

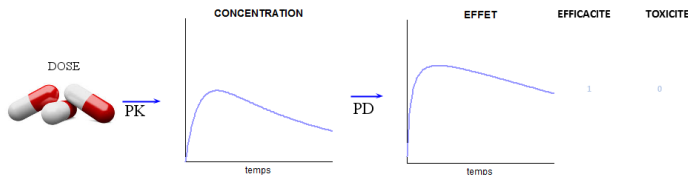
Bayesian approaches for pharmacogenetic models with JAGS and Stan

Julie Bertrand, Maria De Iorio, David J. Balding



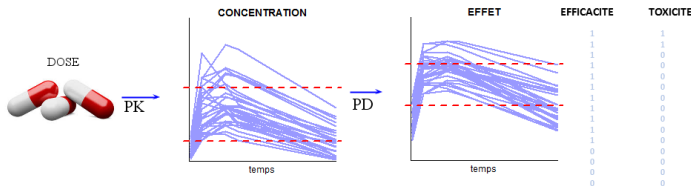
Pharmacological and genetic variability

- Clinical pharmacology: study the interaction between the organism and the drug
 - pharmacokinetics (PK) and pharmacodynamics (PD)



Pharmacological and genetic variability

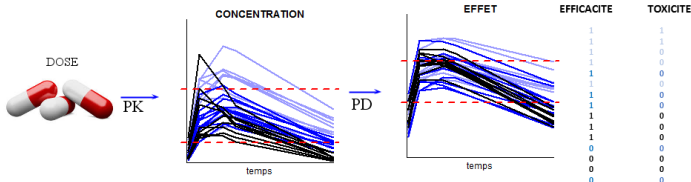
- Clinical pharmacology: study the interaction between the organism and the drug and its variability
 - pharmacokinetics (PK) and pharmacodynamics (PD)



Pharmacological and genetic variability

- Clinical pharmacology: study the interaction between the organism and the drug and its variability
 - pharmacokinetics (PK) and pharmacodynamics (PD)

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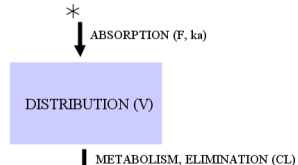
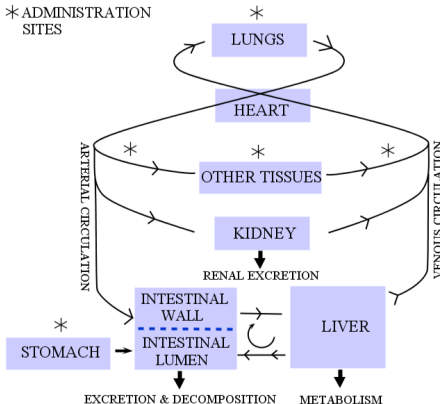


- Pharmacogenetics (PG): genetic part of the variability
 - stratified/precision medicine
 - genes coding for proteins involved in PK/PD processes
 - metabolism enzymes (CYP450, NAT)
 - single nucleotide polymorphism (SNP)

Methodological challenges in PGx

- PK/PD phenotype → not observed
 - data : plasma or insulin concentrations, ...
 - ↪ dynamical model (nonlinear in its parameters)

* ADMINISTRATION
SITES



$$\begin{aligned}
 Q_G(t=0) &= F \times Dose \\
 Q_P(t=0) &= 0 \\
 dQ_G/dt &= -k_a \times Q_G \\
 dQ_P/dt &= k_a \times Q_G - CL/V \times Q_P \\
 C_P(t) &= \frac{Q_P(t)}{V} \\
 &= \frac{F \times Dose}{V} \times \frac{k_a}{k_a - CL/V} (e^{-CL/V \times t} - e^{-k_a \times t})
 \end{aligned}$$

Methodological challenges in PGx

- PK/PD phenotype → not observed
 - data : plasma or insulin concentrations, ...
 - ↪ dynamical model (nonlinear in its parameters)
 - ↪ mixed effect models
 - all patients analyzed simultaneously
 - parameter decomposed in fixed and random effects

Example : $CL_i = \mu_{CL} + \eta_i$ with $\eta_i \sim N(0, \omega_{CL}^2)$

Methodological challenges in PGx

- PK/PD phenotype → not observed
 - data : plasma or insulin concentrations, ...
 - ↪ dynamical model (nonlinear in its parameters)
 - ↪ mixed effect models
 - all patients analyzed simultaneously
 - parameter decomposed in fixed and random effects
 - covariates identification:

Example : $CL_i = \mu_{CL} + \beta \times SNP_i + \eta_i$ with $SNP = \{0, 1, 2\}$

Methodological challenges in PGx

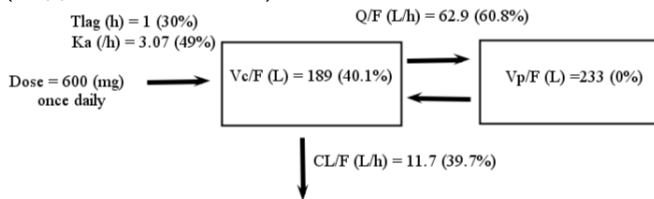
- PK/PD phenotype → not observed
 - data : plasma or insulin concentrations, ...
 - ↪ dynamical model (nonlinear in its parameters)
 - ↪ mixed effect models
- Increased size of genetic data sets up to high throughput screening
 - dimensionality curse $N \ll p$
 - ↪ Stepwise procedure
 - ↪ Penalized regression
 - ↪ Bayesian variable selection

Objectives

- To evaluate, through a realistic simulation study, the performance to detect a pharmacogenetic association
 - stepwise procedure (Monolix software and lm R function)
 - penalized regression (saemix R package and hlasso program)
 - bayesian variable selection (R2jags)

Genetic and PK data

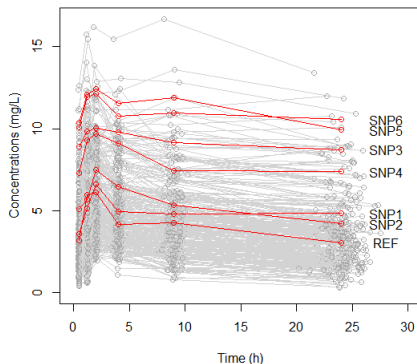
- Generation of genotypes using HAPGEN (Su et al. 2011)
 - $N_s=1227$ snps on 171 genes from the DMET Chip (Daly et al. 2007)
 - 6 [1-56] snps per gene
 - HAPMAP caucasian reference haplotypes
- Pharmacokinetic profiles inspired from real study (Kappelhoff et al. 2005)



- phase II study: $N=300/t=0.5, 1.25, 2, 4, 9, 24h$
- diagonal variance matrix of random effects
- combined residual error model $g = \sigma_{inter} + \sigma_{prop} f(\phi_i, t_{ij})$

Pharmacogenetic effect

- 6 unobserved causal variants explaining 30% of the PK



- $\log Cl_i = \log Cl + \sum_{s=1}^6 \beta_{Cl_s} SNP_{s_i} + \eta_{Cl_i}$ with $SNP_{s_i} = 0, 1, 2$

Stepwise procedure (Lehr et al. 2010)

$$y = \begin{bmatrix} y_{11} \\ y_{ij} \\ y_{Nn_N} \end{bmatrix}$$

$$SNP = \begin{bmatrix} SNP_{11} & SNP_{1s} & SNP_{1N_s} \\ SNP_{i1} & SNP_{is} & SNP_{iN_s} \\ SNP_{N1} & SNP_{Ns} & SNP_{NN_s} \end{bmatrix}$$

Stepwise procedure (Lehr et al. 2010)

Basic dynamic
model: $f(\phi_i, t_{ij})$
 $\phi_i = [ka_i, Cl_i, \dots]$
 $\phi_{pi} = \mu_p + \eta_{pi}$
 $\eta_{pi} \sim N(0, \omega_p)$

Error model:
 $g(f(\phi_i, t_{ij}), \sigma) \varepsilon_{ij}$
 $\varepsilon_{ij} \sim N(0, 1)$

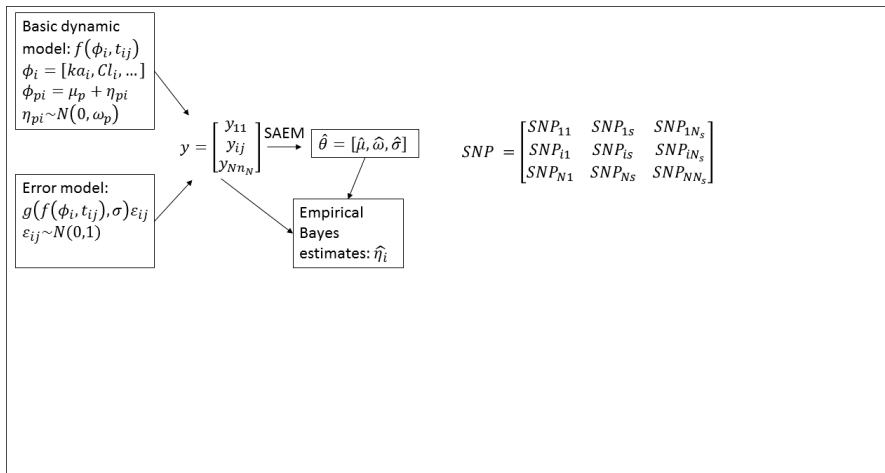
$$y = \begin{bmatrix} y_{11} \\ y_{ij} \\ y_{Nn_N} \end{bmatrix}$$

SAEM

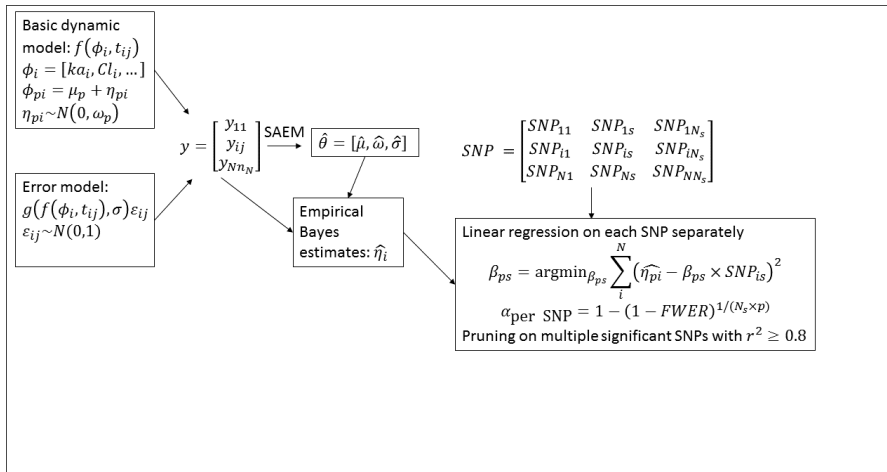
$$\hat{\theta} = [\hat{\mu}, \hat{\omega}, \hat{\sigma}]$$

$$SNP = \begin{bmatrix} SNP_{11} & SNP_{1s} & SNP_{1N_s} \\ SNP_{i1} & SNP_{is} & SNP_{iN_s} \\ SNP_{N1} & SNP_{Ns} & SNP_{NN_s} \end{bmatrix}$$

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Penalized regression (Bertrand et al. 2015)

Dynamic model including parameter-SNP associations:

$$f(\phi_i = \mu + \beta \times \text{SNP}_i + \eta_i, t_{ij})$$

$$\eta_i \sim N(0, \Omega)$$

Error model:

$$g(f(\phi_i, t_{ij}), \sigma) \varepsilon_{ij}$$

$$\varepsilon_{ij} \sim N(0, 1)$$

$$y = \begin{bmatrix} y_{11} \\ y_{ij} \\ y_{Nn_N} \end{bmatrix}$$

$$\text{SNP} = \begin{bmatrix} \text{SNP}_{11} & \text{SNP}_{1s} & \text{SNP}_{1N_s} \\ \text{SNP}_{i1} & \text{SNP}_{is} & \text{SNP}_{iN_s} \\ \text{SNP}_{N1} & \text{SNP}_{Ns} & \text{SNP}_{NN_s} \end{bmatrix}$$

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SAEM

Stochastic Approximation Expectation-step at iteration k

ϕ'_{ik} drawn from $p(\cdot | y; \theta_k)$, Metropolis Hasting algorithm

$s_{ik} = s_{ik} + \tau_k(\phi'_{ik} - s_{ik-1})$, τ_k sequence of decreasing positive numbers

Maximization-step of μ and β at iteration k

$$(\widehat{\mu}, \widehat{\beta}) = \underset{\mu, \beta}{\operatorname{argmin}} \sum_{i=1}^N (s_{ik} - \mu - \beta \times \text{SNP}_i) \Omega^{-1} (s_{ik} - \mu - \beta \times \text{SNP}_i)$$

$$\hat{\theta} = [\hat{\mu}, \hat{\beta}, \hat{\omega}, \hat{\sigma}]$$

Penalized regression (Bertrand et al. 2015)

Dynamic model including parameter-SNP associations:

$$f(\phi_i = \mu + \beta \times \text{SNP}_i + \eta_i, t_{ij})$$

$$\eta_i \sim N(0, \Omega)$$

Error model:

$$g(f(\phi_i, t_{ij}), \sigma) \varepsilon_{ij}$$

$$\varepsilon_{ij} \sim N(0, 1)$$

$$y = \begin{bmatrix} y_{11} \\ y_{ij} \\ y_{Nn_N} \end{bmatrix}$$

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SAEM **modified**

Stochastic Approximation Expectation-step at iteration k

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Maximization-step of μ and β at iteration k

$$(\widehat{\mu}, \widehat{\beta}) = \underset{\mu, \beta}{\text{argmin}} \sum_{i=1}^N (s_{ik} - \mu - \beta \times \text{SNP}_i) \Omega^{-1} (s_{ik} - \mu - \beta \times \text{SNP}_i) + P(\beta)$$

Lasso: $P_{\xi}(\beta)$ approx. a double exponential prior on β

Hlasso: $P_{\gamma, \lambda}(\beta)$ approx. a normal exponential gamma prior on β

ξ and γ set using an asymptotic approximation

$$\hat{\theta} = [\hat{\mu}, \hat{\beta}, \hat{\omega}, \hat{\sigma}]$$

→ Simultaneous SNP selection and estimation of model parameters

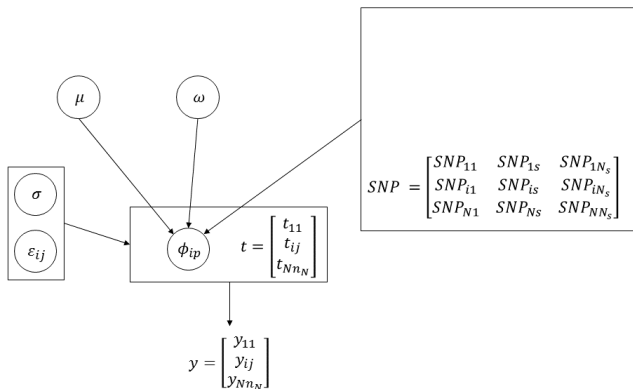
→ All parameter-SNP associations considered simultaneously

Bayesian variable selection (Kuo & Mallick 1998)

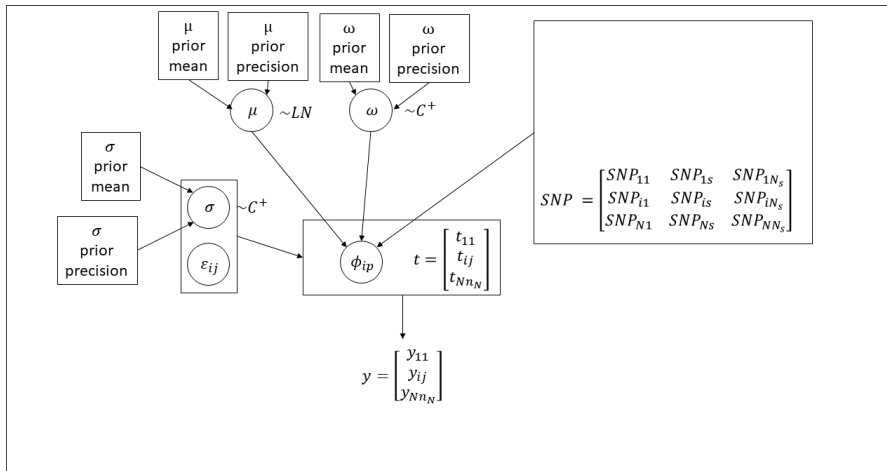
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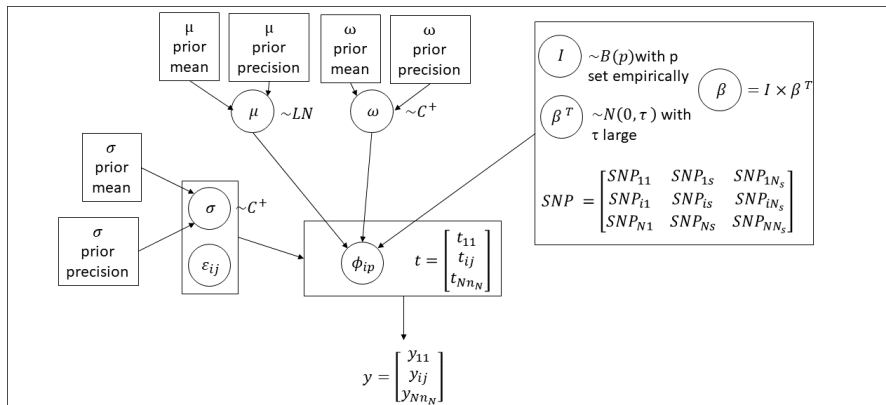
Bayesian variable selection (Kuo & Mallick 1998)



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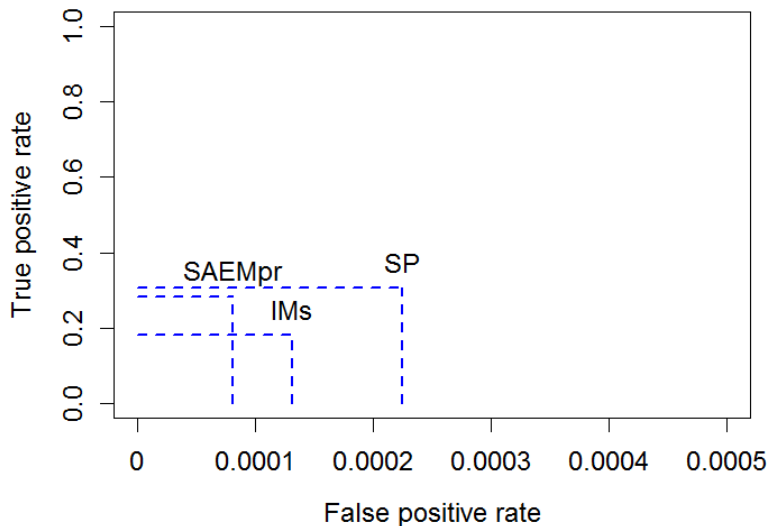


Bayesian variable selection (Kuo & Mallick 1998)



- ↪ Simultaneous SNP selection and estimation of model parameters
- ↪ All parameter-SNP associations considered simultaneously

True and false positive rates



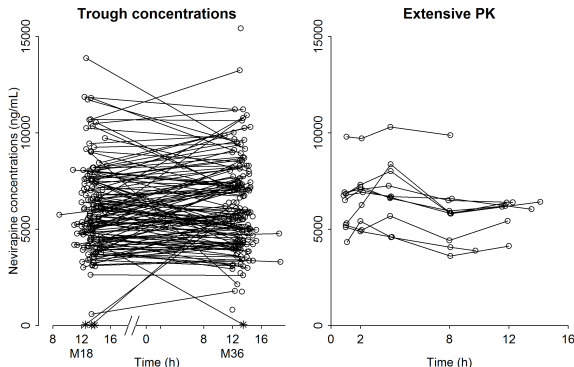
Computing times in hours - mean [range]

Stepwise procedure	0.24 [0.06 - 1.09]
Integrative appr.	1.14[0.83 - 1.61]
Bayesian appr.	19.58 [11.51 - 23.12]

Conclusions

- Feasibility of model-based PGx analyses on a real-case scenario
 - real need of increased sample size compared to classical drug development study designs
- Similar TPR between penalized regression and stepwise procedure
 - cost in computing time non-negligible
 - less so on larger data sets (not shown today)
 - better performance when multiple parameter-SNP associations (not shown today)
- Bayesian appr. performance not yet competitive
 - other indicator model selection (Gibbs variable selection, Stochastic search variable selection)
 - shrinkage priors (Laplacian, Horseshoe)
 - other MCMC algorithms (HMC in the rstan package)

PECAN ANRS 12154 (Dr AM Taburet and Prof D Haas)



- 129 patients on up to 3 occasions with 196 markers

Chromosome	3		7			19	
Gene	<i>NR1I2</i> (PXR)		<i>ABCB1</i> (P-gp)	<i>CYP3A5</i>	<i>CYP3A4</i>	<i>CYP2A6</i>	<i>CYP2B6</i>
Number of markers	49		63	1	36	1	47

- 218 missing SNP with a maximum of 7 per subject

Analyses

- One compartment with 1st-order absorption and elimination
- Inter-individual and -occasion variability on CL/F
 - adjustment on CYP2B6 G516T (rs3745274)
- 1 Stepwise procedure
 - missing data removed
- 2 Indicator Model selection using R2jags
 - missing data imputed using observed allele frequency
- 2 Horseshoe shrinkage prior using Rstan (Piironen et al. 2017)

$$\beta_j | \lambda_j, \tau \sim N(0, \lambda_j^2 \tau^2)$$

$$\lambda_j \sim C^+(0, 1) \text{ and } \tau \sim C^+(0, p_0 / [(N_{SNP} - p_0) \sqrt{M}])$$

- missing data imputed to most common genotype

Stepwise procedure

Table 3 Multivariate linear regression for association between genetic variants and nevirapine clearance estimate

Variants	Gene	β	Statistic	<i>P</i> value
<i>n</i> = 129 (all participants)				
rs3745274	<i>CYP2B6</i>	− 1.606	− 6.66	1.01×10^{-9}
rs2687116	<i>CYP3A4</i>	− 1.913	− 4.68	7.95×10^{-6}
rs2279343	<i>CYP2B6</i>	0.904	4.13	7.07×10^{-5}
rs7251950	<i>CYP2B6</i>	− 0.499	− 4.23	4.77×10^{-5}
<i>n</i> = 128 (without outlier) ^a				
rs3745274	<i>CYP2B6</i>	− 1.568	− 7.66	6.78×10^{-12}
rs2279343	<i>CYP2B6</i>	0.835	4.50	1.66×10^{-5}
rs7251950	<i>CYP2B6</i>	− 0.404	− 4.01	1.09×10^{-4}
rs2032582A	<i>ABCB1</i>	0.456	3.60	4.73×10^{-4}

^aData from one individual outlier (extremely high) nevirapine clearance value were censored from this analysis.

Conclusions

- Similar model PK parameter estimates across methods
- No other snp associations, besides CYP2B6 G516T on CL/F, identified with Bayesian selection methods
- Rstan sampling marginally more efficient than R2jags
- Perspective
 - evaluate adaptive shrinkage in simulation study
 - application to HIV Swiss cohort and EARSII trial on diabetes data

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