BST 210 Applied Regression Analysis



Lecture 16Plan for Today

Extending (binary) Logistic Regression to -

Categorical logistic regression with >2 outcomes:

- Multinomial Logistic Regression (>2 unordered levels)
- Ordinal Logistic Regression (> 2 ordered levels)

Multinomial Outcome Data

• Sometimes we have a categorical outcome variable that can assume *more than two categories* (not ordered):

- Choice of occupation
- RCT outcomes: progression-free survival, disease-free survival, death
- 4 modes of operative delivery
- 3 types (epithelial, germ cell, and stromal) of ovarian tumors
- Diagnosis of 3 bacterial infections seen in children
- 5 forms of meningitis
- 3 prognostic outcomes of elderly after hospitalization
- 6 cost-structure models in hospital reimbursement

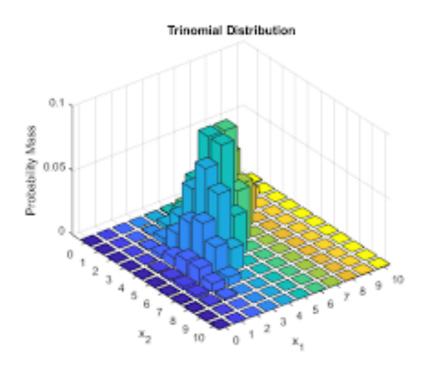
Multinomial Outcome Data

Multinomial Distribution

- The Binomial distribution can be extended to describe number of outcomes in a series of independent trials each having more than 2 possible outcomes.
- If a given trail can result in the k outcomes $E_1, E_2, ..., E_k$ with probabilities $p_1, p_2, ..., p_k$, then the probability distribution of the random variables $X_1, X_2, ..., X_k$, representing the number of occurrences for $E_1, E_2, ..., E_k$ in n independent trials is

$$p_{X_1,...,X_k}(x_1,...,x_k) = \frac{n!}{x_1! x_2! \cdots x_k!} p_1^{x_1} p_2^{x_2} \cdots p_k^{x_k}$$
with $\sum_{i=1}^k x_i = n$, and $\sum_{i=1}^k p_i = 1$.

Multinomial Outcome Data



Recall: (Binary) Logistic Regression

We have been modeling <u>binary outcomes</u> as:

$$logit(p_i) = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + ... + \beta_k x_{ik}$$

always assuming

- (1) we have the correctly specified model to relate Y_i with X_{ik}
- (2) Y_i are independent

Same assumptions must hold for modeling <u>multinomial outcomes</u>

Multinomial Model

- In the <u>multinomial</u> outcome situation, we can use <u>multinomial</u> or polychotomous logistic regression
- We now need proportions p_1 , p_2 , p_3 ,... p_c for an outcome having c levels, instead of simply p_1 and p_2 (all of which still need to sum to 1)
- AND subjects (Y_i) still must be independent of each other
- We'd also need to have enough observations in each category, or >
 10 observations per independent variable (predictor)
- If the above hold, likelihood based methods can be used to estimate model parameters (MLEs), just as in the binary outcome case

Recall: (Binary) Logistic Regression

 For logistic regression with a single covariate X to predict a <u>binary</u> outcome Y, we use the model:

$$logit(p) = log[p / (1-p)] = \beta_0 + \beta_1 x$$

- Here $p = P(Y = 1 \mid X = x)$ and we assume that the relationship between logit(p) and x is linear
- Here we are comparing outcomes for <u>Y=1 to Y=0</u>, for '1-unit changes in X'
- Solving for *p*, we obtain:

$$p = P(Y=1) = \frac{exp^{(\alpha + \beta X)}}{1 + exp^{(\alpha + \beta X)}}$$

and

1-p=P(Y = 0) =
$$\frac{1}{1 + \exp^{(\alpha + \beta X)}}$$

Now: Multinomial model

So how do we develop the multinomial model?

- Suppose there are three possible outcomes, say 1, 2, and 3 (or 0, 1, 2 is another possibility)
- One possible approach to this new problem with a multi-level categorical outcome variable would be to run *two separate logistic regression models* such that:

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In model 1: Compare category 3 vs. category 1 and 2 (combined) \rightarrow ie 0 = {3} and 1 = {1, 2}
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In model 2: Compare category 2 and 3 (combined) vs. category $1 \rightarrow ie 0 = \{2, 3\}$ and $1 = \{1\}$

Multinomial Model

- Fitting two separate models is a valid approach only if it makes sense to combine the categories like this, that is, if the categories are ordered in a logical way (and they may not be)
- Fitting two separate models does not allow us to directly compare the coefficients for the same risk factor in the two separate models
- What to do?

- Instead of fitting two separate models, we will fit a <u>single model</u> with c outcome categories (c may be 3, 4, 5...not 28!)
- We'll call this a *multinomial logistic regression model*
- Standard errors will be smaller this way; model more stable
- Let group 1 be the reference group say (can be any group), and groups 2 through c will be compared to group 1
- Note that c = 2 is simply the special case of a binary logistic regression model
- Let's develop this model →

 Suppose c = 3 and we have a single covariate x. We can then obtain

P(group 1) =
$$\frac{1}{1 + \exp(\alpha_2 + \beta_{21}x) + \exp(\alpha_3 + \beta_{31}x)}$$
P(group 2) =
$$\frac{\exp(\alpha_2 + \beta_{21}x)}{1 + \exp(\alpha_2 + \beta_{21}x) + \exp(\alpha_3 + \beta_{31}x)}$$
P(group 3) =
$$\frac{\exp(\alpha_3 + \beta_{31}x)}{1 + \exp(\alpha_2 + \beta_{21}x) + \exp(\alpha_3 + \beta_{31}x)}$$

• Here the probabilities for all c = 3 outcomes sum to 1.

- How do we derive these probabilities?
- (Remember, we want nice forms for P(1), P(2), and P(3) so that we can begin making interpretations in terms of Risk Ratios, and Relative Risk Ratios.)

Recall: Measures of Effect - Risk Ratio

- The <u>risk ratio</u> (RR) is a way of estimating magnitude of association that we discussed earlier in the course
- It is generally defined as p_1/p_2 and can be estimated by

$$\hat{p}_{1}/\hat{p}_{2}$$
.

It is a measure of relative rather than absolute risk

Now back to our multinomial model setting – let's interpret this model!

- What is the effect for a 1 unit increase in x?
- To answer this question, first we need the <u>risk ratio</u> of being in outcome category 2 versus the reference category 1, which is given by

P(group 2) / P(group 1) =
$$\exp(\alpha_2 + \beta_{21}x)$$

 Then the <u>relative risk ratio</u> of being in outcome category 2 versus the reference category 1 for a one unit increase in x is given by

$$\exp(\alpha_2 + \beta_{21}(x+1)) / \exp(\alpha_2 + \beta_{21}x) = \exp(\beta_{21})$$

How do we derive the above interpretations?

Similarly, the relative risk ratio of being in outcome category 3
versus the reference category 1 for a one unit increase in x is
given by

$$\exp(\beta_{31})$$

- With a bit more thought we can also get the relative risk ratios comparing category 1 to category 2, or category 2 to category 3, etc.
- Note that multinomial models are not easily interpretable in terms of odds ratios

More generally, for k covariates and c outcomes :

P(group 1) =
$$\frac{1}{1 + \sum_{j=2}^{c} \exp(\alpha_j + \sum_{i=1}^{k} \beta_{ji} x_i)}$$

P(group q) =
$$\frac{\exp(\alpha_q + \sum_{i=1}^k \beta_{qi} x_i)}{1 + \sum_{j=2}^c \exp(\alpha_j + \sum_{i=1}^k \beta_{ji} x_i)}$$

for
$$q = 2, 3, ..., c$$
.

Here the probabilities for all c outcomes still sum to 1.

- For each covariate or risk factor x there are c 1
 parameters to be estimated (c = outcome group), one for
 each category of the outcome other than the reference
 group
- For k covariates, there are $k \times (c-1)$ regression parameters, plus the c-1 intercept terms
- And as was noted earlier in the single covariate case, when c = 2, multinomial logistic regression in the general case reduces to (ordinary/binary) logistic regression with k regression parameters plus 1 intercept term

• Here, we can generalize the previous results to say that the <u>relative risk ratio</u> of being in outcome category q versus the reference category 1 for a one unit increase in covariate x_i , i=1,...,k is given by

$$\frac{exp[\beta_{qi}(x_i+1)]}{exp(\beta_{qi}x_i)} = exp(\beta_{qi})$$

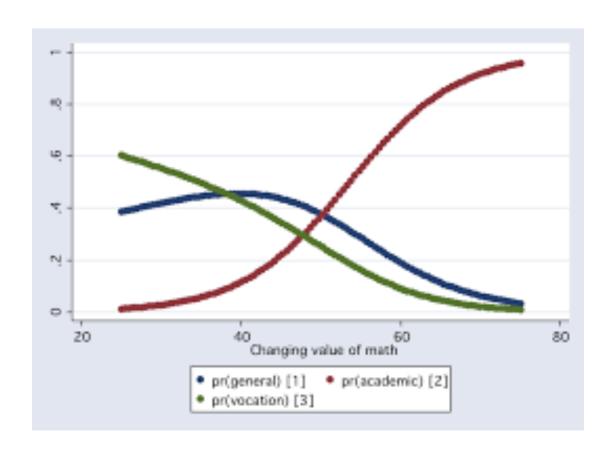
adjusting for the remaining covariates included in the model.

- Probabilities of the outcome(s) can be estimated for subjects with specified combinations of risk factors
- Within each covariate pattern, the estimated probabilities of observing a 1, 2, ..., c will sum to 1
- The probabilities themselves differ depending on the risk factors

Multinomial Regression Model Building

- Once you have a handle on the interpretation of multinomial logistic regression models, assessment of confounding and effect modification as well as model building techniques follow similarly to (ordinary) logistic regression.
- Parameter estimates are found using maximum likelihood methods, and also standard error estimates,
- P-values, and confidence intervals can be obtained.
- Likelihood ratio tests will be especially useful (when testing multiple parameters at a time).

Multinomial Regression Model Building



Example: Hyponatremia

- Data not collected on many (488 out of 14k+ runners); no elite runners
- Self reported: gender, run time, body mass index, weight gain, water consumption, urination freq, and more
- Useful to have clinically meaningful cutoff for sodium variable (≤135 mmol/l)

- Multinomial logistic regression makes no assumptions about the possible ordering of the categories of the outcome variable – the outcome is treated as a nominal variable (and it doesn't even matter which group is chosen to be reference)
- In the hyponatremia example, the outcome is in fact ordinal
- Here we could use <u>ordinal logistic regression</u>

- Suppose that an ordinal outcome Y has c ordered categories (lowest to highest) labeled as j = 1, 2, ... c
- The <u>proportional odds ordinal logistic regression model</u> is based on the probability of success $P(Y \ge j)$ (or cumulative logit):

$$logit[P(Y \ge j)] = log \frac{P(Y \ge j)}{P(Y < j)}$$

$$= \alpha_j + \beta_1 x_1 + \dots + \beta_p x_p$$
where $j = 2, \dots C$

- The β coefficients do not differ across outcome categories (but the intercept terms do) much easier to interpret than multinomial!
- And...we once again have an odds ratio!

- Suppose we have two subjects A and B who differ by 1 unit for the variable x_i , with all other covariates taking the same value
- Then exp(β_i) is just the <u>odds ratio</u> that Y ≥ j versus Y < j for subject A versus subject B
- $\exp(\beta_i)$ is assumed to be the *same* for each possible value of j
- Estimated probabilities do sum to 1 here

- This is called the proportional odds assumption
- If this assumption is valid, there are many fewer parameters to be estimated than in the multinomial logistic regression model
- These betas are also seeming easier to interpret (OR!)
- If c = 2, again the model reduces to (ordinary/binary) logistic regression

Ordinal versus Multinomial

- The ordinal logistic regression model has more assumptions that the multinomial model (ie proportional odds), which may be hard to satisfy (but can be tested)
- Different software packages may have different tests (often approximate tests) of the proportional odds assumption
- This assumption is worth checking, because if the ordinal model is appropriate, model interpretation is simpler (only 1 set of beta coefficients!)

→ Think Parsimony

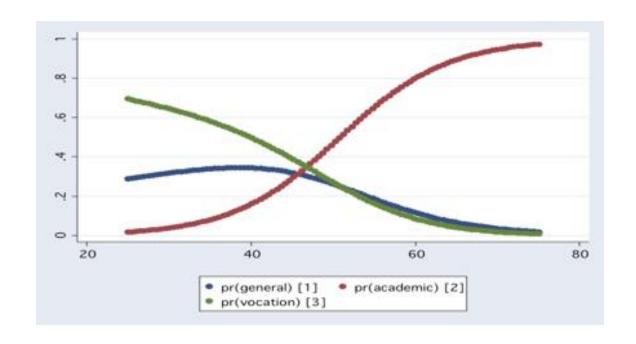
Ordinal versus Multinomial

- Because it has fewer assumptions, the multinomial logistic regression model has more parameters that need to be estimated (so can be more difficult to interpret)
- Because there are so many parameters, it is possible to end up with an over-fitted (not generalizable) multinomial model
- Looking at examples helps!

Note

• Unfortunately, different books and software packages use different notation, and you have to be very careful on how you are interpreting the β coefficients, intercepts, fitted probabilities, and relative risk ratio (for the multinomial model) or the odds ratio (assuming the proportional odds assumption for the ordinal model)...

Recall: Multinomial logistic regression



Ordinal (proportional odds) logistic regression

