

EEE 551 Information Theory (Spring 2022)

Chapter 2: Entropy, Relative Entropy, and Mutual Information

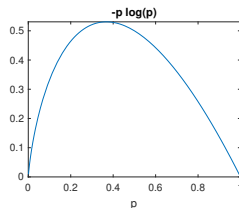
Quick Probability Review and Notation

- Random variable: X, Y, Z, \dots
- Sample value of a random variable: x, y, z, \dots
- Alphabet of a random variable: $\mathcal{X}, \mathcal{Y}, \mathcal{Z}, \dots$
- The alphabet does **not** need to consist of only numbers; e.g. $\mathcal{X} = \{a, b, c\}$
- Probability mass function (PMF): $p(x) = p_X(x) = \Pr\{X = x\}$
- Joint PMF: $p(x, y) = p_{X,Y}(x, y) = \Pr\{X = x, Y = y\}$
- Conditional PMF: $p(x|y) = p_{X|Y}(x|y) = \Pr\{X = x|Y = y\} = \frac{p(x, y)}{p(y)}$
- Variables X and Y are independent iff $p(x, y) = p(x)p(y)$,
or equivalently $p(x|y) = p(x)$
- Expectation: $\mathbb{E}[f(X)] = \sum_{x \in \mathcal{X}} p(x)f(x)$

Entropy

- A measure of the “information” or “uncertainty” in a random variable
- Entropy of a discrete random variable X with PMF $p(x)$:

$$\begin{aligned} H(X) &= - \sum_{x \in \mathcal{X}} p(x) \log p(x) \\ &= \sum_{x \in \mathcal{X}} p(x) \log \frac{1}{p(x)} \\ &= \mathbb{E} \left[\log \frac{1}{p(X)} \right] \end{aligned}$$



- $\log \frac{1}{p(x)}$ measures the “surprisingness” of observing $X = x$, so entropy is the “expected surprisingness” of X
- log is typically base 2: entropy is measured in “bits”¹
- If natural log (denoted \ln), then entropy is measured in “nats”
- By convention, $0 \log 0 = 0$
- We sometimes write

$$H(p_1, p_2, \dots, p_n) = -p_1 \log p_1 - p_2 \log p_2 - \dots - p_n \log p_n$$

i.e. the entropy of the random variable with distribution (p_1, p_2, \dots, p_n)

¹Fun fact! Shannon's 1948 paper was one of the first uses of the term “bit” for “binary digit”.

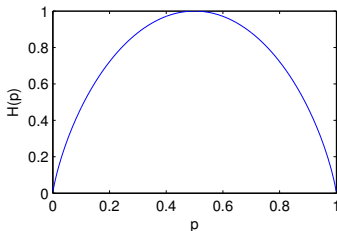
Example 1: Bernoulli random variable

- Let $X = \begin{cases} 1, & \text{with probability } p \\ 0, & \text{with probability } 1 - p \end{cases}$

May equivalently write $X \sim \text{Bern}(p)$ (Bernoulli distribution)

$$H(X) = -p \log p - (1 - p) \log(1 - p)$$

This quantity can be written $H(p, 1 - p)$ or just $H(p)$ (binary entropy function)



- If $p = 0$ or $p = 1$, then $H(X) = 0$ (source is deterministic)
- If $p = 1/2$, then $H(X)$ is maximum (1 bit), since “uncertainty” is largest

Positivity of Entropy

Entropy is non-negative, i.e. $H(X) \geq 0$

Proof:

- Since $0 \leq p(x) \leq 1$, $\log \frac{1}{p(x)} \geq 0$.
- Thus $H(X) = \mathbb{E} \left[\log \frac{1}{p(X)} \right] \geq 0$.

Example 2: Uniform random variable

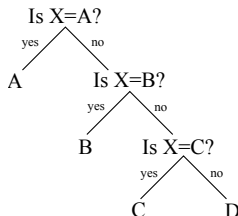
- Alphabet size $|\mathcal{X}| = m$ (could be $\{1, 2, \dots, m\}$, or any other set with m elements)
- $p(x) = \frac{1}{m}$ for all $x \in \mathcal{X}$
- $H(X) = - \sum_{x \in \mathcal{X}} \frac{1}{m} \log \frac{1}{m} = \sum_{x \in \mathcal{X}} \frac{1}{m} \log m = m \cdot \frac{1}{m} \log m = \log m.$
- For example, if $m = 32$, $H(X) = \log 32 = 5$ bits

Example 3

Let $X = \begin{cases} A & \text{with probability } 1/2 \\ B & \text{with probability } 1/4 \\ C & \text{with probability } 1/8 \\ D & \text{with probability } 1/8 \end{cases}$

$$\begin{aligned} H(X) &= \frac{1}{2} \log 2 + \frac{1}{4} \log 4 + \frac{1}{8} \log 8 + \frac{1}{8} \log 8 \\ &= \frac{1}{2}(1) + \frac{1}{4}(2) + \frac{1}{8}(3) + \frac{1}{8}(3) \\ &= 1.75 \text{ bits} \end{aligned}$$

- Entropy is minimum average number of yes/no questions to determine X
- This is exactly equal for this example;
in general it is **approximately** equal²



²We'll get back to this

- Given random variables X, Y with joint PMF $p(x, y)$, the joint entropy is

$$\begin{aligned} H(X, Y) &= - \sum_{x \in \mathcal{X}, y \in \mathcal{Y}} p(x, y) \log p(x, y) \\ &= \sum_{x \in \mathcal{X}, y \in \mathcal{Y}} p(x, y) \log \frac{1}{p(x, y)} \\ &= \mathbb{E} \left[\log \frac{1}{p(X, Y)} \right] \end{aligned}$$

- Similarly define joint entropy for more random variables: e.g. $H(X, Y, Z)$, $H(X_1, X_2, \dots, X_n)$

Joint Entropy of Independent Random Variables

- If X and Y are independent, then $H(X, Y) = H(X) + H(Y)$

Proof:

$$\begin{aligned} H(X, Y) &= - \sum_{x,y} p(x, y) \log p(x, y) \\ &= - \sum_{x,y} p(x)p(y) \log [p(x)p(y)] \\ &= - \sum_{x,y} p(x)p(y) (\log[p(x)] + \log[p(y)]) \\ &= - \sum_{x,y} p(x)p(y) \log p(x) - \sum_{x,y} p(x)p(y) \log p(y) \\ &= - \sum_x p(x) \log p(x) - \sum_y p(y) \log p(y) \\ &= H(X) + H(Y) \end{aligned}$$

- Similarly, if X_1, X_2, \dots, X_n are mutually independent, then

$$H(X_1, X_2, \dots, X_n) = \sum_{i=1}^n H(X_i)$$

Conditional Entropy

- Given random variables X, Y with joint PMF $p(x, y)$, conditional PMF $p(y|x) = \frac{p(x, y)}{p(x)}$, the entropy of Y conditioned on the event that $X = x$ is

$$H(Y|X = x) = - \sum_{y \in \mathcal{Y}} p(y|x) \log p(y|x).$$

- The conditional entropy of Y given X is the above quantity averaged over X :

$$\begin{aligned} H(Y|X) &= \sum_{x \in \mathcal{X}} p(x) H(Y|X = x) \\ &= - \sum_{x \in \mathcal{X}} p(x) \sum_{y \in \mathcal{Y}} p(y|x) \log p(y|x) \\ &= - \sum_{x, y} p(x, y) \log p(y|x) \\ &= \mathbb{E} \left[\log \frac{1}{p(Y|X)} \right]. \end{aligned}$$

- $H(Y|X) \geq 0$.

Chain Rule

- $H(X, Y) = H(X) + H(Y|X)$
- i.e. the uncertainty of X and Y is equal to the uncertainty of X plus the uncertainty of Y given X

■ **Proof:**

$$\begin{aligned}H(X, Y) &= - \sum_{x,y} p(x, y) \log p(x, y) \\&= - \sum_{x,y} p(x, y) \log [p(x) \cdot p(y|x)] \\&= - \sum_{x,y} p(x, y) \log p(x) - \sum_{x,y} p(x, y) \log p(y|x) \\&= - \sum_x p(x) \log p(x) - \sum_{x,y} p(x, y) \log p(y|x) \\&= H(X) + H(Y|X)\end{aligned}$$

- Consequence: $H(Y|X) = H(X, Y) - H(X)$

Further Forms of the Chain Rule

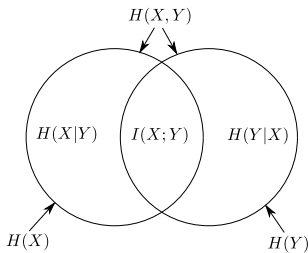
- $H(X, Y) = H(Y) + H(X|Y)$
- $H(X, Y|Z) = H(X|Z) + H(Y|X, Z)$
- $H(X_1, X_2, \dots, X_n|Z) = H(X_1|Z) + H(X_2|X_1, Z) + \dots$
 $\quad\quad\quad + H(X_n|X_1, \dots, X_{n-1}, Z)$
 $\quad\quad\quad = \sum_{i=1}^n H(X_i|X_1, \dots, X_{i-1}, Z)$

Mutual Information

- Given variables X, Y , **mutual information** is the amount of information in X about Y and vice versa:

$$\begin{aligned} I(X; Y) &= \sum_{x \in \mathcal{X}, y \in \mathcal{Y}} p(x, y) \log \frac{p(x, y)}{p(x) p(y)} \\ &= \mathbb{E} \left[\log \frac{p(X, Y)}{p(X) p(Y)} \right] \\ &= H(X) + H(Y) - H(X, Y) \\ &= H(X) - H(X|Y) \\ &= H(Y) - H(Y|X) \end{aligned}$$

- $I(X; Y) = I(Y; X)$
- Can also include multiple variables, e.g. $I(X; Y, Z)$, $I(X, Y; Z_1, \dots, Z_n)$
- Venn diagram representation:



Example 1

$Y \setminus X$	$p(x, y)$		
	1	2	3
1	1/4	1/4	0
2	1/4	0	1/4

$$\blacksquare p(x) = [1/2, 1/4, 1/4] \implies H(X) = \frac{1}{2} \log 2 + \frac{1}{4} \log 4 + \frac{1}{4} \log 4 = \frac{3}{2}$$

$$\blacksquare p(y) = [1/2, 1/2] \implies H(Y) = 1$$

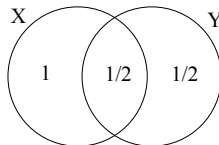
$$\blacksquare H(X, Y) = 4 \cdot \frac{1}{4} \log 4 = 2$$

$$\blacksquare H(Y|X) = H(X, Y) - H(X) = 2 - \frac{3}{2} = \frac{1}{2}$$

$$\text{Alternatively: } H(Y|X) = \frac{1}{2} H(Y|X=1) + \frac{1}{4} H(Y|X=2) + \frac{1}{4} H(Y|X=3) = \frac{1}{2}$$

$$\blacksquare H(X|Y) = H(X, Y) - H(Y) = 2 - 1 = 1$$

$$\blacksquare I(X; Y) = H(Y) - H(Y|X) = 1 - \frac{1}{2} = \frac{1}{2}$$

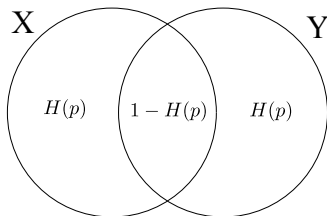


Example 2

- $\mathcal{X}, \mathcal{Y} = \{0, 1\}$, $X \sim \text{Bern}(1/2)$, $p(y|x) = \begin{cases} 1-p & y = x \\ p & y \neq x \end{cases}$
- $H(X) = 1$, $H(Y) = 1$

$$\begin{aligned} H(Y|X) &= \frac{1}{2}H(Y|X=0) + \frac{1}{2}H(Y|X=1) \\ &= \frac{1}{2}H(p) + \frac{1}{2}H(p) \\ &= H(p) \end{aligned}$$

- $H(X, Y) = H(X) + H(Y|X) = 1 + H(p)$
- $I(X; Y) = H(Y) - H(Y|X) = 1 - H(p)$
- $H(X|Y) = H(X) - I(X; Y) = H(p)$



Conditional Mutual Information

$$\begin{aligned} I(X; Y|Z) &= \sum_{x,y,z} p(x,y,z) \log \frac{p(x,y|z)}{p(x|z)p(y|z)} \\ &= \mathbb{E} \left[\log \frac{p(X,Y|Z)}{p(X|Z)p(Y|Z)} \right] \\ &= H(X|Z) - H(X|Y,Z) \\ &= \sum_z p(z) I(X; Y|Z=z) \end{aligned}$$

Chain rule for mutual information

$$\blacksquare I(X; Y, Z) = I(X; Y) + I(X; Z|Y)$$

Proof:

$$\begin{aligned} I(X; Y, Z) &= H(Y, Z) - H(Y, Z|X) \\ &= [H(Y) + H(Z|Y)] - [H(Y|X) + H(Z|X, Y)] \\ &= [H(Y) - H(Y|X)] + [H(Z|Y) - H(Z|X, Y)] \\ &= I(X; Y) + I(X; Z|Y) \end{aligned}$$

■ In general:

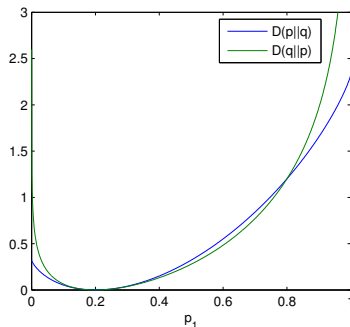
$$\begin{aligned} I(X; Y_1, Y_2, \dots, Y_n) &= I(X; Y_1) + I(X; Y_2|Y_1) + \dots + I(X; Y_n|Y_1, \dots, Y_{n-1}) \\ &= \sum_{i=1}^n I(X; Y_i|Y_1, \dots, Y_{i-1}) \end{aligned}$$

Relative Entropy (or Kullback-Leibler Divergence)

between two distributions $p(x)$ and $q(x)$ on the same alphabet \mathcal{X} :

$$D(p||q) = \sum_{x \in \mathcal{X}} p(x) \log \frac{p(x)}{q(x)} = \mathbb{E}_p \left[\log \frac{p(X)}{q(X)} \right]$$

- $D(p||q)$ is roughly a distance between two distributions, but $D(p||q) \neq D(q||p)$, and it does not satisfy the triangle inequality
- $D(p||q) \geq 0$ with equality iff $p = q$ (we'll prove later)
- **Example:** $p = [p_1 \ 1 - p_1]$, $q = [0.2 \ 0.8]$

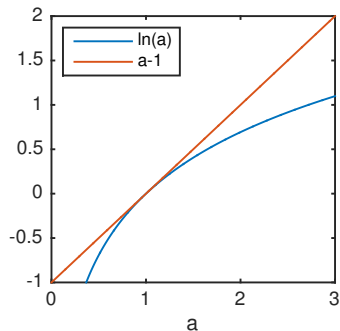


- $I(X; Y) = D(p(x, y) || p(x) p(y))$

A Simple Inequality

For $a > 0$, $\ln a \leq a - 1$, with equality iff $a = 1$.

Proof by picture:



Non-negativity of Relative Entropy

For any $p(x)$, $q(x)$, $D(p||q) \geq 0$ with equality iff $p = q$

Proof:

$$\begin{aligned} D(p||q) &= \sum_{x \in \mathcal{X}} p(x) \log \frac{p(x)}{q(x)} \\ &= (\log e) \sum_{x \in \mathcal{X}} p(x) \ln \frac{p(x)}{q(x)} \\ &= -(\log e) \sum_{x \in \mathcal{X}} p(x) \ln \frac{q(x)}{p(x)} \\ &\geq -(\log e) \sum_{x \in \mathcal{X}} p(x) \left(\frac{q(x)}{p(x)} - 1 \right) \\ &= -(\log e) \sum_{x \in \mathcal{X}} [q(x) - p(x)] \\ &= 0. \end{aligned}$$

Equality iff $\frac{q(x)}{p(x)} = 1$ for all $x \in \mathcal{X}$, i.e. $p = q$.

Non-negativity of Mutual Information

- $I(X; Y) \geq 0$, with equality iff X and Y are independent

Proof:

$$I(X; Y) = \sum_{x,y} p(x, y) \log \frac{p(x, y)}{p(x) p(y)} = D(p(x, y) \| p(x) p(y)) \geq 0$$

Equality iff $p(x, y) = p(x) p(y)$ (i.e. X and Y are independent)

- $I(X; Y|Z) \geq 0$, with equality iff X and Y are independent given Z
i.e. $p(x, y|z) = p(x|z)p(y|z)$

Proof:

$$\begin{aligned} I(X; Y|Z) &= \sum_{x,y,z} p(x, y, z) \log \frac{p(x, y|z)}{p(x|z) p(y|z)} \\ &= \sum_z p(z) \sum_{x,y} p(x, y|z) \log \frac{p(x, y|z)}{p(x|z) p(y|z)} \\ &= \sum_z p(z) D(p(x, y|z) \| p(x|z) p(y|z)) \geq 0 \end{aligned}$$

Additional Properties of Entropy & Mutual Information

- $H(Y|X) \leq H(Y)$ (i.e. conditioning reduces entropy)

Proof: $H(Y) - H(Y|X) = I(X; Y) \geq 0$

Equality iff X and Y are independent

- $H(X_1, X_2, \dots, X_n) \leq \sum_{i=1}^n H(X_i)$

Proof:

$$\begin{aligned} H(X_1, X_2, \dots, X_n) &= \sum_{i=1}^n H(X_i | X_1, \dots, X_{i-1}) \\ &\leq \sum_{i=1}^n H(X_i) \end{aligned}$$

Equality iff X_1, X_2, \dots, X_n all independent

- $H(X) \geq 0$, with equality iff X is deterministic
- $H(Y|X) \geq 0$, with equality iff $Y = g(X)$ for some function g
- $I(X; X) = H(X)$
- $H(X|X) = 0$

Uniform bound

$H(X) \leq \log |\mathcal{X}|$, where $|\mathcal{X}|$ is the number of elements in alphabet \mathcal{X}

Proof:

Let $p(x)$ be the PMF of X , and $q(x) = \frac{1}{|\mathcal{X}|}$ for all $x \in \mathcal{X}$

$$\begin{aligned} 0 &\leq D(p\|q) \\ &= \sum_x p(x) \log \frac{p(x)}{q(x)} \\ &= \sum_x p(x) \log \frac{p(x)}{1/|\mathcal{X}|} \\ &= \sum_x p(x) \log p(x) + \sum_x p(x) \log |\mathcal{X}| \\ &= -H(X) + \log |\mathcal{X}| \end{aligned}$$

Equality iff $p = q$, i.e. X is uniformly distributed on \mathcal{X}

Conditioning DOES NOT Reduce Mutual Information

Example: Let $X \sim \text{Bern}(1/2)$ and $Y \sim \text{Bern}(1/2)$ be independent

Let $Z = X \oplus Y$

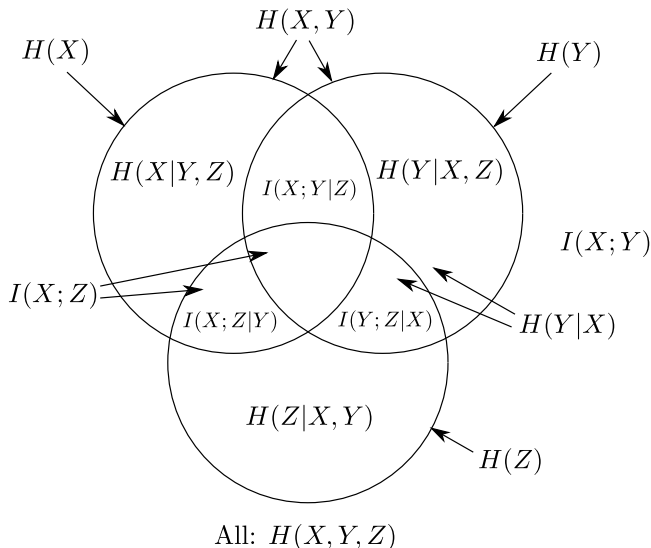
x	y	z	$p(x, y, z)$
0	0	0	1/4
0	0	1	0
0	1	0	0
0	1	1	1/4
1	0	0	0
1	0	1	1/4
1	1	0	1/4
1	1	1	0

$$I(X; Y) = 0$$

$$I(X; Y|Z) = H(X|Z) - H(X|Y, Z) = 1 - 0 = 1$$

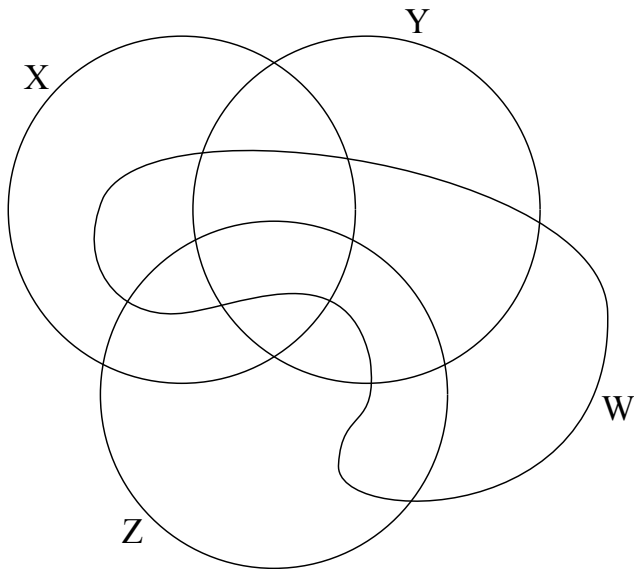
In this case, conditioning **increases** mutual information!

Venn Diagram Representation for 3 Variables



Need to be careful about the center section: it is given by $I(X; Y) - I(X; Y|Z)$ (sometimes written $I(X; Y; Z)$) which can be positive or negative

Venn Diagram Representation for 4 Variables



- Random variables X, Y, Z form a **Markov chain** denoted $X \rightarrow Y \rightarrow Z$ if X and Z are conditionally independent given Y , i.e.

$$p(x, z|y) = p(x|y) p(z|y)$$

- $X \rightarrow Y \rightarrow Z$ iff $p(z|x, y) = p(z|y)$:

$$p(z|x, y) = \frac{p(x, y, z)}{p(x, y)} = \frac{p(y) p(x|y) p(z|y)}{p(x, y)} = \frac{p(x, y) p(z|y)}{p(x, y)} = p(z|y)$$

- $X \rightarrow Y \rightarrow Z$ iff $I(X; Z|Y) = 0$
- $X \rightarrow Y \rightarrow Z$ iff $Z \rightarrow Y \rightarrow X$ (sometimes written $X \leftrightarrow Y \leftrightarrow Z$)
- If $Z = f(Y)$, then $X \rightarrow Y \rightarrow Z$
- $X_1 \rightarrow X_2 \rightarrow X_3 \rightarrow \dots \rightarrow X_n$ if

$$p(x_1, \dots, x_n) = p(x_1) p(x_2|x_1) p(x_3|x_2) \dots p(x_{n-1}|x_{n-2}) p(x_n|x_{n-1})$$

- $X \rightarrow Y \rightarrow Z \rightarrow W$ implies $X \rightarrow Y \rightarrow Z$ and $X \rightarrow Y \rightarrow W$

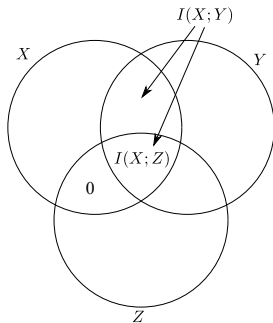
Data Processing Inequality

If $X \rightarrow Y \rightarrow Z$ is a Markov chain, then $I(X; Z) \leq I(X; Y)$
(i.e. shared information cannot **increase** by processing)

Proof: Assuming $X \rightarrow Y \rightarrow Z$,

$$\begin{aligned} I(X; Y) &= I(X; Y, Z) - I(X; Z|Y) \\ &= I(X; Y, Z) \\ &= I(X; Z) + I(X; Y|Z) \\ &\geq I(X; Z) \end{aligned}$$

Proof by Venn diagram:

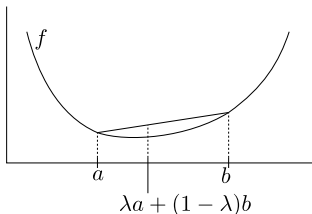


Special case: $I(X; g(Y)) \leq I(X; Y)$ for any function g , since we may take $Z = g(Y)$

Convex and Concave Functions

- A real-valued function $f(a)$ with $a \in \mathbb{R}^n$ is **convex** if for all $0 \leq \lambda \leq 1$ and a, b

$$f(\lambda a + (1 - \lambda)b) \leq \lambda f(a) + (1 - \lambda)f(b)$$



If inequality is strict for all $0 < \lambda < 1$, then f is **strictly convex**

- A function $f(a)$ is **concave** if for all $0 \leq \lambda \leq 1$ and a, b

$$f(\lambda a + (1 - \lambda)b) \geq \lambda f(a) + (1 - \lambda)f(b)$$

- f is convex iff $-f$ is concave
- For scalar a , $f(a)$ is convex iff $f''(a) \geq 0$ for all a , and strictly convex iff $f''(a) > 0$ for all a

Convexity Properties of Entropy and Mutual Information

- $H(X)$ is a concave function of $p(x)$

Proof:

Let $f(p) = -p \log p$. f is strictly concave for $p \geq 0$:

$$f'(p) = -\log p - \frac{p \log e}{p} = -\log p - \log e$$

$$f''(p) = -\frac{\log e}{p} < 0.$$

Thus $H(X) = \sum_x f(p(x))$ is a concave function of the vector $(p(x), x \in \mathcal{X})$

- Since $p(x, y) = p(x)p(y|x)$, think of $I(X; Y)$ as a function of $p(x)$ and $p(y|x)$
 - For a fixed $p(y|x)$, $I(X; Y)$ is a concave function of $p(x)$
 - For a fixed $p(x)$, $I(X; Y)$ is a convex function of $p(y|x)$
(proofs in Cover-Thomas)
- $D(p||q)$ is convex in the pair (p, q) (proof in Cover-Thomas)



Robert Fano (1917–2016)

- Given $X \rightarrow Y \rightarrow \hat{X}$, where \hat{X} is an estimate of X using Y
- Let $P_e = \Pr\{X \neq \hat{X}\}$
- Then

$$H(X|Y) \leq H(P_e) + P_e \log(|\mathcal{X}| - 1)$$

Consequences

- If $P_e = 0$, then $H(X|Y) = 0$ (i.e. X is a function of Y)
- $H(X|Y) \leq 1 + P_e \log |\mathcal{X}|$ (weaker form of Fano's inequality that we will often use)

Proof:

$$\text{Let } E = \begin{cases} 0, & \text{if } X = \hat{X} \\ 1, & \text{if } X \neq \hat{X} \end{cases}$$

$$\begin{aligned} H(X|Y) &= H(X) - I(X; Y) \\ &\leq H(X) - I(X; \hat{X}) \\ &= H(X|\hat{X}) \\ &= H(X, E|\hat{X}) - H(E|X, \hat{X}) \\ &= H(X, E|\hat{X}) \\ &= H(E|\hat{X}) + H(X|E, \hat{X}) \\ &\leq H(E) + H(X|E, \hat{X}) \\ &= H(P_e) + \Pr\{E = 0\}H(X|\hat{X}, E = 0) + \Pr\{E = 1\}H(X|\hat{X}, E = 1) \\ &= H(P_e) + P_e H(X|\hat{X}, E = 1) \\ &\leq H(P_e) + P_e \log(|\mathcal{X}| - 1) \end{aligned}$$