

Joint, Conditional, and Marginal PDFs

Lec. 23

July 30, 2020

Joint PDFs

Let X and Y be two continuous random variables. Then the joint density function $f_{X,Y} : \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}$ satisfies:

$$\mathbb{P}(a \leq X \leq b, c \leq Y \leq d) = \int_a^b \int_c^d f_{X,Y}(x,y) dy dx$$

and

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy dx = 1$$

and

$$f_{X,Y}(x,y) \geq 0 \quad \forall x, y \in \mathbb{R}$$

Joint CDFs

(applies to continuous &
discrete
in general)

Let X and Y be two random variables. Then the joint cumulative distribution function $F_{X,Y} : \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}$ satisfies:

$$F_{X,Y}(x, y) = \mathbb{P}(X \leq x, Y \leq y)$$

$$f_{X,Y}(x, y) = \underbrace{\frac{\partial^2}{\partial x \partial y}}_{\text{In the continuous case, } \rightarrow} F_{X,Y}(x, y)$$

$$\frac{\partial^2}{\partial x \partial y} = \frac{\partial}{\partial x} \cdot \frac{\partial}{\partial y}$$

$$\underbrace{\frac{\partial^2}{\partial y \partial x}}_{\text{assuming the CDF is continuous.}} \text{ is also fine.}$$

Conditional PDFs

Just like $P(A|B) = \frac{P(A \cap B)}{P(B)}$

Let X and Y be two continuous random variables with joint density function $f_{X,Y}$.

For any y with $f_y(y) > 0$, the conditional distribution of X given $Y = y$ is defined as:

$$f_{x|y} := \frac{f_{x,y}(x, y)}{f_y(y)}$$

When Y is continuous, even though $P(Y = y) = 0$, if $f_y(y) > 0$,

$$P(a \leq X \leq b | Y = y) = \int_a^b f_{x|y}(x|y) dx$$

Independence

Let X and Y be two continuous random variables. X and Y are independent if:

$$f_{x,y}(x, y) = f_x(x)f_y(y)$$

for all x, y .

Since $f_{x,y}(x, y) = f_{x|y}(x|y)f_y(y)$, this implies $f_{x|y}(x|y) = f_x(x)$.

Just like
 $P(A \cap B) = P(A|B) \cdot P(B)$
in discrete setting

in order for
 X and Y to
be independent

Marginalization

To recover the individual pdfs from the joint pdf:

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx$$

for $f_X(x)$
to be a valid pdf,
 $\int_{-\infty}^{\infty} f_X(x) dx = 1$

$f_X(x)$ is a marginal pdf
of $f_{X,Y}(x,y)$

Example: $X, Y \stackrel{i.i.d}{\sim} Unif(0, 2)$

independent and identically distributed.

What is $f_{X,Y}(x,y)$ for two uniform rvs on $[0, 2]$?

Uniform on 2×2 square \rightarrow Constant c

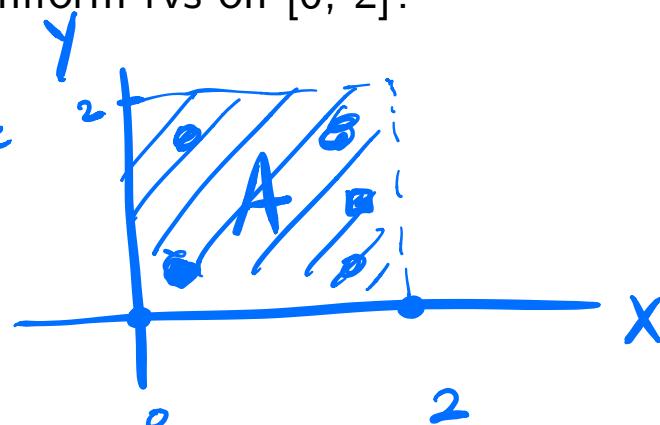
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy dx = 1$$

$$\int_0^2 \int_0^2 c dy dx = 1$$

$$c \int_0^2 \int_0^2 1 \cdot dy dx = 1$$

$$c \cdot 4 = 1$$

$$c = \frac{1}{4}$$



$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{4} & \text{if } (x,y) \in A \\ 0 & \text{o.w.} \end{cases}$$

Example: $X, Y \stackrel{i.i.d}{\sim} Unif(0, 2)$

What is $f_{X,Y}(x, y)$ for two uniform rvs on $[0, 2]$?

We know that it must be a nonzero constant c on the two-by-two square, since all x, y pairs are equally likely. Again, we can use the constraint:

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy dx = 1$$

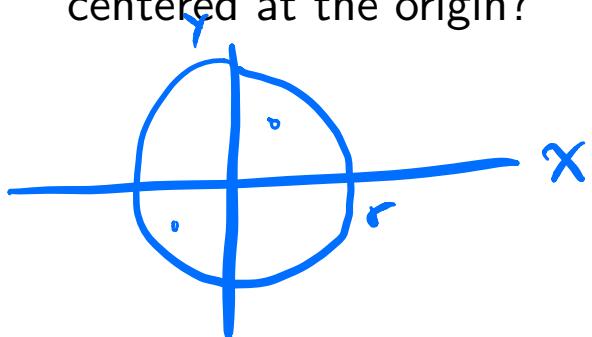
$$\int_0^2 \int_0^2 c dy dx = 1$$

$$4 \cdot c = 1$$

$$\rightarrow c = \frac{1}{4}$$

Uniform Density Over a Disk: Joint

What is $f_{X,Y}(x,y)$ for a uniform density over a disk of radius r centered at the origin?



$$f_{X,Y}(x,y) = \begin{cases} C & \text{if } x^2 + y^2 \leq r^2 \\ 0 & \text{o.w.} \end{cases}$$

$$\begin{aligned} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx dy &= \iint_{\substack{x^2+y^2 \leq r^2}} C dx dy \\ &= C \iint_{\substack{x^2+y^2 \leq r^2}} 1 dx dy \\ &= C \cdot (\text{area of disk}) \\ &= C \cdot \pi r^2 \end{aligned}$$

We know pdf must integrate to 1, so $C\pi r^2 = 1 \Rightarrow C = \frac{1}{\pi r^2}$

Uniform Density Over a Disk: Joint

What is $f_{X,Y}(x,y)$ for a uniform density over a disk of radius r centered at the origin?

$$f(x,y) = \begin{cases} c, & \text{if } x^2 + y^2 \leq r^2 \\ 0, & \text{otherwise} \end{cases}$$

$$\begin{aligned} \iint_{x^2+y^2 \leq r^2} c \cdot dx \cdot dy &= c \cdot \iint_{x^2+y^2 \leq r^2} 1 \cdot dx \cdot dy \\ &= c \cdot [\text{area of disk}] \\ &= c \cdot \pi \cdot r^2 \end{aligned}$$

By definition, the joint must integrate to 1, so

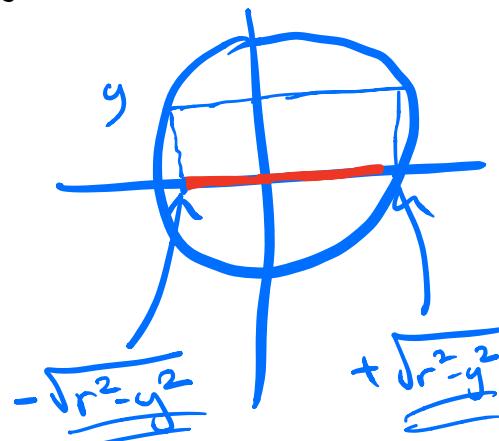
$$1 = c \cdot \pi \cdot r^2 \Rightarrow c = \frac{1}{\pi \cdot r^2}$$

Uniform Density Over a Disk: Marginals

$$x^2 + y^2 \leq r^2$$

What is $f_y(y)$, and $f_x(x)$, now that we know the joint over the disk?

$$\begin{aligned} f_y(y) &= \int_{-\infty}^{\infty} f_{x,y}(x, y) dx \\ &= \int_{-\sqrt{r^2 - y^2}}^{\sqrt{r^2 - y^2}} \frac{1}{\pi r^2} dx \\ &= \frac{1}{\pi r^2} \int_{-\sqrt{r^2 - y^2}}^{\sqrt{r^2 - y^2}} 1 dx \\ &= \frac{2 \cdot \sqrt{r^2 - y^2}}{\pi r^2} \quad (\text{for } -r \leq y \leq r) \\ &\quad 0 \quad \text{o.w.} \end{aligned}$$



$$f_x(x) = \frac{2 \cdot \sqrt{r^2 - x^2}}{\pi r^2} \quad (\text{for } -r \leq x \leq r) \\ 0 \quad \text{o.w.}$$

Uniform Density Over a Disk: Marginals

What is $f_y(y)$, and $f_x(x)$, now that we know the joint over the disk?

$$\begin{aligned}f_y(y) &= \int_{-\sqrt{r^2-y^2}}^{\sqrt{r^2-y^2}} \frac{1}{\pi \cdot r^2} \cdot dx \\&= \frac{1}{\pi \cdot r^2} \int_{-\sqrt{r^2-y^2}}^{\sqrt{r^2-y^2}} \cdot dx \\&= \frac{2\sqrt{r^2 - y^2}}{\pi \cdot r^2}\end{aligned}$$

for $-r \leq y \leq r$

By symmetry,

$$f_x(x) = \frac{2\sqrt{r^2 - x^2}}{\pi \cdot r^2}$$

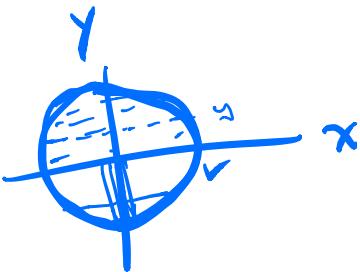
Uniform Density Over a Disk: Conditional PDFs

What is $f_{x|y}(x|y)$?

By def, $f_{x|y}(x|y) = \frac{f_{x,y}(x,y)}{f(y)}$

$$= \frac{1}{\pi r^2} \frac{1}{2\sqrt{r^2-y^2}}$$

$$= \frac{1}{2\sqrt{r^2-y^2}} \pi r^2 \quad \left(\text{for } x^2+y^2 \leq r^2, \text{ also } |y| < r \right)$$



Note : $f_{X|Y}(x|y) \neq \frac{2\sqrt{r^2 - x^2}}{\pi r^2} = f_X(x)$

$\Rightarrow X$ and Y are not independent

What about $\text{Cov}(X, Y)$?

$$\mathbb{E}[x] = 0$$

$$E[Y] = 0$$

$$\mathbb{E}[X | Y=y] = 0$$

by symmetry

$$\mathbb{E}[X|Y=y] = y \cdot \mathbb{E}[X|Y=y] = y \cdot 0 = 0$$

$$\Rightarrow E[x \cdot y] = 0$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y]$$

$$= 0 - 0 = 0$$

X, Y are uncorrelated, but are dependent

Note:

Independent $\Rightarrow E[xy] = E[x]E[y]$
(x and y are not true.)

(the converse is not true as shown in this example.) independent \Rightarrow uncorrelated (but not the other way around)

Uniform Density Over a Disk: Conditional PDFs

What is $f_{x|y}(x|y)$? If $x^2 + y^2 \leq r^2$ and $|y| < r$,

$$f_{x|y}(x|y) = \frac{f_{x,y}(x, y)}{f_y(y)} = \frac{\frac{1}{\pi \cdot r^2}}{\frac{2\sqrt{r^2 - y^2}}{\pi \cdot r^2}} = \frac{1}{2\sqrt{r^2 - y^2}}$$

\Downarrow

$$f_{x|y}(x) \quad f_{x|y}(x, y)$$

Since this differs from $f_x(x)$, we know X and Y are dependent. Furthermore,

$$\mathbb{E}[X] = \mathbb{E}[Y] = 0 \text{ by symmetry}$$

$$\mathbb{E}[X|Y = y] = 0 \text{ by symmetry}$$

$$\mathbb{E}[XY|Y = y] = y \cdot \mathbb{E}[X|Y = y] = y \cdot 0 = 0 \Rightarrow \mathbb{E}[XY] = 0$$

$$\text{Cov}(X, Y) = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y] = 0$$

X, Y are uncorrelated but dependent!!

2D LOTUS

2 min break

Law of the unconscious statistician.

Let X, Y be two random variables with joint PDF $f_{X,Y}(x, y)$, and let $g(x, y)$ be a real-valued function of x, y . Then

$$\mathbb{E}[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X,Y}(x, y) dx dy$$

↑ ↓
use the known
joint pdf of X and Y

You don't need to find the joint
pdf of $g(X, Y)$ in order
to calculate $\mathbb{E}[g(X, Y)]$

Expected Distance Between Two Points

Let $X, Y \stackrel{i.i.d}{\sim} \text{Unif}(0, 1)$. What is $\mathbb{E}[|X - Y|]$?

$$\begin{aligned}\mathbb{E}[|X - Y|] &= \int_0^1 \int_0^1 |x - y| \cdot f_{X,Y}(x,y) dx dy \\ &= \int_0^1 \int_0^1 |x - y| \cdot 1 dx dy \\ &= \iint (x-y) \cdot dx dy + \iint (y-x) dx dy \quad \text{Y>x}\end{aligned}$$

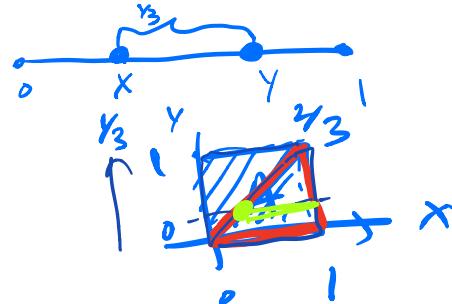
$$= 2 \cdot \iint (x-y) dx dy \quad \text{x>y}$$

$$= 2 \cdot \int_0^1 \int_0^y (x-y) dx dy \quad \text{x>y}$$

$$\rightarrow = 2 \cdot \int_0^1 \left(\frac{x^2}{2} - yx \right) \Big|_{x=y} dy$$

$$\rightarrow = 2 \cdot \int_0^1 \left(\frac{y^2}{2} - y + \frac{1}{2} \right) dy$$

$$= \frac{1}{3}$$



$$f_{XY}(x,y) = \begin{cases} 1 & \text{if } x, y \in \mathbb{R} \\ 0 & \text{otherwise} \end{cases}$$

Observe: $|X - Y| = \max(X, Y) - \min(X, Y)$

$$M = \max(X, Y)$$

$$L = \min(X, Y)$$

$$\begin{aligned}\mathbb{E}[|X - Y|] &= \mathbb{E}[M - L] = \frac{1}{3} \\ &= \mathbb{E}[M] - \mathbb{E}[L] = \frac{1}{3}\end{aligned}$$

Also,

$$\mathbb{E}[M + L] = \mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

$$\mathbb{E}[M] + \mathbb{E}[L] = \frac{1}{2} + \frac{1}{2} = 1$$

Solve the two equations.

$$\mathbb{E}[M] = \frac{2}{3} \quad \mathbb{E}[L] = \frac{1}{3}$$

Expected Distance Between Two Points

Let $X, Y \stackrel{i.i.d}{\sim} \text{Unif}(0, 1)$. What is $\mathbb{E}[|X - Y|]$?

$$\mathbb{E}[|X - Y|] = \int_0^1 \int_0^1 |x - y| f_{X,Y}(x, y) dx dy \quad (1)$$

$$= \int_0^1 \int_0^1 |x - y| \cdot 1 dx dy \quad (2)$$

$$= \int \int_{X > Y} (x - y) dx dy + \int \int_{Y > X} (y - x) dx dy \quad (3)$$

$$= 2 \cdot \int_0^1 \int_y^1 (x - y) dx dy \quad (4)$$

$$= 2 \cdot \int_0^1 \left(\frac{x^2}{2} - yx \right) \Big|_{x=y}^1 dy \quad (5)$$

$$= \frac{1}{3} \quad (6)$$

$$(7)$$

Total Probability Theorem

$$P(A \cap B) = P(A|B) \cdot P(B)$$

Discrete Partition.

Continuous Partition

↓
we know

Partition

probability

	Dis.	Cont.
Dis.	$P(A) = \sum_1^n P(B_i) \cdot P(A B_i)$	$P(A) = \int_{-\infty}^{\infty} f_x(x) \cdot P(A X=x) dx$
Cont.	$f_x(x) = \sum_1^n P(B_i) \cdot f_{x B_i}(x)$	$f_x(x) = \int_{-\infty}^{\infty} f_y(y) \cdot f_{x y}(x y) dy$

we know
(analogous)

$f_x(x) \rightarrow$ continuous

$P(x=x) \rightarrow$ disc.

Total Probability Theorem Examples

Discrete/Continuous:

$$P(\underline{Y} \geq \underline{X}) = \int_{-\infty}^{\infty} f_x(x) \cdot P(\underline{Y} \geq \underline{X} | X = x) dx$$

Top right

$X \sim \text{cont.}$
 $Y \sim \text{discrete.}$

Continuous/Discrete: Bottom left.

Flip a fair coin. If its heads, then $X \sim \text{Exp}(\lambda_1)$, otherwise $X \sim \text{Exp}(\lambda_2)$. Then, for $x > 0$,

$$f_x(x) = \frac{1}{2} \cdot \lambda_1 e^{-\lambda_1 x} + \frac{1}{2} \cdot \lambda_2 e^{-\lambda_2 x}$$

Bayes' Rule

	Dis.	Cont.
Dis.	$P(A_i B) = \frac{P(A_i) \cdot P(B A_i)}{\sum_{j=1}^n P(A_j) \cdot P(B A_j)}$	$P(A_i X = x) = \frac{P(A_i) \cdot f_{X A_i}(x)}{\sum_j^n P(A_j) \cdot f_{X A_j}(x)}$
Cont.	$f_{X A}(x) = \frac{f_X(x) \cdot P(A X=x)}{\int_{-\infty}^{\infty} f_X(t) P(A X=t) dt}$	$f_{X Y}(x y) = \frac{f_X(x) \cdot f_{Y X}(y x)}{\int_{-\infty}^{\infty} f_X(t) f_{Y X}(Y t) dt}$

Discrete/Continuous: *top right*

$$\begin{aligned}
 P(A|X = x) &= P(A|X \in [x, x + \delta]) \\
 &= \frac{P(A) \cdot P(X \in [x, x + \delta]|A)}{P(A) \cdot P(X \in [x, x + \delta]|A) + P(A^c) \cdot P(X \in [x, x + \delta]|A^c)} \\
 &= \frac{P(A) \cdot f_{X|A}(x) \cdot \delta}{P(A) \cdot f_{X|A}(x) \cdot \delta + P(A^c) \cdot f_{X|A^c}(x) \cdot \delta} \\
 &= \frac{P(A) \cdot f_{X|A}(x)}{P(A) \cdot f_{X|A}(x) + P(A^c) \cdot f_{X|A^c}(x)}
 \end{aligned}$$