

# STAT151A Final Exam (Spring 2024)

Please write your full name and email address:

- You have 3 hours for this exam.
- You will be allowed one double-sided cheat sheet.
- Please turn in your cheat sheet with the exam.

In the exam, you will find seven questions. From these, please choose **exactly four** to answer. You will be graded only on the four questions you choose to answer.

Please mark the questions you would like graded with an  $\times$  in the box provided. For example, to select problems 2, 4, 5, and 7, your exam should look like this:

Question 1      ☐  $\leftarrow$  'X' here to grade this question.

Question 2      ☒  $\leftarrow$  'X' here to grade this question.

Question 3      ☐  $\leftarrow$  'X' here to grade this question.

Question 4      ☒  $\leftarrow$  'X' here to grade this question.

Question 5      ☒  $\leftarrow$  'X' here to grade this question.

Question 6      ☐  $\leftarrow$  'X' here to grade this question.

Question 7      ☒  $\leftarrow$  'X' here to grade this question.

If you select more than four questions, we reserve the right to choose which ones we grade.

**Note that each of the seven questions has three parts.** Read the question completely and carefully before answering. Make sure to answer every part to receive full credit.

In this exam, you may use the following results without proof:

- The continuous mapping theorem.
- The law of large numbers and central limit theorems for independent random variables...
  - For univariate random variables,
  - For multivariate random variables,
  - For identically distributed random variables, and
  - For non-identically distributed random variables.
- The OLS estimator is given by  $\hat{\beta} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{Y}$  when  $\mathbf{X}$  is full-rank.
- The OLS estimator satisfies  $(\mathbf{X}^\top \mathbf{X}) \hat{\beta} = \mathbf{X}^\top \mathbf{Y}$ .

## 1 Question 1

○ ← 'X' here to grade this question.

For this question, we'll consider the linear model  $y_n = \beta_0 + \beta_1 w_n + \beta_2 z_n + \varepsilon_n$ .

**(1a)**

Write the set of equations

$$y_n = \beta_0 + \beta_1 w_n + \beta_2 z_n + \varepsilon_n \quad \text{for } n \in \{1, \dots, N\}$$

in matrix form. That is, let  $\mathbf{X}$  denote an  $N \times 3$  matrix,  $\mathbf{Y}$  and  $\boldsymbol{\varepsilon}$  length- $N$  column vectors, and  $\boldsymbol{\beta} = (\beta_0, \beta_1, \beta_2)^\top$  a length-3 column vector. Then express the matrices  $\mathbf{Y}$ ,  $\mathbf{X}$ , and  $\boldsymbol{\varepsilon}$  in terms of the scalars  $y_n$ ,  $w_n$ ,  $z_n$ , and  $\varepsilon_n$  so that  $\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$  is equivalent to the set of regression equations.

**(1b)**

Define the following quantities:

$$\begin{aligned}\bar{z} &:= \frac{1}{N} \sum_{n=1}^N z_n & \bar{w} &:= \frac{1}{N} \sum_{n=1}^N w_n \\ \overline{zz} &:= \frac{1}{N} \sum_{n=1}^N z_n^2 & \overline{ww} &:= \frac{1}{N} \sum_{n=1}^N w_n^2 & \overline{zw} &:= \frac{1}{N} \sum_{n=1}^N z_n w_n \\ \overline{zy} &:= \frac{1}{N} \sum_{n=1}^N z_n y_n & \overline{wy} &:= \frac{1}{N} \sum_{n=1}^N w_n y_n & \bar{y} &:= \frac{1}{N} \sum_{n=1}^N y_n\end{aligned}$$

Write an explicit expressions for  $\frac{1}{N} \mathbf{X}^\top \mathbf{X}$  and  $\frac{1}{N} \mathbf{X}^\top \mathbf{Y}$  in terms of these quantities.

**(1c)**

For this part of the question only, assume that

$$\overline{zw} = \bar{z} = \bar{w} = 0$$

.

In terms of the quantities defined in part (b), write a closed-form expression for  $\hat{\beta}$ , the OLS estimator of the vector  $\beta$ .

**Hint:** The inverse of a  $3 \times 3$  diagonal matrix is given by

$$\begin{pmatrix} v_1 & 0 & 0 \\ 0 & v_2 & 0 \\ 0 & 0 & v_3 \end{pmatrix}^{-1} = \begin{pmatrix} v_1^{-1} & 0 & 0 \\ 0 & v_2^{-1} & 0 \\ 0 & 0 & v_3^{-1} \end{pmatrix}$$

## 2 Question 2

○ ← ‘X’ here to grade this question.

For this question, consider the linear models

$$y_n = \beta^\top \mathbf{x}_n + \varepsilon_n \quad \text{and} \quad y_n = \gamma^\top \mathbf{z}_n + \eta_n$$

with

$$\mathbf{x}_n = \begin{pmatrix} 1 \\ x_n \end{pmatrix} \quad \text{and} \quad \mathbf{z}_n = \begin{pmatrix} 1 \\ z_n \end{pmatrix} \quad \text{where } z_n := 10x_n$$

Assume that  $x_n$  is not a constant (i.e., for at least one pair  $n$  and  $m$ ,  $x_n \neq x_m$ ).

Let  $\mathbf{X}$  denote the  $N \times 2$  matrix whose  $n$ -th row is  $\mathbf{x}_n^\top$ , and  $\mathbf{Z}$  denote the  $N \times 2$  matrix whose  $n$ -th row is  $\mathbf{z}_n^\top$ .

**(2a)**

Using the definitions above, find a  $2 \times 2$  matrix  $\mathbf{A}$  such that  $\mathbf{z}_n = \mathbf{A}\mathbf{x}_n$ . Then, show that  $\mathbf{Z} = \mathbf{X}\mathbf{A}^\top$  for the same matrix  $\mathbf{A}$ .



**(2b)**

Suppose you know that the OLS estimate of  $\beta$  is given by  $\hat{\beta} = (4, 30)$ . What is the value of  $\hat{\gamma}$ , the OLS estimate of  $\gamma$ ? Please use the definitions above and your answer for part (a), and justify your answer.

**(2c)**

Now consider the two models' prediction on a new datapoint with  $\mathbf{x}_{\text{new}} = (1, 50)^\top$  — and so, necessarily,  $\mathbf{z}_{\text{new}} = (1, 500)^\top$  — with respective prediction errors

$$\varepsilon_{\text{new}}^\beta := y_{\text{new}} - \hat{\beta}^\top \mathbf{x}_{\text{new}} \quad \text{and} \quad \varepsilon_{\text{new}}^\gamma := y_{\text{new}} - \hat{\gamma}^\top \mathbf{z}_{\text{new}}.$$

Please select which of (a), (b), (c), or (d) is correct for this particular value of  $\mathbf{x}_{\text{new}}$  and  $\mathbf{z}_{\text{new}}$ :

- a) It is always the case that  $\left| \varepsilon_{\text{new}}^\beta \right| = \left| \varepsilon_{\text{new}}^\gamma \right|$
- b) It is always the case that  $\left| \varepsilon_{\text{new}}^\beta \right| < \left| \varepsilon_{\text{new}}^\gamma \right|$
- c) It is always the case that  $\left| \varepsilon_{\text{new}}^\beta \right| > \left| \varepsilon_{\text{new}}^\gamma \right|$
- d) In general, we cannot determine the relationship between  $\left| \varepsilon_{\text{new}}^\beta \right|$  and  $\left| \varepsilon_{\text{new}}^\gamma \right|$  using the information provided.

Please justify your answer carefully.

### 3 Question 3

○ ← ‘X’ here to grade this question.

For this question, assume that you have access to a function

$$\Phi(z) = \mathbb{P}(\tilde{z} \leq z),$$

as well as its inverse,

$$\Phi^{-1}(p) = z \text{ such that } \mathbb{P}(\tilde{z} \leq z) = p, \text{ for } p \in (0, 1),$$

where  $\tilde{z} \sim \mathcal{N}(0, 1)$  denote a scalar-valued standard normal random variable.

**(3a)**

Suppose that  $\tilde{y} \sim \mathcal{N}(\mu, \sigma^2)$ , where  $\mu$  and  $\sigma$  are known. Using only the functions  $\Phi(\cdot)$ ,  $\Phi^{-1}(\cdot)$ , and the known quantities  $\mu$ , and  $\sigma$ , find a quantity  $a$  such that

$$\mathbb{P}(\tilde{y} \leq a) = 0.90.$$

Note that  $\Phi(\cdot)$  is only for a standard normal random variable. Please **do not assume** that you have direct access to the distribution and quantile functions of generic normal random variables.

Please justify your answer carefully.

**(3b)**

Now, consider the OLS estimator under normal assumptions, so that

$$\hat{\boldsymbol{\beta}} \sim \mathcal{N}\left(\boldsymbol{\beta}, \sigma^2(\mathbf{X}^\top \mathbf{X})^{-1}\right).$$

Note that  $\hat{\boldsymbol{\beta}}$  and  $\boldsymbol{\beta}$  are  $P$ -dimensional vectors,  $\sigma$  is a scalar, and  $(\mathbf{X}^\top \mathbf{X})^{-1}$  is a  $P \times P$  matrix.

Assume that  $\boldsymbol{\beta}$ ,  $\sigma$ , and  $\mathbf{X}$  are all known. In terms of these quantities, find the distribution of  $\hat{\beta}_1$ , the first component of  $\hat{\boldsymbol{\beta}}$ .

**(3c)**

Combining your answers from parts (a) and (b), find a scalar  $b$  such that

$$\mathbb{P}\left(\hat{\beta}_1 \leq b\right) = 0.90.$$

#### 4 Question 4

○ ← ‘X’ here to grade this question.

For this question:

- Let  $\mathbf{x}_n = (1, z_n)^\top$ , where  $z_n \sim \mathcal{N}(0, 2)$ .
- Assume that  $y_n = \boldsymbol{\beta}^\top \mathbf{x}_n + \varepsilon_n$  for each  $n$  and some  $\boldsymbol{\beta} = (\beta_1, \beta_2)^\top$ , where  $\boldsymbol{\beta}$  is a length 2 column vector.
- Assume the residuals  $\varepsilon_n$  are IID with  $\mathbb{E}[\varepsilon_n] = 0$  and  $\mathbb{E}[\varepsilon_n^2] = 1$ , but **not necessarily normal**.
- Assume that the residuals  $\varepsilon_n$  are all independent of all the  $z_n$ .

**(4a)**

Let  $\mathbf{X}$  denote the  $N \times P$  matrix consisting of the observation  $\mathbf{x}_n^\top$  in the  $n$ -th row, and let  $\mathbf{Y}$  denote the  $N$ -vector with  $y_n$  in the  $n$ -th entry.

Write the matrices  $\frac{1}{N}\mathbf{X}^\top\mathbf{X}$  and  $\frac{1}{N}\mathbf{X}^\top\mathbf{Y}$  in terms of  $\beta$ ,  $\varepsilon_n$ ,  $z_n$ ,  $N$ , and constants. Note that  $\mathbf{X}^\top\mathbf{X}$  is a  $2 \times 2$  matrix and  $\mathbf{X}^\top\mathbf{Y}$  is a 2-vector.



**(4b)**

Evaluate the following limits (which will denote convergence in probability):

$$\frac{1}{N} \mathbf{X}^\top \mathbf{X} \rightarrow? \quad \text{and} \quad \frac{1}{N} \mathbf{X}^\top \mathbf{Y} \rightarrow?$$

Note that your answers may depend on  $\beta$ , but should not depend on  $\mathbf{x}_n$  or  $y_n$ , since they are limiting quantities that do not depend on the particular dataset.

Justify your conclusion carefully (state which theorems you use).

**(4c)**

Using your answers from part (a) and (b), find the limit (convergence in probability) of

$$\hat{\beta} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{Y} \rightarrow ?$$

Note that your answer may depend on  $\beta$ , but should not depend on  $\mathbf{x}_n$  or  $y_n$ , since they are limiting quantities that do not depend on the particular dataset.

Justify your conclusion carefully (state which theorems you use).

## 5 Question 5

○ ← ‘X’ here to grade this question.

Let  $\tilde{\mathbf{x}}_n$  denote an IID sequence of random 2-dimensional vectors in  $\mathbb{R}^P$  (**not necessarily normal**), with

$$\tilde{\mathbf{x}}_n = \begin{pmatrix} \tilde{x}_{n1} \\ \tilde{x}_{n2} \end{pmatrix} \quad \text{and} \quad \mathbb{E}[\tilde{\mathbf{x}}_n] = \mathbf{0} \quad \text{and} \quad \text{Cov}(\tilde{\mathbf{x}}_n) =: \mathbf{\Sigma} = \begin{pmatrix} 2 & 1 \\ 1 & 4 \end{pmatrix}.$$

**(5a)**

Find the limiting distribution of the vector

$$\frac{1}{\sqrt{N}} \sum_{n=1}^N \tilde{\mathbf{x}}_n \rightarrow ?$$

Justify your conclusion carefully.

**(5b)**

Using the univariate central limit theorem, find the limiting distributions of the difference between the components of  $\tilde{\mathbf{x}}_n$ :

$$\frac{1}{\sqrt{N}} \sum_{n=1}^N (\tilde{x}_{n1} - \tilde{x}_{n2}) \rightarrow ?$$

Justify your conclusion carefully.

**(5c)**

Find a vector  $\mathbf{v}$  such that  $\mathbf{v}^\top \tilde{\mathbf{x}}_n = \tilde{\mathbf{x}}_{n1} - \tilde{\mathbf{x}}_{n2}$ . Using this vector, show that the solution to (b) also follows from the solution to (a) and the continuous mapping theorem.

Justify your conclusion carefully.

## 6 Question 6

○ ← ‘X’ here to grade this question.

Given a regression on  $\mathbf{X}$  with  $P$  regressors, and the corresponding  $\mathbf{Y}$ ,  $\hat{\mathbf{Y}}$ , and  $\hat{\varepsilon}$ , define the following quantities:

$$RSS := \hat{\varepsilon}^\top \hat{\varepsilon} \quad (\text{Residual sum of squares})$$

$$TSS := \mathbf{Y}^\top \mathbf{Y} \quad (\text{Total sum of squares})$$

$$ESS := \hat{\mathbf{Y}}^\top \hat{\mathbf{Y}} \quad (\text{Explained sum of squares})$$

$$R^2 := \frac{ESS}{TSS}.$$

**6a**

1. Prove that  $RSS + ESS = TSS$ .
2. Express  $R^2$  in terms of  $TSS$  and  $RSS$ .



**6b**

For each of the following questions, **please justify your answer**. An intuitive explanation is enough; a proof or specific example is not necessary.

1. What is  $R^2$  when we include no regressors? ( $P = 0$  and  $\hat{y}_n = 0$  for all  $n$ )
2. What is  $R^2$  when we include  $N$  linearly independent regressors? ( $P = N$ )
3. Can  $R^2$  ever decrease when we add a regressor?
4. Can  $R^2$  ever stay the same when we add a regressor?
5. Can  $R^2$  ever increase when we add a regressor?

**6c**

For each of the following questions, **please justify your answer**. An intuitive explanation is enough; a proof or specific example is not necessary.

These questions will be about the F-test statistic for the null  $H_0 : \boldsymbol{\beta} = \mathbf{0}$ ,

$$\phi = \frac{\hat{\boldsymbol{\beta}}^\top (\mathbf{X}^\top \mathbf{X}) \hat{\boldsymbol{\beta}}}{P \hat{\sigma}^2},$$

where  $\hat{\sigma}^2 := \frac{1}{N-P} \sum_{n=1}^N \hat{\varepsilon}_n^2$ .

1. Write the F-test statistic  $\phi$  in terms of  $TSS$  and  $RSS$ , and  $P$ .
2. Can  $\phi$  ever decrease when we add a regressor?
3. Can  $\phi$  ever stay the same when we add a regressor?
4. Can  $\phi$  ever increase when we add a regressor?

## 7 Question 7

○ ← ‘X’ here to grade this question.

For this question, we will take

$$a_n \sim \mathcal{N}(0, 1) \quad \text{and} \quad b_n = a_n^3.$$

We assume the pairs  $(a_n, b_n)$  are IID, but  $a_n$  and  $b_n$  are not independent. Assume that, for some  $\beta_a$  and  $\beta_b$ ,

$$y_n = \beta_a a_n + \beta_b b_n + \varepsilon_n,$$

where  $\varepsilon_n$  are IID with  $\mathbb{E}[\varepsilon_n] = 0$  and  $\text{Var}(\varepsilon_n) < \infty$ . The residuals  $\varepsilon_n$  and  $a_n$  are all independent of one another. **Note that the residuals are not necessarily normal.**

**(7a)**

Let  $\hat{\alpha}$  denote the OLS estimator of  $y_n \sim \alpha a_n$ , that is, of  $y_n$  regressed on  $a_n$  alone. Note that the regression for  $\hat{\alpha}$  does not include a constant, and does not include  $b_n$ .

Recall that

$$\hat{\alpha} = \frac{\sum_{n=1}^N y_n a_n}{\sum_{n=1}^N a_n^2},$$

and find the limit

$$\hat{\alpha} \rightarrow ?$$

as  $N \rightarrow \infty$ .

The answer may depend on the unknown  $\beta_a$  and  $\beta_b$ .

**Hint:** Standard properties of the normal gives that  $\mathbb{E}[a_n^3] = 0$  and  $\mathbb{E}[a_n^4] = 3$ .

**(7b)**

Letting  $\hat{y}_{\text{new}} = \hat{a}a_{\text{new}}$ , find

$$\mathbb{E}[y_{\text{new}} - \hat{y}_{\text{new}} | a_{\text{new}}, \mathbf{Y}, \mathbf{A}],$$

where  $\mathbf{A} = (a_1, \dots, a_N)^\top$  is the vector of  $a_n$  observations and  $\mathbf{Y} = (y_1, \dots, y_N)^\top$  is the vector of responses in the training set. Note that the expectation is conditional on the training data and on the new regressor, so the only randomness is in  $\varepsilon_{\text{new}}$ .

The answer may depend on the unknown  $\beta_a$  and  $\beta_b$ .

**(7c)**

Assume that  $\beta_b \neq 0$ , that  $N$  is very large.

- What does your result from (a) imply about using  $\hat{a}$  for inference on  $\beta_a$ ?
- What does your result from (b) imply about using  $\hat{a}$  for prediction of  $y_{\text{new}}$ ?