

# Assignment 2

2017-05-04

## Question 1: Inference for the Poisson parameter $\lambda$ :

- (a) Given a general  $\text{Gamma}(a, b)$  prior, derive the posterior distribution of  $\lambda$ , given data  $(x_1, \dots, x_n)$ , followed by the posterior predictive distribution of  $(z_1, \dots, z_m)$ .

The Likelihood is given by:

$$L(\lambda|x) = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} = \frac{e^{-n\lambda} \lambda^{\sum x_i}}{\prod_{i=1}^n (x_i!)}$$

The Prior is a  $\text{Gamma}(a, b)$ :

$$p(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}, \lambda > 0$$

So the posterior is given by:

$$p(\lambda|x) \propto \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\sum x_i + \alpha - 1} e^{-(n+\beta)\lambda}, \lambda > 0,$$

which is equivalent to a  $\text{Gamma}(\sum x_i + \alpha, n + \beta)$ , or  $\text{Gamma}(n\bar{x} + \alpha, n + \beta)$ .

The *Posterior Predictive Distribution* for a Poisson parameter  $\lambda$  with data  $(z_1, \dots, z_m)$  is given by:

$$\begin{aligned} p(z|X) &= \int_0^\infty p(z|\lambda) p(\lambda|X) d\lambda \\ &= \int_0^\infty \text{Poisson}(z|\lambda) \cdot \text{Gamma}(\sum x_i + \alpha, n + \beta) d\lambda \\ &= \int_0^\infty \frac{e^{-\lambda} \lambda^z}{z!} \cdot \frac{(n + \beta)^{\sum x_i + \alpha}}{\Gamma(\sum x_i + \alpha)} \lambda^{\sum x_i + \alpha - 1} e^{-(n+\beta)\lambda} d\lambda \\ &= \frac{(n + \beta)^{\sum x_i + \alpha}}{\Gamma(\sum x_i + \alpha) \Gamma(z + 1)} \int_0^\infty \lambda^{z + \sum x_i + \alpha - 1} e^{-(n+\beta+1)\lambda} d\lambda \\ &= \frac{(n + \beta)^{\sum x_i + \alpha}}{\Gamma(\sum x_i + \alpha) \Gamma(z + 1)} \cdot \frac{\Gamma(z + \sum x_i + \alpha)}{(n + \beta + 1)^{z + \sum x_i + \alpha}} \\ &= \frac{\Gamma(z + \sum x_i + \alpha)}{\Gamma(\sum x_i + \alpha) \Gamma(z + 1)} \left( \frac{n + \beta}{n + \beta + 1} \right)^{\sum x_i + \alpha} \left( \frac{1}{n + \beta + 1} \right)^z \end{aligned}$$

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# Gelman, q. 13 page 69-70
# Poisson with gamma prior
deaths <- c(734,516,754,877,814,362,764,809,223,1066)
lambda.range <- seq(0,3000)
# prior
prior <- dgamma(lambda.range, shape=7, scale=1/.01)
# log likelihood (likelihood values are too large)
log.likelihood <- sum(deaths)*log(lambda.range)-length(deaths)*lambda.range
# posterior
alpha.star <- 7+sum(deaths)
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beta.star <- .01+length(deaths)
post <- dgamma(lambda.range, shape=alpha.star, scale=1/beta.star)
plot(lambda.range, prior, type = "l", lwd = 1.5, xlab = expression(lambda),
ylab = "Density", cex.axis = 1.3, cex.lab = 1.3, bty = "L", ylim = c(0,0.05))
lines(lambda.range, post)
par(new = TRUE)
plot(lambda.range, log.likelihood, type = "l", lty = 2, lwd = 1.5,
axes = FALSE, xlab = "", ylab = "")
axis(side = 4, cex.axis = 1.3)

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- (b) Assuming Gamma(a,b) priors for two Poisson parameters  $\lambda_1$  and  $\lambda_2$ , derive the posterior for  $\phi = \lambda_1/\lambda_2$ . (Hint: use nuisance parameter  $\mu = \lambda_2$ ).

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- (c) Derive the Jeffreys prior for  $(\phi, \mu)$ , and the corresponding marginal posterior for  $\phi$ .
- (d) Derive the reference prior for  $(\phi, \mu)$ , and the corresponding marginal posterior for  $\phi$ .
- (e) Based on (c), consider the posterior for  $\phi$  based on uniform priors for  $\lambda_1$  and  $\lambda_2$ , and comment on when inference based on this posterior could be quite different from that based on the reference posterior from (e). Give a data example, in terms of resulting intervals. (Hint: the above posteriors are related to known pdfs, and by transformation may be simplified even further.).

## Question 2: Inference for variance components:

- (a) Use e.g. SAS (PROC VARCOMP) to perform a classical analysis of the data in Table 5.1.4 of Box & Tiao (1973), based on finding point estimates only.
- (b) Use WinBUGS for a Bayesian analysis of (a), and find reasonable point and interval estimates for  $\sigma_1^2$  and  $\sigma_2^2$ . Include graphs, including one of the joint posterior. [6 marks]
- (c) Box & Tiao also studied a 3-component model (Table 5.3.1).
  - i. Derive central credible intervals, for the 3 individual components, based on Table 5.3.3.
  - ii. Use WinBUGS to do the same, and include graphs.
  - iii. While not going as far as Box & Tiao's Figure 5.3.2, produce a graph of the joint posterior of  $\sigma_2^2$  and  $\sigma_3^2$ , and one of  $\sigma_2^2/\sigma_3^2$
- (d) Box, Hunter & Hunter (1976, Chapter 17.3) studied a pigment paste example with three components, focusing on point estimates only. Use WinBUGS again to perform a Bayesian analysis. Include graphs.