

key issue: censoring: 1) either event happens after sample ends  
2) or the unit/person chooses to leave sample before the event

$T$  = survival time / event time / failure time

$C$  = censoring time  $\rightarrow$  sample ends, or person leaves your sample

observe  $Y = \min(T, C)$

$$\delta = \begin{cases} 1 & \text{if } T \leq C \\ 0 & \text{if } T > C \end{cases}$$

sample  $(y_i, \delta_i)$   $i = 1, \dots, n$  units

assume: 1) only 1 event  
2)  $C + T$  are independent  $\left. \vphantom{\begin{matrix} 1) \\ 2) \end{matrix}} \right\}$  both restrictive, can be relax w/ more complexity

Survival Curve or Function

$$S(t) = \Pr(T > t) \rightarrow \text{declining in } t$$

$\rightarrow$  can't just calculate % of sample "alive" at each  $t$   
b/c of censoring

let  $d_1 < d_2 < \dots < d_K$  be  $K$  unique event dates/times for uncensored observations

$g_k = \#$  units w/ event at  $d_k$

$r_k = \#$  units survived until just before  $d_k$   
"at risk" units

$$\begin{aligned} \Pr(T > d_k) &= \Pr(T > d_k \mid T > d_{k-1}) \cdot \Pr(T > d_{k-1}) \\ &\quad + \underbrace{\Pr(T > d_k \mid T \leq d_{k-1}) \cdot \Pr(T \leq d_{k-1})}_{=0} \end{aligned}$$

$$= \Pr(T > d_k \mid T > d_{k-1}) \cdot \Pr(T > d_{k-1})$$

$$S(d_k) = \underbrace{\Pr(T > d_k \mid T > d_{k-1})}_{=1} \cdot S(d_{k-1})$$

$$S(d_k) = \underbrace{P_r(T > d_k | T > d_{k-1})}_{\text{fraction of units "alive" at } k \text{ who survive just past } k} \cdot S(d_{k-1})$$

$$\hat{P}_r(T > d_k | T > d_{k-1}) = \frac{r_k - q_k}{r_k}$$

fraction of units "alive" at  $k$  who survive just past  $k$

$$\hat{S}(d_k) = \prod_{j=1}^k \left( \frac{r_j - q_j}{r_j} \right) \quad \text{Kaplan Meier Survival Curve}$$

Suppose we have  $(y_i, \delta_i, x_{i1}, \dots, x_{ip})$

and want to predict survival w/  $X$ 's need to account for censoring

use hazard rate/function: likelihood of event happening in the next instant, conditional on having survived to the current instant

$$\begin{aligned} h(t) &= \lim_{\Delta t \rightarrow 0} \frac{P_r(t < T < t + \Delta t | T \geq t)}{\Delta t} \approx \frac{\partial P_r(T = t | T \geq t)}{\partial t} \\ &= \frac{f(t)}{S(t)} = \frac{\cancel{\partial P_r(T = t)} / \partial t}{P_r(T \geq t)} = \frac{\text{unconditional}}{\text{total}} \end{aligned}$$

likelihood of observing  $(y_i, \delta_i)$   $L_i = \begin{cases} f(y_i) & \text{if } i \text{ not censored} \\ S(y_i) & \text{if } i \text{ is censored} \end{cases}$  Prob. of dying at  $y_i$

$$= f(y_i)^{\delta_i} \cdot S(y_i)^{1-\delta_i}$$

$$L = \prod_{i=1}^n f(y_i)^{\delta_i} S(y_i)^{1-\delta_i}$$

$$= \prod_{i=1}^n h(y_i)^{\delta_i} S(y_i) \quad \text{or} \quad h(y_i | x_i)^{\delta_i} \cdot S(y_i | x_i)$$

Cox Proportional Hazards Model:

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$$h(t | \underline{x}_i) = \underbrace{h_0(t)}_{\text{unspecified}} \cdot \underbrace{\exp\left(\sum_{j=1}^p x_{ij} \cdot \beta_j\right)}_{\substack{\text{arguments baseline hazard} \\ \text{"relative risk"} \rightarrow \text{relative to} \\ \underline{x} = (0, 0, \dots, 0)}}$$

$$= \underbrace{\left[ h_0(t) \cdot e^{\beta_0} \right]}_{\substack{\text{unspecified} \\ \text{hazard}}} \cdot \overset{\text{no intercept}}{e^{\left(\sum_{j=1}^p x_{ij} \beta_j\right)}}$$

Suppose  $\delta_i = 1$ ,  $i$  is uncensored  $y_i =$  true event time

$$A_i = \cancel{h_0(y_i)} \exp(\underline{x}_i' \beta) \quad \text{hazard for } i$$

total hazard at time  $y_i$  for all surviving units

$$B_i = \sum_{i': (y_{i'} \geq y_i)} \cancel{h_0(y_i)} \exp(\underline{x}_{i'}' \beta) \quad \leftarrow \begin{array}{l} \text{all others not} \\ \text{yet censored or} \\ \text{not yet had event} \end{array}$$

$$\text{Prob}(i \text{ fails at time } y_i) = \frac{A_i}{B_i} = \frac{\exp(\underline{x}_i' \beta)}{\sum_{i'} (\underline{x}_{i'}' \beta)}$$

hazards cancel

$$\text{Partial Likelihood } \max_{\beta} PL(\beta) = \prod_{i=1}^n \frac{A_i}{B_i} \quad \left( \begin{array}{l} + \text{shrinkage or} \\ \text{regularization} \\ \text{penalty} \end{array} \right)$$