

LECTURE FIVE

Multiple Linear Regression

Data: y_i $x_{i1} \ x_{i2} \ \dots \ x_{im}$, $i=1, \dots, n$

Model: $y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_m x_{im} + \varepsilon_i$
 $\varepsilon_i \text{ iid } N(0, \sigma^2)$

Inference: Prior: $\beta_0, \dots, \beta_m, \log \sigma \text{ iid } \text{Unif}(-C, C)$

posterior:
 $f(\beta_0, \dots, \beta_m) \propto \left[\frac{S(\hat{\beta}_0, \dots, \hat{\beta}_m)}{S(\beta_0, \dots, \beta_m)} \right]^{\frac{n}{2}}$ } peaked at the least squares estimators

$$S(\beta_0, \dots, \beta_m) = \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_{i1} - \dots - \beta_m x_{im})^2$$

$\hat{\beta}_0, \dots, \hat{\beta}_m$: least squares estimators

$$S(\hat{\beta}_0, \dots, \hat{\beta}_m)$$

Multivariate - density

$x_1 \dots x_p$

$$f(x_1 \dots x_p) \propto$$

$$x = \begin{pmatrix} x_1 \\ \vdots \\ x_p \end{pmatrix}$$

$$\left[1 + \frac{1}{v} (x - \mu)^T \Sigma^{-1} (x - \mu) \right]^{-\frac{v+p}{2}}$$

Quadratic Form

$$t_p(\mu, \Sigma, v)$$

\downarrow
p x p matrix scalars
degrees of freedom

① Connection to Multivariate Normal

$$N_p(\mu, \Sigma) \rightarrow p \times p : \text{Covariance Matrix}$$

$$f(x_1, \dots, x_p) = \exp\left[-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu)\right]$$



μ : mean
 Σ : covariance matrix.

Suppose:
 $X \sim N_p(\mu, \Sigma)$, $V \sim \chi^2_\nu$ independent.

Fact:

$$T = \mu + \frac{X - \mu}{\sqrt{V/\nu}}$$

$\hookrightarrow t_p(\mu, \Sigma, \nu)$

$$V \sim \chi^2_\nu$$

$$V = Z_1^2 + \dots + Z_\nu^2$$

$Z_i \stackrel{\text{i.i.d.}}{\sim} N(0, 1)$

$$② \quad T_j = \mu_j + \frac{X_j - \mu_j}{\sqrt{V/\nu}}$$

$$X_j \sim N(\mu_j, \Sigma(j, j))$$

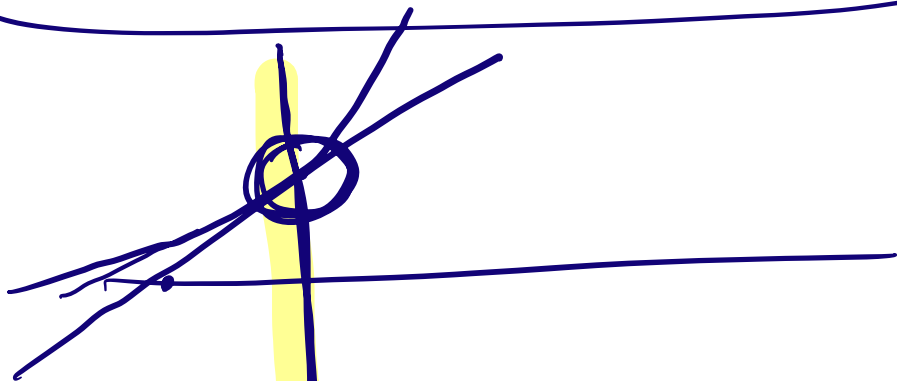
$$T_j \sim t_1(\mu_j, \Sigma(j, j), \nu)$$

③ If the degrees of freedom ν is large, then the t-density is the same as the Normal density.
 (Regression: $\nu = n - m - 1$)

$$\left[\frac{1}{1 + \frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu)} \right]^{\frac{1+p}{2}}$$

$$1 + \frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu) \approx \exp\left(\frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu)\right)$$

$1 + a \approx e^a$ when a is small



Back to Regression

posterior $\propto \left[\frac{S(\hat{\beta})}{S(\beta)} \right]^{\frac{n}{2}}$

$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

\downarrow
 $n \times 1$

$$X = \begin{bmatrix} 1 & x_{11} & \dots & x_{1m} \\ \vdots & x_{21} & & x_{2m} \\ \vdots & \vdots & & \vdots \\ 1 & x_{n1} & & x_{nm} \end{bmatrix}$$

\downarrow
 $n \times (m+1)$

$$\beta = \begin{pmatrix} \beta_0 \\ \vdots \\ \beta_m \end{pmatrix}$$

\downarrow
 $(m+1) \times 1$

$$S(\beta) = \|y - X\beta\|^2$$

$$= \|y - X\hat{\beta}\|^2 + \|X\hat{\beta} - X\beta\|^2$$

$$= S(\hat{\beta}) + (\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta)$$

posterior:

$$\left(\frac{S(\hat{\beta})}{S(\beta)} \right)^{\frac{n}{2}}$$

$$= \left(\frac{S(\hat{\beta})}{S(\hat{\beta}) + (\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta)} \right)^{\frac{n}{2}}$$

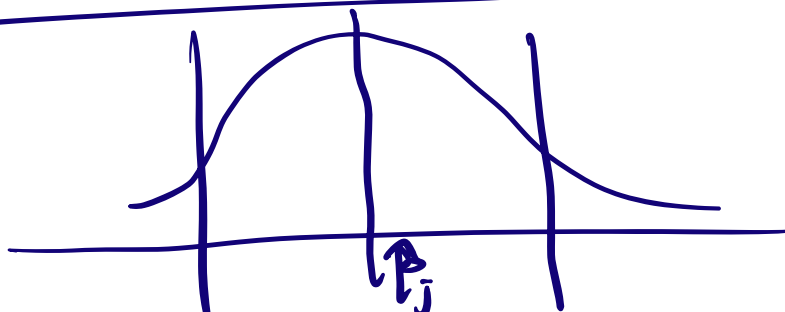
$$= \left[\frac{1}{1 + \frac{1}{S(\hat{\beta})} (\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta)} \right]^{\frac{n}{2}}$$

$$\rightarrow t_{m+1} \left(\hat{\beta}, \frac{S(\hat{\beta})}{n-m-1} (X^T X)^{-1}, n-m-1 \right)$$

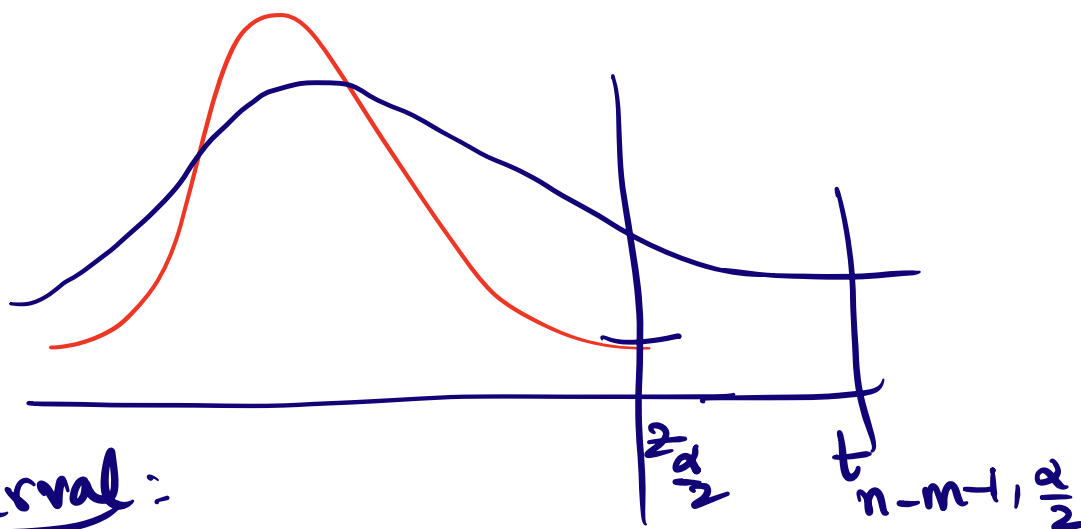
$$\boxed{\beta_j} \mid \beta_j \text{ data} \sim t_1 \left(\hat{\beta}_j, \frac{S(\hat{\beta})}{n-m-1} (X^T X)^{j+1, j+1}, n-m-1 \right)$$

$$(X^T X)^{j+1, j+1} = (j+1, j+1)\text{th diagonal entry of } (X^T X)$$

Uncertainty Interval for β_j :



$t_{n-m-1, \frac{\alpha}{2}}$: $(1-\frac{\alpha}{2})$ quantile of univariate standard t with $n-m-1$ d.f



Interval:

$$\left[\hat{\beta}_j - \sqrt{\frac{S(\hat{\beta})}{n-m-1}} \sqrt{(X^T X)^{j+1, j+1}} t_{n-m-1, \frac{\alpha}{2}}, \right. \\ \left. \hat{\beta}_j + \sqrt{\frac{S(\hat{\beta})}{n-m-1}} \sqrt{(X^T X)^{j+1, j+1}} t_{n-m-1, \frac{\alpha}{2}} \right]$$

Bayesian Inference ($\beta_0, \dots, \beta_m, \log \sigma \sim \text{Unif}(\mathbb{C})$)
C large)

||

Frequentist Inference
 $(X^T X)^{j+1, j+1} = ((X^T X)^{-1})^{(j+1, j+1)}$

① $S(\hat{\beta}) = \sum (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_{i1} - \dots - \hat{\beta}_m x_{im})^2$
↓
Smallest possible sum of squares
RSS (Residual Sum of Squares)

Residual: $y_i - \hat{\beta}_0 - \hat{\beta}_1 x_{i1} - \dots - \hat{\beta}_m x_{im}$
 $\rightarrow i^{\text{th}} \text{ Residual}$

② $n - m - 1$: Residual Degrees of Freedom

③ $\frac{S(\hat{\beta})}{n - m - 1} = \frac{RSS}{\text{Residual DF}} = \hat{\sigma}^2$

$\hat{\sigma} = \sqrt{\frac{RSS}{\text{Residual DF}}}$ } Residual Standard Error

④ $\sqrt{\frac{S(\hat{\beta})}{n - m - 1}} \sqrt{(X^T X)^{-1}}_{(j+1, j+1)}$
 $\rightarrow \text{Standard Error for } \beta_j$

Nonlinear Regression

CA population data

$y_t = \beta_0 + \beta_1 t + \varepsilon_t$
 \uparrow
 log population

$y_t = \beta_0 + \beta_1 t + \beta_2 \text{ReLU}(t - c) + \varepsilon_t$

$\text{ReLU}(z) = \max(z, 0) = z_+$

$$y_t = \beta_0 + \beta_1 t + \beta_2 (t-c)_+ + \varepsilon_t$$

$$t \leq c \rightarrow \beta_0 + \beta_1 t \rightarrow \text{slope: } \beta_1$$

$$t > c \rightarrow \beta_0 + \beta_1 t + \beta_2 (t-c) \rightarrow \text{slope: } \beta_1 + \beta_2$$

① What is c ?

② $\beta_1 \Delta \beta_1 + \beta_2$

Compare to Cubic:

$$y_t = \beta_0 + \beta_1 t + \beta_2 t^2 + \beta_3 t^3 + \varepsilon_t$$

know c :

$$X_c = \begin{bmatrix} 1 & 1 & (1-c)_+ \\ \vdots & \vdots & \vdots \\ 1 & n & (n-c)_+ \end{bmatrix}$$

$$\text{sm-OLS}(y, X_c)$$

$$\text{sm-OLS}(y, X)$$

$$X = \begin{bmatrix} 1 & 1 & 1^2 & 1^3 \\ \vdots & \vdots & \vdots & \vdots \\ 1 & n & n^2 & n^3 \end{bmatrix}$$

$$\varepsilon_t \stackrel{\text{i.i.d.}}{\sim} N(0, \sigma^2)$$

Parameter Estimation

$$\beta_0, \beta_1, \beta_2, c, \sigma$$

↑
Least Squares:

$$\sum_{t=1}^n \left[y_t - \beta_0 - \beta_1 t - \beta_2 \text{ReLU}(t-c) \right]^2$$

$$S(\beta, c)$$

$$\beta = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{pmatrix}$$

minimize over β, c

Strategy

- ① Fix c & minimize $S(\beta, c)$ over β
- Same as doing linear regression with
- $$X_c = \begin{bmatrix} 1 & \frac{1}{2} & \text{ReLU}(1-c) \\ \vdots & \vdots & \vdots \\ 1 & \frac{1}{n} & \text{ReLU}(n-c) \end{bmatrix}$$

Do regression of y on X_c .

Calculate least squares estimate $\hat{\beta}_c$,
& $\text{RSS}(c)$

- ② Minimize $\text{RSS}(c)$ over c .

$$c \in \{1, \dots, n\}$$