

# ADVANCED ECONOMETRICS I

## Practice Exercises (1/2)

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# Using Stata

## How to run commands:

- Use the menu
  - Useful when the command name is not known
- Write the command line
  - Standard approach
- Run a *do* file (contains several command lines)
  - Ideal for running a series of commands
  - Allows implementation of methods not available in Stata

# Stata – Main Windows

The screenshot shows the Stata/SE 12.1 - [Results] window. The interface includes a menu bar (File, Edit, Data, Graphics, Statistics, User, Window, Help) and a toolbar. The main window is divided into several panes:

- Review** (left): A pane for reviewing previous commands. It contains a table with columns for command number and command text. The text "There are no items to show." is displayed.
- Results** (center): A large pane for displaying the results of the current command. It shows the Stata logo, version 12.1, copyright information (1985-2011 StataCorp LP), address (4905 Lakeway Drive, College Station, Texas 77845 USA), contact information (800-STATA-PC, 979-696-4600, http://www.stata.com, stata@stata.co), and a note about the maximum number of variables (5000).
- Variables** (top right): A pane for displaying the variables in the current dataset. It contains a table with columns for variable name and label. The text "There are no items to show." is displayed.
- Input** (bottom): A command window for entering Stata commands. It contains a text area with the prompt "Command" and a cursor.

Four red ovals are overlaid on the screenshot, highlighting the following components:

- Previous commands**: Points to the Review pane.
- Results**: Points to the main Results pane.
- Variables**: Points to the Variables pane.
- Input**: Points to the Command window.

# Stata – Main Menu

## File

- Open – open dataset
- Save – save dataset
- Do – run a *do* file
- Log – register everything that you type and any output that appears in the Results window in a *smcl* file (only accessible in Stata) ou *log* (plain text that can be read by any text processor)
  - Begin – opens a log file
  - Close – closes the log file
  - Suspend – suspends logging
  - Resume – resumes logging

# Stata – Main Menu (cont.)

## File (cont.)

- Import – reads non-Stata data (ASCII, Excel, etc.)
- Export – export Stata data to other formats

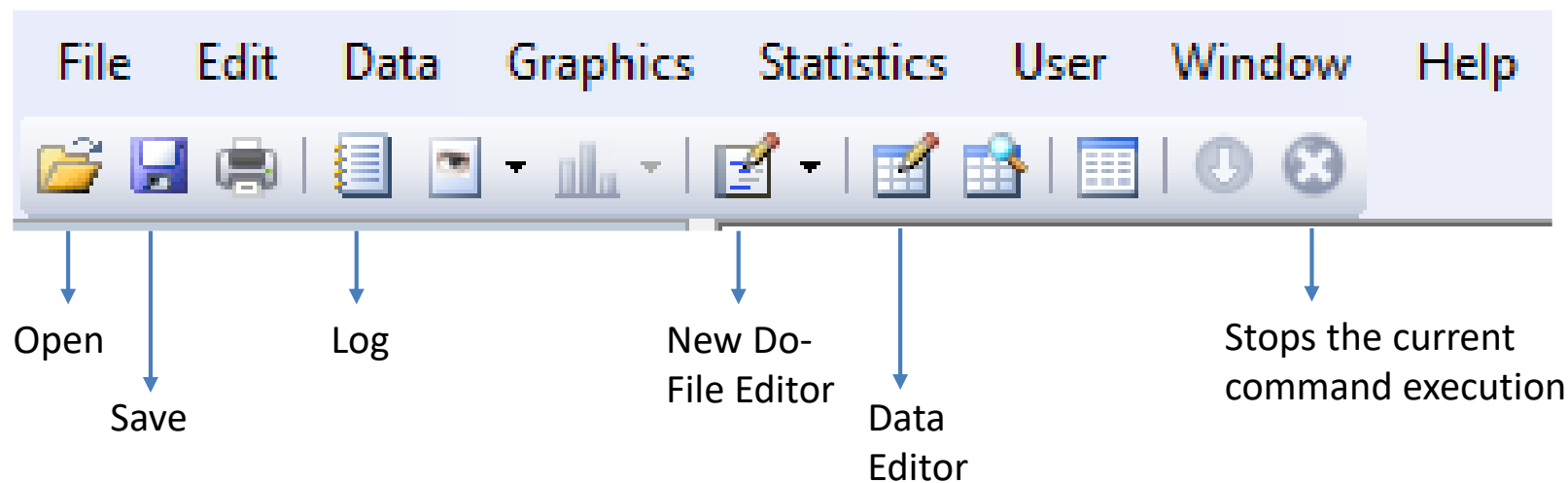
## Window

- Data Editor – brings up a spreadsheet-style editor for entering new data and editing existing data
- Do-file Editor
  - New Do-File Editor – starts a new text editor of *do* files

## Help

- Search – search Stata documentation and other resources
- Stata command – displays help information about the specified command

# Stata – Main Shortcuts



# Stata – Some Commands

## **describe** [*varlist*]

- Produces a summary of the dataset

## **summarize** [*varlist*] [, detail]

- Displays a variety of univariate summary statistics

## **generate** *newvar* = *formula*

- Creates a new variable

## **replace** *varname* = *formula* [if *expression*] [in #/#]

- Replaces contents of an existing variable

## **drop** [*varlist*] [if *expression*] [in #/#]

- Drops variables or observations

# Stata – Some Other Commands

**keep** [*varlist*] [*if expression*] [*in #/#*]

- Keeps variables or observations

**display** *scalar\_expression*

- Displays strings and values of scalar expressions

**scalar** *scalar\_name* = *expression*

- Defines a scalar variable

**return** *list*

- Returns results for general commands, stored in *r()*

**ereturn** *list*

- Returns results for estimation commands, stored in *e()*



# Stata – Others

Missing data:

- .

Conditions:

- Equality: ==
- Different: !=
- Or: |
- And: &

Estimates:

- `predict [, options]`  $\rightarrow \hat{Y}, \hat{u}, X\hat{\beta}, \dots$

# Illustration 1 – Question 1

```
. describe
```

```
    obs:      32,226
   vars:        18
  size:    1,740,204
```

```
-----
      storage  display  value
variable name  type    format  label    variable label
-----
id             int      %8.0g      Firm id
YEAR           int      %8.0g      Year
LEV_ST1        float    %9.0g      Short-term debt / (STD+LTD+Equity)
LEV_LT1        float    %9.0g      Long-term debt / (STD+LTD+Equity)
LEV1           float    %9.0g      Total debt / (STD+LTD+Equity)
COLLAT1        float    %9.0g      Tangible assets / Total assets
COLLAT2        float    %9.0g      (Tangible assets + Stock) / Total assets
SIZE1          float    %9.0g      Log(Total assets)
SIZE2          float    %9.0g      Log(Sales)
PROF1          float    %9.0g      EBIT / Total assets
PROF2          float    %9.0g      EBITDA / Total assets
GROWTH1        float    %9.0g      Sales growth rate
GROWTH2        float    %9.0g      Total assets growth rate
AGE            int      %8.0g      YEAR - Foundation year
LE             byte     %8.0g      =1 if large firm
MicE           byte     %8.0g      =1 if micro firm
SE            byte     %8.0g      =1 if small firm
MedE          byte     %8.0g      =1 if medium firm
-----
```

# Illustration 1 – Question 2

```
. summarize
```

Variable	Obs	Mean	Std. Dev.	Min	Max
-----+-----					
id	32226	2714.09	1597.178	1	5514
YEAR	32226	1998.01	1.989032	1995	2001
LEV_ST1	32226	.1583845	.2185988	0	.9997105
LEV_LT1	32226	.0768912	.1678684	0	.9982489
LEV1	32226	.2352757	.2589438	0	.9997105
-----+-----					
COLLAT1	32226	.3175925	.22441	0	.9982307
COLLAT2	32226	.4127954	.2339423	0	1.081665
SIZE1	32226	13.66189	1.922785	4.976734	22.38044
SIZE2	32226	13.76888	1.851772	5.594711	22.44638
PROF1	32226	.0671496	.1138881	-.822547	10.11724
-----+-----					
PROF2	32226	.1477722	.15208	.0000158	12.47599
GROWTH1	32226	27.98725	559.0894	-99.89224	73055.81
GROWTH2	32226	14.69437	68.35662	-94.02072	7308.115
AGE	32226	20.25119	15.6708	1	212
LE	32226	.0596413	.2368247	0	1
-----+-----					
MicE	32226	.3063986	.4610044	0	1
SE	32226	.4156892	.4928481	0	1
MedE	32226	.218271	.4130787	0	1

# Illustration 1 – Question 3

```
. gen GROUPS="1-Micro" if MicE==1
. replace GROUPS="2-Small" if SE==1
. replace GROUPS="3-Medium" if MedE==1
. replace GROUPS="4-Large" if LE==1

. tabulate GROUPS
```

GROUPS	Freq.	Percent	Cum.
-----+-----			
1-Micro	9,874	30.64	30.64
2-Small	13,396	41.57	72.21
3-Medium	7,034	21.83	94.04
4-Large	1,922	5.96	100.00
-----+-----			
Total	32,226	100.00	

# Illustration 1 – Question 4

```
. tabulate GROUPS YEAR
```

GROUPS	Ano				Total
	1995	1996	1997	1998	
1-Micro	1,418	1,441	1,414	1,431	9,874
2-Small	1,845	1,893	1,907	1,946	13,396
3-Medium	967	968	1,006	1,030	7,034
4-Large	279	267	271	277	1,922
Total	4,509	4,569	4,598	4,684	32,226

GROUPS	Ano			Total
	1999	2000	2001	
1-Micro	1,446	1,392	1,332	9,874
2-Small	1,951	1,950	1,904	13,396
3-Medium	1,024	1,019	1,020	7,034
4-Large	271	283	274	1,922
Total	4,692	4,644	4,530	32,226

# Illustration 1 – Question 5

```
. table GROUPS, contents(mean LEV_ST1 mean LEV_LT1 mean  
LEV1)
```

```
-----  
GROUPS      | mean(LEV_ST1)  mean(LEV_LT1)    mean(LEV1)  
-----+-----  
    1-Micro |      .1288624      .0417675      .1706299  
    2-Small |      .1721035      .0750359      .2471394  
    3-Medium |      .1752302      .1147087      .289939  
    4-Large |      .1527809      .1318627      .2846436  
-----
```

# Illustration 1 – Question 6.1

```
. keep if YEAR==1999
```

```
(27534 observations deleted)
```

```
. regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE
```

Source	SS	df	MS	Number of obs =	4692
Model	7.30759869	8	.913449837	F( 8, 4683) =	30.72
Residual	139.231012	4683	.029731158	Prob > F =	0.0000
Total	146.53861	4691	.031238246	R-squared =	0.0499
				Adj R-squared =	0.0482
				Root MSE =	.17243

LEV_LT1	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE2	.0181417	.0024457	7.42	0.000	.0133469	.0229364
COLLAT2	.0218571	.0111495	1.96	0.050	-1.11e-06	.0437153
PROF1	-.1934835	.0280368	-6.90	0.000	-.2484489	-.1385182
GROWTH2	.0001606	.0000709	2.26	0.024	.0000216	.0002997
AGE	-.0001155	.0001712	-0.67	0.500	-.000451	.0002201
SE	-.0035225	.0073927	-0.48	0.634	-.0180157	.0109707
MedE	.01049	.0112029	0.94	0.349	-.0114729	.0324528
LE	-.0213549	.0177432	-1.20	0.229	-.0561399	.0134301
_cons	-.163565	.0306424	-5.34	0.000	-.2236385	-.1034915

# Illustration 1 – Question 6.2

Explanatory variables which are significant at the 5% level:

- *SIZE2*, *COLLAT2*, *PROF1* and *GROWTH2*

Partial effects:

- When sales increase 1%, the proportion of long-term debt in the firm's capital structure increases 0.018 p.p.
- When the proportion of collateral in total assets increases 1 percentage point (pp), the proportion of long-term debt in the firm's capital structure increases 0.022 pp
- When profitability increases 1 pp, the proportion of long-term debt in the firm's capital structure decreases 0.193 pp
- When the growth rate of total assets increases 1 pp, the proportion of long-term debt in the firm's capital structure increases 0.0002 pp
- Age and size-based group dummies do not influence the proportion of long-term debt in the firm's capital structure



# Illustration 1 – Question 6.2 (cont.)

## Theoretical arguments:

- Positive effects:
  - *SIZE2*: larger firms tend to be more diversified, so their probability of bankruptcy is relatively smaller and it is easier for them to raise debt
  - *COLLAT2*: firms with a greater percentage of their total assets composed of tangible assets have a higher capacity for raising debt since, in case of liquidation, these assets keep their value
  - *GROWTH2*: Firms with more investment opportunities borrow more since their probability of outrunning internally generated funds is larger
- Negative effects:
  - *PROF1*: The more profitable the firm, the greater the availability of internal capital, the less the need for external funds

# Illustration 1 – Question 7.1

Model:

$LEV\_LT1$

$$= \beta_0 + \beta_1 SIZE2 + \beta_2 COLLAT2 + \beta_3 PROF1 + \beta_4 GROWTH2 + \beta_5 AGE + \beta_6 (SE * PROF1) + \beta_7 (MedE * PROF1) + \beta_8 (LE * PROF1) + v$$

Effects of firm profitability:

- Micro firms:  $\beta_3$
- Small firms:  $\beta_3 + \beta_6$
- Medium firms:  $\beta_3 + \beta_7$
- Large firms:  $\beta_3 + \beta_8$

Null hypotheses:

	Small	Medium	Large
Micro	$\beta_6 = 0$	$\beta_7 = 0$	$\beta_8 = 0$
Small		$\beta_6 = \beta_7$	$\beta_6 = \beta_8$
Medium			$\beta_7 = \beta_8$

# Illustration 1 – Question 7.1 (cont.)

```
. gen PROF_SE=PROF1*SE
. gen PROF_MedE=PROF1*MedE
. gen PROF_LE=PROF1*LE
. regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE PROF_SE PROF_MedE PROF_LE
```

Source	SS	df	MS	Number of obs =	4692
-----+-----				F( 8, 4683) =	32.28
Model	7.65730862	8	.957163578	Prob > F =	0.0000
Residual	138.881302	4683	.029656481	R-squared =	0.0523
-----+-----				Adj R-squared =	0.0506
Total	146.53861	4691	.031238246	Root MSE =	.17221

LEV_LT1	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
SIZE2	.0214744	.0016573	12.96	0.000	.0182253	.0247236
COLLAT2	.0246972	.0109637	2.25	0.024	.0032032	.0461912
PROF1	-.0956589	.0365451	-2.62	0.009	-.1673044	-.0240134
GROWTH2	.0001575	.0000708	2.22	0.026	.0000187	.0002963
AGE	-.0001012	.0001687	-0.60	0.548	-.0004319	.0002294
PROF_SE	-.1701758	.0528889	-3.22	0.001	-.273863	-.0664886
PROF_MedE	-.2279809	.0747108	-3.05	0.002	-.3744492	-.0815125
PROF_LE	-.4960758	.1385872	-3.58	0.000	-.767772	-.2243797
_cons	-.2091211	.0228067	-9.17	0.000	-.2538331	-.1644091
-----+-----						

# Illustration 1 – Question 7.1 (cont.)

## Effects of firm profitability:

- When profitability increases 1 percentage point, the proportion of long-term debt in the firm's capital structure decreases, on average:
  - 0.096 p.p. (micro firms)
  - 0.266 p.p. (small firms) →  $\text{display\_b[PROF1]} + \text{b[PROF\_SE]}$
  - 0.324 p.p. (medium firms) →  $\text{display\_b[PROF1]} + \text{b[PROF\_MedE]}$
  - 0.592 p.p. (large firms) →  $\text{display\_b[PROF1]} + \text{b[PROF\_LE]}$

# Illustration 1 – Question 7.1 (cont.)

Profitability effects differ significantly across groups?

```
. test PROF_SE=PROF_MedE
```

```
( 1)  PROF_SE - PROF_MedE = 0
```

```
F( 1, 4683) = 0.62  
Prob > F = 0.4295
```

```
. test PROF_SE=PROF_LE
```

```
( 1)  PROF_SE - PROF_LE = 0
```

```
F( 1, 4683) = 5.80  
Prob > F = 0.0161
```

```
. test PROF_MedE=PROF_LE
```

```
( 1)  PROF_MedE - PROF_LE = 0
```

```
F( 1, 4683) = 3.80  
Prob > F = 0.0514
```

(p-values)	Small	Medium	Large
Micro	0.001***	0.002***	0.000***
Small		0.430	0.016**
Medium			0.051*

# Illustration 1 – Question 7.2

Model:

$$\begin{aligned} &LEV\_LT1 \\ &= \beta_0 + \beta_1 SIZE2 + \beta_2 COLLAT2 + \beta_3 PROF1 + \beta_4 GROWTH2 + \beta_5 AGE \\ &+ \beta_6 LE + \beta_7 (LE * SIZE2) + \beta_8 (LE * COLLAT2) + \beta_9 (LE * PROF1) \\ &+ \beta_{10} (LE * GROWTH2) + \beta_{11} (LE * AGE) + w \end{aligned}$$

Null hypotheses:

- $H_0: \beta_6 = \beta_7 = \beta_8 = \beta_9 = \beta_{10} = \beta_{11} = 0$  (no structural break)

```
. gen SIZE_LE=SIZE2*LE
. gen COLLAT_LE=COLLAT2*LE
. gen GROWTH_LE=GROWTH2*LE
. gen AGE_LE=AGE*LE
```

# Illustration 1 – Question 7.2 (cont.)

```
. regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE LE SIZE_LE COLLAT_LE PROF_LE GROWTH_LE AGE_LE
```

Source	SS	df	MS	Number of obs	=	4692
Model	7.57031284	11	.688210258	F( 11, 4680)	=	23.18
Residual	138.968298	4680	.029694081	Prob > F	=	0.0000
Total	146.53861	4691	.031238246	R-squared	=	0.0517
				Adj R-squared	=	0.0494
				Root MSE	=	.17232

LEV_LT1	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE2	.0201386	.0016462	12.23	0.000	.0169112	.0233659
COLLAT2	.0183107	.01121	1.63	0.102	-.0036663	.0402876
PROF1	-.1875618	.0284758	-6.59	0.000	-.2433877	-.1317359
GROWTH2	.0001402	.0000714	1.96	0.050	2.21e-07	.0002801
AGE	-.0000991	.0001836	-0.54	0.589	-.0004591	.0002609
LE	.1111638	.1700823	0.65	0.513	-.2222776	.4446052
SIZE_LE	-.0112264	.0097198	-1.16	0.248	-.0302819	.007829
COLLAT_LE	.1028742	.0535543	1.92	0.055	-.0021174	.2078658
PROF_LE	-.1782237	.1625401	-1.10	0.273	-.4968789	.1404315
GROWTH_LE	.0012962	.0005855	2.21	0.027	.0001484	.0024441
AGE_LE	.000248	.0004729	0.52	0.600	-.0006792	.0011752
_cons	-.188839	.0224668	-8.41	0.000	-.2328845	-.1447935

# Illustration 1 – Question 7.2 (cont.)

```
. test LE SIZE_LE COLLAT_LE PROF_LE GROWTH_LE AGE_LE
```

```
( 1)  LE = 0  
( 2)  SIZE_LE = 0  
( 3)  COLLAT_LE = 0  
( 4)  PROF_LE = 0  
( 5)  GROWTH_LE = 0  
( 6)  AGE_LE = 0
```

```
F( 6, 4680) = 2.93  
Prob > F = 0.0075
```

The null hypothesis is rejected, which implies that is better to estimate separate models for each size-based group of firms or, equivalently, the model of the previous page, since:

(see the next page)



# Illustration 1 – Question 7.2 (cont.)

```
. regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE if LE==0  
(...)
```

LEV_LT1		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE2		.0201386	.0016489	12.21	0.000	.0169059	.0233713
COLLAT2		.0183107	.0112285	1.63	0.103	-.0037029	.0403242
PROF1		-.1875618	.0285227	-6.58	0.000	-.2434806	-.131643
GROWTH2		.0001402	.0000715	1.96	0.050	-1.16e-08	.0002803
AGE		-.0000991	.0001839	-0.54	0.590	-.0004597	.0002615
_cons		-.188839	.0225038	-8.39	0.000	-.2329578	-.1447202

```
. regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE if LE==1  
(...)
```

LEV_LT1		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
SIZE2		.0089121	.0093123	0.96	0.339	-.0094235	.0272478
COLLAT2		.1211849	.0509079	2.38	0.018	.0209496	.2214202
PROF1		-.3657855	.1555647	-2.35	0.019	-.6720856	-.0594854
GROWTH2		.0014364	.0005649	2.54	0.012	.0003241	.0025487
AGE		.0001489	.0004237	0.35	0.726	-.0006853	.0009831
_cons		-.0776752	.1638915	-0.47	0.636	-.4003704	.24502

# Illustration 1 – Questions 8.1-8.2

```
. quietly regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE
```

```
. estat ovtest
```

Ramsey RESET test using powers of the fitted values of LEV\_LT1

Ho: model has no omitted variables

$F(3, 4680) = 5.74$

Prob > F = 0.0006 → Unsuitable model functional form

```
. estat hettest, rhs fstat
```

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance

Variables: SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE

$F(8, 4683) = 9.27$

Prob > F = 0.0000 → Heteroskedastic model

# Illustration 1 – Questions 8.1-8.2 (cont.)

## RESET test – manual implementation

```
. predict XB
. gen XB2=XB^2
. gen XB3=XB^3
. gen XB4=XB^4
. quietly regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE XB2 XB3 XB4

. test XB2 XB3 XB4

( 1)  XB2 = 0
( 2)  XB3 = 0
( 3)  XB4 = 0

F( 3, 4680) = 5.74
Prob > F = 0.0006
```

# Illustration 1 – Questions 8.1-8.2 (cont.)

## BP test – manual implementation

```
. quietly regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE

. predict uhat, resid

. gen uhat2=uhat^2

. quietly regress uhat2 SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE

. test SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE

( 1)  SIZE2 = 0
( 2)  COLLAT2 = 0
( 3)  PROF1 = 0
( 4)  GROWTH2 = 0
( 5)  AGE = 0
( 6)  SE = 0
( 7)  MedE = 0
( 8)  LE = 0

F( 8, 4683) = 9.27
Prob > F = 0.0000
```

# Illustration 1 – Question 8.3

```
. regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE, vce(robust)
```

Linear regression

```
Number of obs      =      4,692
F(8, 4683)         =      42.14
Prob > F           =      0.0000
R-squared          =      0.0499
Root MSE          =      .17243
```

-----							
		Robust					
LEV_LT1		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----							
SIZE2		.0181417	.0024326	7.46	0.000	.0133726	.0229108
COLLAT2		.0218571	.011979	1.82	0.068	-.0016274	.0453416
PROF1		-.1934835	.0229661	-8.42	0.000	-.2385079	-.1484592
GROWTH2		.0001606	.0000728	2.21	0.027	.000018	.0003033
AGE		-.0001155	.000165	-0.70	0.484	-.000439	.0002081
SE		-.0035225	.0079668	-0.44	0.658	-.0191412	.0120962
MedE		.01049	.0121353	0.86	0.387	-.0133008	.0342808
LE		-.0213549	.0182307	-1.17	0.242	-.0570957	.0143859
_cons		-.163565	.0278539	-5.87	0.000	-.2181718	-.1089582
-----							

# Illustration 1 – Question 8.3 (cont.)

```
. regress LEV_LT1 SIZE2 COLLAT2 PROF1 GROWTH2 AGE SE MedE LE, vce(bootstrap, reps(500)
seed(111))
```

(...)

Linear regression

```
Number of obs      =      4,692
Replications       =         500
Wald chi2(8)       =     366.03
Prob > chi2        =     0.0000
R-squared          =     0.0499
Adj R-squared      =     0.0482
Root MSE          =     0.1724
```

	Observed	Bootstrap	Normal-based			
LEV_LT1	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
SIZE2	.0181417	.0023948	7.58	0.000	.013448	.0228354
COLLAT2	.0218571	.0114746	1.90	0.057	-.0006327	.0443469
PROF1	-.1934835	.0228414	-8.47	0.000	-.2382519	-.1487152
GROWTH2	.0001606	.0000725	2.21	0.027	.0000185	.0003028
AGE	-.0001155	.0001729	-0.67	0.504	-.0004543	.0002234
SE	-.0035225	.0080167	-0.44	0.660	-.019235	.0121899
MedE	.01049	.0119458	0.88	0.380	-.0129234	.0339034
LE	-.0213549	.0185598	-1.15	0.250	-.0577314	.0150216
_cons	-.163565	.0278628	-5.87	0.000	-.2181752	-.1089548

# Illustration 2

Model:

$$\begin{aligned} & \text{Log}(\text{Wage}) \\ &= \beta_0 + \beta_1 \text{Schooling} + \beta_2 \text{Exper} + \beta_3 \text{Exper}^2 + \beta_4 \text{SMSA} \\ &+ \beta_5 \text{South} + u \end{aligned}$$

Aim:

- Estimating  $\beta_1$  (returns to schooling)

Sample:

- 3010 North-American men aged between 24-34 years old

Details:

- Verbeek (2008), pp. 146-150

# Illustration 2

Possible endogenous explanatory variable:

- *Schooling*: there are unobserved characteristics of working individuals (ability, motivation, intelligence) that simultaneously influence their wages and years of schooling

Possible instrumental variables:

- *NearCollege* – may be relevant for attending a college or not (affects the costs of schooling); does not have a direct influence on wages
- *DadCollege, MomCollege* – children's schooling choices are often highly correlated with the schooling characteristics of their parents; does not have a direct influence on wages



# Illustration 2 – Question 1

```
. describe
```

```
      obs:      3,010
    vars:         9
    size:    108,360
```

```
-----
              storage  display  value
variable name  type    format   label    variable label
-----
NearCollege   float    %9.0g           =1 if lived near a college in 1966
Schooling     float    %9.0g           Years of schooling in 1976
DadCollege    float    %9.0g           Father's years of schooling
MomCollege    float    %9.0g           Mother's years of schooling
LWage         float    %9.0g           Log(wage) in 1976
SMSA          float    %9.0g           =1 if lives in a metropolitan area in 1976
South         float    %9.0g           =1 if lives in the south in 1976
Exper         float    %9.0g           Years of experience in 1976
Exper2        float    %9.0g           Experience squared in 1976
-----
```

# Illustration 2 – Question 2

```
. summarize
```

Variable	Obs	Mean	Std. Dev.	Min	Max
-----+-----					
NearCollege	3010	.6820598	.4657535	0	1
Schooling	3010	13.26346	2.676913	1	18
DadCollege	3010	9.988904	3.266501	0	18
MomCollege	3010	10.33663	2.98751	0	18
LWage	3010	6.261832	.4437976	4.60517	7.784889
-----+-----					
SMSA	3010	.7129568	.4524571	0	1
South	3010	.4036545	.4907113	0	1
Exper	3010	8.856146	4.141672	0	23
Exper2	3010	95.57907	84.61831	0	529

# Illustration 2 – Question 3.1

```
. regress LWage Schooling Exper Exper2 SMSA South
```

Source		SS	df	MS	Number of obs = 3010		
-----+-----					F( 5, 3004) = 214.57		
Model		155.959798	5	31.1919596	Prob > F = 0.0000		
Residual		436.681848	3004	.145366794	R-squared = 0.2632		
-----+-----					Adj R-squared = 0.2619		
Total		592.641646	3009	.196956346	Root MSE = .38127		
-----							
LWage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----							
Schooling		.0815797	.003499	23.31	0.000	.0747189	.0884405
Exper		.0838357	.0067735	12.38	0.000	.0705545	.0971169
Exper2		-.0022021	.0003238	-6.80	0.000	-.0028371	-.0015672
SMSA		.1508006	.015836	9.52	0.000	.1197501	.1818511
South		-.1751761	.0146486	-11.96	0.000	-.2038985	-.1464537
_cons		4.611015	.067895	67.91	0.000	4.477889	4.74414

# Illustration 2 – Question 3.2

- Returns to schooling: 8,2%

- Experience:

$$\Delta Exper = 1 \Rightarrow \Delta Wage = \beta_2 + 2\beta_3 Exper = 0.0838 - 0.0044Exper$$

$$0.0838 - 0.0044Exper = 0 \Rightarrow Exper = 19.0$$

- Positive, decreasing effect until 19.0 years of experience
- Negative, increasing (in absolute terms) effect after that point

```
. display _b[Exper]+2*_b[Exper2]*5
.06181454
. display _b[Exper]+2*_b[Exper2]*10
.03979339
(...)
```

<i>Exper</i>	<i>ΔWage</i>
0	8.4%
1	7.9%
5	6.2%
10	4.0%
20	-0.4%
30	-4.8%
45	-11.4%

# Illustration 2 – Question 3.2 (cont.)

- *SMSA:*

- The wages of workers living in metropolitan areas are on average 15.1% higher
- Large firms, which usually offer better salaries, are more common in metropolitan areas

- *South:*

- The wages of workers living in the South are on average 17.5% lower than wages of similar workers living in the North
- The South of USA is a less attractive and poorer region

# Illustration 2 – Question 3.3

- *Schooling, SMSA, South*:

$$H_0: \beta_j = 0, \quad j = 1, 4, 5$$

$$H_1: \beta_j \neq 0$$

$$t = \frac{\hat{\beta}_j}{\hat{\sigma}_{\hat{\beta}_j}} \sim t_{N-p}$$

Variable	<i>t</i>	<i>p-value</i>
<i>Schooling</i>	23.31	0.000
<i>SMSA</i>	9.52	0.000
<i>South</i>	-11.96	0.000

- Critical value (significance level: 5%):

```
.display invt(3010-6, 0.975)
```

```
1.960754
```

$$t_{N-p} = 1.96$$

- p-value (example for *SMSA*):

```
. display 2*ttail(3010-6, 9.52)
```

```
3.435e-21
```

- The null hypothesis is rejected in all cases, implying that the marginal effects of the three variables are relevant

# Illustration 2 – Question 3.3 (cont.)

- *Exper*:

$$H_0: \beta_2 + 2\beta_3 \overline{Exper} = 0$$

$$H_1: \beta_2 + 2\beta_3 \overline{Exper} \neq 0$$

```
. summarize Exper
```

Variable	Obs	Mean	Std. Dev.	Min	Max
-----+-----					
Exper	3010	8.856146	4.141672	0	23

$$t = \frac{\hat{\beta}_2 + 2 * 8.856146 * \hat{\beta}_3}{\hat{\sigma}_{\hat{\beta}_2 + 2 * 8.856146 * \hat{\beta}_3}} \sim t_{N-p}$$

- In Stata, it is easier to apply an  $F$  test (for a simple  $H_0$ ,  $F = t^2$ ):

```
. test Exper+2*Exper2*8.856146=0
```

```
( 1)  Exper + 17.71229*Exper2 = 0
```

```
F( 1, 3004) = 375.17
```

```
Prob > F = 0.0000
```

- The null hypothesis is rejected, implying that the marginal effect of *Exper* is relevant

# Illustration 2 – Question 4.1

Need to test:

- The exogeneity of *NearCollege*:

$$H_0: E(u|NearCollege) = 0$$

$$H_1: E(u|NearCollege) \neq 0$$

- The correlation between *NearCollege* and *Schooling*:

$$H_0: cov(NearCollege, Schooling) = 0$$

$$H_1: cov(NearCollege, Schooling) \neq 0$$



# Illustration 2 – Question 4.1 (cont.)

- Not possible to test whether *NearCollege* is endogenous, since the number of instrumental variables and endogenous explanatory variables is identical → use theoretical arguments
- Testing if  $cov(NearCollege, Schooling) = 0$ :
  - Estimate the reduced form model:  
 $Schooling = \pi_0 + \pi_1 NearCollege + \pi_2 Exper + \pi_3 Exper^2 + \pi_4 SMSA + \pi_5 South + w$
  - Test:  
 $H_0: cov(NearCollege, Schooling) = 0 \Leftrightarrow H_0: \pi_1 = 0$ 
    - The null hypothesis of no correlation between *NearCollege* and *Schooling* is rejected (see the next slide)
- Conclusion: *NearCollege* is a valid instrument for *Schooling*

# Illustration 2 – Question 4.1 (cont.)

```
. regress Schooling NearCollege Exper Exper2 SMSA South
```

Source	SS	df	MS	Number of obs	=	3010
Model	9755.15468	5	1951.03094	F( 5, 3004)	=	496.39
Residual	11806.9254	3004	3.93040126	Prob > F	=	0.0000
				R-squared	=	0.4524
				Adj R-squared	=	0.4515
Total	21562.0801	3009	7.16586243	Root MSE	=	1.9825

Schooling	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
<b>NearCollege</b>	<b>.3456458</b>	<b>.0841956</b>	<b>4.11</b>	<b>0.000</b>	<b>.180559</b>	<b>.5107326</b>
Exper	-.4258437	.0343574	-12.39	0.000	-.4932102	-.3584773
Exper2	.0009774	.0016838	0.58	0.562	-.002324	.0042789
SMSA	.3639914	.086559	4.21	0.000	.1942705	.5337122
South	-.582683	.0763963	-7.63	0.000	-.7324774	-.4328886
_cons	16.68131	.1800091	92.67	0.000	16.32836	17.03426

# Illustration 2 – Question 4.2

Alternative 1: conclude the manual implementation of 2SLS:

- Complete the first step by obtaining an estimate for *Schooling* based on the reduced form model:  

```
. predict schoolhat  
(option xb assumed; fitted values)
```
- Run the second step by estimating by OLS the structural model (with *Schooling* replaced by *schoolhat*)
- Disadvantage: standard deviations are incorrectly estimated because the variability of *schoolhat* is not taken into account

Alternative 2 (recommendable):

- Use the command `ivregress 2sls`, which automatically performs the two steps and computes correct variances

# Illustration 2 – Question 4.2 (cont.)

## Alternative 1

```
. regress LWage schoolhat Exper Exper2 SMSA South
```

Source	SS	df	MS	Number of obs	=	3010
Model	78.1556887	5	15.6311377	F( 5, 3004)	=	91.27
Residual	514.485957	3004	.171266963	Prob > F	=	0.0000
Total	592.641646	3009	.196956346	R-squared	=	0.1319
				Adj R-squared	=	0.1304
				Root MSE	=	.41384

LWage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
schoolhat	.1354201	.0508483	2.66	0.008	.0357191	.2351211
Exper	.1067727	.0228187	4.68	0.000	.0620308	.1515146
Exper2	-.0022553	.000355	-6.35	0.000	-.0029514	-.0015591
SMSA	.1249987	.0297649	4.20	0.000	.066637	.1833603
South	-.1409356	.0359542	-3.92	0.000	-.2114329	-.0704383
_cons	3.703426	.8579277	4.32	0.000	2.021241	5.385611

# Illustration 2 – Question 4.2 (cont.)

## Alternative 2

```
. ivregress 2sls LWage (Schooling = NearCollege) Exper Exper2 SMSA South
```

Instrumental variables (2SLS) regression

Number of obs = 3010  
Wald chi2(5) = 499.36  
Prob > chi2 = 0.0000  
R-squared = 0.2051  
Root MSE = .39562

LWage	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
Schooling	.1354201	.0486085	2.79	0.005	.0401491	.230691
Exper	.1067727	.0218136	4.89	0.000	.0640188	.1495266
Exper2	-.0022553	.0003394	-6.64	0.000	-.0029205	-.00159
SMSA	.1249987	.0284538	4.39	0.000	.0692302	.1807671
South	-.1409356	.0343705	-4.10	0.000	-.2083005	-.0735707
_cons	3.703426	.8201379	4.52	0.000	2.095986	5.310867

Instrumented: Schooling

Instruments: Exper Exper2 SMSA South NearCollege

# Illustration 2 – Question 4.3

- OLS:
  - Returns to schooling: 8.2%
  - Standard error: 0.0035
  - 95% confidence interval: [7.5% ; 8.8%]
- 2SLS:
  - Returns to schooling: 13.5%
  - Standard error: 0.0486
  - 95% confidence interval: [4.0% ; 23.1%]
- OLS yields more precise results, which, however, will be inconsistent in case *Schooling* is endogenous

# Illustration 2 – Question 4.4

- Standard estimation of the variance:

```
. estat endogenous
    Tests of endogeneity
    Ho: variables are exogenous
    Durbin (score) chi2(1)          = 1.33097 (p = 0.2486)
    Wu-Hausman F(1,3003)           = 1.32846 (p = 0.2492)
```

- Robust estimation of the variance:

```
. quietly ivregress 2sls LWage (Schooling = NearCollege) Exper Exper2 SMSA
South, vce(robust)
. estat endogenous
    Tests of endogeneity
    Ho: variables are exogenous
    Robust score chi2(1)            = 1.39046 (p = 0.2383)
    Robust regression F(1,3003)     = 1.38821 (p = 0.2388)
```

- In both cases the null hypothesis of *Schooling* being exogenous is not rejected
- OLS results are consistent

# Illustration 2 – Question 5 (4.2)

```
. ivregress gmm LWage (Schooling = NearCollege) Exper Exper2 SMSA South,  
wmatrix(unadjusted)
```

Instrumental variables (GMM) regression

Number of obs = 3,010

Wald chi2(5) = 499.36

Prob > chi2 = 0.0000

R-squared = 0.2051

Root MSE = .39562

GMM weight matrix: Unadjusted

Needed if you want  
to get the same  
variances as 2SLS.  
Not efficient.

LWage	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
Schooling	.1354201	.0486085	2.79	0.005	.0401491	.230691
Exper	.1067727	.0218136	4.89	0.000	.0640188	.1495266
Exper2	-.0022553	.0003394	-6.64	0.000	-.0029205	-.00159
SMSA	.1249987	.0284538	4.39	0.000	.0692302	.1807671
South	-.1409356	.0343705	-4.10	0.000	-.2083005	-.0735707
_cons	3.703426	.8201379	4.52	0.000	2.095986	5.310867

Instrumented: Schooling

Instruments: Exper Exper2 SMSA South NearCollege



# Illustration 2 – Question 5 (4.4)

```
. estat endogenous
```

```
Test of endogeneity (orthogonality conditions)
```

```
Ho: variables are exogenous
```

```
GMM C statistic chi2(1) = 1.33097 (p = 0.2486)
```

```
. quietly ivregress gmm LWage (Schooling = NearCollege) Exper Exper2 SMSA  
South
```

```
. estat endogenous
```

```
Test of endogeneity (orthogonality conditions)
```

```
Ho: variables are exogenous
```

```
GMM C statistic chi2(1) = 1.39046 (p = 0.2383)
```

# Illustration 2 – Question 6.1

```
. ivregress gmm LWage (Schooling = NearCollege DadCollege MomCollege) Exper  
Exper2 SMSA South
```

```
Instrumental variables (GMM) regression          Number of obs   =       3,010  
                                                Wald chi2(5)     =       690.01  
                                                Prob > chi2      =       0.0000  
                                                R-squared        =       0.2365  
GMM weight matrix: Robust                    Root MSE        =       .38771
```

		Robust					
LWage		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
Schooling		.118024	.0111939	10.54	0.000	.0960844	.1399637
Exper		.0994424	.0082148	12.11	0.000	.0833417	.115543
Exper2		-.0022445	.000339	-6.62	0.000	-.002909	-.0015801
SMSA		.1351776	.0165073	8.19	0.000	.102824	.1675312
South		-.1518082	.016665	-9.11	0.000	-.184471	-.1191455
_cons		3.995256	.1903674	20.99	0.000	3.622143	4.368369

Instrumented:    Schooling

Instruments:    Exper Exper2 SMSA South NearCollege DadCollege MomCollege

# Illustration 2 – Question 6.1 (cont.)

- OLS:
  - Returns to schooling: 8.2%
  - Standard error: 0.0035
  - 95% confidence interval: [7.5% ; 8.8%]
- 2SLS (1 instrument):
  - Returns to schooling: 13.5%
  - Standard error: 0.0486
  - 95% confidence interval: [4.0% ; 23.1%]
- GMM (3 instruments):
  - Returns to schooling: 11.8%
  - Standard error: 0.0112
  - 95% confidence interval: [9.6% ; 14.0%]

# Illustration 2 – Question 6.2

```
. estat endogenous
```

```
Test of endogeneity (orthogonality conditions)
```

```
Ho: variables are exogenous
```

```
GMM C statistic chi2(1) = 11.6718 (p = 0.0006)
```

- The null hypothesis of *Schooling* being exogenous is now rejected
- OLS results are inconsistent

# Illustration 2 – Question 6.3

Test for the endogeneity of the instrumental variables:

```
. estat overid
```

Test of overidentifying restriction:

Hansen's J  $\chi^2(2) = 2.42011$  ( $p = 0.2982$ )

- The hypothesis of the instrumental variables being exogenous is not rejected

# Illustration 2 – Question 6.3 (cont.)

Test for the correlation between instruments and *Schooling* (Stata automatic command):

```
. estat firststage
```

First-stage regression summary statistics

-----					
		Adjusted	Partial	Robust	
Variable		R-sq.	R-sq.	F(3,3002)	Prob > F
-----+-----					
Schooling		0.5070	0.5058	102.032	0.0000
-----					

- The null hypothesis is rejected → instruments and endogenous covariate are correlated

Overall, *NearCollege*, *DadCollege* and *MomCollege* are valid instruments

# Stata – Panel data

Commands for panel data:

- Prefix **xt**

Initial command:

- **xtset** *identvar tempvar*

Examples:

- **xtdescribe** [*varlist*]
- **xtsum** [*varlist*]
- **xttab** *varname*

# Illustration 3

Model:

$$\begin{aligned} & \text{Log}(Wage) \\ &= \beta_0 + \beta_1 \text{Schooling} + \beta_2 \text{Exper} + \beta_3 \text{Exper}^2 + \beta_4 \text{Union} \\ &+ \beta_5 \text{South} + \beta_6 \text{Public} + u \end{aligned}$$

Aim:

- Estimating  $\beta_4$  (impact of collective bargaining on wages)

Sample:

- 545 full-time working males who completed their schooling by 1980 and were observed over the period 1980-1987

Details:

- Verbeek (2008), pp. 375-377



# Illustration 3 – Question 1

```
. xtset NR Year
```

```
    panel variable:  NR (strongly balanced)
```

```
    time variable:  YEAR, 1980 to 1987
```

```
        delta:  1 unit
```

```
. xtdescribe
```

```
    NR:  13, 17, ..., 12548
```

```
    n = 545
```

```
   YEAR: 1980, 1981, ..., 1987
```

```
    T = 8
```

```
    Delta(YEAR) = 1 unit
```

```
    Span(YEAR)  = 8 periods
```

```
    (NR*YEAR uniquely identifies each observation)
```

Distribution of T <sub>i</sub> :	min	5%	25%	50%	75%	95%	max
	8	8	8	8	8	8	8

Freq.	Percent	Cum.		Pattern
-----+-----				
545	100.00	100.00		11111111
-----+-----				
545	100.00			XXXXXXXXXX

# Illustration 3 – Question 1 (cont.)

```
. describe
```

```
Contains data from F:\2-Econometric Methods I\Practice\Data\Data4.dta
```

```
obs:      4,360
vars:      9                               4 Sep 2018 08:56
size:     156,960
```

```
-----
```

variable name	storage type	display format	value label	variable label
NR	float	%9.0g		Observations number
Year	float	%9.0g		Year of observation
Exper	float	%9.0g		Years of experience
Public	float	%9.0g		=1 if working in public administration
South	float	%9.0g		=1 if lives in south
Schooling	float	%9.0g		Years of schooling
Union	float	%9.0g		=1 if wage set by collective bargaining
Exper2	float	%9.0g		Experience squared
Wage	float	%9.0g		Hourly wage

```
-----
```

```
Sorted by: NR  Year
```

# Illustration 3 – Question 2

```
. summarize
```

Variable	Obs	Mean	Std. Dev.	Min	Max
NR	4,360	5262.059	3496.15	13	12548
Year	4,360	1983.5	2.291551	1980	1987
Exper	4,360	6.514679	2.825873	0	18
Public	4,360	.0401376	.1963044	0	1
South	4,360	.3506881	.4772402	0	1
Schooling	4,360	11.76697	1.746181	3	16
Union	4,360	.2440367	.4295639	0	1
Exper2	4,360	50.42477	40.78199	0	324
Wage	4,360	5.919175	3.202225	.0279014	57.50431

# Illustration 3 – Question 2 (cont.)

```
. xtsum Wage Schooling Exper
```

Variable		Mean	Std. Dev.	Min	Max	Observations
-----+-----						
Wage	overall	5.919175	3.202225	.0279014	57.50431	N = 4360
	between		2.455819	1.503564	28.35696	n = 545
	within		2.057397	-16.42349	35.06652	T = 8
Schooling	overall	11.76697	1.746181	3	16	N = 4360
	between		1.747585	3	16	n = 545
	within		0	11.76697	11.76697	T = 8
Exper	overall	6.514679	2.825873	0	18	N = 4360
	between		1.654918	3.5	14.5	n = 545
	within		2.291551	3.014679	10.01468	T = 8

Note: for Min and Max, in the within case, Stata considers the values of  $(Y_{it} - \bar{Y}_i + \bar{Y})$ , not  $(Y_{it} - \bar{Y}_i)$ .

# Illustration 3 – Question 2 (cont.)

```
. xtsum Union South Public
```

Variable		Mean	Std. Dev.	Min	Max	Observations
-----+-----						
Union	overall	.2440367	.4295639	0	1	N = 4360
	between		.3294467	0	1	n = 545
	within		.2759787	-.6309633	1.119037	T = 8
South	overall	.3506881	.4772402	0	1	N = 4360
	between		.4644838	0	1	n = 545
	within		.1111732	-.5243119	1.225688	T = 8
Public	overall	.0401376	.1963044	0	1	N = 4360
	between		.1388064	0	1	n = 545
	within		.1389214	-.8348624	.9151376	T = 8

# Illustration 3 – Question 3

```
. xttab Union
```

Union	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	3296	75.60	511	93.76	80.63
1	1064	24.40	280	51.38	47.50
Total	4360	100.00	791	145.14	68.90

(n = 545)

- Overall, 24,4% of wages resulted from collective bargaining
- For 51,38% of workers (280 out of 545), in at least one year their wages resulted from collective bargaining
- For those 280 workers, 47,5% of the time their wages resulted from collective bargaining

# Illustration 3 – Question 3 (cont.)

```
. xttab South
```

South	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	2831	64.93	374	68.62	94.62
1	1529	35.07	212	38.90	90.15
Total	4360	100.00	586	107.52	93.00

(n = 545)

- Overall, 35,1% of observations concern cases of individuals living in the South
- 38,9% of workers (212 out of 545) lived at least one year in the South
- For those 212 workers, 90,15% of the time they lived in the South

# Illustration 3 – Question 3 (cont.)

```
. xttab Public
```

Public	Overall		Between		Within
	Freq.	Percent	Freq.	Percent	Percent
0	4185	95.99	543	99.63	96.34
1	175	4.01	62	11.38	35.28
Total	4360	100.00	605	111.01	90.08

(n = 545)

- Overall, 4,01 % of observations concern cases of individuals working in Public Administration
- 11,38% of workers (62 out of 545) worked at least one year in Public Administration
- For those 62 workers, 35,28% of the time they worked in Public Administration



# Illustration 3 – Question 4.1 - Pooled

```
. gen LWage=log(Wage)
```

```
. regress LWage Schooling Exper Exper2 Union South Public, vce(cluster NR)
```

Linear regression

Number of obs = 4360  
F( 6, 544) = 60.28  
Prob > F = 0.0000  
R-squared = 0.1704  
Root MSE = .48546

(Std. Err. adjusted for 545 clusters in NR)

-----						
		Robust				
LWage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
Schooling		.1016358	.0090348	11.25	0.000	.0838885 .1193832
Exper		.0995979	.0121631	8.19	0.000	.0757054 .1234904
Exper2		-.0031896	.0008552	-3.73	0.000	-.0048694 -.0015098
Union		.1709018	.0282839	6.04	0.000	.1153428 .2264608
South		-.0622494	.0328083	-1.90	0.058	-.1266958 .002197
Public		.0145377	.046524	0.31	0.755	-.0768511 .1059264
_cons		-.0552731	.1156887	-0.48	0.633	-.2825244 .1719783
-----						

```
. estimates store POOLED
```

# Illustration 3 – Question 4.1 - Between

```
. xtreg LWage Schooling Exper Exper2 Union South Public, be
```

Linear regression

```
Number of obs =    4360
F(   6,    544) =    60.28
Prob > F       =    0.0000
R-squared      =    0.1704
Root MSE      =    .48546
```

(Std. Err. adjusted for 545 clusters in NR)

-----						
		Robust				
LWage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
Schooling	.1016358	.0090348	11.25	0.000	.0838885	.1193832
Exper	.0995979	.0121631	8.19	0.000	.0757054	.1234904
Exper2	-.0031896	.0008552	-3.73	0.000	-.0048694	-.0015098
Union	.1709018	.0282839	6.04	0.000	.1153428	.2264608
South	-.0622494	.0328083	-1.90	0.058	-.1266958	.002197
Public	.0145377	.046524	0.31	0.755	-.0768511	.1059264
_cons	-.0552731	.1156887	-0.48	0.633	-.2825244	.1719783
-----						

```
. estimates store BE
```

# Illustration 3 – Question 4.1 – Random Effects

```
. xtreg LWage Schooling Exper Exper2 Union South Public, re vce(cluster NR)
(...)
```

(Std. Err. adjusted for 545 clusters in NR)

-----							
		Robust					
LWage		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----+-----							
Schooling		.1027451	.0086503	11.88	0.000	.0857907	.1196994
Exper		.1187223	.0104192	11.39	0.000	.0983012	.1391435
Exper2		-.0043045	.0006748	-6.38	0.000	-.005627	-.0029819
Union		.1044636	.0209805	4.98	0.000	.0633425	.1455847
South		-.0166749	.0319162	-0.52	0.601	-.0792294	.0458797
Public		.0333681	.0334151	1.00	0.318	-.0321243	.0988606
_cons		-.1372221	.1113179	-1.23	0.218	-.3554012	.080957
-----+-----							
sigma_u		.33131898					
sigma_e		.35132742					
rho		.47071506	(fraction of variance due to u_i)				
-----							


```
. estimates store RE
```

# Illustration 3 – Question 4.1 – Fixed Effects

```
. xtreg LWage Schooling Exper Exper2 Union South Public, fe vce(cluster NR)
(...)
```

(Std. Err. adjusted for 545 clusters in NR)

-----						
		Robust				
LWage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
Schooling		0	(omitted)			
Exper		.1206647	.0105609	11.43	0.000	.0999194 .1414099
Exper2		-.004429	.0006836	-6.48	0.000	-.0057717 -.0030863
Union		.0833501	.0229404	3.63	0.000	.0382875 .1284127
South		.1038881	.0657487	1.58	0.115	-.0252644 .2330406
Public		.0363333	.0373874	0.97	0.332	-.0371081 .1097748
_cons		1.028155	.0425644	24.16	0.000	.9445442 1.111765
-----+-----						
sigma_u		.41161426				
sigma_e		.35132742				
rho		.57852861	(fraction of variance due to u_i)			
-----						

Ignore

estimates store FE

Ignore

# Illustration 3 – Question 4.1 - LSDV

```
. areg LWage Schooling Exper Exper2 Union South Public, absorb(NR) vce(cluster NR)
note: Schooling omitted because of collinearity
(...)
```

(Std. Err. adjusted for 545 clusters in NR)

-----						
		Robust				
LWage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----						
Schooling	0	(omitted)				
Exper	.1206647	.0112897	10.69	0.000	.0984878	.1428415
Exper2	-.004429	.0007307	-6.06	0.000	-.0058644	-.0029936
Union	.0833501	.0245235	3.40	0.001	.0351778	.1315225
South	.1038881	.070286	1.48	0.140	-.0341772	.2419534
Public	.0363333	.0399676	0.91	0.364	-.0421763	.114843
_cons	1.028155	.0455017	22.60	0.000	.9387742	1.117535
-----						
NR	absorbed		(545 categories)			

```
. estimates store LSDV
```

# Illustration 3 – Question 4.1 – 1st Table

. estimates table POOLED BE RE FE LSDV, b se

Variable	POOLED	BE	RE	FE	LSDV
Schooling	.10163583	.09800416	.10274505	(omitted)	(omitted)
	.00903478	.01092795	.00865032		
Exper	.09959789	-.04653615	.11872235	.12066465	.12066465
	.01216314	.05089296	.01041916	.01056093	.01128974
Exper2	-.00318958	.00523794	-.00430446	-.00442898	-.00442898
	.00085515	.00325606	.00067477	.00068355	.00073072
Union	.17090181	.26016332	.10446355	.08335012	.08335012
	.02828387	.0473469	.02098054	.02294039	.0245235
South	-.06224939	-.07078294	-.01667486	.1038881	.1038881
	.03280828	.03277295	.03191616	.06574874	.07028605
Public	.01453769	-.03308648	.03336815	.03633333	.03633333
	.04652403	.11079747	.03341515	.03738745	.03996755
_cons	-.05527309	.49764241	-.13722212	1.0281548	1.0281548
	.11568873	.22187872	.1113179	.04256437	.04550173

legend: b/se

# Illustration 3 – Question 4.1 – 2nd Table

```
. estimates table POOLED BE RE FE LSDV, b star(0.1 0.05 0.01)
```

Variable	POOLED	BE	RE	FE
Schooling	.10163583***	.09800416***	.10274505***	(omitted)
Exper	.09959789***	-.04653615	.11872235***	.12066465***
Exper2	-.00318958***	.00523794	-.00430446***	-.00442898***
Union	.17090181***	.26016332***	.10446355***	.08335012***
South	-.06224939*	-.07078294**	-.01667486	.1038881
Public	.01453769	-.03308648	.03336815	.03633333
_cons	-.05527309	.49764241**	-.13722212	1.0281548***

Variable	LSDV
Schooling	(omitted)
Exper	.12066465***
Exper2	-.00442898***
Union	.08335012***
South	.1038881
Public	.03633333
_cons	1.0281548***

legend: \* p<.1; \*\* p<.05; \*\*\* p<.01

# Illustration 3 – Question 4.2

```
. quietly xtreg LWage Schooling Exper Exper2 Union South Public, re  
  
. estimates store REs  
  
. quietly xtreg LWage Schooling Exper Exper2 Union South Public, fe  
  
. estimates store FEs
```



# Illustration 3 – Question 4.2

```
. hausman FEs REs
```

---- Coefficients ----				
	(b)	(B)	(b-B)	sqrt(diag(V_b-V_B))
	FEs	REs	Difference	S.E.
Exper	.1206647	.1187223	.0019423	.0013976
Exper2	-.004429	-.0043045	-.0001245	.0001211
Union	.0833501	.1044636	-.0211134	.0073043
South	.1038881	-.0166749	.120563	.0395946
Public	.0363333	.0333681	.0029652	.0124616

b = consistent under  $H_0$  and  $H_a$ ; obtained from xtreg  
 B = inconsistent under  $H_a$ , efficient under  $H_0$ ; obtained from xtreg

Test:  $H_0$ : difference in coefficients not systematic

```
chi2(5) = (b-B)'[(V_b-V_B)^(-1)](b-B)
        = 35.47
Prob>chi2 = 0.0000
```

- Fixed effects estimator is better

# Illustration 3 – Question 4.3

```
. regress D.LWage D.Schooling D.Exper D.Exper2 D.Union D.South D.Public,
vce(cluster NR) nocons
(...)
```

(Std. Err. adjusted for 545 clusters in NR)

		Robust				
D.LWage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
Schooling						
D1.		0	(omitted)			
Exper						
D1.		.119242	.0143533	8.31	0.000	.0910474 .1474366
Exper2						
D1.		-.0040136	.0009438	-4.25	0.000	-.0058676 -.0021596
Union						
D1.		.0436369	.022141	1.97	0.049	.0001446 .0871292
South						
D1.		.0811916	.0782312	1.04	0.300	-.0724807 .2348638
Public						
D1.		.0425582	.0356367	1.19	0.233	-.0274442 .1125606

# Illustration 3 – Question 4.4

```
. xtreg LWage Schooling Exper Exper2 c.Union##i.Year South Public, re vce(cluster NR)
```

(...)

Search help for  
fvvarlist

(Std. Err. adjusted for 545 clusters in NR)

		Robust					
LWage		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
Schooling		.0922807	.0109883	8.40	0.000	.0707441	.1138174
Exper		.1080505	.0164482	6.57	0.000	.0758126	.1402885
Exper2		-.0048138	.0007849	-6.13	0.000	-.0063523	-.0032753
Union		.1734667	.0477046	3.64	0.000	.0799674	.266966
Year							
1981		.0568965	.0328633	1.73	0.083	-.0075144	.1213074
1982		.0524367	.0390681	1.34	0.180	-.0241353	.1290087
1983		.054367	.0474343	1.15	0.252	-.0386024	.1473364
1984		.0742157	.0597359	1.24	0.214	-.0428645	.191296
1985		.0860962	.0679065	1.27	0.205	-.0469982	.2191905
1986		.1385142	.0775591	1.79	0.074	-.0134989	.2905273
1987		.1885931	.0872509	2.16	0.031	.0175845	.3596018

(continues in the next slide)

# Illustration 3 – Question 4.4 (cont.)

		Robust				
LWage	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
Year#c.Union						
1981	-.0455982	.0555045	-0.82	0.411	-.154385	.0631886
1982	-.0560189	.0543852	-1.03	0.303	-.162612	.0505742
1983	-.0893266	.0555898	-1.61	0.108	-.1982806	.0196275
1984	-.0692261	.0542253	-1.28	0.202	-.1755057	.0370534
1985	-.0486523	.0527484	-0.92	0.356	-.1520373	.0547328
1986	-.1301995	.0575233	-2.26	0.024	-.2429431	-.0174558
1987	-.137896	.054057	-2.55	0.011	-.2438458	-.0319463
South	-.0169016	.0319517	-0.53	0.597	-.0795258	.0457227
Public	.0386987	.033165	1.17	0.243	-.0263034	.1037009
_cons	.0002138	.1576026	0.00	0.999	-.3086817	.3091093
sigma_u	.33219311					
sigma_e	.35091184					
rho	.47261814	(fraction of variance due to u_i)				

# Illustration 3 – Question 4.4 (cont.)

Benefit from collective bargaining:

- 1980: 17.3%
- 1987:  $17.3\% - 13.8\% = 3.5\%$

In 1987 wages defined by collective bargaining are still higher, but the differential relative to other wages decreased significantly.

# Illustration 3 – Question 5.1

```
. xtivreg LWage Schooling Exper Exper2 (Union=L(1/4).Union) South Public, fe
vce(cluster NR)
```

		Robust					
LWage	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]		
Union	.0805339	.1312502	0.61	0.539	-.1767117	.3377795	
Schooling	0	(omitted)					
Exper	.1182598	.0293922	4.02	0.000	.0606523	.1758674	
Exper2	-.0035235	.0015683	-2.25	0.025	-.0065973	-.0004498	
South	.0616019	.0860277	0.72	0.474	-.1070092	.230213	
Public	.0297124	.0381017	0.78	0.435	-.0449655	.1043903	
_cons	.9939431	.1367056	7.27	0.000	.7260049	1.261881	
sigma_u	.45429733						
sigma_e	.30701212						
rho	.68648331	(fraction of variance due to u_i)					
Instrumented:	Union						
Instruments:	Schooling Exper Exper2 South Public L.Union L2.Union L3.Union L4.Union						

# Illustration 3 – Question 5.2

```
. xtivreg LWage Schooling Exper Exper2 (Union=L(1/2).Union F(1/2).Union) South
Public, fe vce(cluster NR)
```

		Robust					
LWage	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]		
Union	.043245	.1111023	0.39	0.697	-.1745114	.2610015	
Schooling	0	(omitted)					
Exper	.1167687	.0241845	4.83	0.000	.069368	.1641694	
Exper2	-.0045448	.0016055	-2.83	0.005	-.0076915	-.0013981	
South	.1132399	.1198697	0.94	0.345	-.1217004	.3481802	
Public	.0903547	.0451475	2.00	0.045	.0018673	.1788421	
_cons	1.051376	.1040199	10.11	0.000	.8475009	1.255252	
sigma_u	.45226025						
sigma_e	.30141701						
rho	.69243462	(fraction of variance due to u_i)					
Instrumented:	Union						
Instruments:	Schooling Exper Exper2 South Public L.Union L2.Union F.Union F2.Union						

# Illustration 4 – Model

- Trade-Off theory – optimal / target proportion of debt:

$$MDR_{it}^* = x'_{i,t-1}\beta + \eta_{it}$$

- Target adjustment model:

$$MDR_{it} - MDR_{i,t-1} = (1 - \gamma)(MDR_{it}^* - MDR_{i,t-1})$$

- $0 \leq \gamma \leq 1$
- Adjustment speed:  $1 - \gamma$ 
  - $\gamma = 0 \Rightarrow$  Firms adjust immediately and completely
  - $\gamma = 1 \Rightarrow$  No adjustment

- Econometric model:

$$MDR_{it} = MDR_{i,t-1} + (1 - \gamma)(x'_{i,t-1}\beta + \eta_{it} - MDR_{i,t-1})$$
$$\vdots$$

$$MDR_{it} = \gamma MDR_{i,t-1} + x'_{i,t-1}\theta + \alpha_i + u_{it}$$



# Illustration 4 – Model (cont.)

## Aim:

- Estimate  $\gamma$
- Test whether  $H_0: \gamma = 1$  (trade-off theory not valid)

## Sample:

- 5449 firms observed over the period 1986-2001
- Unbalanced panel
- Source: Compustat Industrial Annual Tapes

## Details:

- Verbeek (2008), pp. 383-388

```
. xtset gvkey yeara
      panel variable:  gvkey (unbalanced)
      time variable:  yeara, 1986 to 2001, but with gaps
                  delta:  1 unit
```

# Illustration 4 – Question 1

```
. xtdescribe
```

```
  gvkey:  1003, 1004, ..., 233397          n =          5449
```

```
  yeara:  1986, 1987, ..., 2001          T =           16
```

```
      Delta(yeara) = 1 unit
```

```
      Span(yeara)  = 16 periods
```

```
      (gvkey*yeara uniquely identifies each observation)
```

```
Distribution of T_i:  min      5%      25%      50%      75%      95%      max
                     1        1        1        3        8        14        16
```

Freq.	Percent	Cum.	Pattern
384	7.05	7.05	.....1
176	3.23	10.28	1111111111111111
167	3.06	13.34	1.....
149	2.73	16.08	.....1.
136	2.50	18.57	11.....
132	2.42	20.99	.....1...
131	2.40	23.40	.....1..
113	2.07	25.47	111.....
98	1.80	27.27	.....1....
3963	72.73	100.00	(other patterns)
5449	100.00		XXXXXXXXXXXXXXXXXX

# Illustration 4 – Question 1 (cont.)

```
. describe
```

```
-----  
              storage   display   value  
variable name  type      format    label      variable label  
-----  
gvkey          long      %12.0g          spc permanent number  
yeara          int       %8.0g          fiscal year  
mdr            float     %9.0g          market debt ratio  
bdr            float     %9.0g          book debt ratio  
ebit_ta        float     %9.0g          earnings before interest payments and taxes,  
                divided by total assets (lag)  
mb             float     %9.0g          ratio of market value to book value of assets  
                (lag)  
dep_ta         float     %9.0g          depreciation expenses as a proportion of fixed  
                assets (lag)  
lnta           float     %9.0g          log of total assets (lag)  
fa_ta          float     %9.0g          proportion of fixed assts (lag)  
rd_dum         byte      %8.0g          =1 if research and development expenditures  
                (lag)  
rd_ta          float     %9.0g          research and development expenditures, divided  
                by total assets (lag)  
indmedian      float     %9.0g          industry median debt ratio (lag)  
rated          byte      %8.0g          =1 if firm has a public debt rating (lag)  
-----
```

# Illustration 4 – Question 2

```
. quietly regress mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, vce(cluster gvkey)
```

```
. estimates store POOLED
```

```
. quietly xtreg mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, vce(cluster gvkey)
```

```
. estimates store RE
```

```
. quietly xtreg mdr L.mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, fe vce(cluster gvkey)
```

```
. estimates store FE
```

# Illustration 4 – Question 2 (cont.)

```
. estimates table POOLED RE FE, b star
```

Variable	POOLED	RE	FE
mdr			
L1.	.8835036***	.78734111***	.53498254***
ebit_ta	-.03233775***	-.03455703***	-.05003294***
mb	.0016432*	.00069052	.00227756*
dep_ta	-.26051795***	-.30137705***	-.12395444
lnta	-.00067042	.00183328*	.03803015***
fa_ta	.02012146***	.02919991***	.05934357***
rd_dum	.00688957**	.0096552***	.00005977
rd_ta	-.12020508***	-.14520778***	-.06567621*
indmedian	.03212249**	.06029162***	.16721793***
rated	.00713406*	.01103821**	.02058981***
_cons	.05818177***	.03798629**	-.60083475***

legend: \* p<0.05; \*\* p<0.01; \*\*\* p<0.001

# Illustration 4 – Question 3.1

```
. xtivreg mdr (L.mdr=L2.mdr) ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta
indmedian rated, fd
(...)
```

D.mdr	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
mdr						
LD.	7.03303	5.494343	1.28	0.201	-3.735684	17.80174
ebit_ta						
D1.	1.207597	.9705551	1.24	0.213	-.6946564	3.10985
mb						
D1.	.244267	.1853757	1.32	0.188	-.1190627	.6075966
dep_ta						
D1.	-1.858345	1.577202	-1.18	0.239	-4.949603	1.232914
lnta						
D1.	-.5214084	.4557998	-1.14	0.253	-1.41476	.3719429
(...)						

Instrumented: L.mdr

Instruments: ebit\_ta mb dep\_ta lnta fa\_ta rd\_dum rd\_ta indmedian rated  
L2.mdr

# Illustration 4 – Question 3.2

```
. xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep vce(robust)
```

```
Arellano-Bond dynamic panel-data estimation   Number of obs       =       15039  
Group variable: gvkey                        Number of groups     =       2996  
Time variable: yeara  
  
Obs per group:   min =           1  
                  avg =    5.019693  
                  max =           14  
  
Number of instruments =       115           Wald chi2(10)       =       212.72  
                                           Prob > chi2         =       0.0000  
  
Two-step results
```

(continues in the next slide)

# Illustration 4 – Question 3.2 (cont.)

(Std. Err. adjusted for clustering on gvkey)

-----							
		WC-Robust					
mdr		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----							
mdr							
L1.		.3819695	.0731919	5.22	0.000	.238516	.525423
ebit_ta		.035684	.0173522	2.06	0.040	.0016744	.0696936
mb		.0147128	.0027217	5.41	0.000	.0093784	.0200472
dep_ta		.0648811	.109432	0.59	0.553	-.1496016	.2793639
lnta		.030107	.0083243	3.62	0.000	.0137916	.0464224
fa_ta		.0150317	.0286987	0.52	0.600	-.0412168	.0712801
rd_dum		-.0178784	.0122991	-1.45	0.146	-.0419841	.0062273
rd_ta		.001471	.035963	0.04	0.967	-.0690151	.0719571
indmedian		.0919917	.0453824	2.03	0.043	.0030439	.1809395
rated		-.0066174	.0099448	-0.67	0.506	-.0261088	.0128741
_cons		-.4235622	.1467658	-2.89	0.004	-.7112179	-.1359065
-----							

Instruments for differenced equation

GMM-type: L(2/.).mdr

Standard: D.ebit\_ta D.mb D.dep\_ta D.lnta D.fa\_ta D.rd\_dum D.rd\_ta D.indmedian  
D.rated

Instruments for level equation

Standard: \_cons



# Illustration 4 – Question 3.3

```
. xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep maxldep(2) vce(robust)
```

```
Arellano-Bond dynamic panel-data estimation   Number of obs       =       15039  
Group variable: gvkey                         Number of groups     =       2996  
Time variable: yeara  
  
Obs per group:   min =           1  
                  avg =    5.019693  
                  max =           14  
  
Number of instruments =       37              Wald chi2(10)        =       179.25  
                                          Prob > chi2          =       0.0000  
  
Two-step results
```

(continues in the next slide)

# Illustration 4 – Question 3.3 (cont.)

(Std. Err. adjusted for clustering on gvkey)

-----							
		WC-Robust					
mdr		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
-----							
mdr							
L1.		.4005973	.0886448	4.52	0.000	.2268567	.574338
ebit_ta		.0445636	.0197946	2.25	0.024	.005767	.0833603
mb		.0165177	.003106	5.32	0.000	.0104301	.0226053
dep_ta		.0832717	.1094867	0.76	0.447	-.1313182	.2978616
lnta		.0252836	.0090865	2.78	0.005	.0074744	.0430928
fa_ta		.0065411	.0306739	0.21	0.831	-.0535787	.0666608
rd_dum		-.0135049	.0122092	-1.11	0.269	-.0374345	.0104248
rd_ta		.0015315	.0361632	0.04	0.966	-.0693471	.07241
indmedian		.0670826	.0493198	1.36	0.174	-.0295825	.1637477
rated		-.0098843	.0101263	-0.98	0.329	-.0297315	.009963
_cons		-.3367156	.1575827	-2.14	0.033	-.6455719	-.0278593
-----							

Instruments for differenced equation

GMM-type: L(2/3).mdr

Standard: D.ebit\_ta D.mb D.dep\_ta D.lnta D.fa\_ta D.rd\_dum D.rd\_ta

D.indmedian D.rated

Instruments for level equation

Standard: \_cons

# Illustration 4 – Question 3.4

```
. xtdpdsys mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta indmedian rated,  
twostep vce(robust)
```

System dynamic panel-data estimation

Group variable: gvkey

Time variable: yeara

Number of obs = 19573

Number of groups = 3777

Obs per group: min = 1

avg = 5.182155

max = 15

Number of instruments = 129

Wald chi2(10) = 2185.15

Prob > chi2 = 0.0000

Two-step results

(continues in the next slide)

# Illustration 4 – Question 3.4 (cont.)

-----						
		WC-Robust				
mdr		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----						
mdr						
L1.		1.030104	.0238744	43.15	0.000	.9833112 1.076897
ebit_ta		.1285762	.018583	6.92	0.000	.0921541 .1649983
mb		.0312589	.0027109	11.53	0.000	.0259457 .0365722
dep_ta		-.1036924	.1374032	-0.75	0.450	-.3729978 .165613
lnta		-.0127849	.0068955	-1.85	0.064	-.0262999 .0007301
fa_ta		-.1085659	.0327925	-3.31	0.001	-.172838 -.0442939
rd_dum		-.0061031	.0153821	-0.40	0.692	-.0362514 .0240452
rd_ta		.0759739	.0441006	1.72	0.085	-.0104617 .1624096
indmedian		-.2372958	.0400293	-5.93	0.000	-.3157518 -.1588397
rated		-.0302876	.0120363	-2.52	0.012	-.0538783 -.006697
_cons		.273204	.1269921	2.15	0.031	.0243041 .522104
-----						

Instruments for differenced equation

GMM-type: L(2/.)mdr

Standard: D.ebit\_ta D.mb D.dep\_ta D.lnta D.fa\_ta D.rd\_dum D.rd\_ta D.indmedian  
D.rated

Instruments for level equation

GMM-type: LD.mdr

Standard: \_cons

# Illustration 4 – Question 4.1

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep vce(robust)
```

```
. estat abond, artests(3)
```

(...)

Arellano-Bond test for zero autocorrelation in first-differenced errors

```
+-----+  
|Order |   z       Prob > z|  
|-----+-----|  
|   1   | -6.206    0.0000 |  
|   2   | -3.587    0.0003 |  
|   3   | -3.3325   0.0009 |  
+-----+  
H0: no autocorrelation
```

There is autocorrelation of order higher than 1: the estimators are not consistent.

# Illustration 4 – Question 4.2

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep
```

```
. estat sargan
```

Sargan test of overidentifying restrictions

H0: overidentifying restrictions are valid

```
chi2(104)      = 436.3939
```

```
Prob > chi2    = 0.0000
```

The hypothesis of exogenous  
instruments is rejected: the estimators  
are not consistent.

# Illustration 4 – Question 4.3

```
. quietly xtabond mdr ebit_ta mb dep_ta lnta fa_ta rd_dum rd_ta  
indmedian rated, twostep vce(robust)
```

```
. test L.mdr=1
```

```
( 1)  L.mdr = 1
```

```
      chi2( 1) =    71.30  
Prob > chi2 =    0.0000
```

The hypothesis that  $\gamma = 1$  is rejected, which implies that firms seem to adjust their debt ratios: there is no evidence against the trade-off theory.