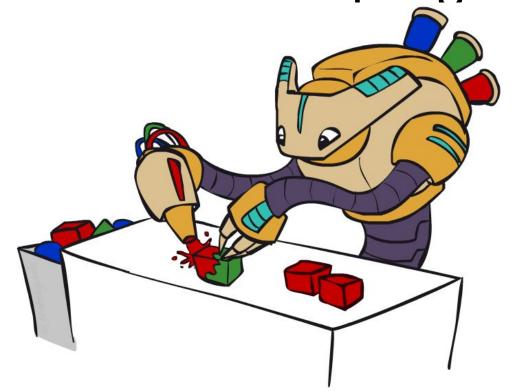
# CS 188: Artificial Intelligence

Supplement on Gibbs Sampling Convergence



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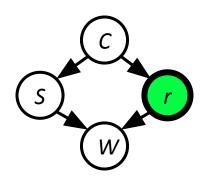
### Gibbs sampling

#### A particular kind of MCMC

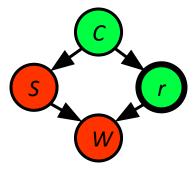
- States are complete assignments to all variables
- Evidence variables E remain fixed, other variables X change
- To generate the next state,
  - pick a variable  $X_i$  with probability  $\rho(i)$
  - sample a value for it conditioned on all the other variables:
    - $-X_{i}' \sim P(X_{i} \mid X_{1},...,X_{i-1},X_{i+1},...,X_{n}, \mathbf{e}) = P(X_{i} \mid \mathbf{x}_{-i}, \mathbf{e})$
    - In a Bayes net,  $P(X_i | \mathbf{x}_{-i}, \mathbf{e}) = P(X_i | markov\_blanket(X_i))$
- Theorem: Gibbs sampling is consistent\*
- Provided all Gibbs distributions are bounded away from 0 and 1 and variable selection is fair

# Gibbs Sampling Example: P(S | r)

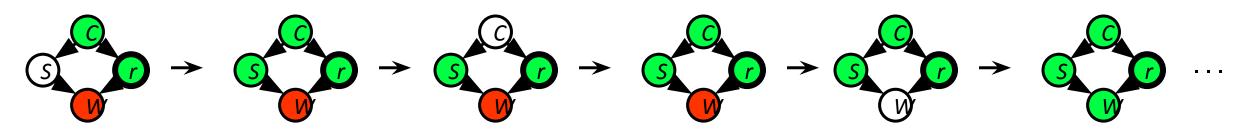
- Step 1: Fix evidence
  - *R* = true



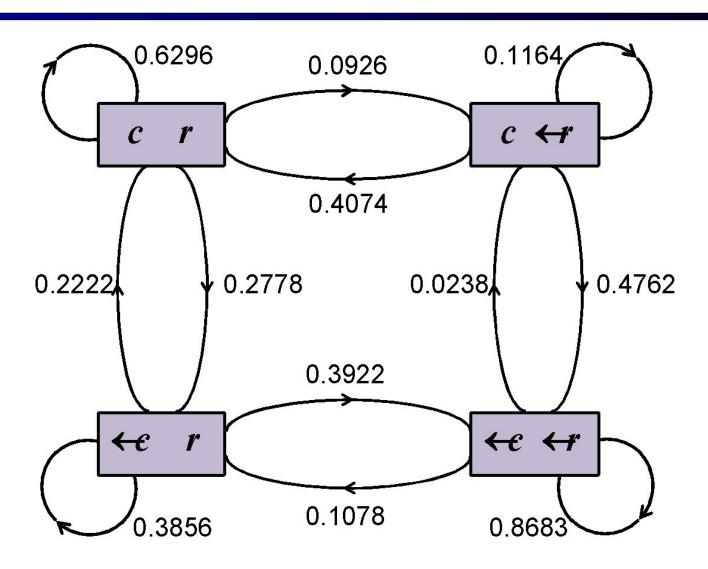
- Step 2: Initialize other variables
  - Randomly



- Step 3: Repeat
  - Choose a non-evidence variable X
  - Resample X from P(X | markov\_blanket(X))



# Markov chain given s, w



#### Why does it work? Overview

- Suppose we run it for a long time and predict the probability of reaching any given state at time t:  $\pi_t(x_1,...,x_n)$  or  $\pi_t(\mathbf{x})$
- Each Gibbs sampling step (pick a variable, resample its value) applied to a state x has a probability  $k(x \rightarrow x')$  of reaching a next state x'
- So  $\pi_{t+1}(\mathbf{x'}) = \sum_{\mathbf{x}} k(\mathbf{x} \to \mathbf{x'}) \; \pi_t(\mathbf{x})$  (standard Markov chain prediction step)
- When the process is in equilibrium  $\pi_{t+1} = \pi_t = \pi$  so  $\pi(\mathbf{x'}) = \sum_{\mathbf{x}} k(\mathbf{x} \to \mathbf{x'}) \pi(\mathbf{x})$
- This has a unique\* solution  $\pi = P(\mathbf{x} | \mathbf{e})$

#### Detailed balance

- More specifically,  $\pi(x) = P(x | e)$  satisfies **detailed balance**:
  - For every pair of states  $\mathbf{x}, \mathbf{x'}, \quad \pi(\mathbf{x'}) \ k(\mathbf{x'} \to \mathbf{x}) = \pi(\mathbf{x}) \ k(\mathbf{x} \to \mathbf{x'})$
- Detailed balance implies that  $\pi$  is the stationary distribution for k:

  - $\pi(\mathbf{x'}) \sum_{\mathbf{x}} k(\mathbf{x'} \to \mathbf{x}) = \sum_{\mathbf{x}} \pi(\mathbf{x}) k(\mathbf{x} \to \mathbf{x'})$
  - $\pi(\mathbf{x'}) = \sum_{\mathbf{x}} \pi(\mathbf{x}) \ k(\mathbf{x} \to \mathbf{x'})$

### What is the transition probability $k(x \rightarrow x')$ ?

- Case 1: x, x' differ in two or more variables
  - Then  $k(\underline{\mathbf{x'}} \mid \underline{\mathbf{x}}) = 0$ ,  $\Rightarrow$  detailed balance

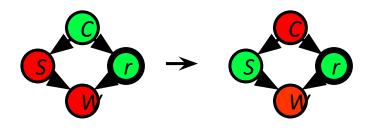


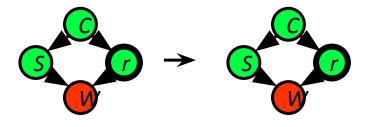
- Case 2: x, x' identical
  - Then detailed balance becomes

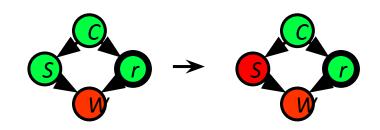
$$\pi \pi(\mathbf{x}) \ k(\mathbf{x} \longrightarrow \mathbf{x}) = \pi(\mathbf{x}) \ k(\mathbf{x} \longrightarrow \mathbf{x})$$



- Case 3:  $\underline{\mathbf{x}}$ ,  $\underline{\mathbf{x}}$  differ in variable  $X_i$ 
  - Then  $k(\mathbf{x} \to \mathbf{x'}) = P(X_i \text{ chosen}) P(X_i \text{ samples } x'_i)$ =  $\rho(i) P(x'_i \mid \mathbf{x}_{i'}, \mathbf{e})$







#### Detailed balance for Case 3

$$\frac{k(\mathbf{x} \to \mathbf{x}')}{k(\mathbf{x}' \to \mathbf{x})} = \frac{\rho(i)P(x'_{i} \mid \mathbf{x}_{-i}, \mathbf{e})}{\rho(i)P(x_{i} \mid \mathbf{x}_{-i}, \mathbf{e})}$$

$$= \frac{P(x'_{i}, \mathbf{x}_{-i} \mid \mathbf{e}) / P(\mathbf{x}_{-i} \mid \mathbf{e})}{P(x_{i}, \mathbf{x}_{-i} \mid \mathbf{e}) / P(\mathbf{x}_{-i} \mid \mathbf{e})}$$

$$= \frac{P(\mathbf{x}' \mid \mathbf{e})}{P(\mathbf{x} \mid \mathbf{e})}$$

- Hence detailed balanced is satisfied by  $\pi(x) = P(x \mid e)$
- So in the limit, a sample generated by Gibbs is drawn from the true posterior  $P(\mathbf{x} | \mathbf{e})$