

# Autocorrelation, White Noise, and Linear Time Series

# This Lecture

By now we know that autocorrelation in the variables and residuals is important. How do we test for it?

1. Autocorrelation Function (ACF)
2. Test individual autocorrelations at a given lag
3. Test joint autocorrelations (Portmanteau, Box-Pierce, Ljung-Box, Q-test)
4. Interpreting an ACF correlogram
5. Partial ACF and correlogram
6. White noise
7. Properties of linear time series

## Recall: Stationarity

- *Weak/Covariance Stationarity*: The mean, variance, and autocovariance are constant through time, i.e.,  $E(r_t) = \mu$  (a constant) and  $Cov(r_t, r_{t-s}) = \gamma_s$  only depends on the distance between the observations  $s$  and not time period  $t$ .
- Note also that in general  $\gamma_0 = Var(r_t)$  and  $\gamma_s = \gamma_{-s}$  (because  $Cov(r_t, r_{t-s}) = Cov(r_{t+s}, r_t)$ ).

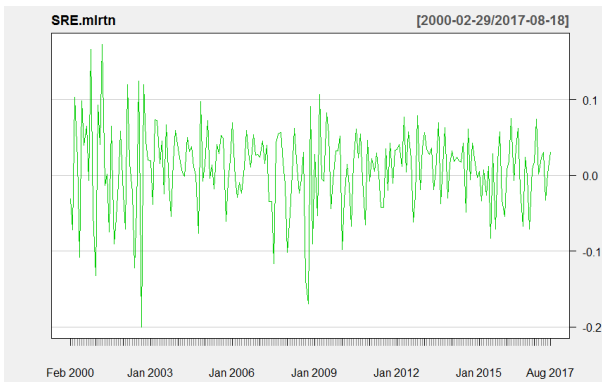
# NONSTATIONARY

Sempra Energy equity price



# STATIONARY

Sempra Energy monthly returns



# Autocorrelation Function

- Correlation between two random variables  $X$  and  $Y$ :

$$\rho_{X,Y} = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \cdot \text{Var}(Y)}} = \frac{E[(X - \mu_X)(Y - \mu_Y)]}{\sqrt{E(X - \mu_X)^2 E(Y - \mu_Y)^2}}$$

- For stationary  $r_t$ , if  $X = r_t$  and  $Y = r_{t-s}$  we have the Autocorrelation Function (ACF):

$$\rho_s = \frac{\text{Cov}(r_t, r_{t-s})}{\sqrt{\text{Var}(r_t) \cdot \text{Var}(r_{t-s})}} = \frac{\text{Cov}(r_t, r_{t-s})}{\text{Var}(r_t)} = \frac{\gamma_s}{\gamma_0}$$

- This is a function of the gap,  $s$ .

# Autocorrelation Function

- We can estimate this for our variable  $r_t$  if we want to build a model with lags of  $r_t$
- We can also do this for our regression residuals if we are worried about time series errors.

$$r_{it} = \hat{\beta}_0 + \hat{\beta}_1 r_{jt} + \hat{e}_t$$

- Calculate and test autocorrelations in  $\hat{e}_t$

$$\frac{\text{Cov}(\hat{e}_t, \hat{e}_{t-s})}{\text{Var}(\hat{e}_t)}$$

## Testing one at a time

- Use t-test for single autocorrelation at a specific lag

$$H_0 : \hat{\rho}_s = 0, \quad H_a : \hat{\rho}_s \neq 0$$

- By Central Limit Theorem,  $\hat{\rho}_s$  converges to Normal.

$$t = \frac{\hat{\rho}_s - 0}{se(\hat{\rho}_s)}$$

- What is  $se(\hat{\rho}_s)$ ?

- If your null is  $r_t \sim i.i.d.$ , then  $se(\hat{\rho}_s) = \sqrt{1/T}$ .
- If your null is  $\gamma_k \neq 0$  for  $k < s$ , then

$$se(\hat{\rho}_s) = \sqrt{\left(1 + 2 \sum_{i=1}^{s-1} \hat{\rho}_i^2\right) / T}$$

- (inflate the variance due to autocorrelation at lower lags)



## Jointly test ACF at multiple lags

- Use a version of a  $\chi^2$ -test for

$$H_0 : \hat{\rho}_1 = \cdots = \hat{\rho}_m = 0, \quad H_a : \hat{\rho}_i \neq 0 \text{ for some } i \in \{1, \dots, m\}$$

- Box-Pierce (Portmanteau):

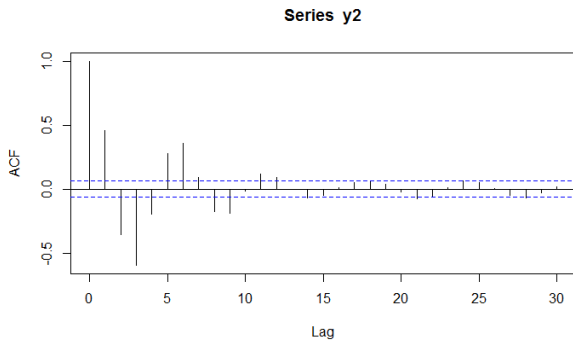
$$Q^*(m) = T \sum_{s=1}^m \hat{\rho}_s^2$$

- Sum of squared Normals  $\sim \chi^2(m)$
- Ljung-Box sample size adjustment (used in practice):

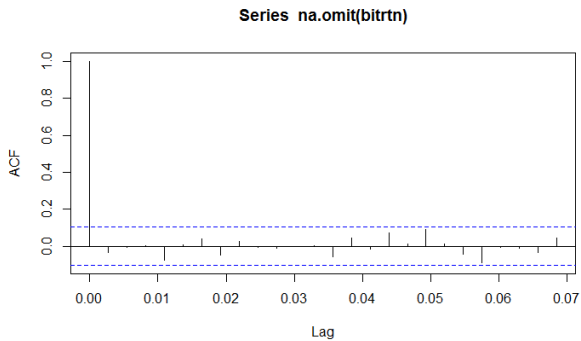
$$Q(m) = T(T+2) \sum_{s=1}^m \frac{\hat{\rho}_s^2}{T-s}$$

- $m$  is whatever your hypothesis is, or  $\approx \ln(T)$  (unless you have seasonality).

# ACF Correlograms



# ACF Correlograms



# Correlograms

- Vertical lines =  $\hat{\rho}_s$  at each lag  $s$ .
  - At  $s = 0$ , always perfectly correlated - spike where  $\hat{\rho}_1 = 1$ .
- Dashed horizontal lines = 95% confidence intervals for individual t-test at lag  $s$ .
- If a vertical line crosses dashed horizontal line, reject the null of zero autocorrelation at that lag.
- High frequency data: often many are just barely significant. Model the important ones.
- Joint significance?

# Partial Autocorrelation Function (PACF)

- ACF measures pairwise correlation at a given time gap, e.g.,  $\text{corr}(r_t, r_{t-3})$  but does not control for the lags in between, e.g.,  $r_{t-1}, r_{t-2}$ .
- PACF can be helpful in choosing number of lags in autoregressive model.
- Sequentially estimate

$$r_t = \phi_{0,1} + \phi_{1,1}r_{t-1} + e_{1,t}$$

$$r_t = \phi_{0,2} + \phi_{1,2}r_{t-1} + \phi_{2,2}r_{t-2} + e_{2,t}$$

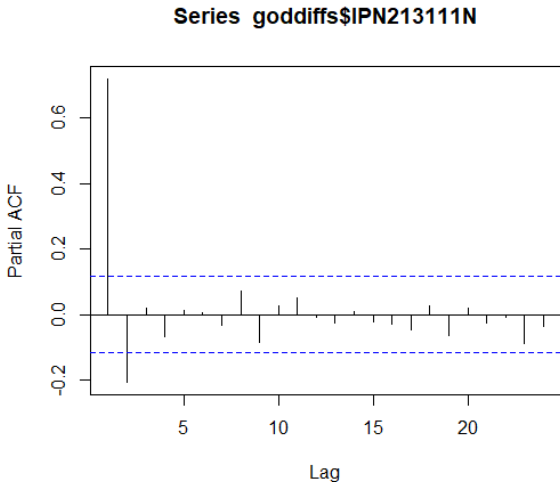
$$r_t = \phi_{0,3} + \phi_{1,3}r_{t-1} + \phi_{2,3}r_{t-2} + \phi_{3,3}r_{t-3} + e_{3,t}$$

etc.

- The PACF is  $(\phi_{1,1}, \phi_{2,2}, \phi_{3,3}, \dots)$ .
- Each successive estimate controls for the lags in between, measures the *additional* explanatory power of the next lag.

## PACF Correlograms

PACF for US oil & gas drilling index. Similar format to ACF, but no 0 coefficient.



# White noise

- *White Noise*: Sequence of i.i.d. (independent and identically distributed) random variables with finite mean and variance.
- *Gaussian White Noise*: Normally distributed white noise with mean 0 and variance  $\sigma^2$ .
- No significant ACFs.

# Linear Time Series

- A linear time series can be written

$$r_t = \mu + a_t + \psi_1 a_{t-1} + \cdots = \mu + \sum_{i=0}^{\infty} \psi_i a_{t-i}$$

- $\mu$  is the mean,  $\psi_0 = 1$
- $a_t$  is mean-zero white noise.
- $a_t$  is the shock/innovation/new information about  $r_t$  that arrives at time  $t$ .
- $r_t$  can be summarized by the cumulative effect of all past shocks.
- $\mu$  can contain a model with other explanatory variables.



# Linear Time Series

$$r_t = \mu + \sum_{i=0}^{\infty} \psi_i a_{t-i}$$

- Weights  $\psi_i$  on past shocks govern behavior.
- If all  $\psi_i = 0$  except  $\psi_0 = 1$ , then  $r_t$  is white noise (no autocorrelation).
- For  $r_t$  to be stationary,  $\psi$ -weights on past shocks must die out.

$$E(r_t) = \mu,$$

$$E[(r_t - \mu)^2] = E\left[\left(\sum_{i=0}^{\infty} \psi_i a_{t-i}\right)^2\right] = \sigma_a^2 \sum_{i=0}^{\infty} \psi_i^2$$

- For a nonstationary series, past shocks are **permanent**.

# Linear Time Series

$$r_t = \mu + \sum_{i=0}^{\infty} \psi_i a_{t-i}$$

- $\psi$ -weights and autocorrelation:

$$\begin{aligned}\gamma_s = \text{Cov}(r_t, r_{t-s}) &= E[(\sum_{i=0}^{\infty} \psi_i a_{t-i})(\sum_{j=0}^{\infty} \psi_j a_{t-s-j})] \\ &= E[\sum_{i,j=0}^{\infty} \psi_i \psi_j a_{t-i} a_{t-s-j}] \\ &= \sum_{j=0}^{\infty} \psi_{j+s} \psi_j E[a_{t-s-j}^2] \\ &= \sigma_a^2 \sum_{j=0}^{\infty} \psi_{j+s} \psi_j\end{aligned}$$

$$\rho_s = \frac{\gamma_s}{\gamma_0} = \frac{\sigma_a^2 \sum_{j=0}^{\infty} \psi_{j+s} \psi_j}{\sigma_a^2 \sum_{j=0}^{\infty} \psi_j^2}$$

- $\psi_j^2$  has to die out enough as  $j$  grows (far enough back in history) to sum to a finite constant
- $\psi_{j+s} \psi_j$  has to die out as  $s$  grows (far enough between observations) for autocorrelation to die out.

# Linear Time Series Takeaways

- Any linear process can be represented this way:
  - A cumulative weighted sum of all past shocks/innovations/information.
- Stationary:
  - Effect of (or weight on) past shocks dies out.
  - Autocorrelations die out as distance between observations grows
- Nonstationary:
  - Shocks are permanent
- White noise:
  - Past shocks get zero weight, autocorrelations are all not significantly different from zero.