Sufficiency

9/5/2023

Outline

1) Review
2) Sufficiency
3) Factorization Theorem

Sufficiency

Motivation: Coin flipping

Suppose $X_1, \ldots, X_n \overset{iid}{\sim} \text{Bernoulli}(\theta)$

$$\Rightarrow X \sim \prod_i \theta^{x_i}(1-\theta)^{1-x_i} \quad \text{on} \quad \{0,1\}^n$$

Then $T(X) = \sum X_i \sim \text{Binom}(n, \theta)$

$$= \theta^t (1-\theta)^{n-t} \binom{n}{t} \quad \text{on} \quad \{0, \ldots, n\}$$

$(X_1, \ldots, X_n) \rightsquigarrow T(X)$ is throwing away data. How do we justify this?

In exp. fam. lingo, $T(X)$ is the "sufficient statistic" for X. Today we'll see why we call it that.

Definition Let $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$ be a statistical model for data X. $T(X)$ is <u>sufficient</u> for $\mathcal{P}$ if $P_\theta(X \mid T)$ does not depend on θ

Example (Cont'd)

$$P_\theta(X = x \mid T = t) = \frac{P_\theta(X = x, T = t)}{P_\theta(T = t)}$$

$$= \frac{\theta^{\sum x_i}(1-\theta)^{n-\sum x_i}\, 1\{\sum x_i = t\}}{\theta^t (1-\theta)^{n-t} \binom{n}{t}}$$

$$= 1\{\textstyle\sum x_i = t\} \Big/ \binom{n}{t}$$

So given $T(X) = t$, X is uniform on all seq.s with $\sum x_i = t$

Factorization Theorem

Often, we can identify sufficient stats by inspecting the density.

Theorem (Factorization Theorem)

Let $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$ be a model with densities $p_\theta(x)$ wrt common measure μ.

$T(X)$ is sufficient iff there exist $g_\theta(t)$, $h(x)$ with

$$p_\theta(x) = g_\theta(T(x))\, h(x)$$

for μ-almost-every x: $\mu(\{x : p_\theta(x) \neq g_\theta(T(x)) \cdot h(x)\}) = 0$

[Avoids counterexamples from changing $p_{\theta_0}(x_0)$ some θ_0, x_0]

Rigorous proof in Keener 6.4

Proof (discrete $\mathcal{X}$): Assume wlog $\mu = \#$ on $\mathcal{X}$

$(\Leftarrow)$ $$\mathbb{P}_\theta(X = x \mid T = t) = \frac{\mathbb{P}_\theta(X = x, T(x) = t)}{\mathbb{P}_\theta(T(x) = t)}$$

$$= \frac{\cancel{g_\theta(t)}\, h(x)\, 1\{T(x) = t\}}{\sum_{T(z)=t} \cancel{g_\theta(t)}\, h(z)}$$

$(\Rightarrow)$ Assume $T(x)$ sufficient.

Take $$g_\theta(t) = \sum_{T(x)=t} p_\theta(x)$$

$$= \mathbb{P}_\theta(T(X) = t)$$

For any $\theta_0 \in \Theta$, let

$$h(x) = p_{\theta_0}(x) \Big/ \sum_{T(z)=T(x)} p_{\theta_0}(z)$$

$$= \mathbb{P}_{\cancel{\theta_0}}(X = x \mid T(X) = T(x))$$

no dep. on θ

Then,

$$g_\theta(T(x))\, h(x) = \mathbb{P}_\theta(T = T(x))\, \mathbb{P}(X = x \mid T = T(x))$$

$$= \mathbb{P}_\theta(X = x)$$

$\boxtimes$

Interpretations of Sufficiency

X is informative about θ only because its distribution depends on θ.

We can think of the data as being generated in two stages:

1) Generate T : distribution dep. on θ
2) Generate $X|T$: does <u>not</u> dep on θ

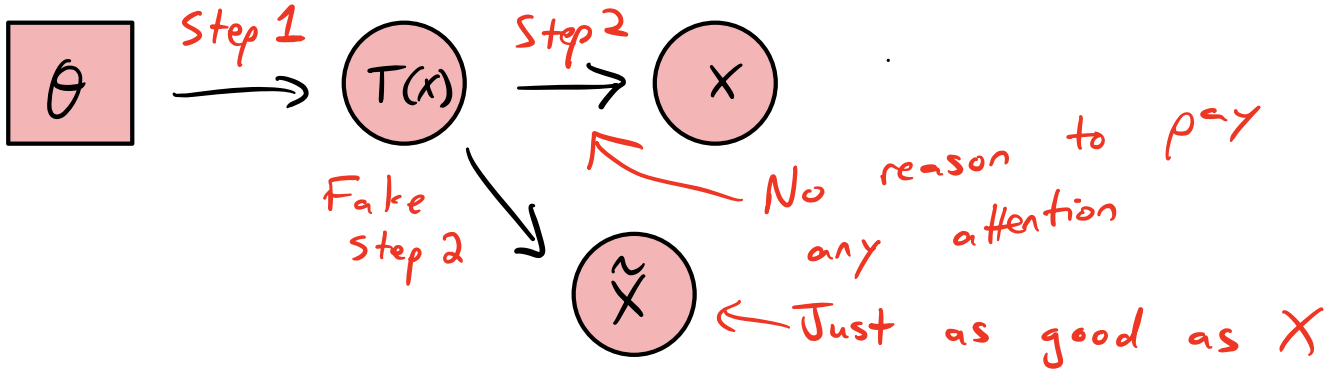
Sufficiency Principle

If $T(X)$ is sufficient for $\mathcal{P}$ then any statistical procedure should depend on X only through $T(X)$

In fact, we could throw away X and generate a new $\tilde{X} \sim P(X|T)$ and it would be just as good as X since $\tilde{X} \sim P_\theta$

(no θ)

In graphical model form:



Examples

Ex. Exponential Families

$$p_\theta(x) = \underbrace{e^{\eta(\theta)' T(x) - B(\theta)}}_{g_\theta(T(x))} \underbrace{h(x)}_{h(x)}$$

Ex. Uniform location family

$$X_1, \dots, X_n \overset{iid}{\sim} U[\theta, \theta+1]$$

$$= 1\{\theta \le x \le \theta+1\}$$

$$p_\theta(x) = \prod_{i=1}^{n} 1\{\theta \le x_i \le \theta+1\}$$

$$= 1\{\theta \le X_{(1)}\}\, 1\{X_{(n)} \le \theta+1\}$$

$\Rightarrow (X_{(1)}, X_{(n)})$ is sufficient.

Order Statistics / Empirical Distribution

Ex. $X_1, \ldots, X_n \overset{iid}{\sim} P_\theta^{(1)}$ for any model

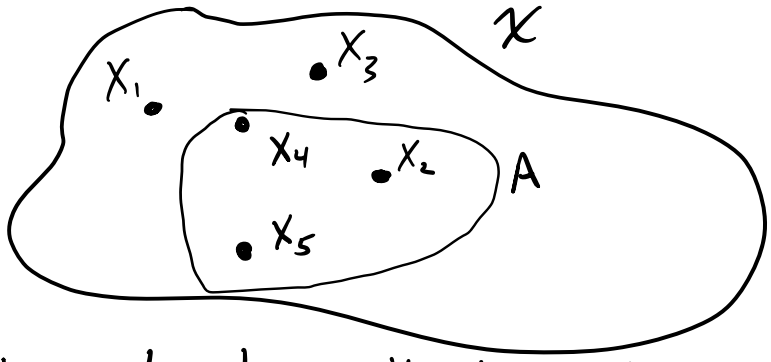
$\mathcal{P}^{(1)} = \{P_\theta^{(1)} : \theta \in \Theta\}$ on $\mathcal{X} \subseteq \mathbb{R}$

P_θ is invariant to perms of $X = (X_1, \ldots, X_n)$

$\Rightarrow$ All permutations of x are equally likely

$\Rightarrow$ **order statistics** $(X_{(i)})_{i=1}^n$ ($X_{(k)} = k^{th}$ smallest are sufficient. [Note $(X_i)_{i=1}^n \rightsquigarrow (X_{(i)})_{i=1}^n$ loses information, specifically the orig. ordering]

For more general $\mathcal{X}$ we can say the **empirical distribution** $\hat{P}_n(\cdot) = \frac{1}{n}\sum_{i=1}^n \delta_{X_i}(\cdot)$ is sufficient, where $\delta_{X_i}(A) = 1\{X_i \in A\}$



$\hat{P}_n(A) = \frac{3}{5}$

[Not important that it's a measure in this context; just keeps track of which values came up how many times]

Minimal Sufficiency

Consider $X_1, \ldots, X_n \overset{iid}{\sim} N(\theta, 1)$

$$p_\theta^{(1)}(x) = \frac{1}{\sqrt{2\pi}} e^{\theta x - \theta^2/2 - x^2/2}$$

exponential family with $T(x) = x$

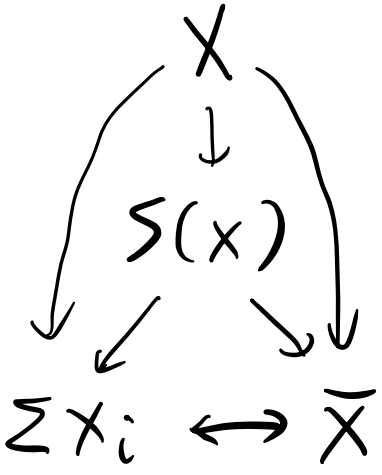
$T(x) = \sum X_i$ sufficient

$\bar{X} = \frac{1}{n}\sum X_i$ also

$S(x) = (X_{(1)}, \ldots, X_{(n)})$ too

$X = (X_1, \ldots, X_n)$ too

Which can be recovered from which others?



these can be compressed further

These are the most compressed. Are they as compressed as possible?

Prop If $T(X)$ is sufficient and $T(x) = f(S(x))$ then $S(X)$ is sufficient

Proof: $p_\theta(x) = g_\theta(T(x))\, h(x)$

$= (g_\theta \circ f)(S(x))\, h(x)$ ☒

Definition: $T(X)$ is *minimal sufficient* if

1) $T(X)$ is sufficient

2) For any other sufficient $S(X)$, $T(X) = f(S(x))$ for some f (a.s. in $\mathcal{P}$)

So, no matter how many more suff. stats we add to our diagram, they will all have arrows pointing to ΣX_i

Likelihood Shape is Minimal

Definition

Assume $\mathcal{P} = \{P_\theta : \theta \in \Theta\}$ has densities $p_\theta(x)$

The <u>likelihood function</u> is the (random) function

$$\text{Lik}(\theta; X) = p_\theta(x)$$

$p_\theta(x)$: function of x with parameter θ

$\text{Lik}(\theta; X)$: function of θ; data X determines which function

The log-likelihood function is its log:

$$\ell(\theta; X) = \log \text{Lik}(\theta; X)$$

The likelihood up to scaling (or ℓ up to vertical shift) is a minimal sufficient statistic

If $T(X)$ is sufficient then

$$\text{Lik}(\theta; x) = \underbrace{g_\theta(T(x))}_{\text{T determines the "shape"}} \underbrace{h(x)}_{\text{scaling}}$$

HW 2: Likelihood ratios $\left(\frac{\text{Lik}(\theta_1; X)}{\text{Lik}(\theta_2; X)} \right)_{\theta_1, \theta_2 \in \Theta}$ minimal suff.

Recognizing Minimal Sufficient Statistics

$T(X)$ is minimal sufficient if

1) $T(X)$ is sufficient (don't forget to check!)

2) $T(x)$ can be recovered from the likelihood shape

Keener Thm 3.11 formalizes condition 2

"$Lik(\cdot; x) \propto Lik(\cdot; y) \Rightarrow T(x) = T(y)$"

equivalently,

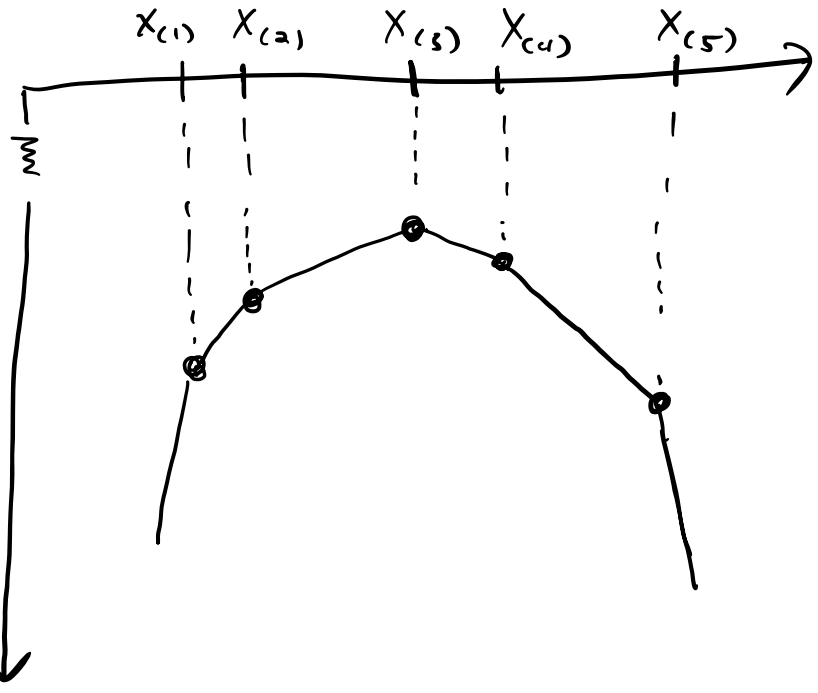
"$\ell(\cdot; x) - \ell(\cdot; y) = const(x, y) \Rightarrow T(x) = T(y)$"

Ex Laplace location family

$$X_1, \ldots, X_n \overset{iid}{\sim} \rho_\theta^{(1)}(x) = \frac{1}{2} e^{-|x-\theta|}$$

$$\ell(\theta; X) = -\sum_{i=1}^{n} |x_i - \theta| - n \log 2$$

Piecewise linear in θ, knots at $x_{(i)}$



On $[x_{(k)}, x_{(k+1)}]$,

slope $= n - 2k$

$\ell(\theta; X) = \ell(\theta; Y) + \text{const} \iff X, Y$ same order statistics

$\Rightarrow$ order stats are minimal suff.

Minimal sufficiency for exp. fam.s

Suppose $p_\eta(x) = e^{\eta' T(x) - A(\eta)} h(x)$

$$\ell(\eta; X) = \underbrace{T(X)'\eta}_{\text{random linear function of } \eta} - \underbrace{A(\eta)}_{\text{deterministic function of } \eta} + \underbrace{\log h(x)}_{\text{(random) const.}}$$

Is $T(X)$ minimal? (always sufficient)

Suppose x and y give same likelihood shape:

$$\ell(\eta; x) - \ell(\eta; y) = \text{const}(x, y)$$

Then $(T(x) - T(y))'\eta = \text{const}(x, y)$ for $\eta \in \Xi$

$\Rightarrow$ $T(x) = T(y)$ or

$$T(x) - T(y) \perp \text{Span}\{\eta_1 - \eta_2 : \eta \in \Xi\}$$

If $\text{Span}\{\dots\} = \mathbb{R}^s$, $T(X)$ is minimal

(That is, if Ξ is not contained in a lower-dim affine space)

Otherwise might not be:

If $s = 2$, $\Xi = \left\{\begin{pmatrix}\theta \\ 0\end{pmatrix} : \theta \in \mathbb{R}\right\}$ then $T_1(X)$ minimal

[Can we conclude $T(X)$ is not minimal?]

Other parameterizations:

$$p_\theta(x) = e^{\eta(\theta)' T(x) - B(\theta)} h(x) \qquad \theta \in \Theta$$

$T(X)$ minimal if $\text{span}\{\eta(\theta_1) - \eta(\theta_2) : \theta_1, \theta_2 \in \Theta\} = \mathbb{R}^s$