

# Inverse Covariance Matrix Age

March 17, 2016

Inspired by:

- [2], gives form of inverse covariance from AR coefficients, univariate.
  - Method based on spectral polynomial and matrix form of lag operator.
  - The method can be extended to multivariate case.
  - It involves calculation of semi-infinite matrix, and the calculation detail is not given.
  - No discussion about the effect of finite truncation or starting point.
- [1], gives form of inverse covariance from AR coefficients, univariate.
  - Method based on spectral polynomial and algebra operations to construct related linear equations.
  - With modification, can be extended to multivariate case to some extent (no close form solution).

## 1 Solving Covariance Matrix And Its Inverse Using Method in [1]

The spectral polynomial of an (stable) AR process satisfies ( $\cdot^H$  means conjugate transpose)

$$\sum_{k=-\infty}^{\infty} R_k z^k = \left( \sum_{k=0}^m A_k z^k \right)^{-1} \Sigma \left( \sum_{k=0}^m A_k^H z^{-k} \right)^{-1}.$$

$$R_k = R_{-k}^H, \quad A_0 = I$$

Match the coefficients of  $z$  in both side of

$$\left( \sum_{k=-\infty}^{\infty} R_k z^k \right) \left( \sum_{k=0}^m A_k^H z^{-k} \right) = \left( \sum_{k=0}^m A_k z^k \right)^{-1} \Sigma,$$

get

$$\sum_{k=0}^m R_{k+j} A_k^H = \begin{cases} B_j \Sigma & j \geq 0 \\ 0 & j < 0 \end{cases}, \quad (1)$$

$$\sum_{k=0}^{\infty} B_k z^k = \left( \sum_{k=0}^m A_k z^k \right)^{-1}. \quad (2)$$

Define ( $V$  is the covariance matrix up to order  $m$ )

$$V = \begin{bmatrix} R_0 & R_1 & \cdots & R_{m-1} \\ R_{-1} & R_0 & & \vdots \\ \vdots & & \ddots & \\ R_{1-m} & \cdots & & R_0 \end{bmatrix}, \quad V_{\star} = \begin{bmatrix} R_{-m} & \cdots & R_{-2} & R_{-1} \\ R_{-m-1} & R_{-m} & & R_{-2} \\ \vdots & & \ddots & \vdots \\ R_{1-2m} & \cdots & & R_{-m} \end{bmatrix}$$

$$A^H = \begin{bmatrix} A_0^H & 0 & \cdots & 0 \\ A_1^H & A_0^H & & \vdots \\ \vdots & \ddots & \ddots & 0 \\ A_{m-1}^H & \cdots & A_1^H & A_0^H \end{bmatrix}, \quad A_{\star}^H = \begin{bmatrix} A_m^H & \cdots & A_2^H & A_1^H \\ 0 & A_m^H & & A_2^H \\ \vdots & & \ddots & \vdots \\ 0 & \cdots & 0 & A_m^H \end{bmatrix}$$

$$B = \begin{bmatrix} B_0 & B_1 & \cdots & B_{m-1} \\ 0 & B_0 & \ddots & \vdots \\ \vdots & & \ddots & B_1 \\ 0 & \cdots & 0 & B_0 \end{bmatrix}, \quad \Gamma = \begin{bmatrix} \Sigma & 0 & \cdots & 0 \\ 0 & \Sigma & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 0 & \cdots & 0 & \Sigma \end{bmatrix}$$

Eq.(2) in matrix form is

$$B = A^{-1}$$

Write Eq.(1) in matrix form:

$$\begin{bmatrix} V & V_\star^H \end{bmatrix} \begin{bmatrix} A^H \\ A_\star^H \end{bmatrix} = B\Gamma \quad (3)$$

$$\begin{bmatrix} V_\star & V \end{bmatrix} \begin{bmatrix} A^H \\ A_\star^H \end{bmatrix} = O \quad (4)$$

From Eq.(3)(4) we get

$$\begin{cases} VA^H + V_\star^H A_\star^H = A^{-1}\Gamma \\ V_\star A^H + V A_\star^H = O \\ AVA^H - A_\star V A_\star^H = \Gamma \end{cases} \quad (5)$$

If define

$$G = A^{-1}A_\star \\ U = A^{-1}\Gamma A^{-1H}$$

Then Eq.(5) become

$$GVG^H - V + U = 0, \quad (6)$$

which is the discrete Lyapunov equation. The solution is

$$V = \sum_{j=0}^{\infty} G^j U G^{jH}.$$

To solve  $V^{-1}$ , start from Eq.(6), and use Binomial inverse theorem

$$\begin{aligned} V^{-1} &= (U + GVG^H)^{-1} \\ &= U^{-1} - U^{-1}G(V^{-1} + G^H U^{-1}G)^{-1}G^H U^{-1} \\ &= A^H \Gamma^{-1} A - A^H \Gamma^{-1} A_\star (V^{-1} + A_\star^H \Gamma^{-1} A_\star)^{-1} A_\star^H \Gamma^{-1} A \end{aligned} \quad (7)$$

Note that by sufficient zero padding to  $A_k$  (increase the size of  $V$  or  $V^{-1}$  but keep the real fitting order), non-zero terms of  $A_\star$  will locate at lower left corner, non-zero terms of  $A$  will locate around (upper) diagonal. Therefore the non-zero terms of  $A^H \Gamma^{-1} A_\star$  will locate at lower left corner, and hence (no matter what  $V^{-1}$  is) the non-zeros of  $A^H \Gamma^{-1} A_\star (V^{-1} + A_\star^H \Gamma^{-1} A_\star)^{-1} A_\star^H \Gamma^{-1} A$  will locate at lower right, the width is (the real)  $m$ .

So as a consequence, except the  $m \times m$  block of the  $V^{-1}$ , it can be calculated as  $A^H \Gamma^{-1} A$

$$V^{-1} \approx A^H \Gamma^{-1} A \quad (8)$$

It happens that only the upper left of  $V^{-1}$  is in our interesting.

Following [2] will get the same result as Eq.(8).

## 2 Comparison to Spectral Density Inversion Approximation

The inverse if spectral density of AR process can be easily calculated as

$$S^{-1}(w) = A^H(w) \Sigma A(w). \quad (9)$$

Then by inverse Fourier transform, we can construct an (another) approximation of  $V^{-1}$

$$Q = \begin{bmatrix} Q_0 & Q_1 & \cdots & Q_{m-1} \\ Q_{-1} & Q_0 & & \vdots \\ \vdots & & \ddots & \\ Q_{1-m} & \cdots & & Q_0 \end{bmatrix}$$

$$Q_j = \mathcal{F}^{-1} [S^{-1}(w)](j)$$

The difference can be expressed as

$$Q - V^{-1} = A_{\#}^H \Gamma^{-1} A_{\#}$$

$$A_{\#} = \begin{bmatrix} A_1 & 0 & \cdots & 0 \\ A_2 & A_2 & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ A_m & \cdots & A_m & A_m \end{bmatrix}$$

### 3 Expression For Certain Block

Define

$$A^{(u)} = \begin{bmatrix} A_{0,[:,u]} & A_{1,[:,u]} & \cdots & A_{m-1,[:,u]} \\ 0 & A_{0,[:,u]} & & \vdots \\ \vdots & & \ddots & A_{1,[:,u]} \\ 0 & \cdots & 0 & A_{0,[:,u]} \end{bmatrix}$$

$$A_{k,[:,u]} = \begin{bmatrix} A_{k,1,u} \\ \vdots \\ A_{k,p,u} \end{bmatrix}$$

$$Q^{(zy)} = A^{(z)H} \Gamma^{-1} A^{(y)}$$

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### References

- [1] E. J. Godolphin and J. M. Unwin. Evaluation of the Covariance Matrix for the Maximum Likelihood Estimator of a Gaussian Autoregressive-Moving Average Process. *Biometrika*, 70(1):279–284, 1983.
- [2] J. Wise. The Autocorrelation Function and the Spectral Density Function. *Biometrika*, 42(1/2):151–159, 1955.