

## WEEK 7

### Week 7 theme

Determine the expression of ordinary moments of order k for a random variable that follows:

- a) Law Gamma
- b) Beta Law
- c) Law  $\chi^2$  (hi-squared)
- d) Legea Student
- e) Legea Fisher-Snedecor

### Moments

- It is called **the initial (regular) moment of order k** of random variable X, the numerical characteristic  $\sigma_k = M(X^k)$
- For k=1 we have  $\sigma_1 = M(X)$  and for k=2,  $D^2(X) = \sigma_2 - \sigma_1^2$
- If X is a discrete type variable with the distribution  $X: \begin{pmatrix} x_i \\ p_i \end{pmatrix}_{i \in I}$ ,  $\sum_{i \in I} p_i = 1$  then  $\sigma_k = \sum_{i \in I} x_i^k p_i$
- If X is variable of continuous type  $X: \begin{pmatrix} x \\ \rho(x) \end{pmatrix}_{x \in R}$ , then  $\sigma_k = \int_R x^k \rho(x) dx$
- It is called **the centered moment of order k** of random variable X, the numerical characteristic  $\mu_k = M[(X - M(X))^k]$ , i.e.

$$\mu_k = \begin{cases} \sum_{i \in I} (x_i - M(X))^k \cdot p_i & , X \text{ discretă} \\ \int_R (x - M(X))^k \rho(x) dx & , X \text{ continuă} \end{cases}$$

- For k=1 we have  $\mu_1 = 0$ , and for k=2,  $\mu_2 = D^2(X)$
- **Theorem** Between centered and initial moments there is the following relationship:

$$\mu_k = \sum_{i=0}^k (-1)^i C_k^i \sigma_{k-i} \sigma_1^i.$$

- In mathematical statistics, the first four centered moments are usually used:  $\mu_1, \mu_2, \mu_3, \mu_4$ .

- The **initial moment of order (r,s)** of the random vector (X,Y) is called the numerical characteristic  $\sigma_{rs} = M(X^r Y^s)$ , i.e.

$$\sigma_{rs} = \begin{cases} \sum_{i \in I} \sum_{j \in J} x_i^r y_j^s p_{ij} & , (X, Y) \text{ discret} \\ \iint_{R^2} x^r y^s \rho(x, y) dx dy & , (X, Y) \text{ continuu} \end{cases}$$

- It is called **the order centered momentum (r,s)** of the random vector (X,Y), the numerical characteristic  $\mu_{rs} = M[(X - M(X))^r (Y - M(Y))^s]$ , i.e.

$$\mu_{rs} = \begin{cases} \sum_{i \in I} \sum_{j \in J} (x_i - M(X))^r (y_j - M(Y))^s p_{ij} & , (X, Y) \text{ discret} \\ \iint_{R^2} (x - M(X))^r (y - M(Y))^s \rho(x, y) dx dy & , (X, Y) \text{ continuu} \end{cases}$$

*Observation*

$$\sigma_{10} = M(X), \sigma_{01} = M(Y), \mu_{20} = D^2(X), \mu_{02} = D^2(Y)$$

### Correlation or covariance

- It's called **the correlation or covariance** of random variables X and Y, the numerical characteristic

$$C(X, Y) = M[(X - M(X))(Y - M(Y))] \text{ adică } C(X, Y) = \mu_{11}$$

*Observation*

$$C(X, Y) = M(XY) - M(X)M(Y), C(X, Y) = \sigma_{11} - \sigma_{10}\sigma_{01}$$

If X, Y independent  $\Rightarrow C(X, Y) = 0$  but not reciprocal.

$$C(X, X) = D^2(X)$$

$$C\left(\sum_{i=1}^m a_i X_i, \sum_{j=1}^n b_j Y_j\right) = \sum_{i=1}^m \sum_{j=1}^n a_i b_j C(X_i, Y_j), \text{ whatever the random variables } X_i \text{ and } Y_j \text{ and}$$

whatever the real constants  $a_i$  and  $b_j$ ,  $1 \leq i \leq m, 1 \leq j \leq n$

$$C(X, Y) = C(Y, X), \text{ whatever } X \text{ and } Y.$$

- It is called **the correlation coefficient** relative to random variables X and Y numerical characteristic

$$r(X, Y) = \frac{C(X, Y)}{\sqrt{D^2(X)} \sqrt{D^2(Y)}}$$

*Observation*

1)  $X, Y \Rightarrow r(X, Y) = 0$  independently of each other is not true;

2) We say that X, Y are uncorrelated if  $r(X, Y) = 0$

**Properties:**

a)  $|r(X, Y)| \leq 1$

b)  $r(X, Y) = +1 \Leftrightarrow Y = aX + b, a > 0$

c)  $r(X, Y) = -1 \Leftrightarrow Y = aX + b, a < 0$

- In practice it is also said that:

1) X and Y are positively perfectly correlated if  $r(X, Y) = 1$ ;

2) X and Y are negatively correlated perfectly if  $r(X, Y) = -1$ ;

3) X and Y are strongly positively (or negatively) correlated if

$$0,75 \leq r(X, Y) < 1 \text{ (or } -1 < r(X, Y) \leq -0,75 \text{);}$$

4) X and Y are weakly positive (or negative) correlated if  $0 < r(X, Y) < 0,25$  (or  $-0,25 < r(X, Y) < 0$ );

The decision-making value margins are conventionally chosen.

- Given the random vector  $Z = (X_1, X_2, \dots, X_n)$   $Z: E \rightarrow R^n$ , it is called **its mean value** and denoted by  $M(Z)$ , if it exists, the n-dimensional vector whose components are the average values of the components of Z, i.e.:  $M(Z) = M(X_1, X_2, \dots, X_n) = (M(X_1), M(X_2), \dots, M(X_n))$ .
- It is called **the covariance** (or correlation) matrix of the vector Z, and is denoted by  $C(Z)$ , if any, the matrix  $C(Z) = (c_{ij})_{\substack{i=1,n \\ j=1,n}} = (C(X_i, Y_j))_{i,j=1,n}$

**Observation**

a) For the case of a two-dimensional random vector, do not confuse the mean of the product of components X and Y, which is  $M(XY)$  with the mean of the vector  $(X, Y)$ , which is  $M(X, Y)$ .

b) Sometimes the correlation matrix  $C(Z)$  is also denoted by  $\Gamma(Z)$ .

c) Unfolded, the covariance matrix  $C(Z)$  has the form:

$$C(Z) = \begin{pmatrix} D^2(X_1) & C(X_1, X_2) & \dots & C(X_1, X_n) \\ C(X_2, X_1) & D^2(X_2) & \dots & C(X_2, X_n) \\ \dots & \dots & \dots & \dots \\ C(X_n, X_1) & C(X_n, X_2) & \dots & D^2(X_n) \end{pmatrix}$$

And as a result of the properties of correlation, we find that the matrix  $C(Z)$  is symmetric.

d) Starting from the definition of correlation coefficient and correlation matrix, if all components of Z are non-constant, then we can introduce the matrix of correlation coefficients  $R(Z)$  whose developed form is:

$$R(Z) = \begin{pmatrix} 1 & r(X_1, X_2) & \dots & r(X_1, X_n) \\ r(X_2, X_1) & 1 & \dots & r(X_2, X_n) \\ \dots & \dots & \dots & \dots \\ r(X_n, X_1) & r(X_n, X_2) & \dots & 1 \end{pmatrix}$$

Both forms of the correlation matrix of the random vector  $Z$  are actually tables measuring the degree of dependence between the components of  $Z$ , considered two by two.

**Application 1.** Let  $(X, Y)$  be a discrete random vector whose probabilistic distribution is given by the table below.

Calculate the correlation coefficient  $r(X, Y)$ .

<i>And</i> <i>X</i>	-1	0	1	2	<i>pi</i>
-1	1/6	1/12	1/12	1/24	9/24
0	1/24	1/6	1/12	1/24	8/24
1	1/24	1/24	1/6	1/24	7/24
<i>qj</i>	6/24	7/24	8/24	3/24	1

*Solution:*

Based on the appropriate formulas, we immediately deduce:

$$M(X) = -1 \frac{9}{24} + 0 \frac{8}{24} + 1 \frac{7}{24} = -\frac{2}{24} = -\frac{1}{12}$$

$$M(X^2) = 1 \frac{9}{24} + 0 \frac{8}{24} + 1 \frac{7}{24} = \frac{16}{24} = \frac{2}{3}$$

$$D(X^2) = M(X^2) - [M(X)]^2 = \frac{2}{3} - \frac{1}{144} = \frac{95}{144}$$

$$M(Y) = -1 \frac{6}{24} + 0 \frac{7}{24} + 1 \frac{8}{24} + 2 \frac{3}{24} = \frac{8}{24} = \frac{1}{3}$$

$$M(Y^2) = 1 \frac{6}{24} + 0 \frac{7}{24} + 1 \frac{1}{24} + 4 \frac{3}{24} = \frac{26}{24} = \frac{13}{12}$$

$$D(Y^2) = M(Y^2) - [M(Y)]^2 = \frac{13}{12} - \frac{1}{9} = \frac{35}{36}$$

$$M(XY) = -1 \cdot \left( -1 \frac{1}{6} + 0 \frac{1}{12} + 1 \frac{1}{12} + 2 \frac{1}{24} \right) + 0 \cdot \left( -1 \frac{1}{24} + 0 \frac{1}{6} + 1 \frac{1}{12} + 2 \frac{1}{24} \right) + 1 \cdot \left( -1 \frac{1}{24} + 0 \frac{1}{24} + 1 \frac{1}{6} + 2 \frac{1}{24} \right) = \frac{5}{24}$$

$$C(X, Y) = M(XY) - M(X)M(Y) = \frac{5}{24} + \frac{1}{36} = \frac{17}{72}$$

$$r(X,Y) = \frac{C(X,Y)}{\sqrt{D^2(X)D^2(Y)}} = \frac{\frac{17}{12}}{\sqrt{\frac{95}{144} \cdot \frac{35}{36}}} \approx 0,295$$

**Application 2.** Let be random variables:

$$X_1: \begin{pmatrix} -1 & 1 \\ p_1 & p_2 \end{pmatrix}; X_2: \begin{pmatrix} 1 & 3 \\ q_1 & q_2 \end{pmatrix}; X_3: \begin{pmatrix} 2 & 4 \\ r_1 & r_2 \end{pmatrix}$$

whose common distribution denoted  $(p_{ijk})$ ,  $1 \leq i, j, k \leq 2$ , is:

$$p_{111} = \frac{1}{16}; p_{112} = \frac{1}{16}; p_{121} = \frac{1}{32}; p_{122} = \frac{3}{32};$$

$$p_{211} = \frac{1}{8}; p_{212} = \frac{1}{4}; p_{221} = \frac{1}{16}; p_{222} = \frac{5}{16}.$$

Determine the two-dimensional and one-dimensional distributions of the three-dimensional random vector  $Z = (X_1, X_2, X_3)$  and correlation matrices  $C(Z)$  and  $R(Z)$ .

**Solution:**

We immediately have two-dimensional distributions

$$\left\{ \begin{array}{l} p_{11\bullet} = p_{111} + p_{112} = \frac{1}{8} \\ p_{21\bullet} = p_{211} + p_{212} = \frac{3}{8} \end{array} \right. ; \quad \left\{ \begin{array}{l} p_{12\bullet} = p_{121} + p_{122} = \frac{1}{8} \\ p_{22\bullet} = p_{221} + p_{222} = \frac{3}{8} \end{array} \right\} \text{ for } (X_1, X_2)$$

$$\left\{ \begin{array}{l} p_{1\bullet 1} = p_{111} + p_{121} = \frac{3}{32} \\ p_{2\bullet 1} = p_{211} + p_{221} = \frac{8}{16} \end{array} \right. ; \quad \left\{ \begin{array}{l} p_{1\bullet 2} = p_{112} + p_{122} = \frac{5}{32} \\ p_{2\bullet 2} = p_{212} + p_{222} = \frac{9}{16} \end{array} \right\} \text{ for } (X_1, X_3)$$

$$\left\{ \begin{array}{l} p_{\bullet 11} = p_{111} + p_{211} = \frac{3}{16} \\ p_{\bullet 21} = p_{121} + p_{221} = \frac{3}{32} \end{array} \right. ; \quad \left\{ \begin{array}{l} p_{\bullet 12} = p_{112} + p_{212} = \frac{5}{32} \\ p_{\bullet 22} = p_{122} + p_{222} = \frac{13}{32} \end{array} \right\} \text{ for } (X_2, X_3)$$

And as a result we can write the following two-dimensional distribution tables:

X2	1	3	pi	X3	2	4	pi	X3	2	4	qi
X1				X1				X2			
-1	1/8	1/8	1/4	-1	3/32	5/32	1/4	1	3/16	5/16	1/2
1	3/8	3/8	3/4	1	3/16	9/16	3/4	3	3/32	13/32	1/2
qj	1/2	1/2	1	rk	9/32	23/32	1	rk	9/32	21/32	1

from which are also observed one-dimensional distributions (distributions of random variables considered  $X_1, X_2, X_3$ ). From these tables we deduce by immediate calculation:

$$M(X_1) = \frac{1}{2}; M(X_1^2) = 1; D^2(X_1) = \frac{3}{4}$$

$$M(X_2) = 2; M(X_2^2) = 5; D^2(X_2) = 1$$

$$M(X_3) = \frac{55}{16}; M(X_3^2) = \frac{101}{8}; D^2(X_3) = \frac{207}{256}$$

$$M(X_1 \cdot X_2) = 1; M(X_1)M(X_2) = 1; C(X_1, X_2) = 0; r(X_1, X_2) = 0$$

$$M(X_1 \cdot X_3) = \frac{29}{16}; M(X_1)M(X_3) = \frac{55}{32}; C(X_1, X_3) = \frac{3}{32}; r(X_1, X_3) = \sqrt{\frac{3}{207}} \approx 0,12$$

$$M(X_2 \cdot X_3) = \frac{113}{16}; M(X_2)M(X_3) = \frac{55}{8};$$

$$C(X_2, X_3) = \frac{3}{16}; r(X_2, X_3) = \frac{3}{\sqrt{207}} \approx 0,21$$

And as a result we can write correlation matrices:

$$C(Z) = \begin{pmatrix} 3/4 & 0 & 3/32 \\ 0 & 1 & 3/16 \\ 3/32 & 3/16 & 207/256 \end{pmatrix} \text{ And } R(Z) = \begin{pmatrix} 1 & 0 & 0,12 \\ 0 & 1 & 0,21 \\ 0,12 & 0,21 & 1 \end{pmatrix}$$

finding that  $X_1$  and  $X_2$  are independent, while between  $X_3$  and  $X_1$  or  $X_3$  and  $X_2$  there is some dependence, even if it is not strong.

### Other numerical characteristics

- It is called **the median** of a random variable  $X$ , the numerical characteristic  $M_e$  that verifies the relationship:  $P(X \geq M_e) \geq \frac{1}{2} \leq P(X \leq M_e)$
- If  $F$  is the distribution function and is continuous  $M_e$ , then it is determined from the equation  $F(M_e) = \frac{1}{2}$ .
- If  $M_e \in (a, b]$  then it is taken  $M_e = \frac{a+b}{2}$
- A **modal value or modulus** of the random variable  $X$  is called any local maximum point of the distribution of  $X$  (in the discrete case) and probability density (in the continuous case).
- If there is only one local maximum, we say that  $X$ 's law is unimodal, otherwise we call it plurimodal.
- The **asymmetry** (Fischer's coefficient) of the random variable  $X$  is called the numerical characteristic defined by  $s = \frac{\mu_3}{\sigma^3}$ .

- It is called **an excess of random variable X, the numerical characteristic defined by**  $e = \frac{\mu_4}{\sigma^4} - 3$

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- If  $e < 0$  then the distribution graph has a flattened appearance and the law is called platycurtic.
- If  $e > 0$  then the distribution graph has a sharp aspect and the law will be called leptocurtic.
- If  $e = 0$  then the distributions are mesocurtic.
- If  $X$  is a random variable with the distribution function  $F(x)$ , the **quartiles (in number of three) of X (or of the distribution of X)** are called the numbers  $q_1$ ,  $q_2$  and  $q_3$  with the properties:

$$\left\{ \begin{array}{l} F(q_1) \leq \frac{1}{4} \\ F(q_1 + 0) \geq \frac{1}{4} \end{array} \right. \quad \left\{ \begin{array}{l} F(q_2) \leq \frac{1}{2} \\ F(q_2 + 0) \geq \frac{1}{2} \end{array} \right. \quad \left\{ \begin{array}{l} F(q_3) \leq \frac{3}{4} \\ F(q_3 + 0) \geq \frac{3}{4} \end{array} \right.$$

We note that  $q_2 = M_e$ .