

# Summary for Hybrid Stochastic Simulation

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Here we summarize some theoretical results from the literature about Hybrid Stochastic Systems and approximation schemes for such systems. We mostly follow the notation in [1] and [3].

## 1 Description of a biochemical stochastic reaction network

We assume a stochastic reaction network consisting of

- $s_0$  species  $S_i$  with copy numbers  $x_i$ , with  $i = 1, \dots, s_0$
- $r_0$  reactions with corresponding stoichiometry vectors  $\xi_k = \nu'_k - \nu_k$  and reaction rates  $\lambda_k(x)$  with  $k = 1, \dots, r_0$

This network can be written as a Markov Jump Process in the random-time-change representation

$$X(t) = x(0) + \sum_{k=1}^{r_0} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (\nu'_k - \nu_k)$$

whereas the  $Y_k$  are independent unit Poisson processes.

For the following we assume a decomposition  $(x_C, x_D)$  of the  $s_0$  species, such that  $x_C$  represents species that are large, scaling with a parameter  $N$ , and  $x_D$  represents species that are of order  $O(1)$ . We define  $\psi_i^C = \frac{1}{N} x_i^C$ .

Additionally we partition the  $r_0$  reactions into  $\mathcal{R}_C, \mathcal{R}_{DC}, \mathcal{R}_D$ , such that

- $\mathcal{R}_C$  corresponds to reactions that involve only species from  $x_C$  and whose reaction rate only depends on  $x_C$
- $\mathcal{R}_D$  corresponds to reactions that involve only species from  $x_D$  and whose reaction rate only depends on  $x_D$
- $\mathcal{R}_{DC}$  corresponds to reactions not belonging to  $\mathcal{R}_C$  or  $\mathcal{R}_D$

## 2 Convergence of Markov Jump Processes to Piecewise Deterministic Markov Processes

The following theorem states the convergence of a partially scaled Markov Jump Process into a PDMP (Piecewise Deterministic Markov Process) [2].

For  $r \in \mathcal{R}_C$  define  $\tilde{\lambda}_r = \frac{1}{N}\lambda_r$  (we can implicitly assume that  $\lambda_r$  is large and scales with  $N$  for  $r \in \mathcal{R}_C$  because otherwise these reactions will not appear in the limit process anyway).

Define  $S_1 \subseteq \mathcal{R}_{DC}$  such that for a reaction  $r \in S_1$  the rate  $\lambda_r$  is large and scales with  $N$  and  $\xi_r^D = 0$ . For each  $r \in S_1$  define  $\tilde{\lambda}_r = \frac{1}{N}\lambda_r$ .

Define  $S_2 \subseteq \mathcal{R}_{DC}$  such that for a reaction  $r \in S_2$  the stoichiometry  $\xi_r^C$  is large and scales with  $N$ . For each  $r \in S_1$  define  $\tilde{\xi}_r^C = \frac{1}{N}\xi_r^C$ .

Define  $S = S_1 \cup S_2$ .

Then we can write the random-time-change representation of the stochastic process corresponding to this reaction network as

$$\begin{aligned}
X^N(t) &= (\Psi_C^N, X_D^N)(t) = x(0) \\
&+ \sum_{k \in \mathcal{R}_C} Y_k \left( \int_0^t \tilde{\lambda}_k(\Psi_C(s)) ds \right) \left( \frac{1}{N}\xi_k^C, 0 \right) \\
&+ \sum_{k \in S_1} Y_k \left( \int_0^t N\tilde{\lambda}_k(X(s)) ds \right) \left( \frac{1}{N}\xi_k^C, 0 \right) \\
&+ \sum_{k \in S_2} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) \left( \tilde{\xi}_k^C, \xi_k^D \right) \\
&+ \sum_{k \in \mathcal{R}_{DC \setminus S}} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) \left( \frac{1}{N}\xi_k^C, \xi_k^D \right) \\
&+ \sum_{k \in \mathcal{R}_D} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (0, \xi_k^D)
\end{aligned}$$

**Theorem 1.** *The limit process of  $X^N(t)$  for  $N \rightarrow \infty$  is*

$$\begin{aligned}
X(t) &\equiv \lim_{N \rightarrow \infty} X^N(t) = (\Psi_C(t), X_D(t)) \\
&= x(0) \\
&+ \sum_{k \in \mathcal{R}_C} \int_0^t \tilde{\lambda}_k(\Psi_C(s)) ds (\xi_k^C, 0) \\
&+ \sum_{k \in S_1} \int_0^t \tilde{\lambda}_k(X(s)) ds (\xi_k^C, 0) \\
&+ \sum_{k \in S_2} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (\tilde{\xi}_k^C, \xi_k^D) \\
&+ \sum_{k \in \mathcal{R}_{DC} \setminus S} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (0, \xi_k^D) \\
&+ \sum_{k \in \mathcal{R}_D} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (0, \xi_k^D)
\end{aligned}$$

which is equivalent to a PDMP.

See [1] for a rigorous proof.

One should note that the distinction between  $\mathcal{R}_C$  and  $S_1$  is just a formality (the dependence of the  $\tilde{\lambda}_r$ ) so we can define  $S_3 = \mathcal{R}_C \cup S_1$  and similarly  $S_4 = \mathcal{R}_D \cup (\mathcal{R}_{DC} \setminus S)$  and write the limit process as

$$\begin{aligned}
X(t) &\equiv \lim_{N \rightarrow \infty} X^N(t) = (\Psi_C(t), X_D(t)) \\
&= x(0) \\
&+ \sum_{k \in S_3} \int_0^t \tilde{\lambda}_k(X(s)) ds (\xi_k^C, 0) \\
&+ \sum_{k \in S_2} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (\tilde{\xi}_k^C, \xi_k^D) \\
&+ \sum_{k \in S_4} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (0, \xi_k^D)
\end{aligned}$$

### 3 Averaging of some discrete variables

Define  $S_5 \subseteq \mathcal{R}_{DC}$  such that for a reaction  $r \in S_5$  the rate  $\lambda_r$  is large and scales with  $N$  and  $\xi_r^D \neq 0$ . For each  $r \in S_5$  define  $\tilde{\lambda}_r = \frac{1}{N} \lambda_r$ . Assume a decomposition  $(x_D^1, x_D^2)$  of the  $x_D$  such that  $\xi_r^{D,1} = 0$  and  $\xi_r^{D,2} \neq 0$  (note that  $x_D^1$  could be empty).

Define  $S = S_1 \cup S_2 \cup S_5$ .

Then the random-time-change representation of the stochastic process can be written as

$$\begin{aligned}
X^N(t) = (\Psi_C^N, X_D^N)(t) = x(0) & \\
& + \sum_{k \in \mathcal{R}_C} Y_k \left( \int_0^t \tilde{\lambda}_k(\Psi_C(s)) ds \right) \left( \frac{1}{N} \xi_k^C, 0, 0 \right) \\
& + \sum_{k \in S_1} Y_k \left( \int_0^t N \tilde{\lambda}_k(X(s)) ds \right) \left( \frac{1}{N} \xi_k^C, 0, 0 \right) \\
& + \sum_{k \in S_5} Y_k \left( \int_0^t N \tilde{\lambda}_k(X(s)) ds \right) \left( \tilde{\xi}_k^C, 0, \xi_k^{D,2} \right) \\
& + \sum_{k \in S_2} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) \left( \tilde{\xi}_k^C, \xi_k^{D,1}, \xi_k^{D,2} \right) \\
& + \sum_{k \in \mathcal{R}_{DC \setminus S}} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) \left( \frac{1}{N} \xi_k^C, \xi_k^{D,1}, \xi_k^{D,2} \right) \\
& + \sum_{k \in \mathcal{R}_D} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) \left( 0, \xi_k^{D,1}, \xi_k^{D,2} \right)
\end{aligned}$$

If  $x_D^2$  only takes a finite number of values (i.e. there exists a finite set  $K$  such that  $X_D^2 \in K$ ) then the process  $X_D^2$  given  $\psi_C, x_D^1$  is ergodic (every state can be reached from a given state).

Define

$$\bar{\lambda}_r(\psi_C, x_D^1) = \int_K \tilde{\lambda}_r(\psi_C, x_D^1, x_D^2) \nu_{\Psi_C, x_D^1}(dx_D^2)$$

for  $r \in S_5$  where  $\nu_{\Psi_C, x_D^1}$  is the unique invariant measure of the ergodic process  $X_D^2$  given  $\psi_C, x_D^1$ .

**Theorem 2.** *The limit process of  $X^N(t)$  for  $N \rightarrow \infty$  is*

$$\begin{aligned}
X(t) &\equiv \lim_{N \rightarrow \infty} X^N(t) = (\Psi_C(t), X_D(t)) \\
&= x(0) \\
&+ \sum_{k \in \mathcal{R}_C} \int_0^t \tilde{\lambda}_k(\Psi_C(s)) ds (\xi_k^C, 0, 0) \\
&+ \sum_{k \in S_1} \int_0^t \tilde{\lambda}_k(X(s)) ds (\xi_k^C, 0, 0) \\
&+ \sum_{k \in S_5} \int_0^t \tilde{\lambda}_k(X(s)) ds (\xi_k^C, 0, 0) \\
&+ \sum_{k \in S_2} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (\tilde{\xi}_k^C, \xi_k^{D,1}, \xi_k^{D,2}) \\
&+ \sum_{k \in \mathcal{R}_{DC \setminus S}} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (0, \xi_k^{D,1}, \xi_k^{D,2}) \\
&+ \sum_{k \in \mathcal{R}_D} Y_k \left( \int_0^t \lambda_k(X(s)) ds \right) (0, \xi_k^{D,1}, \xi_k^{D,2})
\end{aligned}$$

which is again equivalent to a PDMP.

## 4 Multiscale approximations of stochastic reaction networks

For the following we limit ourselves to four types of reactions

$\lambda_k$	Reaction	$\nu_k$
$\kappa'_k$	$\emptyset \rightarrow stuff$	0
$\kappa'_k x_i$	$S_i \rightarrow stuff$	$e_i$
$\kappa'_k V^{-1} x_i (x_i - 1)$	$2S_i \rightarrow stuff$	$2e_i$
$\kappa'_k V^{-1} x_i x_j$	$S_i + S_j \rightarrow stuff$	$e_i + e_j$

whereas  $V$  corresponds to the volume of the system and the  $\kappa'_k$  correspond to molecular reaction rates.

Define

- $z_i^N(t) = N^{-\alpha_i} x_i^N(t)$  with  $\alpha_i \geq 0$  for  $i = 1, \dots, s_0$
- $\kappa_k = N^{-\beta_k} \kappa'_k$  for  $k = 1, \dots, r_0$

with a parameter  $N$  that is assumed to be large.

The rates of the four types of reactions can now be written as

Reaction	$\lambda_k^N$
$\emptyset \rightarrow stuff$	$N^{\beta_k} \kappa_k$
$S_i \rightarrow stuff$	$N^{\beta_k + \alpha_i} \kappa_k z_i$
$2S_i \rightarrow stuff$	$N^{\beta_k + 2\alpha_i} \kappa_k V^{-1} z_i (z_i - N^{-\alpha_i})$
$S_i + S_j \rightarrow stuff$	$N^{\beta_k + \alpha_i + \alpha_j} \kappa_k V^{-1} z_i z_j$

So we have

$$\lambda_k(x) (\nu'_k - \nu_k) = N^{\beta_k + \nu_k \cdot \alpha} \lambda_k^N(z) (\nu'_k - \nu_k)$$

whereas  $\alpha = (\alpha_1, \dots, \alpha_{s_0})^T$ . With this we can write the corresponding multiscale Markov Jump Process as

$$Z_i^N(t) = z(0) + \sum_{k=1}^{r_0} N^{-\alpha_i} Y_k \left( \int_0^t N^{\beta_k + \nu_k \cdot \alpha} \lambda_k^N(Z_i^N(s)) ds \right) (\nu'_{ik} - \nu_{ik})$$

The  $\alpha_i$  and  $\beta_i$  should be chosen in such a way, that the  $z_i$  and  $\kappa_k$  are of order 1 in a loose sense.

In addition we introduce a further scaling parameter  $\gamma$  for the time-scale of the process and write

$$\begin{aligned} Z_i^{N,\gamma}(t) &= Z_i^N(N^\gamma t) \\ &= z(0) + \sum_{k=1}^{r_0} N^{-\alpha_i} Y_k \left( \int_0^t N^{\gamma + \beta_k + \nu_k \cdot \alpha} \lambda_k^N(Z_i^{N,\gamma}(s)) ds \right) (\nu'_{ik} - \nu_{ik}) \end{aligned}$$

See [3] for a more elaborate discussion.

## 5 Constraints for selection of the scaling parameters

In [3] Kang et al. derive constraints for the selection of the scaling parameters  $\alpha_i, \beta_i, \gamma$  that we recapitulate here.

**Condition 1.** Define  $\Gamma_i^+ = \{k : \nu'_{ik} > \nu_{ik}\}$  and  $\Gamma_i^- = \{k : \nu'_{ik} < \nu_{ik}\}$ . Then the condition

$$\max_{k \in \Gamma_i^-} (\beta_k + \nu_k \cdot \alpha) = \max_{k \in \Gamma_i^+} (\beta_k + \nu_k \cdot \alpha)$$

or

$$\gamma \leq \alpha_i - \max_{k \in \Gamma_i^+ \cup \Gamma_i^-} (\beta_k + \nu_k \cdot \alpha)$$

is called the **species balance condition** and ensures that no species explodes or is driven to zero.

We can generalize this condition to linear combinations  $\Theta \cdot z$  of species with  $\Theta \in [0, \infty)^{s_0}$ .

**Condition 2.** Define  $\Gamma_{\Theta}^+ = \{k : \Theta \cdot \xi_k > 0\}$  and  $\Gamma_{\Theta}^- = \{k : \Theta \cdot \xi_k < 0\}$ . Then the condition

$$\max_{k \in \Gamma_{\Theta}^-} (\beta_k + \nu_k \cdot \alpha) = \max_{k \in \Gamma_{\Theta}^+} (\beta_k + \nu_k \cdot \alpha) \quad (1)$$

or

$$\gamma \leq \gamma_i \equiv \max_{i: \Theta_i > 0} \alpha_i - \max_{k \in \Gamma_{\Theta}^+ \cup \Gamma_{\Theta}^-} (\beta_k + \nu_k \cdot \alpha) \quad (2)$$

is called the **collective species balance condition**. It ensures that no linear combination of species explodes or is driven to zero.

Checking this condition is not trivial but in [3] Kang et al. show a significant simplification to check for the **collective species balance condition**:

**Lemma 1.** Let  $G$  be a directed graph where the nodes correspond to the species and a directed edge from  $S_i \rightarrow S_j$  is drawn if there is a reaction consuming  $S_i$  and producing  $S_j$ .

Let  $G_j$  be maximal strongly connected subgraphs such that  $G = \cup_j G_j$  (this decomposition is unique).

Let  $\gamma \in \mathbb{R}$  be fixed. If Condition 2 holds for each  $\Theta \in [0, \infty)^{s_0}$  with  $\text{supp}(\Theta) \subset G_j$  for all maximal strongly connected subgraphs  $G_j$ , then Condition 2 holds for all  $\Theta \in [0, \infty)^{s_0}$ .

This reduces checking Condition 2 for all  $\Theta \in [0, \infty)^{s_0}$  to checking Condition 2 for all  $\Theta$  with  $\text{supp}(\Theta) \subset G_j$  for each of the subgraphs  $G_j$  separately.

**Lemma 2.** Kang et al. continue to show that Condition 2 also holds for  $c_1\Theta^1 + c_2\Theta^2$  for all  $c_1, c_2 > 0$  if either

$$(2) \text{ holds for } \Theta^1 \text{ or } \Theta^2$$

or

$$\begin{aligned} \max_{k \in \Gamma_{\Theta^1}^-} (\beta_k + \nu_k \cdot \alpha) &= \max_{k \in \Gamma_{\Theta^1}^+} (\beta_k + \nu_k \cdot \alpha) \\ &\neq \max_{k \in \Gamma_{\Theta^2}^-} (\beta_k + \nu_k \cdot \alpha) = \max_{k \in \Gamma_{\Theta^2}^+} (\beta_k + \nu_k \cdot \alpha) \end{aligned}$$

**Lemma 3.** If Lemma 2 does not apply but

$$\Gamma_{\Theta^1}^+ \cap \Gamma_{\Theta^2}^- = \emptyset$$

or

$$\Gamma_{\Theta^1}^- \cap \Gamma_{\Theta^2}^+ = \emptyset$$

(i.e. the support of  $\Theta^1$  produces no species that are consumed in the support of  $\Theta^2$  and vice versa), then Condition 2 holds for  $c_1\Theta^1 + c_2\Theta^2$  for all  $c_1, c_2 > 0$ .

**Lemma 4.** *If Lemma 2 does not apply and  $\Gamma_{\Theta^1}^+ \cap \Gamma_{\Theta^2}^- \neq \emptyset$  and  $\Gamma_{\Theta^1}^- \cap \Gamma_{\Theta^2}^+ \neq \emptyset$  suppose that Condition 2 also holds for  $\Theta^1 - \frac{\Theta^1 \cdot \xi_k}{\Theta^2 \cdot \xi_k} \Theta^2$  for all  $k \in (\Gamma_{\Theta^1}^+ \cap \Gamma_{\Theta^2}^-) \cup (\Gamma_{\Theta^1}^- \cap \Gamma_{\Theta^2}^+)$ . Then Condition 2 also holds for  $c_1 \Theta^1 + c_2 \Theta^2$  for all  $c_1, c_2 > 0$ .*

The only case that is left now is that

(2) doesn't hold for both  $\Theta^1$  and  $\Theta^2$  and  $\Gamma_{\Theta^1}^+ \cap \Gamma_{\Theta^2}^- \neq \emptyset$  and  $\Gamma_{\Theta^1}^- \cap \Gamma_{\Theta^2}^+ \neq \emptyset$  and Condition 2 doesn't hold for  $\Theta^1 - \frac{\Theta^1 \cdot \xi_k}{\Theta^2 \cdot \xi_k} \Theta^2$  and

$$\begin{aligned} \max_{k \in \Gamma_{\Theta^1}^-} (\beta_k + \nu_k \cdot \alpha) &= \max_{k \in \Gamma_{\Theta^1}^+} (\beta_k + \nu_k \cdot \alpha) \\ &= \max_{k \in \Gamma_{\Theta^2}^-} (\beta_k + \nu_k \cdot \alpha) = \max_{k \in \Gamma_{\Theta^2}^+} (\beta_k + \nu_k \cdot \alpha) \end{aligned}$$

all hold. We call this case *non-reducible collective species balance*. In this case Condition 2 has to be checked directly (of course this only needs to be done for all  $\Theta \in [0, \infty)^{s_0}$  with support in the subgraphs  $G_j$  with a *non-reducible collective species balance*).

In all other cases it is now sufficient to check Condition 2 on a set  $\Theta^{j,k}$  spanning the subgraph  $G_j$  for all  $j$ .

**Theorem 3.** *For completeness we state that if*

$$\gamma = \min_i \gamma_i = \min_i \left( \alpha_i - \max_{k \in \Gamma_{\Theta}^+ \cup \Gamma_{\Theta}^-} (\beta_k + \nu_k \cdot \alpha) \right)$$

*then each term*

$$N^{-\alpha_i} Y_k \left( \int_0^t N^{\gamma + \beta_k + \nu_k \cdot \alpha} \lambda_k^N \left( Z_i^{N, \gamma}(s) \right) ds \right)$$

*converges to*

$$\begin{cases} 0 & \text{for } \alpha_i > \gamma + \beta_k + \nu_k \cdot \alpha \\ Y_k \left( \int_0^t \lambda_k^N \left( Z_i^{N, \gamma}(s) \right) ds \right) & \text{for } \alpha_i = \gamma + \beta_k + \nu_k \cdot \alpha = 0 \\ \int_0^t \lambda_k^N \left( Z_i^{N, \gamma}(s) \right) ds & \text{for } \alpha_i = \gamma + \beta_k + \nu_k \cdot \alpha \neq 0 \end{cases}$$

*for  $N \rightarrow \infty$  which represents a PDMP of the multiscale network.*

## 6 Combination of convergence results and multi-scale parametrization

To combine the convergence results from 2 with the multiscale Parametrization we note that the membership of a species  $x_i$  in  $x_C, x_D$  depends on  $\alpha_i$ :

$$\begin{cases} x_i \text{ belongs to } x_C & \text{for } \alpha_i > 0 \\ x_i \text{ belongs to } x_D & \text{for } \alpha_i = 0 \end{cases}$$



We also note that  $r \in S_1$  if  $\gamma + \beta_r + \nu_r \cdot \alpha > 0$  and  $\xi_r^D = 0$  (if  $\xi_r^D \neq 0$  one can still take the approach in 3).

We can then use Theorem 1 or Theorem 2 to derive the PDMP that corresponds to the limiting process of the multiscale network.

## References

- [1] Alina Crudu, Arnaud Debussche, Aurélie Muller, and Ovidiu Radulescu. Convergence of stochastic gene networks to hybrid piecewise deterministic processes. *The Annals of Applied Probability*, 22(5):1822–1859, 2012.
- [2] Mark HA Davis. Piecewise-deterministic markov processes: A general class of non-diffusion stochastic models. *Journal of the Royal Statistical Society. Series B (Methodological)*, pages 353–388, 1984.
- [3] Hye-Won Kang and Thomas G. Kurtz. Separation of time-scales and model reduction for stochastic reaction networks. *The Annals of Applied Probability*, 23(2):529–583, 2013.