
$$y \in \mathbb{R}^{C-1}$$

$$y = w^T x$$

$n$  data points  $x^i \in \mathbb{R}^m$

$$\mathbf{X} = \begin{pmatrix} x_1^1 & x_1^2 & & & x_1^n \\ \vdots & \vdots & \ddots & & \vdots \\ x_m^1 & x_m^2 & & & x_m^n \end{pmatrix}$$

$$\mathbf{Y} = \begin{pmatrix} y_1^1 & y_1^2 & & & y_1^n \\ \vdots & \vdots & \ddots & & \vdots \\ y_{C-1}^1 & y_{C-1}^2 & & & y_{C-1}^n \end{pmatrix}$$

$$Y = W^T X$$

$$S = \sum_{i=1}^C S_i$$

$$S_i = \sum_{x \in w_i} (x - \mu_i) (x - \mu_i)^T$$

$$\mu_i = \frac{1}{N_i} \sum_{x \in w_i} x$$

$N_i$ : # of data in class  $i$

$$n = N = \sum_{i=1}^C N_i$$

$$S_B = \sum_{i=1}^C N_i (\mu_i - \mu) (\mu_i - \mu)^T$$

$$\mu = \frac{1}{N} \sum_{i=1}^C \sum_{x \in w_i} x = \frac{1}{N} \sum_{i=1}^C N_i \mu_i$$

$$\tilde{S} = \text{Tr}[W^T S W]$$

$$\tilde{S}_B = \text{Tr}[W^T S_B W]$$

$$\mathcal{J}(W) = \frac{\tilde{S}_B}{\tilde{S}} = \frac{\text{Tr}[W^T S_B W]}{\text{Tr}[W^T S_W W]}$$

need to find  $w_1, \dots, w_{C-1}$  orthonormal  
 that  $\max \mathcal{J}(W)$

$\uparrow \downarrow$   
 $(w_j \perp w_i \forall i \neq j)$

$$\Leftrightarrow S_B w_i = \lambda_i S_W w_i$$

$\downarrow C-1$  largest eigenvalues

$w_i$ :  $(C-1)$  eigenvectors corresponding  
 to the  $C-1$  largest eigenvalues

$$\begin{aligned}
 w_j^T S_B w_i &= \lambda_i w_j^T S_W w_i \xrightarrow{w_j^T S_W w_i = 0} \\
 &= \lambda_i w_j^T S_W w_i
 \end{aligned}$$

# Kernel LDA

2 classes

data in class 1:  $\{x_1^1, \dots, x_{N_1}^1\} = \omega_1$

2:  $\{x_1^2, \dots, x_{N_2}^2\} = \omega_2$

$$\phi: \mathcal{X} \rightarrow \mathcal{H}$$

$\downarrow$   
feature map      feature space

$$m_i^\phi = \frac{1}{N_i} \sum_{j=1}^{N_i} \phi(x_j^i)$$

$$S_B^\phi = (m_1^\phi - m_2^\phi) (m_1^\phi - m_2^\phi)^T$$

$$S^\phi = \sum_{i=1}^2 \sum_{x \in \omega_i} (\phi(x) - m_i^\phi) (\phi(x) - m_i^\phi)^T$$

Need to max.

$$\Sigma(w) = \frac{w^T S_B^\phi w}{w^T S^\phi w} = \langle w, S^\phi w \rangle_{\mathcal{H}}$$

We can find an expansion for  $w$  of the form

$$w = \sum_{i=1}^N \alpha_i \phi(x_i) \quad (N \text{ data points}, \{x_1, \dots, x_N\})$$

kernel  $k(x, x') = \langle \phi(x), \phi(x') \rangle_H$

$$w^\top m_i^\phi = \frac{1}{N_i} \sum_{j=1}^{N_i} \sum_{l=1}^{N_i} \alpha_j k(x_j, x_l)$$

$$= \alpha^\top M_i$$

$$(M_i)_j = \frac{1}{N_i} \sum_{l=1}^{N_i} k(x_j, x_l)$$

$$w^\top S_B^\phi w = \alpha^\top M^B \alpha$$

$$M^B = (M_1 - M_2)(M_1 - M_2)^T$$

Similarly

$$w^\top S_w^\phi w = \alpha^\top M \alpha$$

$$M = \sum_{j=1}^2 K_j \left( I - \frac{1}{N_j} \right) K_j^T$$

I. id. matrix  $(N_j \times N_j)$

$K_j = N \times N_j$  matrix

$$(K_j)_{n,m} = k(x_n, x_m)$$

We can find Fisher's linear discrim. in  $H$  by maximizing

$$\mathcal{J}(\alpha) = \frac{\alpha^T M^B \alpha}{\alpha^T M \alpha}$$

→ solved by finding leading eigenvector

of  $M^B \alpha = \lambda \alpha$

new data point  $x$

$$\begin{aligned} \langle w, \phi(x) \rangle_H &= w^T \phi(x) \\ &= \sum_{i=1}^N \alpha_i k(x_i, x) \end{aligned}$$

In practice: to increase numerical stability

replace  $M$  by  $M + \nu I$

$\downarrow$  id. matrix  
 $\nu > 0$   
small

Rk Kernel LDA  $\Leftrightarrow$  perform Kernel PCA  
then perform LDA  
on resulting coordinates



# Kernel Canonical Correlation Analysis

## Canonical Correlation Analysis

(CCA, Hotelling 1936)

Data  $(X_1, Y_1), \dots, (X_N, Y_N)$

$X_i \in \mathbb{R}^m$  iid  $X_i \sim X$  (r.v.)

$Y_i \in \mathbb{R}^p$  "  $Y_i \sim Y$  (r.v.)

Find  $a \in \mathbb{R}^m, b \in \mathbb{R}^p$  s.t.

the correlation between  $a^\top X$

and  $b^\top Y$  is maximized

$$\rho = \max_{a, b} \text{Cor}[a^\top X, b^\top Y]$$

$$= \max_{a, b} \frac{\text{Cov}[a^\top X, b^\top Y]}{\sqrt{\text{Var}[a^\top X] \text{Var}[b^\top Y]}}$$

Replace unknown distribution of  $X$   
by empirical dist

$$\underbrace{\delta_{x_1} + \dots + \delta_{x_N}}_N$$

Replace  $\mu$   
by  $\bar{y}$

$$\text{of } Y$$

$$\underbrace{\delta_{y_1} + \dots + \delta_{y_N}}_N$$

$$\mu_N^X = \frac{1}{N} \sum_{i=1}^N X_i$$

$$\mu_N^Y = \frac{1}{N} \sum_{i=1}^N Y_i$$

$$C_{XX} = \frac{1}{N} \sum_{i=1}^N (X_i - \mu_N^X) (X_i - \mu_N^X)^T$$

$m \times n$  matrix

$$C_{XY} = \frac{1}{N} \sum_{i=1}^N (X_i - \mu_N^X) (Y_i - \mu_N^Y)^T$$

$m \times p$  matrix

$$C_{YY} = \frac{1}{N} \sum_{i=1}^N (Y_i - \mu_N^Y) (Y_i - \mu_N^Y)^T$$

$$C_{YX} = \frac{1}{N} \sum_{i=1}^N (Y_i - \bar{Y}_N) (X_i - \bar{X}_N)^T$$

Sol. CCA

$$\begin{aligned} * \quad & \max_{\alpha, \beta} \quad \alpha^T C_{XY} \beta \\ \text{s.t.} \quad & \alpha^T C_{XX} \alpha = 1 \quad \beta^T C_{YY} \beta = 1 \end{aligned}$$

Lagrange multiplier method

$\Updownarrow$  generalized eigenpb.

$$\begin{pmatrix} 0 & C_{XY} \\ C_{YX} & 0 \end{pmatrix} \begin{pmatrix} \alpha \\ \beta \end{pmatrix} = \rho \begin{pmatrix} C_{XX} & 0 \\ 0 & C_{YY} \end{pmatrix} \begin{pmatrix} \alpha \\ \beta \end{pmatrix}$$

Solution

$$\alpha = C_{XX}^{-1/2} v_1, \quad \beta = C_{YY}^{-1/2} v_1$$

$v_1$ : left first singular vector of  
the SVD of  $C_{XX}^{-1/2} C_{XY} C_{YY}^{-1/2}$

$v_1$ : right first singular vector of  
the SVD of  $C_{xx}^{-\frac{1}{2}}$   $C_{xy}$   $C_{yy}^{-\frac{1}{2}}$

## Kernel CCA

CCA for feature vectors

$\phi_x(x_1), \dots, \phi_x(x_N) \in \mathcal{H}_x$  (feature space (RKHS))

$\phi_y(y_1), \dots, \phi_y(y_N) \in \mathcal{H}_y$  (RKHS)

$$\max_{\substack{f \in \mathcal{H}_x \\ g \in \mathcal{H}_y}} \frac{\text{Cov}[f(x), g(y)]}{\sqrt{\text{Var}[f(x)]} \sqrt{\text{Var}[g(y)]}}$$

$$\text{Cov}[f(x), g(y)]$$

$$= E_{x,y} \left[ (f(x) - E[f(x)]) (g(y) - E[g(y)]) \right]$$

$$\approx \frac{1}{N} \sum_{i=1}^N \left( f(X_i) - \frac{1}{N} \sum_{j=1}^N f(X_j) \right) \left( g(Y_i) - \frac{1}{N} \sum_{j=1}^N g(Y_j) \right)$$

As in kernel PCA we can assume

$$f = \sum_{i=1}^N \alpha_i \tilde{\phi}_x(X_i) \quad g = \sum_{i=1}^N \beta_i \tilde{\phi}_y(Y_i)$$

$$\tilde{\phi}_x(X_i) = \phi_x(X_i) - \frac{1}{N} \sum_{j=1}^N \phi_x(X_j)$$

$$\tilde{\phi}_y(Y_i) = \phi_y(Y_i) - \frac{1}{N} \sum_{j=1}^N \phi_y(Y_j)$$

$$\max \text{Cov}[f(x), g(y)]$$

$$\frac{\sqrt{\text{Var}[f(x)]}}{\sqrt{\text{Var}[g(y)]}}$$

$$\approx \max_{\substack{f \in \mathcal{H}_x \\ g \in \mathcal{H}_y}} \frac{\sum_{i=1}^N \langle f, \tilde{\phi}_x(X_i) \rangle \langle g, \tilde{\phi}_y(Y_i) \rangle}{\sqrt{\sum_{i=1}^N \langle f, \tilde{\phi}_x(X_i) \rangle^2} \sqrt{\sum_{i=1}^N \langle g, \tilde{\phi}_y(Y_i) \rangle^2}}$$

As in kernel PCA we can assume

$$f = \sum_{i=1}^N \alpha_i \tilde{\phi}_x(x_i) \quad g = \sum_{i=1}^N \beta_i \tilde{\phi}_y(y_i)$$

$\tilde{K}_x$  and  $\tilde{K}_y$ : centered Gram matrices

$$\tilde{K}_{x,i,j} = \langle \tilde{\phi}_x(x_i), \tilde{\phi}_x(x_j) \rangle$$

$$\tilde{K}_{y,i,j} = \langle \tilde{\phi}_y(y_i), \tilde{\phi}_y(y_j) \rangle$$



$$\max_{\alpha \in \mathbb{R}^N} \frac{\alpha^\top \tilde{K}_x \tilde{K}_y \beta}{\sqrt{\alpha^\top \tilde{K}_x \alpha} \sqrt{\beta^\top \tilde{K}_y \beta}}$$

- Regularization

$$\begin{aligned} & \underset{\substack{f \in H_x \\ g \in H_y}}{\max} \sum_{i=1}^N \langle f, \tilde{\Phi}_x(x_i) \rangle \langle g, \tilde{\Phi}_y(y_i) \rangle \\ & \quad \sqrt{\sum_{i=1}^N \langle f, \tilde{\Phi}_x(x_i) \rangle^2 + \epsilon_N \|f\|_{H_x}^2} - \\ & \quad \times \sqrt{\sum_{i=1}^N \langle g, \tilde{\Phi}_y(y_i) \rangle^2 + \epsilon_N \|g\|_{H_y}^2} \end{aligned}$$

Converges as  $N \rightarrow \infty$  (Fukumizu, 2007)

Solution: Generalized eigenproblem

$$\begin{pmatrix} 0 & \tilde{K}_x \tilde{K}_y \\ \tilde{K}_y \tilde{K}_x & 0 \end{pmatrix} \begin{pmatrix} \alpha \\ \beta \end{pmatrix} = \rho \begin{pmatrix} \alpha \\ \beta \end{pmatrix}$$

$$\begin{pmatrix} \tilde{K}_x^2 + \epsilon_N \tilde{K}_x & 0 \\ 0 & \tilde{K}_y^2 + \epsilon_N \tilde{K}_y \end{pmatrix}$$

Application: to image retrieval  
(Haralick, 2004)

$X_i$ : image

$Y_i$ : corresponding text

Idea Use  $d$  eigenvectors  $f_1, \dots, f_d$   
 $g_1, \dots, g_d$

as feature spaces which contain  
the depend. between  $X \& Y$

Given a text  $Y_{\text{new}}$  compute its  
feature vector and the image  $X_i$  whose  
feature maximizes the following inner product

$$X_i \rightarrow \tilde{\phi}_x(X_i) \rightarrow \begin{pmatrix} \langle f_1, \tilde{\phi}_x(X_i) \rangle \\ \vdots \\ \langle f_d, \tilde{\phi}_x(X_i) \rangle \end{pmatrix}. \quad \checkmark$$

$$V = \begin{pmatrix} \langle g_1, \tilde{\phi}_y(y_{\text{new}}) \rangle \\ \vdots \\ \langle g_d, \tilde{\phi}_y(y_{\text{new}}) \rangle \end{pmatrix}$$

$K_x \leftrightarrow$  Gaussian kernel for images

$K_y \leftrightarrow$  Bag of words kernel  
(frequency of words) for texts

## Support Vector Machines

Margin classifier

Training data  $(X_1, Y_1), \dots, (X_N, Y_N)$

$X_i \in \mathbb{R}^d$

$Y_i \in \{-1, 1\}$  (binary labels)

# Linear classifier

$$h(x) = \text{sign}(w^T x + b)$$
$$w \in \mathbb{R}^d, \quad b \in \mathbb{R}$$

Select  $w, b$  so that a new data  
 $x$  is correctly classified

## Hard margin



$$h(x) = \text{sign}(w^T x + b)$$

Assumption: data is linearly separable  
 $(w, b)$  and  $(cw, cb)$  ( $c > 0$ )  
 give the same classifier

Fix the scale

$$\begin{cases} \min (w^T X_i + b) = 1 & \text{if } Y_i = 1 \\ \max (w^T X_i + b) = -1 & \text{if } Y_i = -1 \end{cases}$$

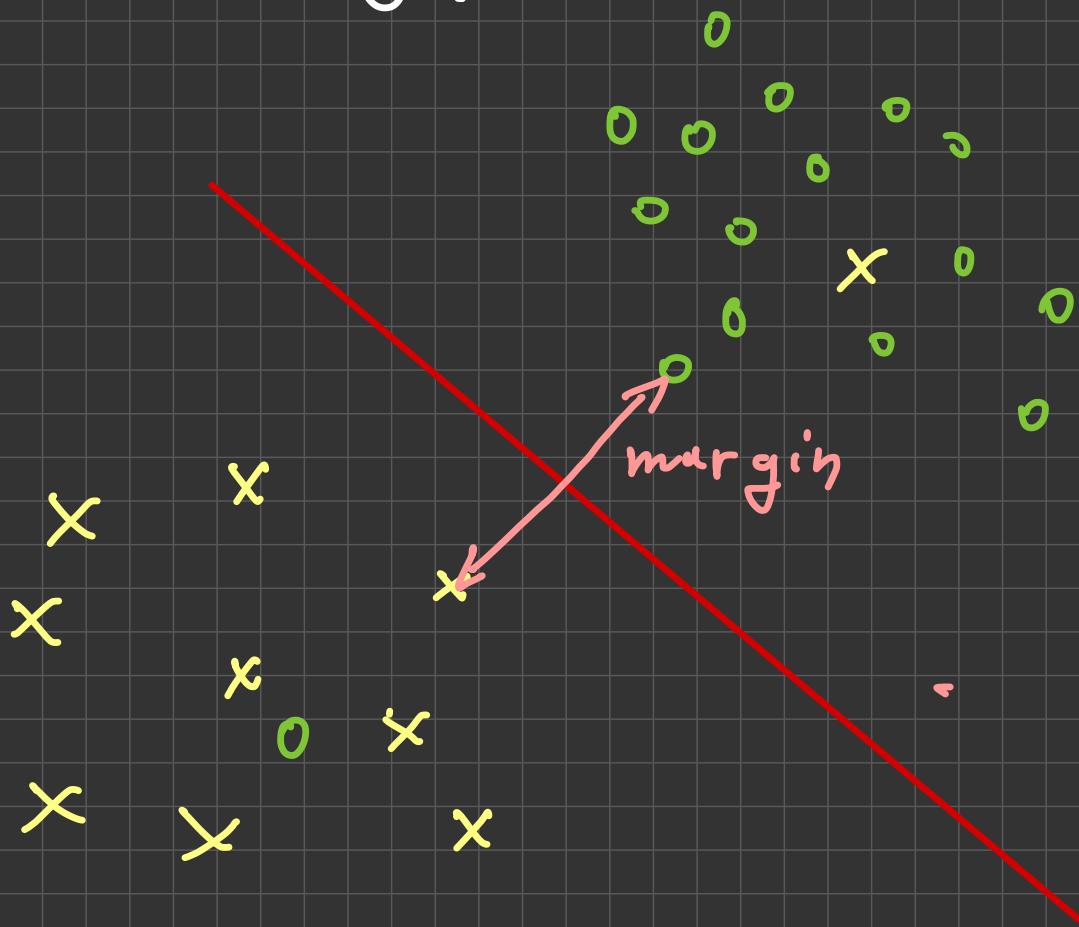
$\Rightarrow$  margin =  $\frac{2}{\|w\|}$

$$\begin{cases} \max \frac{2}{\|w\|} \\ \text{s.t. } w^T X_i + b \geq 1 \quad \text{if } Y_i = 1 \\ \quad w^T X_i + b \leq -1 \quad \text{if } Y_i = -1 \end{cases}$$

$$\begin{cases} \min \frac{\|w\|}{2} \\ \text{s.t. } Y_i (w^T X_i + b) \geq 1 \quad \forall i \end{cases}$$

→ quadratic program

Soft margin

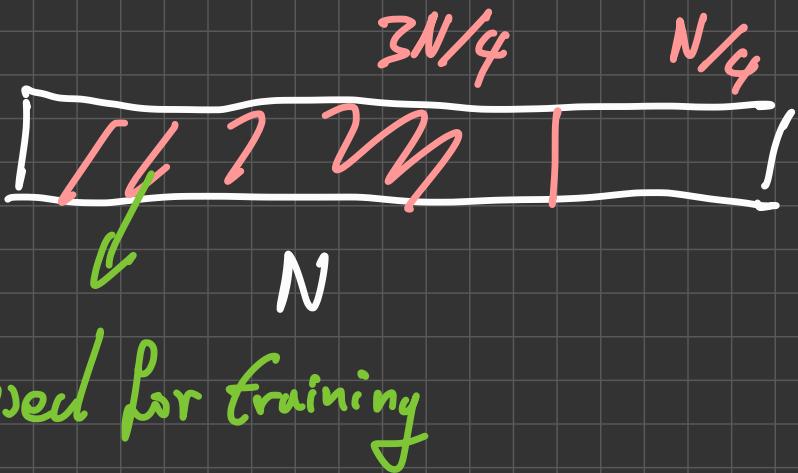


$$\begin{aligned} & \min \frac{\|w\|^2}{2} + C \sum_{i=1}^N \xi_i \\ \text{s.t. } & y_i (w^\top X_i + b) \geq 1 - \xi_i \quad \forall i \\ & \xi_i \geq 0 \end{aligned}$$

Quadratic program

$C \rightarrow$  found via cross-validation

Training data:



Regularization

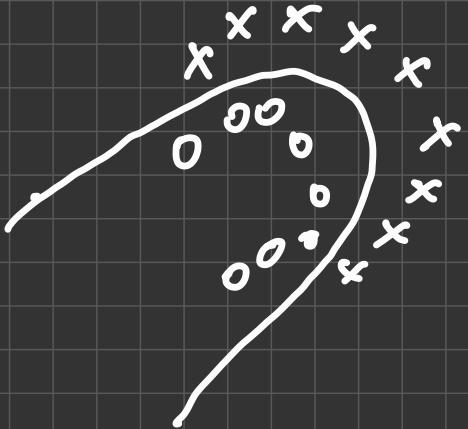
$$(*) \min_{w,b} \sum_{i=1}^n \left( 1 - Y_i (w^\top X_i + b) \right)_+ + \lambda \underbrace{\|w\|^2}_{\text{regul.}}$$

Hinge loss

Proof The cost  $\frac{\|w\|^2}{2} + C \sum_i \xi_i$   
can only increase with  $\xi_i \nearrow$

$$\Rightarrow \text{opt sol } \xi_i^+ = \left( 1 - Y_i (w^\top X_i + b) \right)_+$$

$\Leftrightarrow (*)$



## Kernelization of SVM

Training data  $(X_1, Y_1), \dots, (X_n, Y_n)$

$X_i \in \mathcal{X}$  (arbitrary)

$Y_i \in \{-1, 1\}$

Feature space  $\mathcal{H}$  (RKHS)

map  $\phi: \mathcal{X} \rightarrow \mathcal{H}$

kernel  $K(x, x') = \langle \phi(x), \phi(x') \rangle_{\mathcal{H}}$

Linear classifier on  $\mathcal{H} \Leftrightarrow$  nonlinear classifier in  $\mathcal{X}$

$$h(x) = \text{sign}(\langle \rho, \phi(x) \rangle_{\mathcal{H}} + b)$$

$f \in H, b \in \mathbb{R}$

$$h(x) = \text{sign}(f(x) + b)$$

Nonlinear SVM

$$\min_{f, b} \frac{\|f\|_H^2}{2} + C \sum_{i=1}^n \xi_i$$

$$\xi_i \quad \text{s.t. } Y_i (\langle f, \phi(x_i) \rangle_H + b) \geq 1 - \xi_i$$

$$\xi_i \geq 0$$



$$\begin{aligned} \min_{f, b} & \sum_{i=1}^n (1 - Y_i (f(x_i) + b))_+ \\ & + \lambda \|f\|_H^2 \end{aligned}$$

Representer theorem

$$f = \sum_{i=1}^n w_i \phi(x_i) = \sum_{i=1}^n w_i K(\cdot, x_i)$$

$K(X, X)$ :  $n \times n$  matrix

$$(K(X, X))_{i,j} = k(x_i, x_j) \\ = \langle \phi(x_i), \phi(x_j) \rangle_H$$

$$\Leftrightarrow \min_{w, b, \xi} \frac{1}{2} w^T K(X, X) w + C \sum_{i=1}^n \xi_i$$

s.t.  $y_i ((k(X, X)w)_i + b) \geq 1 - \xi_i$

$$\xi_i \geq 0$$

Lagrange multiplier method

The dual prob is

$$\max_{\alpha} \sum_{i=1}^n \alpha_i - \frac{1}{2} \sum_{i,j=1}^n \alpha_i \alpha_j y_i y_j (K(X, X))_{i,j}$$

s.t.  $0 \leq \alpha_i \leq C$

$$\sum_{i=1}^n y_i \alpha_i = 0$$

$\uparrow$

$\alpha^* = \arg \max$

only data  
with  $0 < \alpha_i^* \leq C$

The classifier

$$f^*(x) + b^* = \sum_{i=1}^n \alpha_i^* y_i k(x, x_i) + b^*$$

$$b^* = y_i - \sum_{j=1}^n \alpha_j^* y_j k(x_i, x_j)$$

$\forall i / 0 < \alpha_i^*$

Indeed the Lagrangian associated with

$$\min_{w, b, \xi} \frac{1}{2} w^\top k(X, X)w + C \sum_{i=1}^n \xi_i$$

s.t.  $y_i ((k(X, X)w)_i + b) \geq 1 - \xi_i$

$\xi_i \geq 0$

is

$$\Lambda(w, b, \xi, \alpha, \beta) =$$

$$\frac{1}{2} w^T k(X, X) w + C \sum_{i=1}^n \xi_i$$

$$+ \sum_i \alpha_i (1 - \xi_i - y_i (k(X, X) w)_i + b))$$

$$+ \sum_i \beta_i (-\xi_i)$$

The optimal  $w$

$$\nabla_w \Lambda = 0 \Leftrightarrow k(X, X) w - \sum_i \alpha_i y_i k(X, X)_{:, i} = 0$$

$$\Rightarrow w_i = \alpha_i y_i$$

$$\nabla_b \Lambda = 0 \Rightarrow \sum_i \alpha_i y_i = 0 \quad 0 \leq \alpha_i \leq C$$

$$\nabla_{\xi_i} \Lambda = 0 \Rightarrow C - \alpha_i - \beta_i = 0$$

at the opt. in  $w, b, \xi_i$

$$\min_{w, b, \xi} \Lambda = -\frac{1}{2} \sum_{i,j} (\alpha_i y_i k(x_i, x_j) \alpha_j y_j) + \sum_i \alpha_i$$

(KKT equations)

Thm The sol. of primal and dual SVN  
is given by

$$(1) 1 - y_i (f^*(x_i) + b^*) - \xi_i^* \leq 0 \quad \forall i$$

$$(2) \xi_i^* \geq 0 \quad \forall i$$

$$(3) 0 \leq \alpha_i^* \leq C \quad \forall i$$

$$(4) \alpha_i^* (1 - y_i (f^*(x_i) + b^*) - \xi_i^*) = 0 \quad \forall i$$

$$(5) \xi_i^* (C - \alpha_i^*) = 0$$

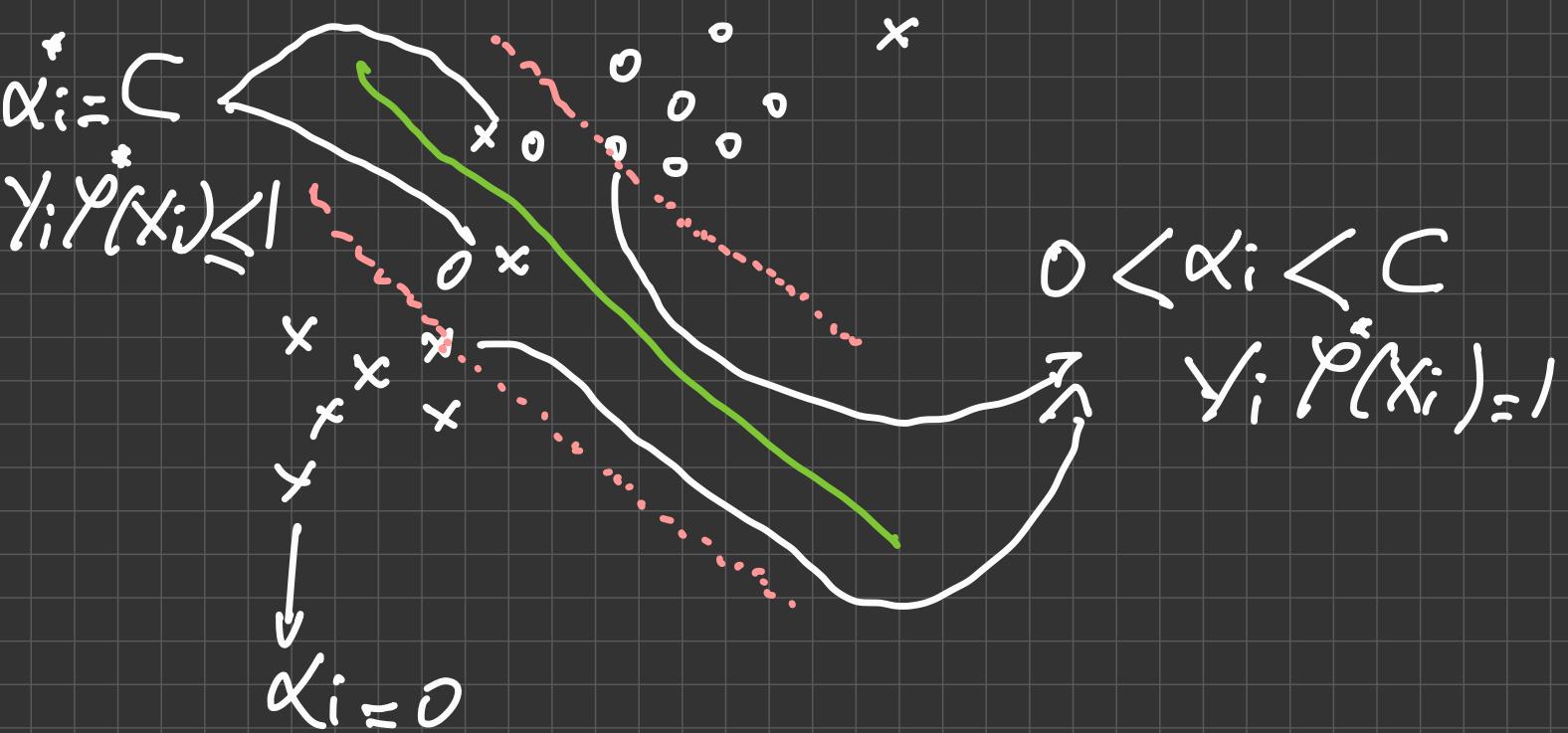
$$(6) \sum_{j=1}^n k(x_i, x_j) w_j^*$$

$$- \sum_{j=1}^n \alpha_j^* y_j k(x_i, x_j) = 0$$

$$(7) \sum_{j=1}^n \alpha_j^* y_j = 0$$

Sparse expression

$$y^*(x) = f^*(x) + b^* = \sum_{x_i: \text{support vectors}} \alpha_i^* y_i k(x, x_i) + b^*$$



# Operator-Valued Kernels

Pb  $f^*: \mathcal{X} \rightarrow \mathcal{Y}$

$f^*$ : unknown

Given  $f^*(X) = Y$  approximate  $f^*$   
 $(X, Y) \in \mathcal{X}^N \times \mathcal{Y}^N$

$$X = (X_1, \dots, X_N)$$

$$Y = (Y_1, \dots, Y_N)$$

$$f^*(X) = (f^*(X_1), \dots, f^*(X_N))$$

$\mathcal{X}$ : arbitrary

$\mathcal{Y}$ : separable Hilbert space

$\dim(\mathcal{Y}) < \infty \rightarrow$  matrix-valued kernel  
(Alvarez et al 2012)

$\dim(\mathcal{Y}) = \infty \rightarrow$  operator-valued kernel

(Kacidi et al., 2016)

$\langle \cdot, \cdot \rangle_y$ : inner product on  $\mathcal{Y}$

$\mathcal{L}(\mathcal{Y})$ : set of bounded linear operators mapping  $\mathcal{Y}$  to  $\mathcal{Y}$

Def We call

$K: \mathcal{X} \times \mathcal{X} \rightarrow \mathcal{L}(\mathcal{Y})$

an operator named kernel if

(i)  $K$  is Hermitian, i.e.

$$K(x, x') = K(x', x)^T \quad \forall x, x' \in \mathcal{X}$$

$A^T$ : adjoint of  $A$  w.r.t.  $\langle \cdot, \cdot \rangle_y$

$$\langle Ay, y' \rangle_y = \langle y, A^T y' \rangle_y$$

$$(ii) \sum_{i,j=1}^m \langle y_i, K(x_i, x_j) y_j \rangle_y > 0$$

$$\forall (x_i, y_i) \in X \times Y, m \in \mathbb{N}$$

$$c(\lim(Y)) = S$$

$K(x_i, x_j)$ :  $S \times S$  matrix

$$K_{ij} = \left( \begin{array}{c} \downarrow \\ \square \\ \downarrow \end{array} \right)$$

$K$  non degenerate  $\Leftrightarrow \sum_{i,j=1}^m \langle y_i, k(x_i, x_j) y_j \rangle \neq 0$

$y_i = 0 \quad \forall i$   
if  $x_i \neq x_j \quad \forall i, j$

$$\underline{\mathbb{E}_X} \quad \lim_{\epsilon \rightarrow 0} \langle Y \rangle = S$$

$$K(x, x') = k(x, x') \underbrace{I_S}_{A \in \mathcal{S}}$$

entangled kernels (2021)

## RKHS

$K$  non-degenerate, locally bounded  
separately continuous



$x \mapsto K(x, x')$   
continuous  
 $\forall x'$

RKHS:  $\mathcal{H}$

$$\mathcal{H} = \left\{ f: X \rightarrow Y \right\}$$

$$= \text{Closure Span } \left\{ z \mapsto K(z, x) y \mid \begin{array}{l} x \in X \\ y \in Y \end{array} \right\}$$

Closure w.r.t. to inner product identified  
by reproducing property

$$\left\langle f, K(\cdot, x)y \right\rangle_{\mathcal{H}} = \left\langle f(x), y \right\rangle_{\mathcal{Y}}$$

## Feature Maps

$\tilde{\mathcal{F}}$ : separable Hilbert space  
 $(\langle \cdot, \cdot \rangle_{\tilde{\mathcal{F}}})$

$\mathcal{L}(Y, \mathcal{F})$ : set of bounded linear operators from  $Y$  to  $\mathcal{F}$

Def We say that  $\tilde{\mathcal{F}}$  and

$\psi: X \rightarrow \mathcal{L}(Y, \mathcal{F})$

are a feature space and a feature map for the kernel  $K$  if

$$\forall x, x', y - y' \quad y^\top \Psi(x) \Psi(x') y'$$

$$y^\top K(x, x') y' = \langle \Psi(x) y, \Psi(x') y' \rangle_{\mathcal{F}}$$

$\Psi^T(x)$ : adjoint of  $\Psi(x) \in \mathcal{L}(\mathcal{F}, Y)$

$$\langle \Psi(x) y, \alpha \rangle_{\mathcal{F}} = \langle y, \Psi^T(x) \alpha \rangle_Y$$

$$\alpha^\top \alpha' = \langle \alpha, \alpha' \rangle_{\mathcal{F}}$$

$$K(x, x') = \Psi^T(x) \Psi(x')$$

$$clim(Y) = s$$

$$clim(\mathcal{F}) = s'$$

$$\Psi: \mathcal{X} \rightarrow \mathcal{L}(Y, \mathcal{F})$$

$\Psi(x)$ :  $s' \times s$  matrix

$$\Psi_{i,j}(x) \quad \begin{array}{l} 1 \leq i \leq s' \\ 1 \leq j \leq s \end{array}$$

$$K(x, x') = \begin{matrix} \Psi^T(x) & \Psi(x') \\ S \times S' & S' \times S \\ \curvearrowright & \curvearrowright \\ S \times S \text{ matrix} \end{matrix}$$

$$K(x, x') = \sum_{i=1}^{S'} \Psi_i^T(x) \Psi_i(x')$$

\$S' = 1\$

$$\lim (\mathcal{F}) = 1$$

$$\Psi: \mathcal{X} \rightarrow \mathcal{L}(Y, \mathbb{R})$$

$$(K(x, x')) = \Psi(\cdot)$$

- For  $\alpha \in \mathcal{F}$

$$\psi^T \alpha: \mathcal{X} \rightarrow Y$$

$x \rightarrow y /$

$$\langle y; y' \rangle_y = \langle y'; \psi^T(x) \alpha \rangle_y \quad \forall y'$$

Thm The RKHS  $\mathcal{H}$  defined

$$\text{by } K(x, x') = \Psi^T(x) \Psi(x')$$

is the linear span of  $\Psi^T \alpha$

over  $\alpha \in \mathcal{F}$  s.t.  $\|\alpha\|_{\mathcal{F}} < \infty$

Furthermore

$$\langle \Psi(\cdot) \alpha, \Psi(\cdot) \alpha' \rangle_{\mathcal{H}}$$

$$= \langle \alpha, \alpha' \rangle_{\mathcal{F}}$$

$$\|\Psi(\cdot) \alpha\|_{\mathcal{H}}^2 = \|\alpha\|_{\mathcal{F}}^2 \quad \forall \alpha, \alpha' \in \mathcal{F}$$

—

$$\Psi: \mathcal{X} \rightarrow \mathcal{L}(Y, \mathcal{F})$$

$$\Psi^T: \mathcal{X} \rightarrow \mathcal{L}(\mathcal{F}, Y)$$

$$\mathcal{H} = \left\{ x \mapsto \Psi^T(x) \alpha \right\}$$

$$\dim(Y) = s$$

$$\dim(\mathcal{F}) = s'$$

$$f_i = \left\{ \left( \sum_j \Psi_{i,j}^\top(x) \alpha_j \right) \right\}_i$$

$$\left( \begin{array}{l} \sum_j \Psi_{i,j}^\top(w) \alpha_j \\ \sum_j \Psi_{s,r,j}^\top(x) \alpha_j \end{array} \right)$$