A. Robust Quickest Change Detection

 In some scenarios, the linear relationship of the system, which is denoted as $\mathbf{H} \in \mathbb{R}^{M \times N}$ in our paper, cannot be fully determined. In practice, \mathbf{H} is usually estimated by some methods, which may cause the inaccuracy of \mathbf{H} . In this section, by exploiting the robust counterpart optimization techniques, we formulate the quickest change detection problem with uncertain \mathbf{H} also as a MIQP problem. And therefore, both our detectors can be used in this case.

Before considering the uncertainty of \mathbf{H} , we first derive an equivalent form of our MIQP problem (16). In (16) the constraint $(\mathbf{I} - \mathbf{P}_{\mathbf{H}}^{\perp})(\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}) = \mathbf{0}$ is equivalent to $\mathbf{H}^{T}(\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}) = \mathbf{0}$, since they are essentially ensuring that $(\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)})$ lie in the complementary space of $\mathcal{R}(\mathbf{H})$. And based on this constraint and the definition of $\tilde{\boldsymbol{x}}^{(t)}$ in (32), the item $(\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)})^{T} \tilde{\boldsymbol{x}}^{(t)}$ in the objective function is equivalent to $(\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)})^{T} \boldsymbol{x}^{(t)}$. Thus, an equivalent formulation of (16) can be written as follows:

$$\min \frac{1}{2\sigma^{2}} \left\{ \left\| (\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}) \right\|_{2}^{2} - 2(\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)})^{T} \boldsymbol{x}^{(t)} \right\}$$

$$s.t. \begin{cases} \mathbf{H}^{T} (\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}) = \mathbf{0} \\ \rho_{L} \boldsymbol{u}^{(t)} \leq \boldsymbol{\mu}_{+}^{(t)} + \boldsymbol{\mu}_{-}^{(t)} \leq \rho_{U} \boldsymbol{u}^{(t)} \\ \boldsymbol{\mu}_{+}^{(t)} \geq \mathbf{0} \\ \boldsymbol{\mu}_{-}^{(t)} \geq \mathbf{0} \\ \boldsymbol{\mu}_{+}^{(t)} + \rho_{U} \boldsymbol{b}^{(t)} \leq \rho_{U} \mathbf{1} \\ \boldsymbol{\mu}_{-}^{(t)} - \rho_{U} \boldsymbol{b}^{(t)} \leq \mathbf{0} \end{cases}$$

$$\operatorname{var}: \boldsymbol{u}^{(t)} \in \{0, 1\}^{M}, \boldsymbol{b}^{(t)} \in \{0, 1\}^{M}, \boldsymbol{\mu}_{\perp}^{(t)}, \boldsymbol{\mu}_{-}^{(t)}.$$

$$(A.1)$$

In this paper, we assume that the uncertainty set of \mathbf{H} is constraint-wise. Specifically, the uncertainties between different rows of \mathbf{H}^T are decoupled. This assumption is commonly used and can be applied to many practical problems (Yang et al., 2014) (Bertsimas & Sim, 2004). Let h_i denote the *i*-th column of \mathbf{H} . And each h_i belongs to an uncertainty set, i.e. $h_i \in \mathcal{U}_i$. By simply relaxing the constraint $\mathbf{H}^T(\boldsymbol{\mu}_+^{(t)} - \boldsymbol{\mu}_-^{(t)}) = \mathbf{0}$, we get the new constraints as follows:

$$-\varepsilon_i \le \boldsymbol{h}_i^T(\boldsymbol{\mu}_+^{(t)} - \boldsymbol{\mu}_-^{(t)}) \le \varepsilon_i, \ \forall \ \boldsymbol{h}_i \in \mathcal{U}_i, \ i = 1, \cdots, N,$$
(A.2)

where ε_i represents a small positive constant. And the above constraints can be equivalently transformed into several linear constraints, not only when \mathcal{U}_i represents some simple uncertain sets, including interval-based uncertainty set (Duchi, 2018), but also some very general uncertain sets (Yang et al., 2014), including polyhedron uncertainty set and D-norm uncertainty set. We take the general polyhedron uncertainty set as an example, i.e. $\mathcal{U}_i = \{\mathbf{h}_i \mid \mathbf{D}_i \mathbf{h}_i \leq \mathbf{d}_i\}$, $i = 1, \dots, N$. According to (Ben-Tal et al., 2009), by exploiting the robust counterpart optimization techniques, the constraints A.2 under the polyhedron uncertainty set can be equivalently reformulated as

$$\mathbf{p}_{i}^{T} \mathbf{d}_{i} \leq \varepsilon_{i}, \quad \mathbf{D}_{i}^{T} \mathbf{p}_{i} = \boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}, \quad \mathbf{p}_{i} \geq 0, \quad i = 1, \cdots, N,$$

$$\mathbf{n}_{i}^{T} \mathbf{d}_{i} \leq \varepsilon_{i}, \quad -\mathbf{D}_{i}^{T} \mathbf{n}_{i} = \boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}, \quad \mathbf{n}_{i} \geq 0, \quad i = 1, \cdots, N.$$
(A.3)

Thus, our MIQP problem where every column of H belongs to the polyhedral uncertainty set is as follows:

$$\min \frac{1}{2\sigma^{2}} \left\{ \left\| (\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}) \right\|_{2}^{2} - 2(\boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)})^{T} \boldsymbol{x}^{(t)} \right\} \\
= \left\{ \begin{aligned}
& \boldsymbol{p}_{i}^{T} \boldsymbol{d}_{i} \leq \varepsilon_{i}, & \boldsymbol{D}_{i}^{T} \boldsymbol{p}_{i} = \boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}, & \boldsymbol{p}_{i} \geq 0, & i = 1, \cdots, N \\
& \boldsymbol{n}_{i}^{T} \boldsymbol{d}_{i} \leq \varepsilon_{i}, & -\boldsymbol{D}_{i}^{T} \boldsymbol{n}_{i} = \boldsymbol{\mu}_{+}^{(t)} - \boldsymbol{\mu}_{-}^{(t)}, & \boldsymbol{n}_{i} \geq 0, & i = 1, \cdots, N \\
& \boldsymbol{\rho}_{L} \boldsymbol{u}^{(t)} \leq \boldsymbol{\mu}_{+}^{(t)} + \boldsymbol{\mu}_{-}^{(t)} \leq \boldsymbol{\rho}_{U} \boldsymbol{u}^{(t)} \\
& \boldsymbol{\mu}_{+}^{(t)} \geq \boldsymbol{0} \\
& \boldsymbol{\mu}_{+}^{(t)} \geq \boldsymbol{0} \\
& \boldsymbol{\mu}_{+}^{(t)} + \boldsymbol{\rho}_{U} \boldsymbol{b}^{(t)} \leq \boldsymbol{\rho}_{U} \boldsymbol{1} \\
& \boldsymbol{\mu}_{-}^{(t)} - \boldsymbol{\rho}_{U} \boldsymbol{b}^{(t)} \leq \boldsymbol{0} \\
\end{aligned} \tag{A.4}$$

$$\text{var}: \boldsymbol{u}^{(t)} \in \{0,1\}^{M}, \boldsymbol{b}^{(t)} \in \{0,1\}^{M}, \boldsymbol{\mu}_{+}^{(t)}, \boldsymbol{\mu}_{-}^{(t)}, \{\boldsymbol{p}_{i}\}, \{\boldsymbol{n}_{i}\}.$$

And the linear constraints introduced by the uncertainty of **H** do not change the nature of the problem, so that both our detectors can still be used. For convenient subsequent comparisons, we rename the SDPCUSUM and BBCUSUM considering the uncertainty of **H** as RSDPCUSUM and RBBCUSUM, repectively.

B. Numerical Results

We compare the performance of different detectors on $Data\ I$, and the results show that RSDPCUSUM and RBBCUSUM achieve better performance when the matrix H is inaccurate.

We assume the **H** with the value given in (24) models the true linear relationship and is denoted as \mathbf{H}_r . The inaccurate estimate of **H** is denoted as \mathbf{H}_e , and its value in our experiment is set to be

$$\mathbf{H} = \begin{bmatrix} 1.4 & 0 \\ 0 & 1.3 \\ -1.2 & 1.2 \\ 0.8 & 0.6 \end{bmatrix}. \tag{1}$$

And we use the polyhedron uncertainty set to describe the uncertainty of H, i.e,

$$\mathbf{D}_{i}\mathbf{h}_{i} \leq \mathbf{d}_{i}, \ i = 1, 2,$$

and

$$\mathbf{D}_1 = \mathbf{D}_2 = \left[egin{array}{c} \mathbf{I} \\ -\mathbf{I} \end{array}
ight],$$

where I represents the 4×4 identity matrix. d_1 and d_2 are set to be

$$\mathbf{d}_1 = [1.5, 0.5, -0.5, 1.5, -0.5, 0.5, 1.5, -0.5]^T, \mathbf{d}_2 = [0.5, 1.5, 1.5, 1.5, 0.5, -0.5, -0.5, -0.5]^T.$$

The experimental results are shown in TABLE 3. It is seen from TABLE 3 that under different (ρ_L , ρ_H), the average detection delay of RSDPCUSUM and RBBCUSUM are smaller, which implies that RSDPCUSUM and RBBCUSUM achieve better performance with inaccurate **H**.

Table 3. Performance Comparison on Simulated Data

	BBCUSUM	RSDPCUSUM	RBBCUSUM
	AVERAGE DETECTION DELAY	AVERAGE DETECTION DELAY	AVERAGE DETECTION DELAY
$ \rho_L = 0.5, \rho_H = 3.0 $ $ \sigma^2 = 2.0 $			
$\sigma^2 = 2.0$	461.5	419.6	406.7
FALSE ALARM PERIOD = 500.0			
$ \rho_L = 0.1, \rho_H = 1.5 $ $ \sigma^2 = 2.0 $			
$\sigma^2 = 2.0$	478.2	402.7	408.8
FALSE ALARM PERIOD = 500.0			

C. Performance Analysis

In our paper, we provide a guaranteed sub-optimality gap for SDPCUSUM. And similar conclusion exists in RSDPCUSUM.

Theorem C.1. Let s_M denote the optimal solution to problem A.1 and s_P denote the suboptimal solution generated by the RSDPCUSUM. Let the function $g: \mathbb{R}^M \to \mathbb{R}$ denote the objective function of problem A.1. The objective function value of the suboptimal solution s_P can be bounded as follows:

$$g(\mathbf{s}_M) \le g(\mathbf{s}_P) \le g(\mathbf{s}_M) + \frac{2\sqrt{M}\rho_H}{\sigma^2} \|\mathbf{x}\|_2,$$
 (2)

where σ is the variance of noise, M is the dimension of the observation vector, and ρ_H is the upper bound of the detectable part of the injected anomaly.

The proof is similar to that of **Theorem 1**. References Ben-Tal, A., El Ghaoui, L., and Nemirovski, A. Robust optimization, volume 28. Princeton university press, 2009. Bertsimas, D. and Sim, M. The price of robustness. *Operations research*, 52(1):35-53, 2004. Duchi, J. Optimization with uncertain data. Notes, 2018, 2018. Yang, K., Huang, J., Wu, Y., Wang, X., and Chiang, M. Distributed robust optimization (dro), part i: Framework and example. Optimization and Engineering, 15:35-67, 2014.