

Statistical Theory Problem Set 2

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(i) Then: (since the meaning of y , $f(y; \theta) = 0$ for $y < 0$)

$$\begin{aligned} \text{pr}(Y^{\text{obs}}, I | \theta, \alpha) &= \prod_{i: I_i=1} f(Y_i, \theta) \pi(Y_i, \alpha) \prod_{i: I_i=0} \int f(Y_i, \theta) (1 - \pi(Y_i, \alpha)) dY_i \\ &= \prod_{i: I_i=1} f(Y_i, \theta) \frac{Y_i}{\alpha} \prod_{i: I_i=0} \int_0^{\alpha} f(Y_i, \theta) (1 - \frac{Y_i}{\alpha}) dY_i \\ &= \left(\prod_{i: I_i=1} \frac{Y_i f(Y_i, \theta)}{\alpha} \right) \cdot \left[\int_0^{\alpha} f(y, \theta) (1 - \frac{y}{\alpha}) dy \right]^{(n - \sum_{i=1}^n I_i)} \\ &= \left[\prod_{i: I_i=1} \frac{Y_i f(Y_i, \theta)}{\alpha} \right] \cdot \left[1 - \frac{E[Y_i | \theta]}{\alpha} \right]^{(n - \sum_{i=1}^n I_i)} \end{aligned}$$

(ii) $\text{Pr}(Y^{\text{obs}} | I, n_1=500, \theta, \alpha)$

$$= \text{Pr}(Y^{\text{obs}} | I, \theta, \alpha)$$

$$= \text{Pr}(Y_1, Y_2, \dots, Y_{n_1}, I_1=1, I_2=1, \dots, I_{n_1}=1, I_{n_1+1}=0, I_{n_1+2}=0, \dots, I_n=0 | \theta, \alpha)$$

$$\text{Pr}(I_1=1, I_2=1, \dots, I_{n_1}=1, I_{n_1+1}=0, I_{n_1+2}=0, \dots, I_n=0 | \theta, \alpha)$$

$$= \frac{\left(\prod_{i=1}^{n_1} \frac{Y_i f(Y_i, \theta)}{\alpha} \right) \cdot \left[\int_0^{\alpha} f(y, \theta) (1 - \frac{y}{\alpha}) dy \right]^{n-n_1}}{\left[\prod_{i=1}^{n_1} \int_0^{\alpha} \frac{y f(y, \theta)}{\alpha} dy \right] \cdot \left[\int_0^{\alpha} f(y, \theta) (1 - \frac{y}{\alpha}) dy \right]^{n-n_1}}$$

$$= \frac{\left[\prod_{i=1}^{n_1} \frac{Y_i f(Y_i, \theta)}{\alpha} \right] \cdot \left[\int_0^{\alpha} f(y, \theta) (1 - \frac{y}{\alpha}) dy \right]^{n-n_1}}{\left[\int_0^{\alpha} f(y, \theta) \frac{y}{\alpha} dy \right]^{n_1} \left[\int_0^{\alpha} f(y, \theta) (1 - \frac{y}{\alpha}) dy \right]^{n-n_1}} = \frac{\prod_{i=1}^{n_1} [Y_i f(Y_i, \theta)]}{\left[\int_0^{\alpha} f(y, \theta) y dy \right]^{n_1}}$$

where $n_1=500$

(iii) Actually here α will be $+\infty$ since Gamma is not bounded
 then $\pi(Y_i, \alpha) = \frac{Y_i}{\alpha}$ will not be a "regular" density

But If we ignore this and regard α a normal constant.

$$\text{Then: } \Pr[Y_i | I_i=1, \theta, \alpha] = \frac{\Pr[Y_i, I_i=1 | \theta, \alpha]}{\Pr[I_i=1 | \theta, \alpha]}$$

$$= \frac{f(Y_i, \theta) \frac{Y_i}{\alpha}}{\int_0^{+\infty} f(Y_i, \theta) \frac{Y_i}{\alpha} dY_i} = \frac{f(Y_i, \theta) \cdot Y_i}{E[Y]}$$

$$\text{So } E[Y_i | I_i=1, \theta, \alpha] = \frac{E[Y_i^2]}{E[Y_i]} = \frac{\theta_2 + \theta_1^2}{\theta_1}$$

(iv) In the likelihood in (ii)

$$L = \frac{\prod_{i=1}^{n_1} [Y_i \cdot f(Y_i, \theta)]}{\left[\int_0^{\alpha} f(y, \theta) y dy \right]^{n_1}}$$

where $n_1 = 500$ is considered as a constant.

Since there's no further factorization (not rigorous enough)

the minimal sufficient statistic is just $(Y_1, Y_2, \dots, Y_{500})$

In the assumption in (iii) again consider α a normal constant $\in \mathcal{R}$

$$L = \prod_{i=1}^{n_1} \left[\frac{(\theta_1/\theta_2)^{\theta_1/\theta_2}}{\Gamma(\theta_1/\theta_2)} \cdot Y_i^{\theta_1/\theta_2} e^{-\frac{\theta_1 Y_i}{\theta_2}} \right] \cdot \frac{1}{\alpha^{\sum_{i=1}^{n_1} I_i}}$$

(and likelihood in (i))

$$L = \prod_{i=1}^n \left[\frac{(\theta_1/\theta_2)^{\theta_1/\theta_2}}{\Gamma(\theta_1/\theta_2)} Y_i^{\theta_1/\theta_2} e^{-\frac{\theta_1 Y_i}{\theta_2}} \right] \cdot \frac{1}{\alpha^{\sum_{i=1}^n I_i}} \cdot \left[1 - \frac{\theta_1}{\alpha} \right]^{n - \sum_{i=1}^n I_i}$$

$$= h(\theta_1, \theta_2, \alpha) \cdot \left(\prod_{i=1}^n Y_i \right)^{\theta_1/\theta_2} e^{-\frac{\theta_1 \sum_{i=1}^n Y_i}{\theta_2}} \cdot (\alpha - \theta_1)^{-\sum_{i=1}^n I_i}$$

There's no further factorization, so minimal sufficient statistic is

$$\left(\prod_{i=1}^n Y_i^{I_i}, \sum_{i=1}^n Y_i \cdot I_i, \sum_{i=1}^n I_i \right)$$

(V) Again assume α is a constant $\in \mathbb{R}$.

$$\begin{aligned} & \text{then } \Pr(Y^{\text{obs}} | I^{\text{obs}}, n_1=500, \theta, \alpha) \\ &= \frac{\prod_{i=1}^{n_1} \left[\frac{(\theta_1/\theta_2)^{\theta_1^2/\theta_2}}{\Gamma(\theta_1^2/\theta_2)} Y_i^{\theta_1^2/\theta_2} e^{-\frac{\theta_1 Y_i}{\theta_2}} \right]}{\theta_1^{n_1}} \\ &= h(\theta_1, \theta_2) \cdot \left[\prod_{i=1}^{n_1} Y_i \right]^{\theta_1^2/\theta_2} e^{-\frac{\theta_1}{\theta_2} \sum_{i=1}^{n_1} Y_i} / \theta_1^{n_1} \quad n_1=500 \text{ is a constant} \end{aligned}$$

There's no further factorization

so minimal statistic is $(\sum_{i=1}^{n_1} Y_i, \prod_{i=1}^{n_1} Y_i)$

If $\pi(Y_i, \alpha)$ is not a function of Y_i $\pi(Y_i, \alpha) = \pi(\alpha)$
consider also the situation in (ii) and assumption in (iii)

$$\begin{aligned} & \text{Then } \Pr(Y^{\text{obs}} | I^{\text{obs}}, n_1=500, \theta, \alpha) \\ &= \frac{\prod_{i=1}^{n_1} [f(Y_i, \theta) \pi(\alpha)]}{\prod_{i=1}^{n_1} \left[\int_0^{+\infty} \underline{f(Y, \theta)} \cdot \pi(\alpha) dy \right]} = \frac{\prod_{i=1}^{n_1} f(Y_i, \theta)}{\prod_{i=1}^{n_1} \left[\int_0^{+\infty} \underline{f(Y, \theta)} \cdot \pi(\alpha) dy \right]} \\ &= h(\theta_1, \theta_2) \cdot \left[\prod_{i=1}^{n_1} Y_i \right]^{\theta_1^2/\theta_2 - 1} e^{-\frac{\theta_1}{\theta_2} \sum_{i=1}^{n_1} Y_i} \end{aligned}$$

The likelihood ratio of two situation is $\left[\prod_{i=1}^{n_1} Y_i \right] / \theta_1^{n_1}$

So they will give the same inference for θ_2

but will not give the same inference for θ_1

necessary

consider their specific method.