

Supplementary Document for BNB

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1 Proof of Theorem 6.1

1.1 Preliminaries Given a tree \mathcal{T}_k , the tree partitions the data space, mapping each data point X_t into a leaf of \mathcal{T}_k : let $\text{leaf}(X_t)$ map X_t to its leaf. Define \mathcal{X} as the set of leaves, and $d(x, y)$ as the shortest path distance (in terms of number of graph hops) between leaves x and y .

LEMMA 1.1. (\mathcal{X}, d) is a metric space.

1.2 Separation Depth

THEOREM 1.1. *The separation depth $s_{k,t}$ can be rewritten in terms of distances in this metric space:*

$$(1.1) \quad s_{k,t} = d_{lim} - \frac{1}{2} \min_{x \in L, y \in R} d(\text{leaf}(x), \text{leaf}(y)) + 1$$

Proof. Consider the pair of points $x \in L$ and $y \in R$ minimizing Eq. (6.2). Their shortest path distance is $d(\text{leaf}(x), \text{leaf}(y))$, so their deepest common ancestor (call it z) must be $d(\text{leaf}(x), \text{leaf}(y))/2$ steps above them. Since $\text{leaf}(x)$ and $\text{leaf}(y)$ are at depth d_{lim} , thus z is at depth $d_{lim} - d(\text{leaf}(x), \text{leaf}(y))/2$. Moreover, since x and y were the closest pair of points, z must be the deepest node with descendants in both L and R . Thus, L and R are fully separated exactly at depth $d_{lim} - d(\text{leaf}(x), \text{leaf}(y))/2 + 1$, but not at any lower depth (due to node z). ■

2 Proof of Theorem 6.2

THEOREM 2.1. *For any $\lambda > 0$, letting $p_\lambda = 2^{\lambda-1}$, we have¹:*

$$(2.2) \quad P(s_{k,t} \leq \lambda) \leq (1/2)^{w-p_\lambda}.$$

Proof. Fix an arbitrary ordering $<$ on the data space (e.g. lexicographic ordering), and let X^1, \dots, X^{2w} be the original data points (X_1, \dots, X_{2w}) sorted according to the order $<$. Since the data points X_1, \dots, X_{2w} are i.i.d., if we condition on X^1, \dots, X^{2w} , by symmetry, the conditional probability that the original data is mapped

to X^1, \dots, X^{2w} by any particular permutation is equal, which is $1/(2w)!$.

Define Z_i as a 0-1 random variable, taking value 0 if X^i is one of X_1, \dots, X_w , and 1 otherwise. Then exactly w of the random variables Z_1, \dots, Z_{2w} are 1, and since we earlier showed that every permutation of the X^1, \dots, X^{2w} is equally likely, thus now each assignments of 0s and 1s to Z_1, \dots, Z_{2w} with exactly w 1s is also equally likely. Thus, by symmetry, each such assignment has probability $1/\binom{2w}{w}$.

For any λ , $P(s_{k,t} \leq \lambda)$ is the probability that the sets L and R were fully separated by the time the tree reached level λ . Note that at level λ , our tree partitions the data space into $p_\lambda = 2^{\lambda-1}$ parts. The number of assignments to the Z variables for which L and R are fully separated is then at most 2^{p_λ} , since each of the p_λ parts we have to choose to assign either 0 or 1 to all variables in that part. Moreover, we earlier showed that each of the $\binom{2w}{w}$ assignments has the same probability. Thus, the probability that L and R are fully separated at level λ is at most

$$\begin{aligned} (2.3) \quad P(s_{k,t} \leq \lambda) &\leq \frac{2^{p_\lambda}}{\binom{2w}{w}} \\ (2.4) \quad &= \frac{2^{p_\lambda}(1) \dots (w)}{(w+1) \dots (2w)} \\ (2.5) \quad &\leq 2^{p_\lambda}(1/2)^w \\ (2.6) \quad &\leq (1/2)^{w-p_\lambda}. \end{aligned}$$

In fact, this can be tightened significantly by using stronger bounds for central binomial coefficients: it is known that $\binom{2w}{w} \geq 4^w/\sqrt{4w}$ for any positive integer w [1]. Substituting this instead gives

$$\begin{aligned} (2.7) \quad P(s_{k,t} \leq \lambda) &\leq \frac{2^{p_\lambda}}{\binom{2w}{w}} \\ (2.8) \quad &\leq \frac{2^{p_\lambda} \sqrt{4w}}{4^w} \\ (2.9) \quad &= 2^{p_\lambda-1-2w} \sqrt{w} \\ (2.10) \quad &= (1/2)^{2w-1-p_\lambda} \sqrt{w}. \end{aligned}$$

¹Note that some theorem statements have been amended from their previous versions. We will correct this fully in the final version of the paper. ■

3 Proof of Theorem 6.3

The next theorem concerns our final change score c_t . For any error threshold $\varepsilon > 0$, we have:

THEOREM 3.1.

(3.11)

$$P(c_t \geq d_{lim} + 1 - \log(1 + \log \frac{\varepsilon}{N} + w)) \leq N(1/2)^{w-p_\lambda}$$

Proof. Applying union bound on the (weaker) result of the previous theorem gives

$$(3.12) \quad P(s_{k,t} \leq \lambda) \leq N(1/2)^{w-p_\lambda} \quad \forall k \in \{1, \dots, N\}$$

Recalling that $c_t = d_{lim} + 1 - \frac{1}{N} \sum_{j=1}^N s_{j,t}$, substituting this gives:

$$(3.13) \quad P(c_t \geq d_{lim} + 1 - \lambda) \leq N(1/2)^{w-p_\lambda}$$

Finally, if $p_\lambda = \log \frac{\varepsilon}{N} + w$, then $N(1/2)^{w-p_\lambda} \leq \varepsilon$. Substituting this into (3.13) (and using the fact that $\lambda = \log p_\lambda + 1$) gives the result. ■

References

- [1] N. D. Kazarinoff. *Analytic inequalities*. Courier Corporation, 2014.