Pour se remettre dans le bain



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- 1. Citer les quatre estimateurs bayésiens les plus couramment utilisés.
- 2. Qu'est-ce qu'une région de crédibilité ?
- 3. Comment prédire quand on est bayésien(ne) ?
- 4. Décrire l'approche quasi-bayésienne.
- 5. Pourquoi l'approche quasi-bayésienne peut-elle être vue comme une généralisation de l'apprentissage bayésien ?
- 6. Illustrer la provenance du quasi-posterior au moyen d'une formulation variationnelle.
- 7. Citer les quatre estimateurs quasi-bayésiens les plus couramment utilisés.
- 8. Rappeler le principe de l'ERM et de l'agrégation à poids convexe (EWA). Quel est le lien entre EWA et apprentissage quasi-bayésien ?



Assessing the performance: the oracle approach

Oracle:

$$\phi^{\star} \in \underset{\phi \in \mathcal{Y}^{\mathfrak{X}}}{\arg \min} \ R(\phi).$$

Ultimate goal: do almost as well as the oracle.



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Excess risk:

$$\mathcal{E}(\cdot) = R(\cdot) - R^* \ge 0, \qquad R^* = R(\phi^*).$$





Let R^* denote the oracle risk. For any $\epsilon > 0$,

$$\mathbb{P}\left(R\left(\widehat{\phi}_{\lambda}\right) - R^{\star} \leq \spadesuit \inf_{\phi \in \mathfrak{F}} \left\{R(\phi) - R^{\star} + \frac{\Delta(\phi, \epsilon)}{n^{\alpha}}\right\}\right) \geq 1 - \epsilon,$$

where $\spadesuit \geq 1$ and $\lambda \propto n$. If $\spadesuit = 1$, the inequality is *exact* or *sharp*.

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The remainder term grows with d and the size of \mathcal{F} . It decreases with n.

Hoeffding inequality

Let V_1, \ldots, V_n be independent real-valued random variables such that $a_i \leq V_i \leq b_i$ a.s. Let $\bar{V}_n = \frac{1}{n} \sum_{i=1}^n V_i$.

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$$\mathbb{P}(\bar{V}_n - \mathbb{E}\bar{V}_n > t) \leq \exp\left(-\frac{2n^2t^2}{\sum_{i=1}^n(b_i - a_i)^2}\right), \forall t > 0.$$



Lemma (Csiszar, 1975; Catoni, 2004)

Let (A, A) be a measurable space. For any probability μ on (A, A) and any measurable function $h: A \to \mathbb{R}$ such that $\int (\exp \circ h) \mathrm{d}\mu < \infty$,

$$\log \int (\exp \circ h) d\mu = \sup_{m \in \mathcal{M}_{\pi}(A,A)} \left\{ \int h dm - \mathcal{K}(m,\mu) \right\},\,$$

with the convention $\infty-\infty=-\infty$. Moreover, as soon as h is upper-bounded on the support of μ , the supremum with respect to m on the right-hand side is reached for the Gibbs distribution g given by

$$\frac{\mathrm{d}g}{\mathrm{d}\mu}(a) = \frac{\exp \circ h(a)}{\int (\exp \circ h) \mathrm{d}\mu}, \quad a \in A.$$



The PAC-Bayesian theory



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- Shawe-Taylor and Williamson (1997). A PAC analysis of a Bayes estimator, COLT
- McAllester (1998). Some PAC-Bayesian theorems, COLT
- McAllester (1999). PAC-Bayesian model averaging, COLT
- Catoni (2004). Statistical Learning Theory and Stochastic Optimization, Springer
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Machine Learning

Alquier and Lounici (2011). PAC-Bayesian theorems for sparse regression estimation with exponential weights,

Flectronic Journal of Statistics



A flexible and powerful framework



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Numerous models addressed by the PAC-Bayesian literature.

- Alquier and Wintenberger (2012). Model selection for weakly dependent time series forecasting, Bernoulli
- Seldin, Laviolette, Cesa-Bianchi, Shawe-Taylor and Auer (2012). PAC-Bayesian inequalities for martingales,
 IEEE Transactions on Information Theory
- Alquier and Biau (2013). Sparse Single-Index Model, Journal of Machine Learning Research
- Guedj and Alquier (2013). PAC-Bayesian Estimation and Prediction in Sparse Additive Models, Electronic Journal of Statistics
- Guedj and Robbiano (2017). PAC-Bayesian High Dimensional Bipartite Ranking, Journal of Statistical Planning
 and Inference
- Alquier and Guedj (2017). An Oracle Inequality for Quasi-Bayesian Non-Negative Matrix Factorization, Mathematical Methods of Statistics
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- Li, Guedj and Loustau (2018). A Quasi-Bayesian Perspective to Online Clustering, arXiv preprint

Towards (almost) no assumptions to derive powerful results

- 🛢 Bégin, Germain, Laviolette and Roy (2016). PAC-Bayesian bounds based on the Rényi divergence, AISTATS
- Alquier and Guedj (2017). Simpler PAC-Bayesian bounds for hostile data, Machine Learning



Sampling



Monte Carlo integration

Objective: approximation of an integral

$$\mathcal{J} = \int h(x)f(x)\mathrm{d}x.$$

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$$\mathcal{J} = \int h(x)f(x)\mathrm{d}x.$$

Key idea: exploit the fact that $\mathcal{J} = \mathbb{E}_{X \sim f}[h(X)]$.





Sample a sequence $x_1, \ldots, x_m \sim f$.



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Then use

$$\hat{\mathcal{J}}_m = \frac{1}{m} \sum_{i=1}^m h(x_i)$$

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Justification: by the Strong Law of Large Numbers,

$$\hat{\mathcal{J}}_m \to \mathcal{J}$$
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Approximation evaluation



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Estimate the variance with

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Approximation evaluation

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$$\nu_m = \frac{1}{m} \frac{1}{m-1} \sum_{i=1}^m (h(x_i) - \hat{\mathcal{J}}_m)^2,$$

and recall that for m large,

$$\frac{\hat{\mathcal{J}}_m - \mathbb{E}[h(X)]}{\sqrt{\nu_m}} \approx \mathcal{N}(0, 1).$$



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$$\mathbb{E}_{X \sim f}[h(X)] = \int h(x)f(x)dx = \int h(x)\frac{f(x)}{g(x)}g(x)dx$$
$$= \mathbb{E}_{X \sim g}\left[h(X)\frac{f(X)}{g(X)}\right],$$

which allows us to use other distributions.





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2. Instrumental distribution g may be chosen among distributions easy to simulate.

 The same sample generated from g can be used repeatedly, not only for different functions h but also for different densities f.





The optimal choice is

$$g^{\star}(x) = \frac{|h(x)|f(x)}{\int |h(x)|f(x)dx},$$

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The instrumental function may be π (the prior). But often inefficient if data informative, and impossible is π is improper...

Sampling random variables

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In practice, this theorem has a very limited scope since the pseudo-inverse F^- is usually unknown/not analytically tractable.



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The algorithm:

- 1. Sample $z \sim g$ and $u \sim \mathcal{U}(0, Mg(z))$
- 2. If $u \le f(z)$, take x = z, otherwise go back to 1.



How should we choose g?

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Nice fact: no need to know the normalizing constant of f!

