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Exercise: If X_1 is uniform on $[0, 1]$ ✓

and

$X_2 | X_1$

also uniform

on $[0, X_1]$ ✓

find the joint and marginal distribution of X_1 and X_2 .

Solution:

Joint distribution of X_1 and X_2 :

$$f_{X_1, X_2}(x_1, x_2) = f_{X_1}(x_1) \cdot f_{X_2 | X_1}(x_2 | x_1)$$

$$f_{X_1}(x_1) = \begin{cases} 1, & 0 < x_1 < 1 \\ 0, & \text{otherwise} \end{cases}$$

$$0 < x_1 < 1$$

otherwise

$$= 1 * \frac{1}{x_1}$$

$$= \frac{1}{x_1}$$

$$f_{X_2 | X_1}(x_2 | x_1) = \begin{cases} \frac{1}{x_1}, & 0 < x_2 < x_1 \\ 0, & \text{otherwise} \end{cases}$$

$$0 < x_2 < x_1$$

$$0 < x_1 < 1$$
$$0 < x_2 < x_1$$

Aside:

① $X \sim \text{Unif.}([a, b])$

$$f(x) = \frac{1}{b-a}, \text{ in } [a, b]$$

Solution:

$$f_{X_1, X_2}(x_1, x_2) = \frac{1}{x_1}$$

$$0 < x_1 < 1$$
$$0 < x_2 < x_1$$

Marginal
distribution
for X_1 :

Given
 $f_{X_1}(x_1) = 1, \quad 0 < x_1 < 1$

Another way:

$$\begin{aligned} f_{X_1}(x_1) &= \int_0^{x_1} f_{X_1, X_2}(x_1, x_2) dx_2 \\ &= \int_0^{x_1} \frac{1}{x_1} dx_2 = 1 \end{aligned}$$

when $0 < x_1 < 1$

Marginal
distribution
for X_2 :

$$\begin{aligned} f_{X_2}(x_2) &= \int_{x_2}^1 f(x_1, x_2) dx_1 \\ &= \int_{x_2}^1 \frac{1}{x_1} dx_1 \\ &= \ln(x_1) \Big|_{x_2}^1 \\ &= \ln(1) - \ln(x_2) \end{aligned}$$

$$f_{X_2}(x_2) = -\ln(x_2), \quad 0 < x_2 < 1$$

Example: Let X and Y be independent std. normal variables.

Find the density of $Z = X + Y$.

Solution:

$$f_Z(z)$$

Answer

$$Z \sim N(0, 2)$$

(done in class before!)

Aside!

$$Z = \frac{Y}{X}$$

and $X, Y \sim N(0, 1)$

Then $Z \sim$ Std. Cauchy

Solution:

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x) \cdot f_Y(z-x) dx$$

$$(\because X \perp Y)$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(z-x)^2} dx$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2}(x^2 + (z-x)^2)\right) dx$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2}(2x^2 - 2xz + z^2)\right) dx$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2}\left(2\left(x - \frac{1}{2}z\right)^2 + \frac{z^2}{2}\right)\right) dx$$

$$\exp(x) = e^x$$

$$= \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{4}z^2\right) \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2}\left(\frac{x-\frac{1}{2}z}{\frac{1}{\sqrt{2}}}\right)^2\right) dx$$

$$= \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{4}z^2\right) \cdot \sqrt{2\pi} \left(\frac{1}{\sqrt{2}}\right)^2$$

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx = 1$$

pdf of Normal

$$\int_{-\infty}^{\infty} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx = \sqrt{2\pi}\sigma^2$$

~~$N(\mu, \sigma^2)$~~

~~$$\frac{1}{\sqrt{2\pi}(\sqrt{2})^2} \exp\left(-\frac{1}{2}\left(\frac{z}{\sqrt{2}}\right)^2\right)$$~~

$$= \frac{1}{\sqrt{2\pi}(\sqrt{2})^2} \exp\left(-\frac{1}{2}\left(\frac{z}{\sqrt{2}}\right)^2\right)$$

So, this is the pdf of Normal $(0, (\sqrt{2})^2)$
 \Rightarrow Normal $(0, 2)$

Expected Value

Definition: (Discrete)

If X is a discrete random variable with probability mass function (p.m.f) $p(x)$, the expected value of X is:

$$E(X) = \sum_i x_i p(x_i) \quad \text{--- (1) } \checkmark$$

provided $\sum_i |x_i| p(x_i) < \infty \quad \text{--- (2) } \checkmark$

Aside: (2) $\Rightarrow E(X) < \infty$

$$\begin{aligned} |E(X)| &= \left| \sum_i x_i p(x_i) \right| \leq \sum_i |x_i| p(x_i) \\ &= \sum_i |x_i| p(x_i) < \infty \end{aligned}$$

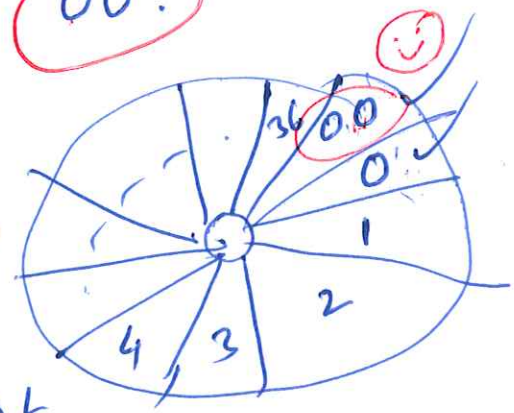
Triangle inequality by (2)

Hence $|E(X)| < \infty$
 $\Rightarrow E(X)$ is finite.

Example: Roulette

A roulette wheel number 1 through 36 as well as 0 and 00.

If you bet \$1 that an odd number comes up, you win or lose \$1 according to whether or not that event occurs.



X : ~~net~~ gain

$X = +1$ with prob. $= \frac{18}{38}$
 $X = -1$ with prob. $= \frac{20}{38}$

The expected value of X :

$$E(X) = (+1) \frac{18}{38} + (-1) \frac{20}{38} = \left(-\frac{1}{19} \right)$$

Expected value = "Average"
(Statistical)

(7/2)

1, 2, 3, 4

$$\frac{1+2+3+4}{4}$$

X	Prob.
1	$\frac{1}{4}$
2	$\frac{1}{4}$ $\frac{1}{8}$
3	$\frac{1}{4}$
4	$\frac{3}{8}$

$$E(X) = 1\left(\frac{1}{4}\right) + 2\left(\frac{1}{8}\right) + 3\left(\frac{1}{4}\right) + 4\left(\frac{3}{8}\right)$$

$$\frac{1}{2} - \frac{1}{8} = \frac{3}{8}$$

Expectation of a geometric random variable

Discrete
 $k=1, 2, 3, \dots$

$$P(X=k) = q^{k-1} p \quad (q=1-p)$$

$$E(X) = \sum_{k=1}^{\infty} \underbrace{k}_{\text{value}} \underbrace{q^{k-1} p}_{\text{prob.}} = p \sum_{k=1}^{\infty} k q^{k-1}$$

$$= p \left(\sum_{k=1}^{\infty} \frac{d}{dq} (q^k) \right)$$

$$= p \cdot \frac{d}{dq} \left(\sum_{k=1}^{\infty} q^k \right)$$

$$= p \cdot \frac{d}{dq} \left(\frac{1}{1-q} \right)$$

$$= p \cdot \frac{1}{(1-q)^2} = p \cdot \frac{1}{p^2}$$

$$= \boxed{\frac{1}{p}}$$

$$S = \sum_{k=1}^{\infty} k q^{k-1}$$

$$= 1 + 2q + 3q^2 + 4q^3 + \dots$$

$$q + 2q^2 + 3q^3 + \dots$$

(-)

$$(1-q)S = 1 + q + q^2 + q^3 + \dots$$

$$(1-q)S = \frac{1}{1-q}$$

$$S = \frac{1}{(1-q)^2}$$

$$\text{Hence, } E(X) = p \cdot S = p \cdot \frac{1}{(1-q)^2} = p \cdot \frac{1}{p^2} = \boxed{\frac{1}{p}}$$

Summary: $X \sim \text{Geometric}$ (prob of success = p)
 $E(X) = \frac{1}{p}$

Poisson distribution

$$P(X=k) = \frac{e^{-\lambda} \lambda^k}{k!}$$

$$k = 0, 1, 2, \dots$$

$$E(X)$$

$$= \sum_{k=0}^{\infty} k \cdot \left(\frac{e^{-\lambda} \lambda^k}{k!} \right)$$

$$= e^{-\lambda} \sum_{k=1}^{\infty} \frac{\lambda^k}{(k-1)!}$$

$$= e^{-\lambda} \lambda \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!}$$

$$u = k-1 \Rightarrow e^{-\lambda} \lambda \sum_{u=0}^{\infty} \frac{\lambda^u}{u!}$$

$$= e^{-\lambda} \cdot \lambda \cdot e^{\lambda} = \lambda$$

Summary:

$X \sim \text{Poisson}(\lambda)$

$$E(X) = \lambda$$

$$X \sim \text{Binomial}(n, p)$$

$$E(X) = np$$

$$P(X=k) = \binom{n}{k} p^k (1-p)^{n-k}$$

~~$$E(X) =$$~~

$$E(X) = \sum_{k=0}^n k \cdot \binom{n}{k} p^k (1-p)^{n-k}$$

$$= \dots$$

$$= np$$

(Homework)

Definition (Continuous random variable):

If X is a continuous random variable with density $f(x)$, then

$$E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx \dots (1)$$

provided $\int_{-\infty}^{\infty} |x| \cdot f(x) dx < \infty$

If (1) diverges, the expectation is undefined.

Normal distribution

$$f(x) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}, \quad -\infty < x < \infty$$

$$E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx = \int_{-\infty}^{\infty} x \cdot \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx$$

Sub:
 $u = x - \mu$
 $du = dx$

$$= \int_{-\infty}^{\infty} (u + \mu) \cdot \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{u^2}{2\sigma^2}} du$$
$$= \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} u \cdot e^{-\frac{u^2}{2\sigma^2}} du + \frac{\mu}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{u^2}{2\sigma^2}} du$$

Aside:

$$\int_{-\infty}^{\infty} f(x) dx = 0$$

$f(x)$: odd function

$$f(-x) = -f(x)$$

$$f(u) = u \cdot e^{-\frac{u^2}{2\sigma^2}}$$

$$f(-u) = -f(u)$$

Hence $f(u)$ is odd

$$= \mu \cdot \frac{1}{\sigma \sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{u^2}{2\sigma^2}} du$$

$$= \boxed{\mu}$$

Standard
Cauchy distribution

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We found
pdf of
Standard
Cauchy

$$f(x) = \frac{1}{\pi} \cdot \left(\frac{1}{1+x^2} \right),$$
$$-\infty < x < \infty$$

$$E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx$$

odd function.

$$= \int_{-\infty}^{\infty} x \cdot \left(\frac{1}{\pi} \cdot \frac{1}{1+x^2} \right) dx$$

$$= 0 \quad \checkmark$$

But

$$\int_{-\infty}^{\infty} |x| \cdot f(x) dx = \int_{-\infty}^{\infty} |x| \cdot \left(\frac{1}{\pi} \cdot \frac{1}{1+x^2} \right) dx$$

Even function.

$$\int_{-\infty}^{\infty} f(x) dx = 2 \int_0^{\infty} f(x) dx$$

If $f(x)$ is even
i.e., $f(-x) = f(x)$

$$= \frac{2}{\pi} \int_0^{\infty} |x| \cdot \frac{1}{1+x^2} dx$$
$$= \frac{2}{\pi} \int_0^{\infty} \frac{x}{1+x^2} dx$$

$$= \frac{2}{\pi} \int_1^{\infty} \frac{1}{u} \frac{du}{2x}$$

$$= \frac{1}{\pi} \ln u \Big|_1^{\infty}$$

$$\rightarrow \infty$$

Sub,

$$u = 1 + x^2$$

$$du = 2x dx$$

When $x=0, u=1$
 $x=\infty, u=\infty$

Hence, $\int_{-\infty}^{\infty} |x| f(x) dx \rightarrow \infty$

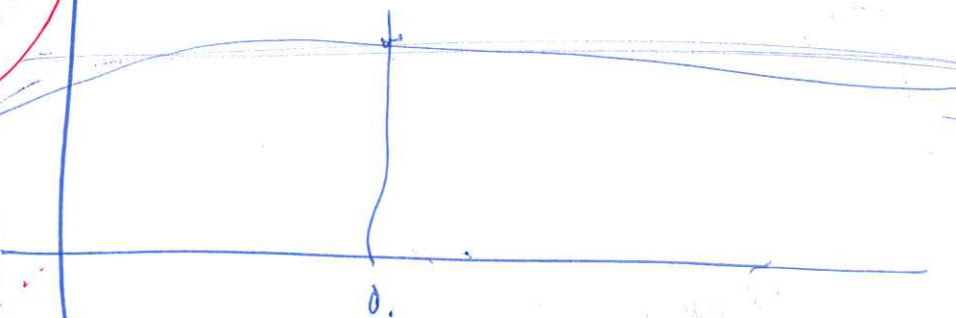
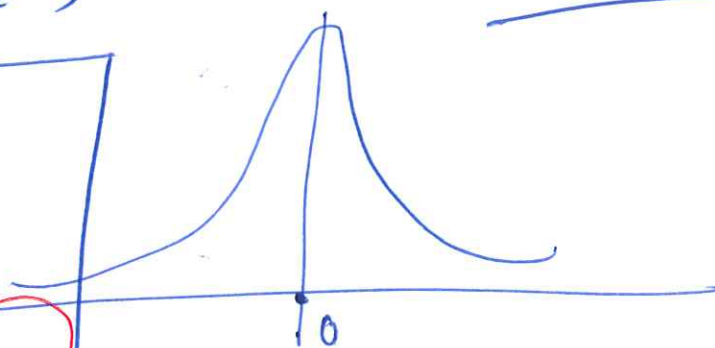
(Though, $E(X) = \int_{-\infty}^{\infty} x f(x) dx = 0$)

We say $E(X)$ fails to exist.

P.V. $\left(\int_{-\infty}^{\infty} x^3 dx \right)$

$$= \lim_{R \rightarrow \infty} \int_{-R}^R x^3 dx$$

$$= \lim_{R \rightarrow \infty} 0 = 0$$



$$\int_{-\infty}^{\infty} x^3 dx = \int_{-\infty}^0 x^3 dx + \int_0^{\infty} x^3 dx$$

$\rightarrow (\infty + \infty) \cdot ?$
