

Probability Homework 2

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Ex 5.3.12: We need $Z_0=1$ here

If $P(\xi_i^m=1)=1$, then it's obvious

Else if $\mu \leq 1$: $P\left[\lim_n Z_n/\mu^n = 0\right] = 1$

So it can only be the situation $\mu > 1$

Then $P(Z_n=0 \text{ for some } n) = P$

Also let $r = P(\lim Z_n/\mu^n = 0)$

Since Z_n/μ^n is a martingale that > 0

$Z_n/\mu^n \xrightarrow{\text{a.s.}} \infty$ as $n \rightarrow +\infty$

Then $r = P(Z=0) = \sum_k P(Z=0 | Z_1=k) P(Z_1=k)$

Consider $P(Z=0 | Z_1=k) = P(Z^{(1)}+Z^{(2)}+\dots+Z^{(k)}=0 | Z_1=1, Z_2=1, \dots)$

where $Z_n^{(i)}$ is k independent copy of branching process
same

and $Z^{(i)} = \lim_n Z_n^{(i)}/\mu^n$

Since $Z^{(i)} \geq 0$ and $P(Z^{(i)} | Z_1=1) = P(Z^{(i)} | Z_0=1)$

$\Rightarrow P(Z=0 | Z_1=k) = P(Z=0)^k$

$$\text{Therefore } r = \sum_k p_k P(Z=0)^k = \varphi[P(Z=0)] = \varphi(r)$$

and since $r < 1$

$$\text{therefore } r = p$$

$$\text{We have } \left\{ \lim_{n \rightarrow \infty} Z_n / \mu^n \geq 0 \right\} \subset \left\{ Z_n \geq 0 \text{ a.s.} \right\}$$

and now they all have probability $1-p$
therefore they are equal a.s.

Ex. 5.3.13.

$$\varphi = p_0 + p_1 z + p_2 z^2 + p_3 z^3 = \frac{1}{8} + \frac{3}{8}z + \frac{3}{8}z^2 + \frac{1}{8}z^3$$

$$\text{then solve } \varphi(z) = z$$

$$\Rightarrow z^3 + 3z^2 - 5z + 1 = 0$$

$$\Rightarrow (z^2 + 4z - 1)(z - 1) = 0$$

$$\text{So the root in } [0, 1] \text{ is } \frac{\sqrt{20} - 4}{2} = \sqrt{5} - 2 \approx 0.236068$$

So the die out rate is 0.236068

Ex 5.4.5: $(X_n + C)^2$ is a submartingale a.s.

$$P(\max_{1 \leq m \leq n} X_m > \lambda) = P(\max_{1 \leq m \leq n} X_m + C > \lambda + C)$$

$$\leq P(\max_{1 \leq m \leq n} (X_m + C)^2 > (\lambda + C)^2) \leq \frac{E(X_n + C)^2}{(\lambda + C)^2}$$

$$\text{let } \lambda + C = -\frac{E(X_n - \lambda)^2}{E(X_n - \lambda)} \Rightarrow P(\max_{1 \leq m \leq n} X_m > \lambda) \leq \frac{\text{Var}(X_n)}{E(X_n - \lambda)^2} = \frac{E X_n^2}{E X_n^2 + \lambda^2}$$

$$\text{since } E X_n = E X_0 = 0$$

Ex 5.4.6.

$$\begin{aligned}
 \text{(i)} \quad E(\bar{X}_n | M) &= \int_0^{+\infty} P(\bar{X}_n | M \geq \lambda) d\lambda \\
 &= \int_0^1 P(\bar{X}_n | M \geq \lambda) d\lambda + \int_1^{+\infty} P(\bar{X}_n | M \geq \lambda) d\lambda \\
 &\leq 1 + \int_1^{+\infty} \frac{1}{\lambda} \int X_n^+ \mathbb{1}_{\{\bar{X}_n | M \geq \lambda\}} d\mu d\lambda \\
 &= 1 + \int X_n^+ \int_1^{+\infty} \frac{1}{\lambda} \mathbb{1}_{\{\bar{X}_n | M \geq \lambda\}} d\lambda d\mu \\
 &= 1 + \int X_n^+ \log^+(\bar{X}_n | M) d\mu
 \end{aligned}$$

(ii) for $a \log^+ b : a > 0, b > 0$

if $b \leq 1$ then $a \log^+ b \leq 0 \leq a \log^+ a + \frac{b}{e}$

if $b > 1$ $a \log^+ b = a \log b = a \log b - a \log a + a \log a$

for $a \log b - a \log a$ use derivative we find.

$$\begin{aligned}
 a \log b - a \log a &\leq b/e \log b - \frac{b}{e} \log \frac{b}{e} = \frac{b}{e} \\
 \text{therefore } a \log^+ b &\leq a \log a + \frac{b}{e} \leq a \log^+ a + \frac{b}{e}
 \end{aligned}$$

$$\Rightarrow a \log^+ b \leq a \log^+ a + \frac{b}{e}$$

$$\Rightarrow \int X_n^+ \log^+(\bar{X}_n | M) d\mu \leq \int X_n^+ (\log^+ X_n^+) d\mu + \frac{1}{e} \int \bar{X}_n | M d\mu$$

$$\Rightarrow E(\bar{X}_n | M) \leq \left(1 - \frac{1}{e}\right)^{-1} E[X_n^+ \log^+ X_n^+]$$

$$\text{let } M \rightarrow +\infty \Rightarrow E[\bar{X}_n] \leq (1 - e^{-1})^{-1} E[X_n^+ \log^+ X_n^+]$$

5: If X_n is not a \mathcal{F}_n -martingale, there must exist N such that

$$E[X_{n+1} | \mathcal{F}_N] \neq X_N$$

Choose N to be inf of those such value

$$N = \inf \{ n : E[X_{n+1} | \mathcal{F}_N] \neq X_N \}$$

Since $E[X_{n+1} | \mathcal{F}_N] \neq X_N$, without any loss we suppose

$$P[E[X_{n+1} | \mathcal{F}_N] > X_N] > 0$$

$$\text{and let } T = \begin{cases} N+1 & E(X_{n+1} | \mathcal{F}_N) > X_N \\ N & E(X_{n+1} | \mathcal{F}_N) \leq X_N \end{cases}$$

Then $\{T \leq n\} = \Omega$ if $n \geq N+1$

$\{T \leq n\} = \emptyset$ if $n < N$

and $\{T \leq N\} = \{T = N\} = \{E[X_{n+1} | \mathcal{F}_N] \leq X_N\} \in \mathcal{F}_N$

therefore T is a \mathcal{F}_n -stopping time.

$$\begin{aligned} E[X_T] &= E[X_{N+1} \mathbb{1}_{E(X_{n+1} | \mathcal{F}_N) > X_N} + X_N \mathbb{1}_{E(X_{n+1} | \mathcal{F}_N) \leq X_N}] \\ &= E[E[X_{n+1} | \mathcal{F}_N] \mathbb{1}_{\dots > X_N} + X_N \mathbb{1}_{\dots \leq X_N}] \end{aligned}$$

$$\begin{aligned} \text{Since } P(E[X_{n+1} | \mathcal{F}_N] > X_N) > 0 &\quad E\{E[X_{n+1} | \mathcal{F}_N] \mathbb{1}_{E(X_{n+1} | \mathcal{F}_N) > X_N}\} \\ &> E\{X_N \mathbb{1}_{E(X_{n+1} | \mathcal{F}_N) > X_N}\} \end{aligned}$$

$$\text{therefore } EX_T > E\{X_N [1_{\dots > X_N} + 1_{\dots \leq X_N}]\} = EX_N$$

Since N is the inf, we have $E[X_{i+1} | \mathcal{F}_i] = X_i$ for $i = 0, 1, \dots, N-1$

$$\Rightarrow EX_N = EX_0 \Rightarrow EX_T > EX_0$$

but T here is bounded, \mathcal{F}_n -stopping time, contradiction

$$\Rightarrow X_n \text{ is } \mathcal{F}_n\text{-martingale}$$

Ex 5.5.1

$$\varphi(x)/x \rightarrow +\infty \Rightarrow \forall G, \exists M \text{ s.t. } x > M \Rightarrow \varphi(x) > G \cdot x$$

Therefore $\forall \varepsilon > 0$ let G be large enough such that $\frac{C}{G} < \varepsilon$
and M is related to G .

Then

$$\begin{aligned} E[|X_i| \mathbf{1}_{|X_i| > M}] &\leq E\left[\frac{\varphi(X_i)}{G} \mathbf{1}_{|X_i| > M}\right] \\ &\leq \frac{1}{G} E[\varphi(X_i)] \leq \frac{C}{G} < \varepsilon. \quad \text{Hence} \end{aligned}$$

$$\Rightarrow \sup_i E[|X_i| ; |X_i| > M] < \varepsilon$$

$$\Rightarrow \limsup_m \sup_i E[|X_i| ; |X_i| > M] = 0$$

$\Rightarrow \{X_i : i \in I\}$ is uniformly integrable.