**Bridging Gaussian Markov Random Fields and** 

Copulas: A Fast Algorithm for Efficient Gaussian

Copula Density Computation with Matérn-like

**Precision Matrices** 

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Abstract

Gaussian Markov Random Fields (GMRFs) have long been a powerful tool for

modeling spatial and temporal dependencies in various fields. Similarly, copulas have

proven invaluable for modeling complex dependency structures in multivariate data.

However, the combination of these two approaches - using GMRFs within copula mod-

els - has historically been computationally inefficient, limiting their joint application

to smaller datasets or simpler models. This work presents a novel algorithm that over-

comes these limitations, allowing for fast and efficient computation of Gaussian Cop-

ula densities using GMRF precision structures. By bridging the gap between GMRFs

and copulas, this method opens up new possibilities for analyzing large-scale, high-

dimensional spatial and spatio-temporal data with complex dependency structures.

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# Introduction

Gaussian Markov Random Fields (GMRFs) and Copulas are two powerful tools in modern statistics, each with their own strengths in modeling complex data structures. GMRFs have been widely used for modeling spatial and temporal dependencies, particularly in fields such as environmental science, epidemiology, and image analysis. Their ability to capture local dependencies through a sparse precision matrix makes them computationally attractive for high-dimensional problems.

Copulas, on the other hand, provide a flexible framework for modeling multivariate dependencies, allowing for the separate specification of marginal distributions and their joint behavior. The Gaussian copula, in particular, has gained popularity due to its interpretability and connection to the multivariate normal distribution.

Let  $\mathbf{X} = (X_1, X_2, \dots, X_n)$  be a multivariate random vector with marginal distribution functions  $F_i$  for  $i = 1, 2, \dots, n$  and dependence that is governed by a GMRF copula. The joint distribution function of  $\mathbf{X}$  can be written as:

$$F_{\mathbf{X}}(\mathbf{x}) = C(F_1(x_1), F_2(x_2), \ldots, F_n(x_n)),$$

where C is the Gaussian copula defined by the GMRF precision matrix Q.

The Gaussian copula C is given by:

$$C(u_1,u_2,\dots,u_n)=\Phi_Q(\Phi^{-1}(u_1),\Phi^{-1}(u_2),\dots,\Phi^{-1}(u_n)),$$

where  $\Phi_Q$  is the joint cumulative distribution function of a multivariate normal distribution with mean vector  $\mathbf{0}$  and precision matrix Q, and  $\Phi^{-1}$  is the inverse of the standard normal cumulative distribution function. Its density is It is imperative that the precision matrix Q governing the GMRF Copula, C, has marginal variance equal to 1 so that is it on the same scale as the transformed data,  $\Phi^{-1}(u_i)$ . This can be troublesome because GMRFs are defined in terms of their precision matrices, Q, which more often than not have marginal variances that are different from 1.

This paper presents a fast and efficient algorithm for creating a Matérn-like precision matrix,  $\mathbf{Q}$ , with unit marginal variance, and computing the multivariate Gaussian copula density of  $\mathbf{Z} = \Phi^{-1}(\mathbf{u})$  where  $\mathbf{u} \sim \text{Uniform}(0,1)$ . The method leverages the special structure of the precision matrix and employs efficient eigendecomposition techniques to avoid explicit formation and inversion of the large precision matrix  $\mathbf{Q}$ .

# Methods

# **GMRF** Copula

Let  $\mathbf{X} = (X_1, X_2, \dots, X_n)$  be a multivariate random vector with marginal distribution functions  $F_i$  for  $i = 1, 2, \dots, n$  and dependence that is governed by a Gaussian copula. We create the vector

The matrix, Q, is defined as

$$Q = Q_1 \otimes I + I \otimes Q_1,$$

where  $Q_1$  is the precision matrix of a standardized one-dimensional AR(1) process and  $\otimes$  is the kronecker product.

### Theory

## One-Dimensional AR(1) Matrix

The core of our approach is based on the eigendecomposition of an AR(1) precision matrix, which forms the building block of our Matérn-like precision structure. For a one-dimensional AR(1) process with parameter  $\rho$ , the precision matrix  $Q_1$  has a tridiagonal structure:

$$Q = \frac{1}{1 - \rho^2} \begin{bmatrix} 1 & -\rho & 0 & \cdots & 0 & 0 \\ -\rho & 1 + \rho^2 & -\rho & \cdots & 0 & 0 \\ 0 & -\rho & 1 + \rho^2 & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & \cdots & 1 + \rho^2 & -\rho \\ 0 & 0 & 0 & \cdots & -\rho & 1 \end{bmatrix}$$

### Matérn-like Precision Matrix

For two-dimensional spatial fields, we construct a Matérn-like precision matrix Q using Kronecker products:

$$Q = Q_1 \otimes I + I \otimes Q_1$$

where I is the identity matrix and  $\otimes$  denotes the Kronecker product The eigenvalues  $(\lambda_k)$  and eigenvectors  $(v_k)$  of  $Q_1$  can be calculated numerically and used to evaluate the multivariate Gaussian density, letting us skip out on forming Q entirely. This is due to the following theorem:

### Eigendecomposition of Kronecker Sums

#### Theorem

Let  $A \in \mathbb{R}^{n \times n}$  have eigenvalues  $\lambda_i, i \in \{1, \dots, n\}$ , and let  $B \in \mathbb{R}^{m \times m}$  have eigenvalues  $\mu_j, i \in \{1, \dots, m\}$ . Then the Kronecker sum  $A \oplus B = (I_m \otimes A) + (B \otimes I_n)$  has eigenvalues  $\lambda_i + \mu_j, i \in \{1, \dots, n\}, j \in \{1, \dots, m\}$ .

Moreover, if  $x_1,\ldots,x_p$  are linearly independent right eigenvectors of A corresponding to  $\lambda_1,\ldots,\lambda_p$   $(p\leq n)$ , and  $z_1,\ldots,z_q$  are linearly independent right eigenvectors of B corresponding to  $\mu_1,\ldots,\mu_q$   $(q\leq m)$ , then  $z_j\otimes x_i\in\mathbb{R}^{mn}$  are linearly independent right eigenvectors of  $A\oplus B$  corresponding to  $\lambda_i+\mu_j,\ i\in\{1,\ldots,p\},\ j\in\{1,\ldots,q\}.$ 

#### Discussion

This theorem provides a crucial insight that allows us to efficiently construct the eigendecomposition of the full precision matrix Q for two-dimensional spatial fields by leveraging the eigendecomposition of the one-dimensional precision matrix  $Q_1$ , avoiding the computationally intensive process of explicitly forming and inverting the large matrix Q.

Furthermore, we can compute the log-density of an observation x directly by using just the eigenvalues and eigenvectors og  $Q_1$ , thereby enabling efficient computation of the log-density of the multivariate normal distribution even for large spatial fields.

# Example: Calculating Eigenvalues and Eigenvectors for Q Using $Q_1$

To illustrate the application of Theorem 13.16, consider the one-dimensional precision matrix  $Q_1$  with known eigendecomposition. Suppose  $Q_1$  has eigenvalues  $\lambda_i$  and corresponding eigenvectors  $v_i$ . For a two-dimensional spatial field, we construct the Matérn-like precision matrix Q using Kronecker products as follows:

$$Q = Q_1 \otimes I + I \otimes Q_1$$

where I is the identity matrix and  $\otimes$  denotes the Kronecker product.

### **Eigenvalues**

The eigenvalues of Q can be determined from the eigenvalues of  $Q_1$ . If  $Q_1$  has eigenvalues  $\lambda_i$  for i = 1, 2, ..., n, then the eigenvalues of Q are given by:

$$\lambda_{ij} = \lambda_i + \lambda_j$$
 for  $i, j = 1, 2, \dots, n$ .

### **Eigenvectors**

Similarly, the eigenvectors of Q can be constructed from the eigenvectors of  $Q_1$ . If  $v_i$  and  $v_j$  are eigenvectors of  $Q_1$  corresponding to eigenvalues  $\lambda_i$  and  $\lambda_j$  respectively, then the eigenvectors of Q are given by:

$$v_{ij} = v_i \otimes v_j \quad \text{for} \quad i,j = 1,2,\dots,n.$$

Here,  $v_{ij}$  is the Kronecker product of  $v_i$  and  $v_j$ .

These relationships allow us to efficiently compute the eigendecomposition of the full precision matrix Q using the eigendecomposition of the smaller matrix  $Q_1$ , significantly reducing the computational complexity.

# Efficient Density Calculation Using Eigendecomposition

Given the eigendecomposition of Q1, we can efficiently compute the multivariate normal density with respect to the precision matrix Q. Let  $\lambda_i$  and  $v_i$  be the eigenvalues and eigenvectors

of Q1, respectively.

The log-density of a multivariate normal distribution with precision matrix Q is given by:

$$\log p(x) = -\frac{1}{2}(n\log(2\pi) - \log|Q| + x^TQx)$$

where n is the dimension of x, |Q| is the determinant of Q, and  $x^TQx$  is the quadratic form.

### Log-Determinant Calculation

The log-determinant of Q can be computed efficiently using the eigenvalues of Q1:

$$\log |Q| = \sum_{i=1}^d \sum_{j=1}^d \log(\lambda_i + \lambda_j)$$

where d is the dimension of Q1.

### **Quadratic Form Calculation**

First, we define the product of the eigenvector and the data vector x as:

$$y_{ij} = (v_i \otimes v_j)^T x,$$

where  $v_i$  and  $v_j$  are the eigenvectors of  $Q_1$  and  $\otimes$  denotes the Kronecker product.

Next, we define the eigenvalue sum as:

$$\mu_{ij} = \lambda_i + \lambda_j,$$

where  $\lambda_i$  and  $\lambda_j$  are the eigenvalues of  $Q_1$ . Using these definitions, the quadratic form can be expressed as:

$$x^{T}Qx = \sum_{i=1}^{d} \sum_{j=1}^{d} \mu_{ij} y_{ij}^{2}.$$

In this way, we calculate the quadratic form without having to form the matrix Q or its full set of eigenvectors or values.

### Scaling the Input x

The input vector  $x = \Phi^{-1}(u)$  has zero mean and unit marginal variance. Instead of scaling the precision matrix to have  $(Q^{-1})_{ii} = 1$ , we calculate the marginal standard deviations,  $\sigma_k$ , implied by Q using the eigenstructure of  $Q_1$ , then scale the input vector x using those standard deviations. First we compute the marginal standard deviations:

$$\sigma_k = \sqrt{\sum_{i=1}^d \sum_{j=1}^d \frac{(v_i \otimes v_j)_k^2}{\lambda_i + \lambda_j}}$$

where  $(v_i \otimes v_j)_k$  is the k-th element of the Kronecker product  $v_i \otimes v_j$ .

# Algorithm Implementation

The density calculation is implemented as follows:

- Compute the eigendecomposition of  $Q_1$ .
- Calculate the marginal standard deviations  $\sigma_k$ .
- For each observation x:
  - a. Standardize x by element-wise multiplication with  $\sigma_k$ .
  - b. Compute the log-determinant and quadratic form using the formulas above.
  - c. Combine the terms to get the log-density.

This approach avoids explicit formation and inversion of the full precision matrix Q, allowing for efficient computation even for large spatial fields. This method has been implemented in C++ code that can perform a full set of computations for a  $200 \times 200$  spatial grid (i.e. Q would be  $40.000 \times 40.000$ ) in just over one second.