

Random Variables and Random Process/Digital Communications/Linear Algebra Review (6hrs)

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Review of Random Variable (R.V.)

- A probability space is a collection (Ω, F, P) .
 Ω : sample space containing all possible outcomes.
 F : event space, subspace of Ω with properties.
 P : probability measure.
Random variable is a function
(outcomes mapping to real numbers) $\Omega \rightarrow R$
- Data channel noise interference are all R.V.s

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Properties of Prob.

- $P(\phi) = 0, P(\Omega) = 1$.
- $0 \leq P(A) \leq 1, \forall A \in \Omega$.
- If A_1, A_2, A_3, \dots are disjoint, then
$$P\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} P(A_i) .$$

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Properties of R.V.

- Expectation is linear.
- $\text{var}(ax + b) = a^2 \text{var}(x)$.
- $\text{cov}(x, y) = E[xy] - E[x]E[y]$.
- Covariance=0=> uncorrelated

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Discrete R.V.

- $\Omega \rightarrow 1, 2, 3, \dots$

- Ex. Binomial distribution

$$P(n = k) = \binom{n}{k} P^k (1 - p)^{n-k}$$

$k = 0, 1, 2, \dots, n$; n : the number of trials.

Bernoulli : special case, $n=1$.

- Ex. Poisson distribution

$$P(X = k) = e^{-\lambda} \cdot \frac{\lambda^k}{k!}, \quad k = 0, 1, \dots, \infty$$

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Continuous R.V.

- Ex. Exponential distribution

$$f(x) = \frac{1}{\lambda} \cdot e^{-x/\lambda}$$

- Ex. Gaussian distribution

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{-\frac{(x-m_x)^2}{2\sigma^2}} \sim N(m_x, \sigma^2)$$

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- Ex. Multivariate Gaussian R.V.

$$f(\vec{x}) = \frac{1}{(2\pi)^{\frac{N}{2}} |R|^{\frac{1}{2}}} \cdot \exp\left[-\frac{1}{2} (\vec{x} - \vec{m})^T R^{-1} (\vec{x} - \vec{m})\right]$$

Covariance matrix :

$$R = E[\vec{x} \vec{x}^T] - \vec{m} \vec{m}^T$$

mean :

$$\vec{m} = E[\vec{x}]$$

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- Ex. Chi [kai]- square distribution
 x_n^2 with n degree of freedom.

$$x_n^2 = \sum_{i=1}^n y_i^2; \quad y_i \sim N(0,1) \text{ i.i.d.}$$

$$f_{x_n^2}(x) = \frac{1}{2^{n/2} \Gamma(n/2)} \cdot x^{\frac{n}{2}-1} \cdot e^{-\frac{x}{2}}; \quad x \geq 0;$$

where $\Gamma(\cdot)$ is gamma function .

$$\Gamma(n+1) = n\Gamma(n) = \int_0^\infty y^n e^{-y} dy; \quad n > -1$$

$$\Gamma(n+1) = n! \text{ if } n \text{ is integer}$$

$$\Gamma(1) = 1, \quad \Gamma(1/2) = \sqrt{\pi}$$

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Marginal p.d.f.

- Conditional p.d.f.

$$f(x|y) = f(x, y) / f(y) \xrightarrow{\text{if } x, y \text{ is indep.}} f(x)$$

- Conditional expectation

$$E[x|y] = \int x \cdot f(x|y) dx;$$

$$E_y[E_x[x|y]] = E[x] = E[E[x|y]]$$

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- Ex.

coin1	coin2	P
H	H	0
H	T	0.25
T	H	0.25
T	T	0.5

$$P(\text{coin1} = H | \text{coin2} = H) \stackrel{?}{=} P(\text{coin1} = H)$$

$$\Rightarrow 0 / 0.25 \neq 0.25$$

$\Rightarrow \text{coin1} = H$ and $\text{coin2} = H$ not independent.

note: if x, y are independent,

then $f(x|y) = f(x) \forall x, y$

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- Ex. Bivariate Gaussian distribution

$$f(x, y) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-r^2}} \cdot \exp\left\{\frac{-1}{2(1-r^2)}\left[\frac{(x-m_1)^2}{\sigma_1^2} - 2r\frac{(x-m_1)(y-m_2)}{\sigma_1\sigma_2} + \frac{(y-m_2)^2}{\sigma_2^2}\right]\right\} \sim N(m_1, m_2, \sigma_1, \sigma_2, r)$$

$$f(x | y) \sim N\left(r\frac{\sigma_1}{\sigma_2}(y-m_2) + m_1, \sigma_1\sqrt{1-r^2}\right)$$

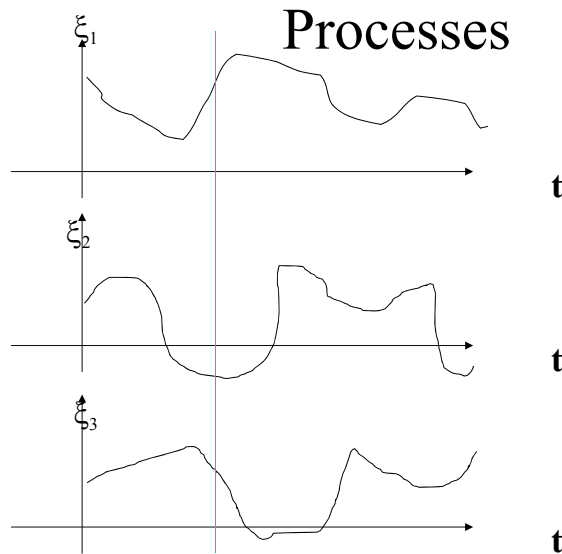
(proof is left for the readers)

$$\Rightarrow E[x | y] = r\frac{\sigma_1}{\sigma_2}(y-m_2) + m_1$$

$$R = \begin{bmatrix} \sigma_1^2 & r\sigma_1\sigma_2 \\ r\sigma_1\sigma_2 & \sigma_2^2 \end{bmatrix}$$

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Random Variables and Random Processes



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- Random variable (r.v): One outcome $\xi_i \Rightarrow$ (maps to) one real number $x(\xi_i)$
- Random process (r.p): One outcome $\xi_i \Rightarrow$ one function of time $x(t, \xi_i)$
- t, ξ fixed \Rightarrow a real number
- t fixed, ξ variable \Rightarrow a r.v. (In this sense, a r.p. can be thought as time-varying r.v.'s)
- t variable, ξ fixed \Rightarrow a function of time
- t, ξ variable \Rightarrow a r. p.

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- We define mean, autocorrelation, autocovariance for r.p $x(t)$
- Mean $\mu_x(t) = E[x(t)]$
- Autocorrelation $R_x(t, s) = E[x(t)x(s)]$. (multiplied by, times)
- Autocovariance $C_x(t, s) = E[(x(t) - \mu_x(t))(x(s) - \mu_x(s))]$
- $= E[x(t)x(s)] - \mu_x(t) \mu_x(s)$
- **Note: expectation is over outcome not time.**

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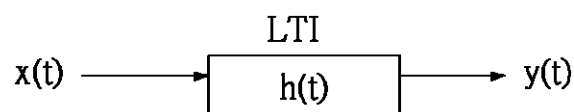
r.p. $x(t)$ is wide - sense stationary (WSS) if

(1) $\mu_x(t) = \mu_x(0)$

(2) $R_x(t, t + \tau) = R_x(\tau)$ for all t , (2nd - 1st time index)

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Linear Filtering of Random Processes



$\tilde{h}(t) \stackrel{\Delta}{=} h(-t)$, $x(t)$ is WSS with autocorrelation

$$R_x(\tau) \stackrel{\Delta}{=} E\{x(t)x(t+\tau)\}$$

(a) $\mu_y(t) = (\mu_x * h)(t)$, where $\mu_x(t) = E\{x(t)\}$

(b) $R_y(\tau) = (R_x * h * \tilde{h})(\tau)$

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Proof of (a) $\mu_y(t) = E\{y(t)\} = E\{x(t) * h(t)\} = E\left\{\int_{-\infty}^{\infty} x(u)h(t-u)du\right\}$

$$= \int_{-\infty}^{\infty} E\{x(u)\}h(t-u)du = \int_{-\infty}^{\infty} u_x(u)h(t-u)du = (u_x * h)(t)$$

Proof of (b) $E\{y(t)y(t+\tau)\}$

$$= \iint E\{x(t-\alpha)x(t+\tau-\beta)\}h(\alpha)h(\beta)d\alpha d\beta$$

$$= \iint R_x(\tau+\alpha-\beta)h(\alpha)h(\beta)d\alpha d\beta$$

$$= \iint R_x(\tau-\gamma-\beta)\tilde{h}(\gamma)h(\beta)d\gamma d\beta$$

$$= (R_x * \tilde{h} * h)(\tau)$$

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Spectral density for WSS r.p. x(t) and y(t) Ex. Noise and filtered noise

Fourier transform pair

$$R_x(\tau) \leftrightarrow S_x(f) \text{ (NWTT1)}$$

$$h(t) \leftrightarrow H(f)$$

$$\tilde{h}(t) \leftrightarrow H^*(f)$$

$$\text{Thus, } R_y(\tau) = (R_x * \tilde{h} * h)(t) \leftrightarrow S_y(f) = |H(f)|^2 S_x(f)$$

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Important Results

(1) If $x(t)$ is WSS, then $y(t)$ is WSS.

(2) If $x(t)$ is Gaussian, then $y(t)$ is Gaussian.

Proof of (1): from (a) and (b) above.

Proof of (2): Use the fact that linear combination of Gaussian r.v.'s is also Gaussian r.v.

$$\left(\int_{-\infty}^{\infty} g(\alpha) \cdot x(\alpha) d\alpha \text{ is also Gaussian}\right)$$

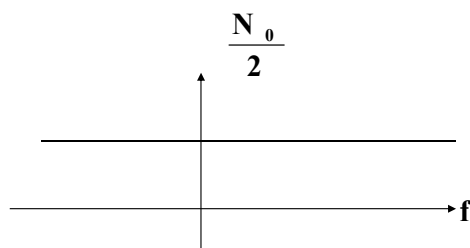
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Thermal Noise

Represented by additive white Gaussian noise (AWGN) random process $x(t)$

$$S_x(f) = N_0 / 2$$

$$R_x(\tau) = \frac{N_0}{2} \delta(\tau)$$



White light is composed of many colors

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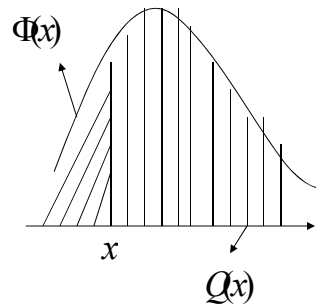
Probability density function (pdf) for $N(0,1)$

$$f_x(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$

cumulative distribution function (cdf) for $N(0,1)$ r.v.

$$F_x(X) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}} du = P_r\{x \leq x\} = \Phi(x) = 1 - Q(x)$$

Note: $\Phi\left(\frac{x-\mu}{\sigma}\right)$: cdf of $N(\mu, \sigma^2)$



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Comparison:

White noise: noise rvs at different times are uncorrelated

Gaussian noise: at one time, noise rv is Gaussian distributed

1)+2) noise rvs at different times are independent Gaussian

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Relationship of Two r.p. (optional)

Def.: crosscorrelation $R_{xy}(t,s)=E[x(t)y(s)]$

Def.: crosscovariance $C_{xy}(t,s)=E[(x(t)-\mu_x(t))(y(s)-\mu_y(s))]$

Lemma: If $R_{xy}(t,s)=\mu_x(t)\mu_y(s)$, or equivalently $C_{xy}(t,s)=0$, for all t and s , then $x(t)$ and $y(t)$ are uncorrelated.

Def.: $x(t)$ and $y(t)$ are independent iff for all $n, t_1, \dots, t_n, s_1, \dots, s_n$, the random vector $\underline{x}=[x(t_1), \dots, x(t_n)]$ and $\underline{y}=[y(s_1), \dots, y(s_n)]$ are independent. That is, $x(t)$ and $y(t)$ are independent iff $F_{xy}(\underline{x}, \underline{y})=F_x(\underline{x})F_y(\underline{y})$ or equivalently $f_{xy}(\underline{x}, \underline{y})=f_x(\underline{x})f_y(\underline{y})$.

Facts : (1) indep r.p. are uncorrelated

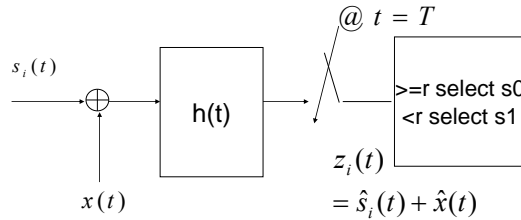
(2) jointly Gaussian+uncorrelated= \Rightarrow indep.

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Digital Communication Review

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General Binary Comm. System



Without loss of generality, assume $(s_0 * h)(T) > (s_1 * h)(T)$

$$z_i(t) = \int h(t-\alpha)s_i(\alpha)d\alpha + \int h(t-\alpha)x(\alpha)d\alpha = \hat{s}_i(t) + \hat{x}(t)$$

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$z_i(T)$ is a Gaussian r.v. with mean $\hat{s}_i(T) = (s_i * h)(T)$

and variance $R_{\hat{x}}(0) = (h * \hat{h} * R_x)(0)$ p.16

$$P(\text{error}|0 \text{ sent}) = P(z_k(T) < r) = \Phi\left(\frac{r - \hat{s}_0(T)}{\sqrt{R_{\hat{x}}(0)}}\right) \quad \text{p.21}$$

$$P(\text{error}|1 \text{ sent}) = Q\left(\frac{r - \hat{s}_1(T)}{\sqrt{R_{\hat{x}}(0)}}\right)$$

Now add more constraints

(1) Assume BPSK : $s_i(t) = (-1)^i \cdot s(t)$, $s(t)$ arbitrary

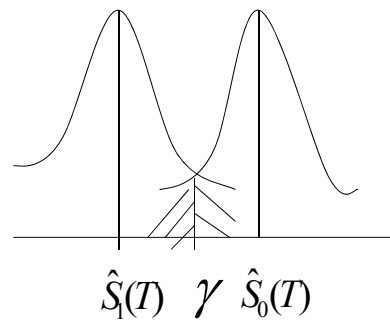
(2) Assume AWGN : Let $\tilde{h}(\mu) \stackrel{\Delta}{=} h(-\mu)$

$$R_{\hat{x}}(\tau) = \iint R_x(\tau - \gamma - \beta) \tilde{h}(\gamma) h(\beta) d\gamma d\beta$$

$$R_{\hat{x}}(0) = \iint R_x(-\gamma - \beta) \tilde{h}(\gamma) h(\beta) d\gamma d\beta = \frac{N_0}{2} \int \tilde{h}(-\beta) h(\beta) d\beta$$

$$= \frac{N_0}{2} \int h^2(\beta) d\beta = \frac{N_0}{2} \|h\|^2 = \frac{N_0}{2} \int |H(f)|^2 df \quad (\text{Parseval theorem})$$

(3) assume $P_{e,1} = P_{e,0}$, then $r = 0$



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Example in a IEEE T COM paper [Wan13]

Since we assume flat fading for each subcarrier, the lowpass equivalent received signal of user k on subcarrier m is given by

$$y_{k,m}(t) = \sqrt{P_{k,m}} H_{k,m} X_{k,m}[l] \exp\left(\frac{j2\pi mt}{T}\right) + n_{k,m}(t) \quad (2)$$

where $n_{k,m}(t)$ is Additive White Gaussian Noise (AWGN) with two-sided power spectral density $2N_0$.

To detect the signal on subcarrier m , a correlation operation is performed $Y_{k,m} = \frac{1}{T} \int_0^T y_{k,m}(t) \exp(-j2\pi mt/T) dt$.

The noise power can be calculated as $P_N = E[|N_{k,m}|^2] = 2N_0/T$ and the power for the desired signal is $P_{k,m} |H_{k,m}|^2$.

$$h(t) = \begin{cases} \frac{1}{T} & 0 \leq t \leq T \\ 0 & \text{otherwise} \end{cases} \quad \text{so } \|h\|^2 = T \left(\frac{1}{T}\right)^2 = \frac{1}{T}$$

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Matched Filter (M.F.) is Optimal in AWGN

First we introduce Schwartz inequality (recall inner product in geometry in high school)
For any functions f and g on (a, b) for which

$$\|f\| = \sqrt{\int_a^b f^2(u) du} < \infty, \quad \|g\| = \sqrt{\int_a^b g^2(u) du} < \infty,$$

$$(f, g) = \int_a^b f(u)g(u) du, \quad \text{we have}$$

$$(f, g)^2 \leq \|g\|^2 \|f\|^2 \quad \text{equality holds iff there is a real number } \lambda \text{ such that } f = \lambda g$$

Assume BPSK : $s_0(t) = s(t)$ and $s_1(t) = -s(t)$,

let $s_{T_0}(u) = s(T-u)$, matched filter

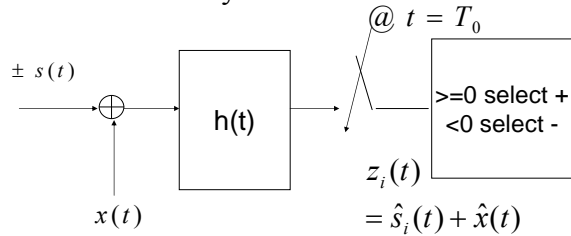
$$\text{Note } r - \hat{s}_1(T) = 0 + (s * h)(T) = \int h(u)s(T-u) du$$

$$P_e = Q\left(\sqrt{\frac{(\int h(u)s(T-u) du)^2}{\frac{N_0}{2} \|h\|^2}}\right) \geq Q\left(\sqrt{\frac{\|s_{T_0}\|^2 \|h\|^2}{\frac{N_0}{2} \|h\|^2}}\right) = Q\left(\sqrt{\frac{\|s_T\|^2}{\frac{N_0}{2}}}\right)$$

with equality holds iff $h(t) = \lambda s_T(t) = \lambda s(T-t)$ matched filter

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Summary



$$P_{e,\min} \geq Q\left(\sqrt{\frac{\|s\|^2}{N_0}}\right) = Q\left(\sqrt{\frac{E}{N_0}}\right) = Q(\sqrt{SNR})$$

where

$$\text{received SNR} = \frac{E\|h\|^2}{\frac{N_0}{2}\|h\|^2} = \frac{E}{\frac{N_0}{2}} \text{ (previous page)}$$

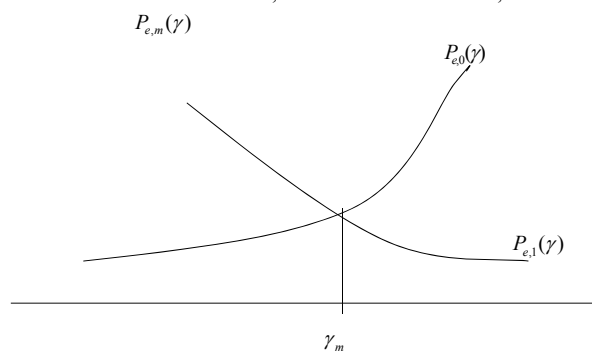
Note : unit J = (W/(1/s))

Q: Why Tx energy
Divided by psd =
received SNR?

A: h norm square is
ignored

Minimax Approach

- minimize the maximum of $P_{e,0}$ and $P_{e,1}$ select r to give minimum of $P_{e,m} = \max\{P_{e,0} \text{ and } P_{e,1}\}$
- As r increases, $P_{e,0}$ increases but $P_{e,1}$ decreases.



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Choose $r = r_m$, where r_m is the solution to $P_{e,0} = P_{e,1}$

$$\Phi\left(\frac{r - \hat{s}_0(T)}{\sqrt{R_{\hat{x}}(0)}}\right) = \Phi\left(\frac{\hat{s}_1(T) - r}{\sqrt{R_{\hat{x}}(0)}}\right) \Rightarrow r_m = \frac{\hat{s}_0(T) + \hat{s}_1(T)}{2}$$

$$P_{e,m} = 1 - \Phi\left(\frac{\frac{1}{2}(\hat{s}_0(T) - \hat{s}_1(T))}{\sqrt{R_{\hat{x}}(0)}}\right)$$

Let $s(t) = s_0(t) - s_1(t)$ and $s_T(u) = s(T-u)$. Then

$$P_{e,m} = 1 - \Phi\left(\sqrt{\frac{\frac{1}{4}(h, s_T)^2}{\frac{N_0}{2}\|h\|^2}}\right) \geq 1 - \Phi\left(\sqrt{\frac{\frac{1}{4}\|s_T\|^2}{\frac{N_0}{2}}}\right) \text{ with equality holds iff}$$

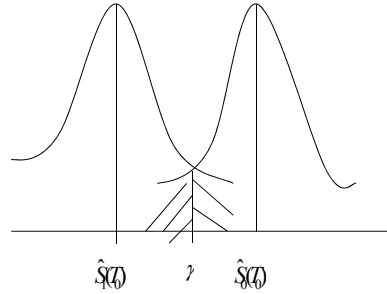
$h(t) = s_T(t) = s_0(T-t) - s_1(T-t)$ matched filter

$$\|s_T(t)\|^2 = \int [s_0(T-t) - s_1(T-t)]^2 dt = \|s_0\|^2 + \|s_1\|^2 - 2 \int s_0(t) \cdot s_1(t) dt$$

$$= \varepsilon_0 + \varepsilon_1 - 2(\varepsilon_0, \varepsilon_1) = 2\bar{\varepsilon}(1 - \rho) \text{ where } \bar{\varepsilon} = \frac{\varepsilon_0 + \varepsilon_1}{2}, \rho = \frac{(\varepsilon_0, \varepsilon_1)}{\bar{\varepsilon}} \text{ correlation coefficient.}$$

$$P_{e,m} = 1 - \Phi\left(\sqrt{\frac{\bar{\varepsilon}(1 - \rho)}{N_0}}\right) \text{ AWGN, MF, minimax threshold}$$

$\rho = -1$ best \Rightarrow BPSK best



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Receiver Structure

The received signal is $y(t) = s_i(t) + x(t)$

(a) M.F. implementation

(b) Correlator implementation (lower hardware cost)

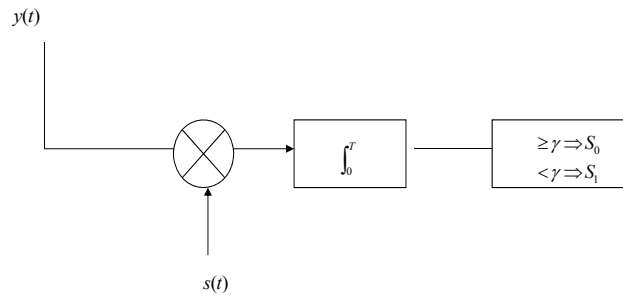
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$$\begin{aligned}
\hat{y}(T) &= \int_0^T y(u) \cdot h(T-u) du \\
&= \int_0^T y(u) \cdot s(T-(T-u)) du \\
&= \int_0^T y(u) \cdot s(u) du
\end{aligned}$$

$$u \sim [0, T], -u \sim [-T, 0], T-u \sim [0, T]$$

$$\hat{y}(T) = \int_0^T y(u) \cdot s(u) du$$

when s is time limited to $[0, T]$, we have



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Quadriphase modulation

Transmitted signal is $R_e[Z(t)]$

where $Z(t) = \sigma(t)e^{j\omega_c t}$ and $\sigma(t) = [b_I(t-t_0) + jb_Q(t)] \cdot e^{j\theta}$

General form : $R_e[Z(t)] = b_I(t-t_0)\cos(W_c t + \theta) - b_Q(t)\sin(W_c t + \theta)$

$$b_I(t) = \sum_{\lambda=-\infty}^{\infty} b_{\lambda}^{(I)} \psi(t - \lambda T)$$

$$b_Q(t) = \sum_{\lambda=-\infty}^{\infty} b_{\lambda}^{(Q)} \psi(t - \lambda T)$$

$\psi(t)$: bit waveform

$$\left. \begin{aligned}
&\text{BPSK : } t_0 = 0, b_Q(t) \equiv 0 \\
&\text{QPSK : } t_0 = 0 \\
&\text{OQPSK : } t_0 = \frac{T}{2}
\end{aligned} \right\} \quad \psi(t) = P_T(t)$$

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Signal Space

orthonormal set of basis function

$\phi_q(t), q = 1, 2, \dots, N$ over interval $(0, T)$

$$\int_0^T \phi_i(t) \phi_k(t) dt = \delta_{ik}$$

$$S_i(t) = \sum_{j=1}^N S_{ij} \phi_j(t), 0 \leq t \leq T, i = 1, \dots, M$$

$$S_{ij} = \int_0^T S_i(t) \phi_j(t) dt = (S_i, \phi_j)$$

Generalized Fourier transform/series coefficients

Digital communication problem \Rightarrow Geometry problem

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Gram Schmit procedure

find orthonormal basis

$$1) \text{ choose } \phi_1(t) = \frac{S_1(t)}{\|S_1\|}$$

$$2) V_2 = S_2 - (S_2, \phi_1) \cdot \phi_1(t) \quad \phi_2(t) = \frac{V_2(t)}{\|V_2\|}$$

$$3) V_j(t) = S_j - (S_j, \phi_1) \cdot \phi_1(t) - \dots - (S_j, \phi_{j-1}) \cdot \phi_{j-1}(t)$$

$$\phi_j(t) = \frac{V_j(t)}{\|V_j\|}$$

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Geometry Approach for Computing P_e

use S_{ij} to represent energy, correlation, error prob

(a) Signal energy

$$\varepsilon_i = \int_0^T S_i^2(t) dt = \int_0^T \sum_{j=1}^N \sum_{k=1}^N S_{ij} S_{ik} \phi_j(t) \phi_k(t) dt = \sum_{j=1}^N S_{ij}^2 = \|S_i\|^2 \quad \text{where } S_i = [S_{i1}, \dots, S_{iN}]$$

(b) correlation coefficient

$$\rho_{ij} = \frac{\Delta(S_i, S_j)}{\bar{\varepsilon}}, \quad \text{where } \bar{\varepsilon} = \frac{\varepsilon_0 + \varepsilon_1}{2}$$

(c) minimax error prob.

$$d_{ij}^2 = \int_0^T [S_i(t) - S_j(t)]^2 dt = \sum_{k=1}^N (S_{ik}^2 + S_{jk}^2 - 2S_{ik}S_{jk})$$

$$= \varepsilon_i + \varepsilon_j - 2(S_i, S_j) = 2\bar{\varepsilon}(1 - \rho_{ij})$$

$$\therefore P_{ij} = Q\left(\frac{\sqrt{\bar{\varepsilon}(1-\rho)}}{\sqrt{N_0/2}}\right) = Q\left(\frac{d_{ij}/2}{\sqrt{N_0/2}}\right) \quad \text{symbol error prob. } i \leftrightarrow j \text{ depend on distance in constellation}$$

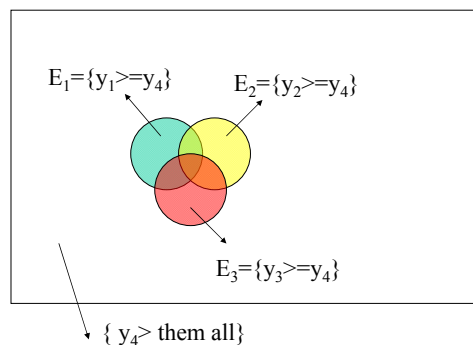
(pp. 32) BPSK is best

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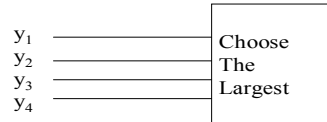
Union Bound (good for high E_b/N_0) on BER

SNR=0, three circles totally overlap

SNR=infinity, three circles become three dots



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M=4 branches
Suppose symbol 4 is sent

$$\begin{aligned}
 \text{Probability Error} &= P(E_1 \cup E_2 \cup E_3) \\
 &= P(E_1) + P(E_2) + P(E_3) - P(E_1 E_2) - P(E_2 E_3) - P(E_3 E_1) + P(E_1 E_2 E_3) \\
 &\leq P(E_1) + P(E_2) + P(E_3)
 \end{aligned}$$

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Approximate union bound

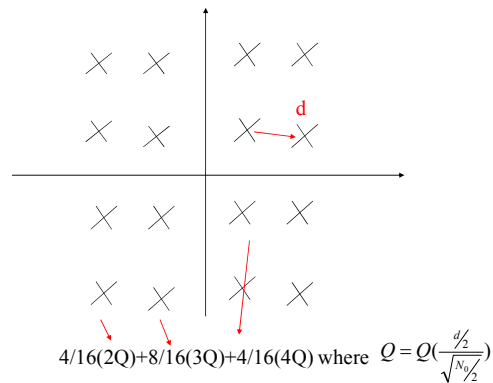
$$\text{Prob. of symbol error} \approx N_{NB} Q\left(\frac{d/2}{\sqrt{N_0/2}}\right)$$

$d \uparrow$, Q function \downarrow , so we ignore all but the largest terms

N_{NB} : number of nearest neighbors

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Example: 16 Quadrature Amplitude Modulation(QAM)



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Bit and symbol error prob.

Problem :

How to relate bit error prob. to symbol error prob.

M - ary signals (symbols)

0 0 0
 0 0 1
 0 1 0
 0 1 1
 1 0 0
 1 0 1
 1 1 0
 1 1 1

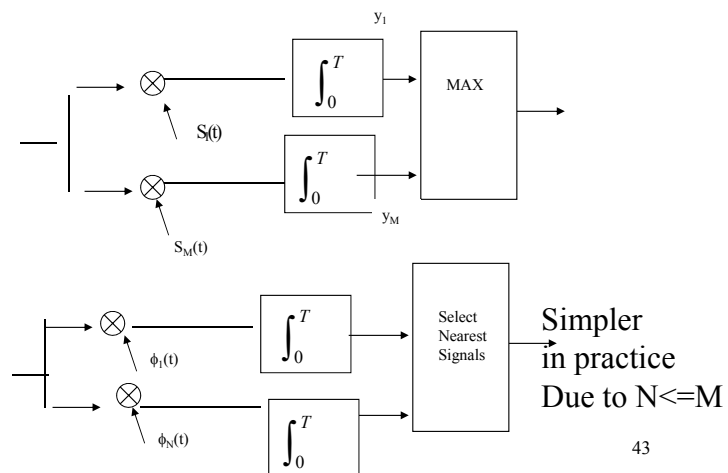
ex : (M - 1) possible symbol errors. $\frac{M}{2}$ bit errors for each bit position

$\frac{M/2}{M-1}$ = conditional prob. of bit error given a symbol error

$P_e = \frac{M/2}{M-1} P(\varepsilon)$ valid if the distance between any two symbols is equal
 (e.g. orthogonal signals)

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Coherent detection of M-ary signals



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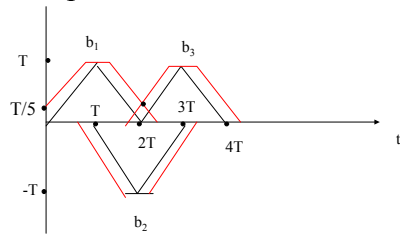
Intersymbol interference

Causes:

- Band-limited channel
- Frequency selective fading (delay spread > symbol interval)- see next notes

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Example



Sample at $t = 2T$

$$b_2 T + b_1 \cdot \frac{T}{5} + b_3 \cdot \frac{T}{5}$$

$$\text{signal} \quad -T \pm \frac{T}{5} \pm \frac{T}{5} = \frac{-3T}{5}, -T, -T, \frac{-7T}{5}$$

$$\overline{P_e} = \frac{1}{4} [Q\left(\frac{3T/5}{\sigma}\right) + 2Q\left(\frac{T}{\sigma}\right) + Q\left(\frac{7T/5}{\sigma}\right)] \quad (\text{p.26})$$

$$= E\left\{Q\left(\frac{|\text{signal}|}{\sigma}\right)\right\} \neq Q\left(\frac{T}{\sigma}\right) = Q\left(\frac{E\{|\text{signal}|\}}{\sigma}\right)$$

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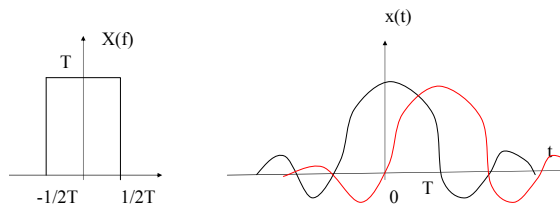
Condition for No ISI

$x(t)$: pulse shape after filtering in Rx

$$x(t) = \sum_{n=-\infty}^{\infty} x(nT) \frac{\sin \pi(t - nT)/T}{\pi(t - nT)/T}$$

to avoid ISI $\Rightarrow x(nT) = 0 \quad \forall n \neq 0$

$$\text{Example} \quad X(f) = \begin{cases} T & |f| < \frac{1}{2T} \\ 0 & \text{otherwise} \end{cases}$$



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Nyquist Criterion for No ISI

problem : if sampling time is not correct $\sum \left| \frac{\sin(\frac{\pi t}{T})}{(\frac{\pi t}{T})} \right| \approx \sum \frac{1}{(\frac{\pi t}{T})}$ diverge

Let $\frac{1}{T} < 2W$ ($X(f)$ is bandlimited to $[-W, W]$)

Let $X_{eq}(f) = \sum_{n=-N}^N X(f + \frac{n}{T})$ $|f| < \frac{1}{2T}$ aliasing DTFT

If $x(kT) = \int_{-\frac{1}{2T}}^{\frac{1}{2T}} X_{eq}(f) \cdot e^{j2\pi f kT} df = 0$ for $k \neq 0$

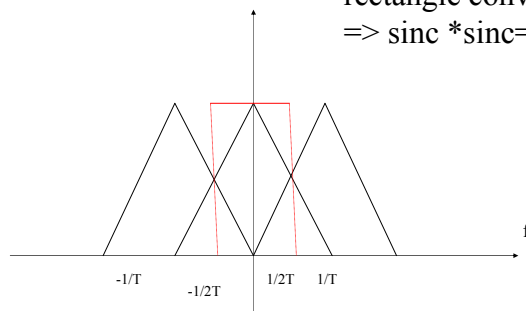
then

$X_{eq}(f) = \begin{cases} \cos \pi f T & |f| \leq \frac{1}{2T} \\ 0 & \text{otherwise} \end{cases}$ (Nyquist criterion for No ISI)

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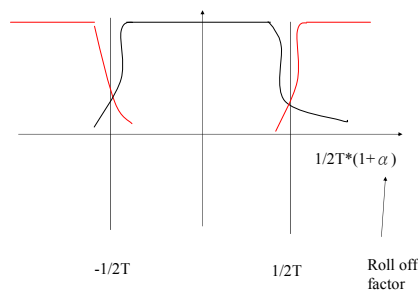
Example 1: sinc²

rectangle convol rectangle in freq domain
=> sinc * sinc = sinc² in time domain



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Example 2: raised cosine(RC)



Commonly used
Ex. Vector signal analyzer
Can set 1)TX RC 2) RX RC
3) TX sqrt(RC) RC sqrt(RC)

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Linear Algebra Review

國中解聯立方程式的延伸

- 1) Pseudo inverse- NWT8 MIMO spatial multiplexing
- 2) QR decomposition
- 3) Singular Value Decomposition (SVD)- optimal MIMO precoding for rank deficiency
NWT8 p.31 (4G TMs)

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Review of linear algebra

- $\vec{X} = (X_1, X_2, \dots, X_n) \in \mathbb{R}^n$
 $\vec{Z} = (Z_1, Z_2, \dots, Z_n) \in \mathbb{C}^n$
- N vectors are linearly independent satisfying following :

$$\sum_{i=1}^n a_i \vec{X}_i = 0 \text{ iff } a_i = 0 \quad \forall i$$
- Span of a set of vectors

$$\{\vec{S} \mid \vec{S} = \sum_{i=1}^n a_i \vec{X}_i, \forall a_i \in \mathbb{R}\}$$

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- A basis for a vector space is a set of vectors with properties :
 - * It is linear independent. (not nec orthogonal)
 - * It spans the space.
- Orthonormal basis
 $\vec{X}_1, \vec{X}_2, \dots, \vec{X}_n$ is orthogonal and length 1.
- $n \times m$ matrix : a mapping $\mathbb{R}^m \rightarrow \mathbb{R}^n$

$$\begin{bmatrix} a_{11} & \cdots & a_{1m} \\ \vdots & \ddots & \vdots \\ a_{n1} & \cdots & a_{nm} \end{bmatrix} \begin{bmatrix} X_1 \\ \vdots \\ X_m \end{bmatrix} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}$$

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domain of the matrix $M = R^m$

range of the matrix $M = \{\vec{y} \mid M \vec{X} = \vec{y}, \forall \vec{X} \in R^m\}$

- Null space of $M = \{\vec{X} \mid M \vec{X} = 0\}$
nullity = dimension of nullspace.
- Rank = dimension of the range space.
Rank + nullity = dimension of domain.
- An $n \times n$ matrix M is invertible iff it has full rank : $\det M \neq 0$, *eigenvalues* $\neq 0$.

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- Ex.

$$\begin{bmatrix} 1 & 1 & 1 \\ 0 & 1 & 0 \\ 2 & 3 & 2 \end{bmatrix} \begin{bmatrix} X_1 \\ X_2 \\ X_3 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix} \Rightarrow \begin{cases} X_1 + X_3 = 0 \\ X_2 = 0 \end{cases}$$

$$\Rightarrow \begin{bmatrix} X_1 \\ 0 \\ -X_1 \end{bmatrix} \Rightarrow \vec{X} = X_1 \begin{bmatrix} 1 \\ 0 \\ -1 \end{bmatrix}$$

Only one variable, nullity = 1.

Rank = 2.

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- Inner product- operation on 2 vectors with the following properties:

$$\vec{X} \cdot \vec{X} \geq 0, \text{equality holds iff } \vec{X} = 0$$

$$\vec{X} \cdot \vec{Y} = \vec{Y} \cdot \vec{X} (X, Y \in C \Rightarrow \text{add complement: } \vec{X} \cdot \vec{Y}^*)$$

$$(\vec{X} + \vec{Y}) \cdot \vec{Z} = \vec{X} \cdot \vec{Z} + \vec{Y} \cdot \vec{Z}$$

$$(r \vec{X}) \cdot \vec{Y} = r(\vec{X} \cdot \vec{Y})$$

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- Ex.
 $\vec{X} \cdot \vec{Y} = \vec{X}^T \cdot M \cdot \vec{Y}$ is an inner product if M is positive definite ($\vec{X}^T \cdot M \cdot \vec{X} \geq 0$, equality holds iff $\vec{X} = 0$)

- Ex.
 $\vec{X} \cdot \vec{Y} = \sum_{i=0}^n X_i Y_i$

- Ex.
 $\vec{X} \cdot \vec{Y} = \int_0^T X(t)Y(t)dt$

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- Norm $\|\vec{X}\|$ - vector operation with the following properties
 $\|\vec{X}\| \geq 0$, equality holds iff $\vec{X} = 0$.

$$\|r \vec{X}\| = r \|\vec{X}\|.$$

$$\|\vec{X} + \vec{Y}\| \leq \|\vec{X}\| + \|\vec{Y}\|.$$

Ex. $\|\vec{X}\| = \sqrt{\vec{X} \cdot \vec{X}}$

Ex. $\sqrt[4]{\sum_{i=1}^n |X_i|^4}$

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Approaches to solve $Mx \approx b$

- Normal equations-quick and dirty: pseudo inverse, numerically unstable
- QR- standard in libraries uses orthogonal decomposition
- SVD - decomposition which also gives indication how linear independent columns are

Quick and Dirty Approach

Multiply by M^T to get the **normal equations**:

$$M^T M x = M^T b$$

However, sometimes $M^T M$ can be *nearly singular or singular*.

Consider the matrix $M = \begin{bmatrix} 1 & 1 \\ e & 0 \\ 0 & e \end{bmatrix}$

The matrix $M^T M = \begin{bmatrix} 1+e^2 & 1 \\ 1 & 1+e^2 \end{bmatrix}$

becomes singular if e is less than the square root of the machine precision.

Goal of QR: a more stable approach

$$M = \underbrace{Q}_{\substack{\text{mxn} \\ \text{mxm}}} \underbrace{\begin{bmatrix} R \\ O \end{bmatrix}}_{\text{mxn}} = Q \begin{bmatrix} ? & ? & \dots & ? \\ 0 & \ddots & & \vdots \\ \vdots & 0 & \ddots & \vdots \\ \vdots & & 0 & ? \\ 0 & 0 & \dots & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}$$

R :
 $n \times n$,
 upper tri.

O :
 $(m-n) \times n$,
 all zeros

Ex. 3 (n) unknowns 5 (m) equations because of noise⁶⁰

Reformulating Least Squares using QR

$$\|r\|^2 = \|b - Mx\|^2 = \left\| b - Q \begin{bmatrix} R \\ 0 \end{bmatrix} x \right\|^2 = \left\| Q^T b - Q^T Q \begin{bmatrix} R \\ 0 \end{bmatrix} x \right\|^2$$

because Q is orthonormal ($Q^T Q = I$) $\|Q^T x\|^2 = \|x\|^2$ (see next page)

$$= \left\| \begin{bmatrix} c_1 - Rx \\ c_2 \end{bmatrix} \right\|^2 \text{ because } Q^T b = \begin{bmatrix} c_1 \\ c_2 \end{bmatrix}$$

$$= \|c_1 - Rx\|^2 + \|c_2\|^2$$

$\geq \|c_2\|^2$ minimum error if we choose $Rx = c_1$

Numerically stable because R is upper triangular (even if M has no inverse)

Can use 國中數學 解聯立方程式 消去法

$$2x + 3y = 5$$

$$x + y = 2$$

$$\Rightarrow 2x + 3y = 5$$

$$y = 1$$

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QR for least square

A matrix Q is **orthonormal** if $Q^T Q = I$.

If Q is orthogonal then for any x, $\|Qx\|^2 = \|x\|^2$,

that is an orthogonal matrix preserves the 2 norm.

Examples of orthonormal matrices:

$$\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \quad \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix} \quad \begin{pmatrix} \cos y & \sin y \\ -\sin y & \cos y \end{pmatrix} \quad \begin{matrix} \text{Givens} \\ \text{rotations} \end{matrix}$$

for some angle y

Most good least squares solvers use the QR approach.

In Matlab: $x = M \backslash b$.

Singular Value Decomposition

The singular value decomposition(SVD) of a matrix A is given by

$$M = U \Sigma V^T$$

$$m \times n = m \times m \quad m \times n \quad n \times n$$

where U and V are unitary matrices (座標轉換 e.g. 圓座標 \Leftrightarrow 直角座標) and S (座標值) is a rectangular diagonal matrix.

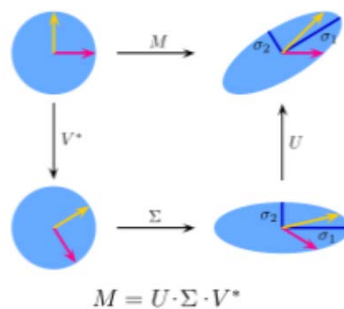
The columns of U are left singular vectors (orthonormal basis of K^m , $K=R$ real or C complex)

The columns of V are right singular vectors (orthonormal basis of K^n)

The diagonal entries of Σ are singular values of M.

Visualization of the SVD of a 2D matrix M

- The SVD decomposes M into three simple transformations: an initial rotation V^* , a scaling Σ along the coordinate axes, and a final rotation U . The lengths σ_1 and σ_2 of the semi-axes of the ellipse (橢圓) are the singular values of M.



- The left-singular vectors of \mathbf{M} are a set of orthonormal eigenvectors of $\mathbf{M}\mathbf{M}^*$.
- The right-singular vectors of \mathbf{M} are a set of orthonormal eigenvectors of $\mathbf{M}^*\mathbf{M}$.
- The non-zero singular values σ_i of \mathbf{M} (found on the diagonal entries of $\mathbf{\Sigma}$) are the square roots of the non-zero eigenvalues of both $\mathbf{M}^*\mathbf{M}$ and $\mathbf{M}\mathbf{M}^*$.

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- The linear transformation

$$\begin{cases} T : K^n \rightarrow K^m \\ x \mapsto \mathbf{M}x \end{cases}$$
- has a particularly simple description with respect to these orthonormal bases: we have

$$T(\mathbf{V}_i) = \sigma_i \mathbf{U}_i, \quad i = 1, \dots, \min(m, n),$$

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where σ_i is the i -th diagonal entry of Σ , and $T(V_i) = 0$ for $i > \min(m, n)$.

- The geometric content of the SVD theorem can thus be summarized as follows: for every linear map $T : K^n \rightarrow K^m$ one can find orthonormal bases of K^n and K^m such that T maps the i -th basis vector of K^n to a non-negative multiple of the i -th basis vector of K^m , and sends the left-over ($|m-n|$) basis vectors to zero. With respect to these bases, the map T is therefore represented by a diagonal matrix with non-negative real diagonal entries.

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SVD for least squares

If M is an $m \times n$ matrix of rank n , then

$$M = U \Sigma V^T = \begin{bmatrix} U_1 & U_2 \end{bmatrix} \begin{bmatrix} \Sigma_1 \\ 0 \end{bmatrix} \quad V^T = U_1^T \Sigma_1^{-1} V^T$$

$m \times m$
 $m \times n$
 $n \times n$

where U_1 has the first n columns of U and Σ_1^{-1} is $n \times n$.

The solution to the least squares problem $Ax \approx b$ is given by

$$x = V \Sigma_1^{-1} U_1^T b$$

Requires 4 to 10 times more work than QR but shows dependencies in model.

Example

$$\mathbf{M} = \begin{bmatrix} 1 & 0 & 0 & 0 & 2 \\ 0 & 0 & 3 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 2 & 0 & 0 & 0 \end{bmatrix}$$

$$\mathbf{U} = \begin{bmatrix} 0 & 0 & 1 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & -1 & 0 \\ 1 & 0 & 0 & 0 & 0 \end{bmatrix}$$

$$\mathbf{\Sigma} = \begin{bmatrix} 2 & 0 & 0 & 0 & 0 \\ 0 & 3 & 0 & 0 & 0 \\ 0 & 0 & \sqrt{5} & 0 & 0 \\ 0 & 0 & 0 & \mathbf{0} & 0 \end{bmatrix}$$

$$\mathbf{V}^* = \begin{bmatrix} 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ \sqrt{0.2} & 0 & 0 & 0 & \sqrt{0.8} \\ 0 & 0 & 0 & 1 & 0 \\ -\sqrt{0.8} & 0 & 0 & 0 & \sqrt{0.2} \end{bmatrix}$$

- Notice $\mathbf{\Sigma}$ is zero outside of the diagonal (grey italics) and one diagonal element is zero (red bold)

Application

- [Ali17] MIMO-NOMA:
- Choose the left singular vector corresponding to largest singular value to represent the MIMO channel vectors of multiple users in a NOMA cluster,
- Max replaces sum (reduced dimension)

ML/MAP Estimation and Detection

Example: $x = \theta + n$, $\theta = 1$ or -1 , $n = \text{AWGN}$
MAP and ML

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Introduction

- Estimation theory plays a major role in the design of statistical signal processing systems. Estimation problems are often delineated into **deterministic** (or classical) parameter estimation and **random** (or Bayesian) parameter estimation.
- 1) **classical estimation**, observations are random and parameters are regarded as unknown constant (deterministic) values.
- 2) **Bayesian parameter estimation**, the parameters are also viewed as random, and our prior knowledge of their behavior is expressed through an a priori density.
 - **Maximum A Posteriori (MAP) – Bayesian (with prior)**
 - **Maximum Likelihood (ML) – Classical**

Maximum Likelihood => minimum Euclidean distance in AWGN (demodulation, Viterbi decoding etc.)

Assume AWGN channel with variance σ^2

($y = x + \text{AWGN}$)

log likelihood function

$$= \log \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y-x)^2}{2\sigma^2}}$$

$$= \log \frac{1}{\sqrt{2\pi\sigma^2}} - \frac{1}{2\sigma^2} (y-x)^2$$

bit metric = $(y-x)^2$

(Euclidean 歐幾里德 distance) 2 !

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Maximum a posterior probability (MAP) including prior probability (used in turbo decoding, soft-input soft-output turbo principle etc.)

$$p(x = +1 | y) = p(y | x = +1)p(x = +1) / p(y)$$

$$p(x = -1 | y) = p(y | x = -1)p(x = -1) / p(y)$$

a posterior probability (APP)

= likelihood function * prior probability

If prior probability = constant (1/2)

then maximum a posterior probability = maximum likelihood

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