

Lecture Notes on Advanced Econometrics

Lecture 10: GLS, WLS, and FGLS

Generalized Least Square (GLS)

So far, we have been dealing with heteroskedasticity under OLS framework. But if we knew the variance-covariance matrix of the error term, then we can make a heteroskedastic model into a homoskedastic model.

As we defined before

$$E(uu') = \sigma^2 \Omega = \Sigma.$$

Define further that

$$\Omega^{-1} = P'P$$

P is a “ $n \times n$ ” matrix

Pre-multiply P on a regression model

$$Py = PX\beta + Pu$$

or

$$\tilde{y} = \tilde{X}\beta + \tilde{u}$$

In this model, the variance of \tilde{u} is

$$E(\tilde{u}\tilde{u}') = E(Puu'P') = PE(uu')P' = P\sigma^2\Omega P' = \sigma^2 P\Omega P' = \sigma^2 I$$

Note that $P\Omega P' = I$, because define $P\Omega P' = A$, then $P'P\Omega P' = P'A$. By the definition of P , $\Omega^{-1}\Omega P' = P'A$, thus $P' = P'A$. Therefore, A must be I .

Because $E(\tilde{u}\tilde{u}') = \sigma^2 I$, the model satisfies the assumption of homoskedasticity. Thus, we can estimate the model by the conventional OLS estimation.

Hence,

$$\begin{aligned}\hat{\beta} &= (\tilde{X}'\tilde{X})^{-1} \tilde{X}'\tilde{y} \\ &= (X'P'PX)^{-1} X'P'Py \\ &= (X'\Omega^{-1}X)^{-1} X'\Omega^{-1}y\end{aligned}$$

is the efficient estimator of β . This is called the **Generalized Least Square (GLS)** estimator. Note that the GLS estimators are unbiased when $E(\tilde{u} | \tilde{X}) = 0$. The variance of GLS estimator is

$$\text{var}(\hat{B}) = \sigma^2 (\tilde{X}'\tilde{X})^{-1} = \sigma^2 (X' \Omega^{-1} X)^{-1}.$$

Note that, under homoskedasticity, i.e., $\Omega^{-1} = I$, GLS becomes OLS.

The problem is, as usual, that we don't know $\sigma^2 \Omega$ or Σ . Thus we have to either assume Σ or estimate Σ empirically. An example of the former is Weighted Least Squares Estimation and an example of the later is Feasible GLS (FGLS).

Weighted Least Squares Estimation (WLS)

Consider a general case of heteroskedasticity.

$$\text{Var}(u_i) = \sigma_i^2 = \sigma^2 \omega_i.$$

Then,

$$E(uu') = \sigma^2 \begin{bmatrix} \omega_1 & 0 & 0 \\ 0 & \omega_2 & 0 \\ 0 & 0 & \omega_n \end{bmatrix} = \sigma^2 \Omega, \text{ thus } \Omega^{-1} = \begin{bmatrix} \omega_1^{-1} & 0 & 0 \\ 0 & \omega_2^{-1} & 0 \\ 0 & 0 & \omega_n^{-1} \end{bmatrix}.$$

Because of $\Omega^{-1} = P'P$, P is a $n \times n$ matrix whose i -th diagonal element is $1/\sqrt{\omega_i}$. By pre-multiplying P on y and X , we get

$$y_* = Py = \begin{bmatrix} y_1 / \sqrt{\omega_1} \\ y_2 / \sqrt{\omega_2} \\ y_n / \sqrt{\omega_n} \end{bmatrix} \quad \text{and} \quad X_* = PX = \begin{bmatrix} 1/\sqrt{\omega_1} & x_{11}/\sqrt{\omega_1} & \dots & x_{1k}/\sqrt{\omega_1} \\ 1/\sqrt{\omega_2} & x_{21}/\sqrt{\omega_2} & \dots & x_{2k}/\sqrt{\omega_2} \\ 1/\sqrt{\omega_n} & x_{n1}/\sqrt{\omega_n} & \dots & x_{nk}/\sqrt{\omega_n} \end{bmatrix}.$$

The OLS on y_* and X_* is called the Weighted Least Squares (WLS) because each variable is weighted by $\sqrt{\omega_i}$. The question is: where can we find ω_i ?

Feasible GLS (FGLS)

Instead of assuming the structure of heteroskedasticity, we may estimate the structure of heteroskedasticity from OLS. This method is called Feasible GLS (FGLS). First, we estimate $\hat{\Omega}$ from OLS, and, second, we use $\hat{\Omega}$ instead of Ω .

$$\hat{\beta}_{FGLS} = (X' \hat{\Omega}^{-1} X)^{-1} X' \hat{\Omega}^{-1} y$$

There are many ways to estimate FGLS. But one flexible approach (discussed in Wooldridge page 277) is to assume that

$$\text{var}(u | X) = u^2 = \sigma^2 \exp(\delta_0 + \delta_1 x_1 + \delta_2 x_2 + \dots + \delta_k x_k)$$

By taking log of the both sides and using \hat{u}^2 instead of u^2 , we can estimate

$$\log(\hat{u}^2) = \alpha_0 + \delta_1 x_1 + \delta_2 x_2 + \dots + \delta_k x_k + e.$$

The predicted value from this model is $\hat{g}_i = \log(\hat{u}^2)$. We then convert it by taking the exponential into $\hat{\omega}_i = \exp(\hat{g}_i) = \exp(\log(\hat{u}^2)) = \hat{u}^2$. We now use WLS with weights $1/\hat{\omega}_i$ or $1/\hat{u}^2$.

Example 1

```
. * Estimate the log-wage model by using WAGE1.dta with WLS
. * Weight is educ

. * Generate weighted variables
. gen w=1/(educ)^0.5
. gen wlogwage=logwage*w
. gen wfemale=female*w
. gen weduc=educ*w
. gen wexper=exper*w
. gen wexpsq=expsq*w

. * Estimate weighted least squares (WLS) model

. reg wlogwage weduc wfemale wexper wexpsq w, noc
```

Source	SS	df	MS	Number of obs =	524
Model	113.916451	5	22.7832901	F(5, 519) =	1660.16
Residual	7.12253755	519	.013723579	Prob > F	= 0.0000
				R-squared	= 0.9412
				Adj R-squared	= 0.9406
				Root MSE	= .11715
Total	121.038988	524	.230990435		

wlogwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]

weduc		.080147	.006435	12.455	0.000	.0675051	.0927889
wfemale		-.3503307	.0354369	-9.886	0.000	-.4199482	-.2807133
wexper		.0367367	.0045745	8.031	0.000	.0277498	.0457236
wexpsq		-.0006319	.000099	-6.385	0.000	-.0008264	-.0004375
w		.4557085	.0912787	4.992	0.000	.2763872	.6350297

End of Example 1

Example 2

```
. * Estimate reg
. reg logwage educ female exper expsq
(Output omitted)
. predict e, residual
. gen logesq=ln(e*e)

. reg logesq educ female exper expsq
(output omitted)
. predict esqhat
(option xb assumed; fitted values)
. gen omega=exp(esqhat)

. * Generate weighted variables
. gen w=1/(omega)^0.5
. gen wlogwage=logwage*w
. gen wfemale=female*w
. gen weduc=educ*w
. gen wxper=exper*w
. gen wxpsq=expsq*w

. * Estimate Feasible GLS (FGLS) model
. reg wlogwage weduc wfemale wxper wxpsq w, noc
```

Source	SS	df	MS	Number of obs = 524			
Model	31164.1981	5	6232.83962	F(5, 519) = 1569.72			
Residual	2060.77223	519	3.97065941	Prob > F = 0.0000			
Total	33224.9703	524	63.406432	R-squared = 0.9380			
				Adj R-squared = 0.9374			
				Root MSE = 1.9927			

wlogwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
weduc	.0828952	.0069779	11.880	0.000	.0691868	.0966035
wfemale	-.2914609	.0349884	-8.330	0.000	-.3601971	-.2227246
wexper	.0376525	.004497	8.373	0.000	.0288179	.0464872
wexpsq	-.0006592	.0001008	-6.540	0.000	-.0008573	-.0004612
w	.3848487	.0950576	4.049	0.000	.1981038	.5715936

End of Example 2