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High Dimensional Statistics: Exercise Sheet 4

Exercise 3. Let X_1, \dots, X_n be a sequence of i.i.d. random variables with common distribution $\mathcal{N}_k(\mu, \Sigma)$. Assume that the mean vector $\mu \in \mathbb{R}^k$ is known. Show that the maximum likelihood estimator of $\Sigma \in \mathbb{R}^{k \times k}$ is given by:

$$\hat{\Sigma}_{ML} = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)(X_i - \mu)^T$$

Solution. For a given density function f of random variables X_1, \dots, X_n , the likelihood function w.r.t parameter under study, i.e., Σ is defined as follows:

$$l_{\Sigma}(X_1, \dots, X_n) = \log f_{\mu, \Sigma}^n(X_1, \dots, X_n)$$

As $X_1, \dots, X_n \sim \mathcal{N}_k(\mu, \Sigma)$, the density function is defined as:

$$\begin{aligned} f_{\mu, \Sigma}^n(X_1, \dots, X_n) &= \prod_{i=1}^n \left\{ (\det(\Sigma))^{-1/2} (2\pi)^{-k/2} \exp\left(-\frac{1}{2}(X_i - \mu)^T \Sigma^{-1} (X_i - \mu)\right) \right\} \\ &= (\det \Sigma)^{\frac{n}{2}} (2\pi)^{\frac{kn}{2}} \exp\left(-\frac{1}{2} \sum_{i=1}^n (X_i - \mu)^T \Sigma^{-1} (X_i - \mu)\right) \end{aligned}$$

Thus, substituting the density function in the log-likelihood function yields:

$$\log f_{\mu, \Sigma}^n(X_1, \dots, X_n) = -\frac{kn}{2} \log(2\pi) - \frac{n}{2} \log(\det(\Sigma)) - \frac{1}{2} \sum_{i=1}^n (X_i - \mu)^T \Sigma^{-1} (X_i - \mu)$$

The maximum likelihood estimator (MLE) $\hat{\Sigma}$ of the parameter Σ is the following:

$$\hat{\Sigma}_{ML} = \arg \max_{\Sigma} l_{\Sigma}(X_1, \dots, X_n) = \arg \max_{\Sigma} \log f_{\mu, \Sigma}^n(X_1, \dots, X_n)$$

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$$= \arg \max_{\Sigma} \left[-\frac{kn}{2} \log(2\pi) - \frac{n}{2} \log(\det(\Sigma)) - \frac{1}{2} \sum_{i=1}^n (X_i - \mu)^T \Sigma^{-1} (X_i - \mu) \right]$$

We aim at finding the MLE by finding the critical points of $\log f_{\mu, \Sigma}^n$. For this purpose, we define a new variable $\Sigma' = \Sigma^{-1}$ and instead of computing the derivative of the function w.r.t Σ , we do it w.r.t Σ' . We will also reformulate $\log f_{\mu, \Sigma}^n$ based on $\Sigma' = \Sigma^{-1}$. This changing of variable yields the same critical point due to the following observation:

$$\max_{\Sigma} l(\Sigma) = \max_{\Sigma'} l(\Sigma')$$

In order to reformulate, we define $A(\Sigma)$ and $B(\Sigma)$ as the following:

$$\begin{aligned} A(\Sigma) &= -\frac{n}{2} \log(\det(\Sigma)) \\ &= \frac{n}{2} \log(\det(\Sigma))^{-1} \\ &= \frac{n}{2} \log(\det(\Sigma^{-1})) \end{aligned}$$

$$B(\Sigma) = -\frac{1}{2} \sum_{i=1}^n (X_i - \mu)^\top \Sigma^{-1} (X_i - \mu)$$

We can rewrite $A(\Sigma)$ and $B(\Sigma)$ based on $\Sigma' = \Sigma^{-1}$:

$$A(\Sigma') = \frac{n}{2} \log(\det(\Sigma'))$$

$$B(\Sigma') = -\frac{1}{2} \sum_{i=1}^n (X_i - \mu)^\top \Sigma' (X_i - \mu)$$

We now reformulate $\log f_{\mu, \Sigma}^n$ as following:

$$\begin{aligned} \log f_{\mu, \Sigma}^n(X_1, \dots, X_n) &= -\frac{kn}{2} \log(2\pi) + A(\Sigma) + B(\Sigma) \\ \log f_{\mu, \Sigma'}^n(X_1, \dots, X_n) &= -\frac{kn}{2} \log(2\pi) + A(\Sigma') + B(\Sigma') \end{aligned}$$

We compute the derivative of log-likelihood function w.r.t Σ' as following:

$$\begin{aligned} \frac{d}{d\Sigma'} \log f_{\mu, \Sigma'}^n(X_1, \dots, X_n) &= \frac{d}{d\Sigma'} (A(\Sigma') + B(\Sigma')) \\ &= \frac{d}{d\Sigma'} A(\Sigma') + \frac{d}{d\Sigma'} B(\Sigma') \end{aligned}$$

In order to compute the first term, $\frac{d}{d\Sigma'} A(\Sigma')$, we proceed as following:

$$\begin{aligned} \frac{\partial}{\partial \Sigma'_{ij}} A(\Sigma') &= \frac{\partial}{\partial \Sigma'_{ij}} \frac{n}{2} \log(\det(\Sigma')) \\ &= \frac{n}{2} \frac{1}{\det(\Sigma')} \frac{\partial \det(\Sigma')}{\partial \Sigma'_{ij}} \\ &= \frac{n}{2 \det(\Sigma')} \text{adj}(\Sigma')_{ij} && \text{(Jacobi's formula)} \\ &= \frac{n}{2 \det(\Sigma')} \det(\Sigma') (\Sigma'^{-1})_{ji} && (\text{adj}(A) = \det(A) A^{-1}) \\ &= \frac{n}{2} (\Sigma'^{-1})_{ji} \\ &= \frac{n}{2} (\Sigma'^{-1})_{ij} && (\Sigma \text{ and therefore } \Sigma' \text{ is symmetric}) \end{aligned}$$

The third equality is a result of Jacobi's formula, which expresses the derivative of the determinant of a matrix A in terms of the adjugate of A and the derivative of A as the following:

$$\frac{d}{dt} \det(A(t)) = \text{tr}(\text{adj}(A(t)) \frac{dA(t)}{dt})$$

We conclude that

$$\frac{d}{d\Sigma'} A(\Sigma') = \frac{n}{2} \Sigma'^{-1} = \frac{n}{2} \Sigma$$

To compute the second term, we have:

$$\frac{d}{d\Sigma'} B(\Sigma') = \frac{d}{d\Sigma'} \frac{1}{2} \sum_{i=1}^n (X_i - \mu)^\top \Sigma' (X_i - \mu)$$

We now make this observation that each term of the sum, $(X_i - \mu)^\top \Sigma' (X_i - \mu)$, is a scalar, hence the trace of this quantity equates itself, i.e.,

$$(X_i - \mu)^\top \Sigma' (X_i - \mu) = \text{tr}((X_i - \mu)^\top \Sigma' (X_i - \mu))$$

Consequently,

$$\sum_{i=1}^n (X_i - \mu)^\top \Sigma' (X_i - \mu) = \sum_{i=1}^n \text{tr}((X_i - \mu)^\top \Sigma' (X_i - \mu))$$

Going back to computing derivative of $B(\Sigma')$:

Consequently,

$$\begin{aligned} \frac{d}{d\Sigma'} B(\Sigma') &= -\frac{d}{d\Sigma'} \frac{1}{2} \sum_{i=1}^n (X_i - \mu)^\top \Sigma' (X_i - \mu) \\ &= -\frac{1}{2} \sum_{i=1}^n \frac{d}{d\Sigma'} (X_i - \mu)^\top \Sigma' (X_i - \mu) \\ &= -\frac{1}{2} \sum_{i=1}^n \frac{d}{d\Sigma'} \text{tr}((X_i - \mu)^\top \Sigma' (X_i - \mu)) \\ &= -\frac{1}{2} \sum_{i=1}^n \frac{d}{d\Sigma'} \text{tr}(\underbrace{(X_i - \mu)^\top}_P \underbrace{\Sigma' (X_i - \mu)}_Q) \\ &= -\frac{1}{2} \sum_{i=1}^n \frac{d}{d\Sigma'} \text{tr}(P \Sigma' Q) \\ &= -\frac{1}{2} \sum_{i=1}^n Q P \quad \quad \quad d\text{tr}(AXB)/dX = BA \\ &= -\frac{1}{2} \sum_{i=1}^n (X_i - \mu)^\top (X_i - \mu) \end{aligned}$$

We feed the obtained derivatives of the two terms into the derivative of the log-likelihood function as following:

$$\begin{aligned} \frac{d}{d\Sigma'} \log f_{\mu, \Sigma'}^n(X_1, \dots, X_n) &= \frac{d}{d\Sigma'} A(\Sigma') + \frac{d}{d\Sigma'} B(\Sigma') \\ &= \frac{n}{2} \Sigma - \frac{1}{2} \sum_{i=1}^n (X_i - \mu)^\top (X_i - \mu) \end{aligned}$$

In order to find the critical points of the log-likelihood function, we set its derivative to zero:

$$\frac{n}{2} \Sigma - \frac{1}{2} \sum_{i=1}^n (X_i - \mu)^\top (X_i - \mu) = 0 \implies \Sigma = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^\top (X_i - \mu)$$

As the log-likelihood is a concave function, we conclude that this critical point is minimum and hence we have obtained the MLE.