## Exponential families

### Outline

- i) Exponential families
- 2) Differential identities
- 3) MGF

### Exponential Families

An s-parameter exponential family is a family 
$$\overline{y} = \{P_m : z \in \Xi_i\}$$
 with densities of the form  $P_2(x) = e^{z^i T(x)} - A(z^i) h(x)$  with base measure  $M$  on sample space  $X$ 

- · y E = E Rs called natural parameter
- T(x) is s-dimensional sufficient statistic factorization theorem:  $g_{\eta}(T(x)) = e^{\gamma'T(x)-A(\eta)}$
- · h(x) ≥0 called base density or carrier density

  can be absorbed into base measure M
  - A(y) called log-partition function or normalizing const.

    A(.) determined by T, h, m:

$$A(\gamma) = \log \left[ \int_{\chi} e^{\gamma' T(\chi)} h(\chi) d\mu(\chi) \right] \leq \infty$$

The natural parameter space is the set of all 7 that give us normalizable  $p_2$   $\Xi_1 = \{ \gamma : A(\gamma) < \infty \}$ 

A(y) is always a convex function => =, convex set

If we absorb h into M, log-densities almost linear log  $\rho_{\mathcal{T}}(x) = \chi^{1} T(x) - A(\chi)$  (with hiden) Can think of T(x) as basis

Very nice structure when we multiply densities

- · Combining evidence from independent obs.
- · Prior \* likelihood in Bayesian calculations

or divide them

- · Calculating conditional probabilities
  - · Likelihood ratios
  - · Relative densities

### Exemples

$$\rho_{\lambda}(x) = \prod_{i=1}^{n} e \times \rho \left\{ (\log \lambda) \times_{i} - \lambda \right\} \frac{1}{\times_{i}!}$$

$$= e \times \rho \left\{ (\log \lambda) \left( \sum_{i} x_{i} \right) - n \right\} \right\} \frac{1}{\times_{i}!}$$

$$\gamma(\lambda) = \log \lambda \qquad T(x) = \sum_{i} x_{i}$$

$$A(\eta) = ne^{\eta} \qquad h(x) = \prod_{i} \frac{1}{\times_{i}!}$$

Generic (n obs) 
$$X_{1,...,} X_{n} \stackrel{iid}{\sim} \rho_{\chi}^{(i)}(x) = e^{\chi^{(i)}(x) - A^{(i)}(x)}$$

$$\rho_{\chi}(x) = \prod_{i=1}^{n} e^{x} \rho \left\{ \chi' T''(x_{i}) - A'(\chi) \right\} h''(x_{i})$$

$$= e^{x} \rho \left\{ \chi' \left( \sum_{i} T''(x_{i}) \right) - \gamma A''(\chi) \right\} \prod_{i=1}^{n} h'(x_{i})$$

$$= A(\chi) \qquad A(\chi)$$

=> Dimension of T(x) doesn't grow with n

# Differential Identities

Write 
$$e^{A(\gamma)} = \int e^{\gamma'T(x)} h(x) d\mu(x)$$
 (\*)

We can derive lots of useful identities
by differentiating (X) on both sides,
pulling derivative inside [not always allowed]

Then  $g(\gamma) = \int f e^{\gamma'T} h d\mu$  has cts partial derivatives of all orders for  $\gamma \in \Xi_{\Gamma}$ . If we can get them by differentiating under the  $\int sign$ .  $\Rightarrow$  on  $\Xi_{\Gamma}^{\circ} A(\gamma)$  has all partial derivatives

Différentiate once:

$$\frac{\partial}{\partial \eta_{i}} e^{A(\eta)} = \frac{\partial}{\partial \eta_{i}} \int e^{\eta'(T(x))} h(x) d\mu(x)$$

$$e^{A(\eta)} \frac{\partial A}{\partial \eta_{i}} (\eta) = \int T_{i}(x) e^{\eta'(T(x)) - A(\eta)} h(x) d\mu(x)$$

$$\Rightarrow \frac{\partial A}{\partial \eta_{i}} (\eta) = \mathbb{E}_{\eta} [T_{i}(x)]$$

$$\nabla A(z) = \mathbb{E}_z \left[ T(x) \right]$$

$$\frac{\partial^2}{\partial \gamma_i \partial \gamma_k} e^{A(\gamma_i)} = \frac{\partial^2}{\partial \gamma_i \partial \gamma_k} \int_{\gamma_i} e^{\gamma_i' T} d\mu$$

$$e^{A(2)}\left(\frac{\partial^{2}A}{\partial \gamma_{i}\partial \gamma_{k}} + \frac{\partial A}{\partial \gamma_{i}} \frac{\partial A}{\partial \gamma_{k}}\right) = \int T_{j}T_{k} e^{\gamma_{j}T_{k}} d\mu$$

$$E[T_{j}] E[T_{k}]$$

$$E[T_{j}T_{k}]$$

$$\frac{\partial^2 A}{\partial \gamma_i \partial \gamma_k} (\gamma) = Cov_{\chi} (T_i, T_k)$$

$$abla^2 A(\eta) = Var_{\eta}(T(x)) \in \mathbb{R}^{5\times 5}$$

$$= \frac{1}{2} \times e^{-x} = \frac{1}{2} \times$$

$$\mathbb{E}_{\gamma}[X] = \frac{d}{d\gamma} e^{\gamma} = e^{\gamma} = \lambda$$

$$V_{ar_2}(x) = \frac{d^2}{dx^2} e^2 = e^2 = \lambda$$

NB: We would get wrong answer by differentiating

# Moment-generating function can get kth order moments of T(X) by 1) Differentiating (X) k times, then 2) Dividing by $e^{A(z)}$ That is because $M_{\chi}^{T}(n) = e^{A(\chi+n)} - A(\chi)$ is the moment-generating function (mgf) of T(x) when X~P2 $M_{\chi}^{\tau(\chi)}(u) = \mathbb{E}_{\chi}\left[e^{u'\tau(\chi)}\right]$ $= \int e^{u'T} e^{\eta'T} - A(\eta) h d\eta$ $= e^{A(\eta+u)-A(\eta)} \begin{cases} (\eta+u)^{T}-A(\eta+u) \\ e^{T} & \text{hd} \end{cases}$ Useful for · finding moments · finding dist. of sums of indep. RVs

Cumulant-generating function

$$K_{\chi}^{T}(n) = \log M_{\chi}^{T}(n) = A(\chi + n) - A(\chi)$$

(A is sometimes colled egf)

#### Other Parameterizations

Sometimes it is more convenient to use a different parameterization:  $\rho_{\theta}(x) = e^{\gamma(\theta)} T(x) - B(\theta) h(x)$ 

$$\rho_{\theta}(x) = e^{\gamma(\theta)'T(x)} - \beta(\theta) h(x)$$

$$\beta(\theta) = A(\gamma(\theta))$$

Many, many examples, sometimes requires massaging to see that they are exp. fam.s:

Let 
$$\theta = (m, \sigma^2)$$
  
 $\rho(x) = \frac{1}{\sqrt{2\pi}\sigma^2} e^{-(m-x)^2/2\sigma^2}$ 

$$= \exp \left\{ \frac{m}{\sigma^2} \times - \frac{1}{2\sigma^2} \times^2 - \frac{m^2}{2\sigma^2} - \frac{1}{2} \log (2\pi \sigma^2) \right\}$$

$$\gamma(\theta) = \binom{m/\sigma^2}{-1/2\sigma^2} \qquad T(x) = \binom{x}{x^2} \qquad h(x) = 1$$

$$B(\theta) = \frac{n^2}{2\sigma^2} + \frac{1}{2}\log(2\pi\sigma^2)$$

Natural parameterization

$$P_{3}(x) = e^{\gamma'(x^{2})} - A(3)$$

More examples

$$\frac{M \text{ ore examples}}{X \sim \text{Binom}(n, \theta)}$$

$$\frac{f_{\theta}(x)}{f_{\theta}(x)} = \frac{G^{x}(1-\theta)^{n-x}\binom{n}{x}}{(1-\theta)^{n-x}\binom{n}{x}} \qquad x = 0, ..., n$$

$$= \left(\frac{G}{1-\theta}\right)^{x}(1-\theta)^{n}\binom{n}{x}$$

$$= exp\left\{\log\left(\frac{G}{1-\theta}\right) \cdot x + n\log\left(1-\theta\right)\right\}\binom{n}{x}$$

Beta 
$$X \sim Beta(\alpha, \beta)$$

$$P_{\alpha,\beta}(x) = x^{\alpha-1} (1-x)^{\beta-1}/B(\alpha,\beta) \leftarrow Beta function$$

$$= \exp\{\alpha \log x + \beta \log(1-x) - \log B(\alpha,\beta)\} \frac{1}{x(1-x)}$$

$$\frac{1}{3} = {\alpha \choose \beta} \quad T(x) = {\log x \choose \log(1-x)} \quad h(x) = \frac{1}{x(1-x)}$$

 $\gamma(0) = \log(\frac{0}{1-0})$  "log odds ratio"

Practically everything else on wikipedia too:
Beta, Gamma, Multinom., Dirichlet, Pareto, Wishart...

### Interpretation: Exponential tilting

Can think of 
$$p_{2}(x) = e^{y'T(x)-A(y)}h(x)$$
 as

an exponential tilt of the carrier  $h(x)$ 

- 1) Start with carrier h(x)
- 2) Multiply by e ?'T(x)
- 3) Re-normalite by e

$$T(x) = (T_1(x), ..., T_s(x))$$
 gives linear space of directions in which we can tilt  $h(x)$ 

- 1) Only span (T, ..., Ts) matters
- 2) Could absorb h into  $\mu$  ( $d\nu(x) = h d\mu(x)$ )
  (wlog  $h(x) \equiv 1$  if we want)
- 3) Can add constant to T(x)

: many others

Suppose 
$$X \sim \rho_n(x) = e^{\gamma'T(x) - A(\eta)}$$
 with  $n$  (wlog  $h \equiv 1$ )

Then  $T(X) \sim q_n(t) = e^{\gamma't - A(\eta)}$  with  $v$ ,

where v is the measure u "pushed forward"

through  $T: X \to \mathbb{R}^S$   $v(B) \triangleq u(\{x: T(x) \in B\})$ 

$$P_{2}(T(x) \in B) = \int 1_{B}(\tau(x))e^{2^{t}T(x)-A(\tau_{1})} d_{M}(x)$$

$$= \int 1_{B}(t)e^{2^{t}t-A(\tau_{1})} d_{v}(t)$$

Simplest in discrete case: (drop h=1 assumption)

$$P_{\eta}(T(x) = t) = \sum_{x:T(x)=t} e^{\eta(T(x)-A(\eta)} h(x) \mu(\{x\})$$

$$= e^{\eta'(t-A(\eta))} \sum_{x:T(x)=t} h(x) \mu(\{x\})$$

$$= e^{\eta'(T(x)-A(\eta))} h(x) \mu(\{x\})$$

#### Canonical Form

The structure is most evident when:

• 
$$h(x) = 1$$
 (wlog: absorb h into u)

# Minimal form

Form of  $\rho_{\alpha}(x) = e^{2^{i}T(x)} - A(x)h(x)$  minimal if  $\gamma \in \Xi$  and  $\tau(x)$  satisfy no linear constraints: no  $\alpha \neq 0$ ,  $b \in \mathbb{R}$  s.t.  $\gamma' = b$  for all  $\gamma \in \Xi$  or  $\tau(x)' = b$   $\gamma = b$ 

Otherwise we can represent P as an r-dim. ex.fam. for some r < s

If  $P_{\pi}$  minimal, then T(x) is minimal suff.  $|P_{\pi}| = T(x) = T(x) = T(x) = T(x) = T(x)$ 

Need to show  $l(\cdot;x) = l(\cdot;y) + c_{xy} \Rightarrow T(x) = T(y)$ (a holds by safe)

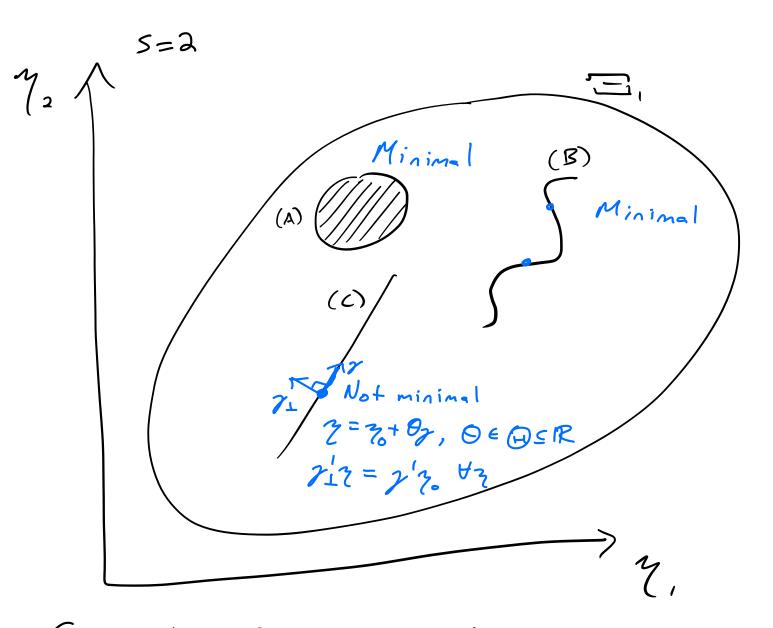
 $l(\gamma;x)-l(\gamma;y)=\gamma'(\tau(x)-\tau(y))$ 

Can find y, S = = s.t. y a + 5 a unless a=0

 $\Rightarrow T(x) = T(y)$ 

Important: Converse is not true: reducing dimension of T(x) may or may not be a data reduction (See multinomial problem in hw)

$$\rho_{\eta}(x) = e^{\frac{\eta}{2}T(x) - A(\eta)}h(x)$$



$$e^{\gamma'T(x)-A(\gamma)} = e^{(\gamma'T(x))-A(\gamma_0+\theta_{\gamma'})} h(x)$$