You can work together, but you must write up your solutions independently. Be sure to show how you derived your solutions.

For the simple linear regression model  $Y_i = \alpha + \beta X_i + \epsilon_i$ ,

- 1. Find the method of moments estimator of  $\beta$ .
- 2. Find the MLE of  $\beta$ .
- 3. To test  $H_0: \beta = 0$ ,
  - (a) derive a LRT test.
  - (b) derive a Wald test.
  - (c) derive a score test.

Be explicit about the assumptions you are making and include any equations you solve in addition to the final answer.

## Statistical Theory Final Solution

## Bohao Tang

## October 10, 2017

1. Suppose  $\alpha, \beta$  is parameters and  $\epsilon_i, X_i$  are samples from two uncorrelated random variables, hence  $Y_i$  are also samples from a random variables. Also we suppose  $\epsilon_i$  have mean 0 and  $\frac{\sum X_i^2}{n} > (\frac{\sum X_i}{n})^2$ .

Then we have moments:

$$\mathbb{E}[Y] = \alpha + \beta \mathbb{E}[X] + \mathbb{E}[\epsilon] = \alpha + \beta \mathbb{E}[X] \tag{1}$$

$$\mathbb{E}[XY] = \alpha \mathbb{E}[X] + \beta \mathbb{E}[X^2] + \mathbb{E}[\epsilon X] = \alpha \mathbb{E}[X] + \beta \mathbb{E}[X^2]$$
 (2)

Replace the expectation by the sample mean we have the MME of  $\alpha$ ,  $\beta$  are the solution of equation (n is the sample size):

$$\frac{\sum Y_i}{n} = \alpha + \beta \frac{\sum X_i}{n} \tag{3}$$

$$\frac{\sum X_i Y_i}{n} = \alpha \frac{\sum X_i}{n} + \beta \frac{\sum X_i^2}{n} \tag{4}$$

Then we solve that  $\hat{\beta} = \frac{\overline{XY} - \overline{X} \cdot \overline{Y}}{\overline{X^2} - \overline{X}^2}$ , where  $\overline{X}, \overline{Y}, \overline{XY}, \overline{X^2}$  are the sample means of  $X, Y, XY, X^2$ .

2. Suppose the sample size is n and  $(\epsilon_1, \epsilon_2, \dots, \epsilon_n) \sim N(\mathbf{0}, \sigma^2 \mathbf{I})$ , where  $\mathbf{I}$  is the identity matrix of size  $n \times n$ . And  $X_i$  are given numbers,  $\sigma$  is known,  $\alpha, \beta$  are parameters. Also suppose that  $\frac{\sum X_i^2}{n} > (\frac{\sum X_i}{n})^2$ .

Then we have the likelihood of the sample is:

$$L(\alpha, \beta) = \prod_{1}^{n} \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(Y_i - \alpha - \beta X_i)^2}{2\sigma^2}\right\}$$

Calculate the derivtives of  $-\log L(\alpha, \beta)$ , we get the regular equation for the MLE:

$$\sum_{i=1}^{n} Y_i - \alpha - \beta X_i = 0 \tag{5}$$

$$\sum_{1}^{n} X_i (Y_i - \alpha - \beta X_i) = 0 \tag{6}$$

And the unique solution is indeed the minimal point of  $-\log L$ . Then we get that  $\hat{\beta} = \frac{\overline{XY} - \overline{X} \cdot \overline{Y}}{\overline{X^2} - \overline{X}^2}$ , where  $\overline{X}, \overline{Y}, \overline{XY}, \overline{X^2}$  are the sample means of  $X, Y, XY, X^2$ .

- 3. Here, we use the same assumption and notation as in question 2.
  - (a) In this situation, the likelihood ratio is:

$$\Lambda = \frac{\sup_{\alpha} L(\alpha, 0)}{\sup_{\alpha, \beta} L(\alpha, \beta)}$$

The rejection field will be  $\{\Lambda < C\}$ , where C is choosen by solve:

$$\{\sup_{\alpha} \mathbf{P}(\Lambda < C | \alpha = \alpha, \beta = 0)\} = l_{\alpha}$$

Where  $l_{\alpha}$  is the level of the test you need.

(b) In this situation, the test statistics will be  $T = \frac{(\hat{\beta} - 0)^2}{var(\hat{\beta})}$ . And the rejection field will be  $\{T > C\}$ , where C is choosen by solve:

$$\{\sup_{\alpha} \mathbf{P}(T > C | \alpha = \alpha, \beta = 0)\} = l_{\alpha}$$

Where  $l_{\alpha}$  is the level of the test you need.

(c) In this situation, denote  $\hat{\alpha}_0$  be the MLE of  $\alpha$  when  $\beta = 0$ . And  $U(\alpha, \beta)$  be the derivtive vector  $\frac{\partial \log L(\alpha, \beta)}{\partial(\alpha, \beta)}$ . And  $I(\alpha, \beta)$  be the Fisher information matrix of this model.

Then the test statistics is

$$S = U^{T}(\hat{\alpha}_{0}, 0)I^{-1}(\hat{\alpha}_{0}, 0)U(\hat{\alpha}_{0}, 0)$$

The rejection field is  $\{S > C\}$ , where C is choosen by solve:

$$\{\sup_{\alpha} \mathbf{P}(S > C | \alpha = \alpha, \beta = 0)\} = l_{\alpha}$$

Where  $l_{\alpha}$  is the level of the test you need.