Estimating the evolution of scoring rates in soccer match: Model selection and variable selection with ???-LASSO

Mladen Laudanovic and Frank Wood

MODEL 0: $y_{id} \mid X_{id}, \beta \sim \text{Pois}(\mu \alpha_{id})$ MODEL 1: $y_{id} \mid X_{id}, \beta \sim \text{Pois}(\lambda_{id})$ MODEL 2: $y_{id} \mid X_{id}, \beta \sim \text{Pois}(\alpha_{i}\lambda_{id})$

where α_i are market intensities and $\lambda_{id} = g(\beta^T X_{id})$ and also

MODEL 3':
$$y_{id} \mid X_{id}, \beta \sim \text{Pois}(\alpha_i \lambda_{id})$$

where α_i are unknown intensities. In this case likelihood is

$$\prod_{i} \prod_{d} \alpha_{i}^{y_{id}} \lambda_{id}^{y_{id}} e^{-\alpha_{i}\lambda_{id}} / y_{id}! = \prod_{i} \alpha_{i}^{\sum_{d} y_{id}} \prod_{d} \lambda_{id}^{y_{id}} e^{-\alpha_{i}\lambda_{id}} / y_{id}!$$

so $n_i = \sum_d y_{id}$ is sufficient statistic for α_i , with the distribution $n_i \sim \text{Pois}(\alpha_i \sum_d \lambda_{id})$. Likelihood for n_i is

$$\prod_{i} \alpha_{i}^{n_{i}} \left(\sum_{d} \lambda_{id} \right)^{n_{i}} e^{-\alpha_{i} \sum_{d} \lambda_{id}} / n_{i}!$$

hence after conditioning we get:

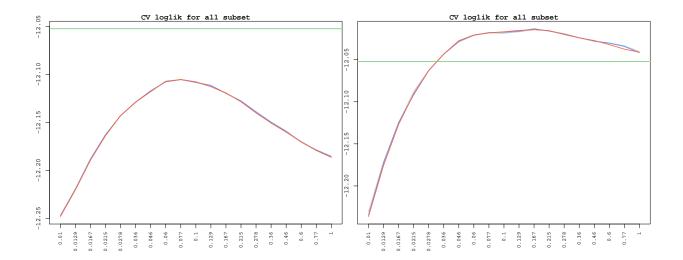
MODEL 3:
$$y_{i.} \mid \beta, n_i, X_{i.} \sim M\left(n_i, \left\{\frac{\lambda_{id}}{\sum_{d^*} \lambda_{id^*}}\right\}_{d=1}^D\right)$$

- -d are minutes in a match (1-44, 46-89 for both teams; D = 176 per match)
- -i are matches (I=302)
- -g is a link function. There are three of them: $e^x, \log(1+e^x), \log(\frac{1}{1+e^{-x}})$

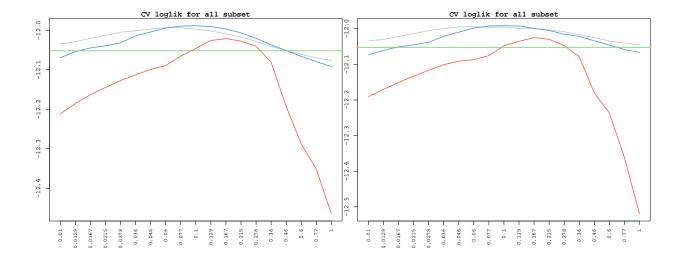
Three models with three different link functions were tested. There were two subsets of variables, each one having 100 of them (selected by "quality", one without other with market variables)

-K number of variables (K=100)

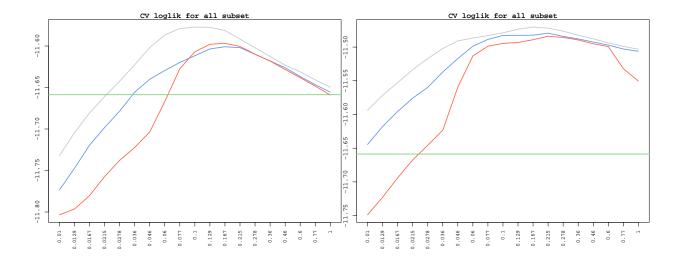
MODEL1: $y_{id} \mid X_{id}, \beta \sim \text{Pois}(\lambda_{id})$



MODEL2: $y_{id} \mid X_{id}, \beta \sim \text{Pois}(\alpha_i \lambda_{id})$



MODEL3:
$$y_{i.} \mid \beta, n_{i}, X_{i.} \sim M\left(n_{i}, \left\{\frac{\lambda_{id}}{\sum_{d^{\star}} \lambda_{id^{\star}}}\right\}_{d=1}^{D}\right)$$



INTRODUCTION: BASIC MODEL

$$y_{id} \mid X_{id}, \beta \sim \text{Pois}\left(g(\beta^T X_{id})\right)$$

$$P(\beta \mid X, Data) \propto P(\beta) \prod_{i} \prod_{d} g(\beta^{T} X_{id})^{y_{id}} e^{-g(\beta^{T} X_{id})} / y_{id}!$$

- -d are minutes in a match (1-44, 46-89 for both teams; D = 176 per match)
- -i are matches (I=302)
- -g is a link function $(\exp(x))$ is default
- K number of variables (length of beta)

$$\ell(\beta) = \log(P(\beta)) + \sum_{l=(i,d)} \left[y_l \log \left(g(\beta^T X_l) \right) - g(\beta^T X_l) \right] + \text{const}$$

is it convex with respect to beta and lambda???

In coordinatewise gradient descent algorithm in each step $k \in 1, ... K$ we are maximizing ℓ as a function of one parameter $z = \beta_k$. It is fine for convex functions.

So goal is to minimize

$$h(z) = \sum_{l} \left[g(r_l + (z - \beta_k)x_{lk}) - y_l \log \left(g(r_l + (z - \beta_k)x_{lk}) \right) \right] + z^2/2\tau$$

where $r_l = \beta^T X_l$ and τ is regularization parameter.

One good way to find the maximum of ℓ [David's paper, at al.] is that in each step we find maximum of 2nd degree Tailor polynomial of h, so suggested shift for β_k will be $\Delta u_k = -h'(\beta_k)/h''(\beta_k)$ where

$$h'(\beta_k) = \sum_{l} \left[g'(r_l) x_{lk} - y_l x_{lk} g'(r_l) / g(r_l) \right] + \beta_k / \tau$$
$$h''(\beta_k) = \sum_{l} \left[g''(r_l) x_{lk}^2 - y_l x_{lk}^2 \frac{g''(r_l) g(r_l) - g'(r_l)^2}{g(r_l)^2} \right] + 1 / \tau$$

SHAWN'S MODEL

Pros and cons

$$y_{id} \mid X_{id}, \beta \sim \text{Pois}\left(\alpha_i g(\beta^T X_{id})\right)$$

$$P(\beta \mid X, \text{Data}) \propto P(\beta) \prod_{i} \prod_{d} (\alpha_{i} g(\beta^{T} X_{id}))^{y_{id}} e^{-\alpha_{i} g(\beta^{T} X_{id})} / y_{id}!$$

Since $n_i = \sum_d y_{id}$ is sufficient statistic for α_i we should maximize

$$P(\beta \mid X, \mathrm{Data}, n) = P(\beta) \prod_{i} \prod_{d} \left(\frac{g(\beta^T X_{id})}{\sum_{d^{\star}} g(\beta^T X_{id^{\star}})} \right)^{y_{id}}$$

$$\ell(\beta) = \log(P(\beta)) + \sum_{i} \sum_{d} y_{id} \log g(\beta^{T} X_{id}) - \sum_{i} n_{i} \log \sum_{d^{\star}} g(\beta^{T} X_{id^{\star}}) + \text{const}$$

$$h(z) = \sum_{i} n_{i} \log \sum_{d} g(r_{id} + (z - \beta_{k})x_{idk}) - \sum_{i,d} y_{id} \log g(r_{id} + (z - \beta_{k})x_{idk}) + z^{2}/2\tau$$

$$h'(\beta_k) = \sum_{i} n_i \frac{\sum_{d} g'(r_{id}) x_{idk}}{\sum_{d} g(r_{id})} - \sum_{i,d} y_{id} \frac{g'(r_{id}) x_{idk}}{g(r_{id})} + \beta_k / \tau$$

$$h''(\beta_k) = \sum_{i} n_i \frac{\left(\sum_{d} g''(r_{id}) x_{idk}^2\right) \left(\sum_{d} g(r_{id})\right) - \left(\sum_{d} g'(r_{id}) x_{idk}\right)^2}{\left(\sum_{d} g(r_{id})\right)^2} - \sum_{i,d} y_{id} x_{idk}^2 \frac{g''(r_{id}) g(r_{id}) - g'(r_{id})^2}{g(r_{id})^2} + 1/\tau$$