

# Completeness

## Outline

- 1) Completeness
- 2) Ancillarity
- 3) Basu's Theorem

# Completeness

Def  $T(X)$  is complete for  $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$

$$\text{if } \mathbb{E}_\theta f(T(X)) = 0 \quad \forall \theta$$

$$\Rightarrow f(T) \stackrel{a.s.}{=} 0 \quad \forall \theta$$

[Name comes from a prior notion that

$\mathcal{P}^T = \{P_\theta^T : \theta \in \Theta\}$  is "complete basis"

wrt inner product  $\langle f, P_\theta^T \rangle = \int f(t) dP_\theta^T(t)$

(see HW 3)

Ex. (Cont'd) Laplace location family has  
minimal suff. stat.  $S = (X_{(i)})_{i=1}^n$ . Complete?

No: Let  $M(S) = \text{median}(X)$

$$\bar{X}(S) = \frac{1}{n} \sum X_i$$

$$\mathbb{E}_\theta \bar{X} = \mathbb{E}_\theta M = \theta \quad (\text{by symmetry})$$

$$\mathbb{E}_\theta [\bar{X}(S) - M(S)] = 0 \quad \forall \theta$$

$S(X)$  still has "a lot of extra fluff"

Ex  $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} U[0, \theta] \quad \theta \in (0, \infty)$

Can show  $T(X) = X_{(n)}$  min. suff. Complete?

Find density of  $T(X)$ :

$$P_{\theta}(T \leq t) = \left(\frac{t}{\theta} \wedge 1\right)^n = \left(\frac{t}{\theta}\right)^n \wedge 1$$

$$\Rightarrow p_{\theta}(t) = \frac{d}{dt} P_{\theta}(T \leq t)$$

$$= n \frac{t^{n-1}}{\theta^n} \mathbb{1}_{\{t \leq \theta\}}$$

Suppose  $0 = E_{\theta} f(T) \quad \forall \theta > 0$

$$= \frac{n}{\theta^n} \int_0^{\theta} f(t) t^{n-1} dt \quad \forall \theta > 0$$

$$\Rightarrow \int_0^{\theta} f(t) t^{n-1} dt = 0 \quad \forall \theta > 0$$

$$\Rightarrow f(t) t^{n-1} = 0 \quad \text{a.e. } t > 0$$

Def Assume  $\mathcal{P} = \{P_\eta : \eta \in \Xi\}$  has densities

$$p_\eta(x) = e^{\eta' T(x) - A(\eta)} h(x)$$

If  $T(x)$  satisfies no linear constraint  $\begin{pmatrix} \beta \neq 0, \alpha: \\ \beta' T(x) \stackrel{a.s.}{=} \alpha \end{pmatrix}$   
and  $\Xi$  contains an open set, we say  
 $\mathcal{P}$  is full-rank

If  $\mathcal{P}$  is not full-rank we say it is curved

[Note: If  $T(x)$  satisfies linear constraint, then  
 $\mathcal{P}$  might still be full-rank for a lower-dim.  
sufficient statistic]

Proof in Lehmann & Romano, Thm. 4.3.1

Theorem If  $\mathcal{P}$  is full rank then  
 $T(x)$  is complete sufficient

Proof idea wlog  $T(x) = x$ ,  $p_\eta(x) = e^{\eta' x - A(\eta)}$ ,  $0 \in \overset{\text{interior}}{\Xi}^\circ$

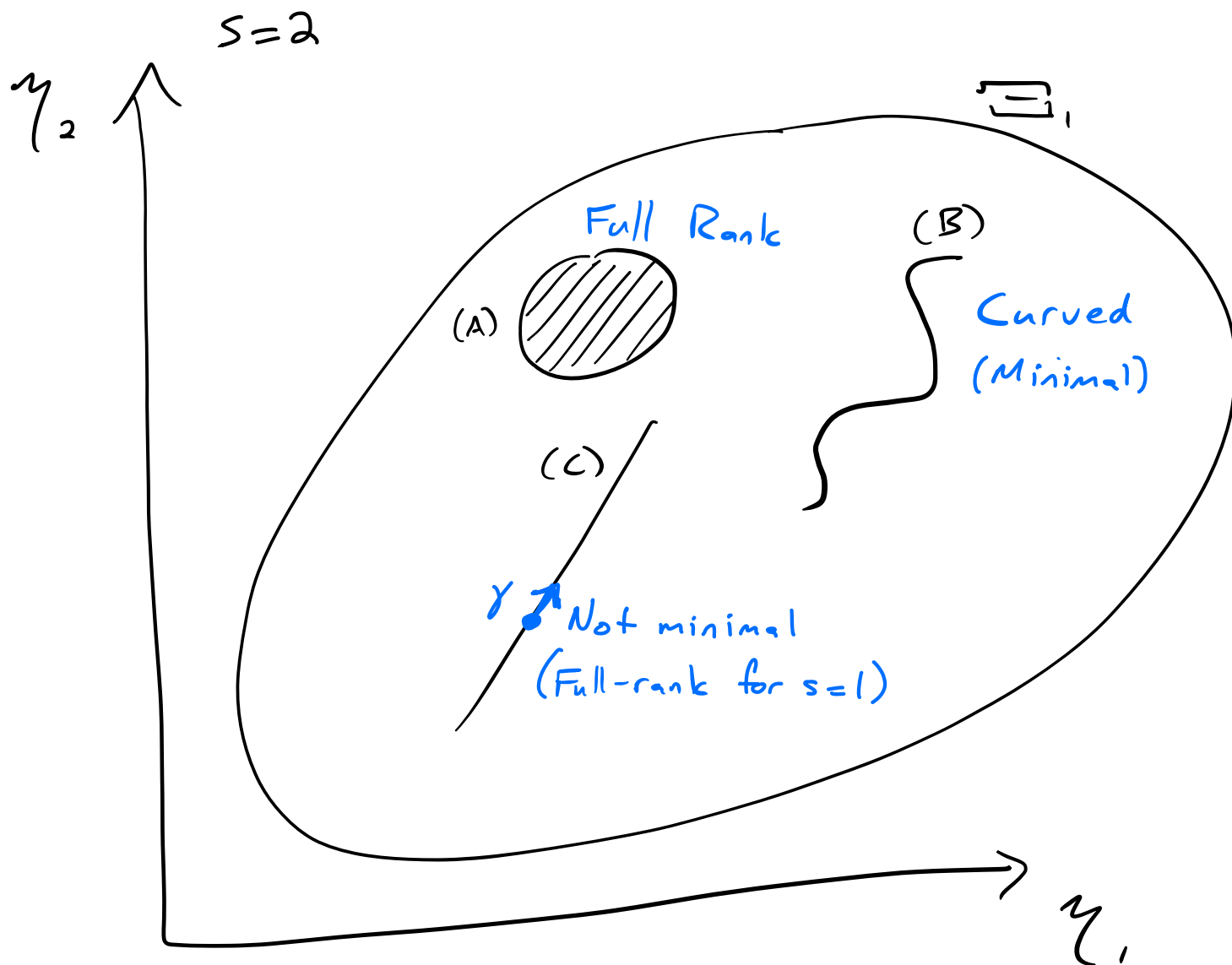
Write  $f(x) = f^+(x) - f^-(x)$ , for  $f^+, f^- \geq 0$

$$\int e^{\eta' x} f^+(x) d\mu(x) = \int e^{\eta' x} f^-(x) d\mu(x)$$

$$\text{" MGF for } Y^+ \sim \frac{f^+(x)}{\int f^+ d\mu} \quad \text{" MGF for } Y^- \sim \frac{f^-(x)}{\int f^- d\mu}$$

Uniqueness of MGFs  $\Rightarrow Y^+ \stackrel{D}{=} Y^- \Rightarrow f^+ \stackrel{a.s.}{=} f^-$

## Diagram again



$T(x)$  definitely complete for (A)

Maybe not for (B), (C)

Theorem If  $T(X)$  complete sufficient  
for  $\mathcal{P}$  then  $T(X)$  is minimal

Game plan for completeness proofs: show two things are  
a.s. equal by showing they have = expectation.

Proof Assume  $S(X)$  is minimal suff

$$\text{Let } \bar{T}(S(X)) = \mathbb{E}_{\theta} [T(X) \mid S(X)]$$

~~$\theta$~~   $\leftarrow S \text{ suff.}$

$$\text{Claim: } \bar{T}(S(X)) \stackrel{\text{a.s.}}{=} T(X)$$

$$\text{We have } S(X) \stackrel{\text{a.s.}}{=} f(T(X)) \quad (S \text{ minimal suff})$$

$$\text{Let } g(t) = t - \bar{T}(f(t))$$

$$\begin{aligned} \mathbb{E}_{\theta} [g(T(X))] &= \mathbb{E}_{\theta} T(X) - \mathbb{E}_{\theta} [\bar{T}(S(X))] \\ &= \mathbb{E}_{\theta} T(X) - \mathbb{E}_{\theta} [\mathbb{E}[T \mid S]] \\ &= 0 \end{aligned}$$

$$\Rightarrow g(T(X)) \stackrel{\text{a.s.}}{=} 0 \quad (\text{completeness}) \quad \square$$

# Ancillarity

Two reasons to care about completeness:

1) Uniqueness of unbiased estimators using  $T$

$$\text{If } \mathbb{E}_\theta \delta_1(T) = \mathbb{E}_\theta \delta_2(T) = g(\theta), \forall \theta \in \Theta$$

$$\text{Then } \mathbb{E}_\theta[\delta_1 - \delta_2] = 0 \Rightarrow \delta_1 \stackrel{a.s.}{=} \delta_2$$

[We will explore this further next time]

2) Basu's theorem: neat way to show independence

Def  $V(X)$  is ancillary for  $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$   
if its distribution does not depend  
on  $\theta$ . ( $V$  carries no info. about  $\theta$ )

(Aside:) Conditionality Principle

If  $V(X)$  is ancillary then all inference  
should be conditional on  $V(X)$

[will return to this in testing & CI unit]

# Basu's Theorem

## Theorem (Basu)

If  $T(X)$  is complete sufficient and  $V(X)$  is ancillary for  $\mathcal{J}$ , then

$$V(X) \perp\!\!\!\perp T(X) \quad \text{for all } \theta \in \Theta$$

## Proof

Want  $P_\theta(V \in A, T \in B) = P_\theta(V \in A) P_\theta(T \in B)$  all  $A, B, \theta$

$$\text{Let } q_A(T(X)) = P_\theta(V \in A \mid T) \quad \leftarrow T \text{ suff.}$$

$$\rho_A = P_\theta(V \in A) \quad \leftarrow V \text{ ancillary}$$

$$E_\theta[q_A(T) - \rho_A] = \rho_A - \rho_A = 0, \quad \forall \theta$$

$$\Rightarrow q_A(T) \stackrel{\text{a.s.}}{=} \rho_A \quad \forall \theta$$

$$\begin{aligned} P_\theta(V \in A, T \in B) &= \int q_A(t) 1\{t \in B\} dP_\theta^T(t) \\ &= \rho_A \int 1\{t \in B\} dP_\theta^T(t) \\ &= P(V \in A) P_\theta(T \in B) \quad \square \end{aligned}$$



## Using Basu's Theorem

Ancillarity, Completeness, Sufficiency are all properties wrt a family  $\mathcal{P}$

Independence is a property of a distribution

If you can't verify the thm's hypotheses for one family, try a different family!

Ex.  $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2) \quad \mu \in \mathbb{R}, \sigma^2 > 0$

Sample mean  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$

Sample variance  $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

Want to show  $\bar{X} \perp\!\!\!\perp S^2$

But neither stat. is ancillary or sufficient in the full family with  $\mu, \sigma^2$  unknown

To apply Basu, use family with  $\sigma^2$  known:

$$\mathcal{P} = \{ N(\mu, \sigma^2)^n : \mu \in \mathbb{R} \}$$

In  $\mathcal{P}$ ,  $\bar{X}$  is complete sufficient

and  $S^2$  is ancillary since

$$S^2 = \sum (z_i - \bar{z})^2 \text{ for } z_i = X_i - \mu \stackrel{i.i.d.}{\sim} N(0, \sigma^2)$$

↖ not statistics  
but doesn't matter

Therefore  $\bar{X} \perp\!\!\!\perp S^2$

[ Conclusion has nothing to do with "known"  
or "unknown" parameters ]