Confidence Intervals

using mtcars dataset

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I) Sample and population variance

```
library(prettyR) # for method valid.n()
```

The two following estimators for σ^2 are used to build confidence intervals (CI).

I.1) Sample variance

• If we work on a sample of the whole population, expectation μ is unknown and estimated by $\overline{X_n}$. In this case, sample variance $\hat{\sigma_n}^2$ is an unbiased and consistent estimator for σ^2 :

$$\hat{\sigma_n}^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \overline{X_n})^2$$

```
X <- mtcars$wt  # random variable
var(X, na.rm=TRUE)  # sample variance
## [1] 0.957379</pre>
```

```
sd(X, na.rm=TRUE) # sample standard deviation
```

[1] 0.9784574

I.2) Population variance

• If expectation μ is known, population variance σ^2 is an unbiased estimator for σ^2 :

$$\tilde{\sigma_n}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$$

```
X <- mtcars$wt  # random variable
mu <- 3.2  # expectation (arbitrary value)
pop_var <- 1/(valid.n(X, na.rm=TRUE))*sum((X - mu)^2)  # population variance
pop_sd <- sqrt(pop_var)  # population standard deviation
pop_var

## [1] 0.9277584

## [1] 0.9632022</pre>
```

I) Confidence interval of the expectation μ

```
library(prettyR) # for method valid.n()
```

- some terms:
 - X: random variable
 - $-\overline{X_n}$: empirical mean for n individuals
 - $-\sigma^2$: variance
 - $-\hat{\sigma}^2$: sample variance (estimator for σ^2)
 - $-\tilde{\sigma}^2$: population variance (estimator for σ^2)
 - $-1-\alpha$: confidence level
 - $-Z_{(1-\alpha/2)}$: quantile of order $1-\alpha/2$ for a Normal random variable
 - $-t_{(1-\alpha/2;n-)}$: quantile of order $1-\alpha/2$ for a Student random variable with n-1 degree of freedom

I.1) If X is not assumed to be gaussian, and variance σ^2 unknown (most frequent case):

Asymptotic symmetric CI, thanks to the Central Limit Theorem, is:

$$\left[\overline{X_n} - \frac{\hat{\sigma_n}}{\sqrt{n}} Z_{(1-\alpha/2)}; \overline{X_n} + \frac{\hat{\sigma_n}}{\sqrt{n}} Z_{(1-\alpha/2)}\right]$$

With $\overline{X_n}$ the empirical mean and $\hat{\sigma_n}$ the population standard deviation

```
X <- mtcars$wt  # random variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
IC_min <- mean(X, na.rm=TRUE)-sd(X, na.rm=TRUE)/sqrt(valid.n(X, na.rm=TRUE))*qnorm(1-alpha/2)
IC_max <- mean(X, na.rm=TRUE)+sd(X, na.rm=TRUE)/sqrt(valid.n(X, na.rm=TRUE))*qnorm(1-alpha/2)
#Asymptotic CI:
mean(X)</pre>
```

[1] 3.21725

```
IC_min

## [1] 2.878238

IC_max
```

[1] 3.556262

I.2) If X is assumed to be gaussian, and variance σ^2 known:

Asymptotic symmetric CI is:

$$\left[\overline{X_n} - \frac{\sigma}{\sqrt{n}} Z_{1-\alpha/2}; \overline{X_n} + \frac{\sigma}{\sqrt{n}} Z_{1-\alpha/2}\right]$$

With $\overline{X_n}$ the empirical mean

```
X <- mtcars$wt  # random variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
sd <- 0.9  # standard deviation (arbitrary value)
IC_min <- mean(X, na.rm=TRUE)-sd/sqrt(valid.n(X, na.rm=TRUE))*qnorm(1-alpha/2)
IC_max <- mean(X, na.rm=TRUE)+sd/sqrt(valid.n(X, na.rm=TRUE))*qnorm(1-alpha/2)
#Asymptotic CI:
mean(X)</pre>
```

[1] 3.21725

IC_min

[1] 2.905422

IC_max

[1] 3.529078

I.3) If X is gaussian, and variance σ^2 unknown:

Exact symmetric CI is:

$$\left[\overline{X_n} - \frac{\hat{\sigma_n}}{\sqrt{n}} t_{(1-\alpha/2;n-1)}; \overline{X_n} + \frac{\hat{\sigma_n}}{\sqrt{n}} t_{(1-\alpha/2;n-1)}\right]$$

With $\overline{X_n}$ the empirical mean and $\hat{\sigma_n}$ the sample standard deviation

```
X <- mtcars$wt  # random variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
IC_min <- mean(X, na.rm=TRUE)-sd(X, na.rm=TRUE)/sqrt(valid.n(X, na.rm=TRUE))*qt((1-alpha/2), (valid.n(X IC_max <- mean(X, na.rm=TRUE)+sd(X, na.rm=TRUE)/sqrt(valid.n(X, na.rm=TRUE))*qt((1-alpha/2), (valid.n(X #Exact CI: mean(X))</pre>
```

```
## [1] 3.21725
```

```
IC_min
```

[1] 2.864478

```
IC_max
```

[1] 3.570022

I.4) If X is gaussian, and variance σ^2 known:

Exact symmetric CI is:

$$\left[\overline{X_n} - \frac{\sigma}{\sqrt{n}} Z_{1-\alpha/2}; \overline{X_n} + \frac{\sigma}{\sqrt{n}} Z_{1-\alpha/2}\right]$$

With $\overline{X_n}$ the empirical mean

```
X <- mtcars$wt  # random variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
sd <- 0.9  # standard deviation (arbitrary value)
IC_min <- mean(X, na.rm=TRUE)-sd/sqrt(valid.n(X, na.rm=TRUE))*qnorm((1-alpha/2), (valid.n(X, na.rm=TRUE)
IC_max <- mean(X, na.rm=TRUE)+sd/sqrt(valid.n(X, na.rm=TRUE))*qnorm((1-alpha/2), (valid.n(X, na.rm=TRUE)
#Exact CI:
mean(X)</pre>
```

[1] 3.21725

```
IC_min
```

[1] -2.026648

```
IC_max
```

[1] 8.461148

II) Confidence interval of the variance σ^2

(Gaussian distribution only)

```
library(prettyR) # for method valid.n()
```

II.1) If expectation μ is unknown (most often case):

We use sample variance $\hat{\sigma_n}^2$ (see I.1)). Confidence interval is:

$$\left[\frac{\hat{\sigma_n}^2.n}{C_{(1-\alpha_2;n-1)}}; \frac{\hat{\sigma_n}^2.n}{C_{(\alpha_1;n-1)}}\right]$$

With $\alpha_1 + \alpha_2 = \alpha$

```
X <- mtcars$wt  # random variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
n <- (valid.n(X, na.rm=TRUE))
var(X)  # sample variance</pre>
```

[1] 0.957379

```
# exact CI
IC_min <- var(X)*n/qchisq((1-alpha/2), n-1)
IC_max <- var(X)*n/qchisq((alpha/2), n-1)
IC_min</pre>
```

[1] 0.635184

```
IC_max
```

[1] 1.746769

II.2) If expectation μ is known:

We use **population variance** $\tilde{\sigma_n}^2$ (see I.2)). Confidence interval is:

$$\left[\frac{\tilde{\sigma_n}^2.n}{C_{(1-\alpha_2;n)}}; \frac{\tilde{\sigma_n}^2.n}{C_{(\alpha_1;n)}}\right]$$

With $\alpha_1 + \alpha_2 = \alpha$

```
X <- mtcars$wt  #variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
mu <- 3.2  # expectation (arbitrary value)
pop_var <- 1/(valid.n(X, na.rm=TRUE))*sum((X - mu)^2)  # population variance

# exact CI
IC_min <- pop_var*n/qchisq((1-alpha/2), n)
IC_max <- pop_var*n/qchisq((alpha/2), n)
pop_var</pre>
```

[1] 0.9277584

```
IC_min
```

[1] 0.6000001

 IC_{max}

[1] 1.623129

III) Confidence interval of a proportion

Several methods exists, most often used are:

II.1) Asymptotic confidence interval

$$\left[\overline{X_n} - \frac{1}{2\sqrt{n}}Z_{1-\alpha/2}; \overline{X_n} + \frac{1}{2\sqrt{n}}Z_{1-\alpha/2}\right]$$

With $\overline{X_n}$ the estimation of the proportion

```
X <- mtcars$wt  # random variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
X_n_bar <- 0.3  # estimation of the proportion
n <- (valid.n(X, na.rm=TRUE))
IC_min <- X_n_bar - 1/(2*sqrt(n)*qnorm(1-alpha/2))
IC_max <- X_n_bar + 1/(2*sqrt(n)*qnorm(1-alpha/2))
X_n_bar</pre>
```

[1] 0.3

 ${\tt IC_min}$

[1] 0.2549031

IC_max

[1] 0.3450969

II.2) Agresti-coull Confidence interval

$$\left[\overline{X_n} - \sqrt{\frac{\overline{X_n}.(1 - \overline{X_n})}{n}} Z_{1-\alpha/2}; \overline{X_n} + \sqrt{\frac{\overline{X_n}.(1 - \overline{X_n})}{n}} Z_{1-\alpha/2}\right]$$

With $\overline{X_n}$ the estimation of the proportion

```
X <- mtcars$wt  # random variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
X_n_bar <- 0.3  # estimation of the proportion
n <- (valid.n(X, na.rm=TRUE))
IC_min <- X_n_bar - sqrt(X_n_bar*(1-X_n_bar)/n)*qnorm(1-alpha/2)
IC_max <- X_n_bar + sqrt(X_n_bar*(1-X_n_bar)/n)*qnorm(1-alpha/2)
X_n_bar</pre>
```

```
## [1] 0.3
```

 ${\tt IC_min}$

[1] 0.1412248

IC_max

[1] 0.4587752

IV) Confidence interval of the parameter λ of an exponential distribution

$$\left[\frac{1 - \frac{Z_{1-\alpha/2}}{\sqrt{n}}}{\overline{X_n}}; \frac{1 + \frac{Z_{1-\alpha/2}}{\sqrt{n}}}{\overline{X_n}}\right]$$

With $\overline{X_n}$ the empirical mean

```
X <- seq(from = 0, to = 8, by = 0.02)  # (sequence, for simulation)
X <- dexp(X, rate = 2)# simulation of a random variable
alpha <- 0.05  # confidence level = 1-alpha = 0.95
lambda_hat <- 1/mean(X) # estimation of lambda
n <- (valid.n(X, na.rm=TRUE))
IC_min <- (1-qnorm(1-alpha/2)/sqrt(n))/mean(X)
IC_max <- (1+qnorm(1-alpha/2)/sqrt(n))/mean(X)
lambda_hat</pre>
```

[1] 7.861718

 ${\tt IC_min}$

[1] 7.092245

IC_max

[1] 8.631191