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Probability + Stochastic Process
✓ ✓

Chapter 4

30] Find $E\left(\frac{1}{X+1}\right)$, where $X \sim \text{Poisson}(\lambda)$

(Done in HW)

48] Let X, Y and Z be uncorrelated random variables with variances σ_X^2, σ_Y^2 and σ_Z^2 , respectively.

$$\text{Let } U = Z + X$$

$$V = Z + Y$$

Find $\text{Cov}(U, V)$ and ρ_{UV} .

Solution: $\text{Cov}(U, V) \stackrel{\text{def}}{=} E(UV) - E(U)E(V)$

$$\begin{aligned} \text{Cov}(U, V) &= E(U - E(U))(V - E(V)) \\ &= E((Z+X)(Z+Y)) - E(Z+X)E(Z+Y) \\ &= E(Z^2 + ZY + ZX + XY) \\ &\quad - (E(Z)^2 + E(X)E(Z) + E(Y)E(Z) + E(X)E(Y)) \end{aligned}$$

$$\begin{aligned}
 &= E(z^2) + E(zY) + E(zX) + E(XY) \\
 &\quad - E(z)^2 - E(X)E(z) - E(Y)E(z) - E(X)E(Y) \\
 &= \cancel{E(z^2)} + \cancel{E(zY)} + \cancel{E(zX)} + \cancel{E(XY)} \\
 &\quad - \cancel{E(z)^2} - \cancel{E(X)E(z)} - \cancel{E(Y)E(z)} - \cancel{E(X)E(Y)} \\
 &= \text{Var}(z) + \text{Cov}(z, Y) + \text{Cov}(X, z) \\
 &\quad + \text{Cov}(X, Y)
 \end{aligned}$$

$$= \text{Var}(z)$$

$$= \sigma_z^2 //$$

$\therefore X, Y, z$ are uncorrelated
 $\text{Cov}(z, Y) = \text{Cov}(X, z) = \text{Cov}(X, Y) = 0$

$$\checkmark \checkmark \text{Cov}(z, X, z+Y) = \sigma_z^2$$

$$\rho_{UV} = \frac{\text{Cov}(U, V)}{\sqrt{\text{Var}(U) \text{Var}(V)}} = \frac{\sigma_z^2}{\sqrt{(\sigma_z^2 + \sigma_X^2)(\sigma_z^2 + \sigma_Y^2)}}$$

$$\begin{aligned}
 \text{Var}(U) &= \text{Var}(z + X) = \text{Var}(z) + \text{Var}(X) \\
 &= \sigma_z^2 + \sigma_X^2 \quad (\because \text{Cov}(z, X) = 0)
 \end{aligned}$$

$$\begin{aligned}
 \text{Var}(V) &= \text{Var}(z + Y) = \sigma_z^2 + \sigma_Y^2 \quad (\because \text{Cov}(z, Y) = 0)
 \end{aligned}$$

58] Let X and Y be jointly distributed random variables with correlation coefficient ρ_{XY} .

Define (standardize) random variables,
 $\tilde{X} = \frac{X - E(X)}{\sqrt{\text{Var}(X)}}$, $\tilde{Y} = \frac{Y - E(Y)}{\sqrt{\text{Var}(Y)}}$

Show that $\text{Cov}(\tilde{X}, \tilde{Y}) = \rho_{XY}$.

Solution:

$$\begin{aligned}
 \text{Cov}(\tilde{X}, \tilde{Y}) &= \text{Cov}\left(\frac{X - E(X)}{\sqrt{\text{Var}(X)}}, \frac{Y - E(Y)}{\sqrt{\text{Var}(Y)}}\right) \\
 &= \frac{1}{\sqrt{\text{Var}(X)\text{Var}(Y)}} \text{Cov}(X - E(X), Y - E(Y)) \\
 &= \frac{1}{\sqrt{\text{Var}(X)\text{Var}(Y)}} \left[\text{Cov}(X, Y) - \text{Cov}(X, E(Y)) - \text{Cov}(E(X), Y) + \text{Cov}(E(X), E(Y)) \right] \\
 &= \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}} = \rho_{XY}
 \end{aligned}$$

Aside

$$\begin{aligned}
 \text{Cov}(X, a) &= E((X - E(X))(a - a)) \\
 &= 0
 \end{aligned}$$

67] A fair coin is tossed n times, and the number of heads, N , is counted. The coin is then tossed N more times.

Find the expected total # of heads generated by this process.

$$n = 100$$

70 Heads

$$N = 70$$

$$\text{Total} = 100 + 70 = 170$$

Question: How many heads

$$70 + X_1$$

$$n = 100$$

40 Heads

$$N = 40$$

$$\text{Total} = 100 + 40 = 140$$

Question: How many heads

$$40 + X$$

N

Solution: Let X denote the # of heads in the 2nd stage of the process

So, we want to compute

$$E(N + X) = E(N) + E(X)$$

$$\frac{N}{2}$$

For the first stage (n toss with N heads) (5)

of heads are binomially distributed with parameters $(n, \frac{1}{2})$

$$E(N) = n \cdot \frac{1}{2} = \frac{n}{2} \dots (1)$$

Now, $E(N+X) = E(E(N+X|N))$

fix.

(law of total expectation)

$$= E(E(N|N) + E(X|N))$$

$$= E(N + E(X|N))$$

fix.

$$= E\left(N + N \cdot \frac{1}{2}\right)$$

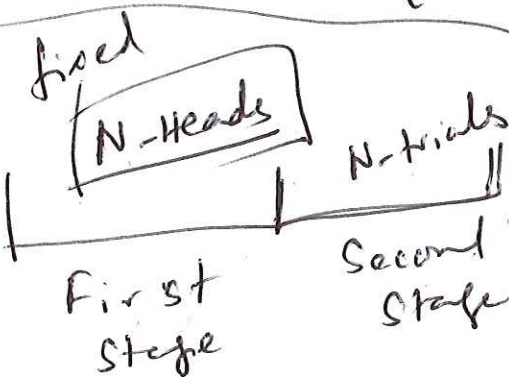
$$= E\left(\frac{3N}{2}\right)$$

$$= \frac{3}{2} E(N)$$

$$= \frac{3}{2} \left(\frac{n}{2}\right) \text{ (by (1))}$$

$$= \frac{3}{4} n$$

$$\begin{aligned} E(N|N) &\uparrow \\ &\text{(fixed)} \\ &= E(N) \\ &\uparrow \\ &\text{fixed} \\ &= N \end{aligned}$$



47] Show that

$$\text{Cov}(\bar{X}, Y) \leq \sqrt{\text{Var}(\bar{X}) \text{Var}(Y)}$$

Solution:

$$-1 \leq \rho_{\bar{X}Y} \leq 1 \quad (\text{Done in class})$$

$$\frac{\text{Cov}(\bar{X}, Y)}{\sqrt{\text{Var}(\bar{X}) \text{Var}(Y)}} \leq 1$$

$$\Rightarrow \text{Cov}(\bar{X}, Y) \leq \sqrt{\text{Var}(\bar{X}) \text{Var}(Y)}$$

$$| \langle f, g \rangle | \leq \|f\| \|g\|$$

Distribution Derived
from the Normal
Distribution

* $Z \sim N(0, 1)$

$U = Z^2$

$U =$ chi-squared distribution
with 1 degree of freedom

χ^2 -distribution

Suppose $\bar{X} \sim N(\mu, \sigma^2)$

$$\frac{(\bar{X} - \mu)}{\sigma} \sim N(0, 1)$$

$$\left(\frac{\bar{X} - \mu}{\sigma} \right)^2 \sim \chi_1^2$$

independent

* If U_1, U_2, \dots, U_n are independent
 χ_1^2 (a chi-square distribution with 1 degree of freedom)

then

$$V = U_1 + U_2 + \dots + U_n$$

V : chi-square distribution with n -degrees of freedom.

$$(\chi_n^2)$$

Definition: If ① $Z \sim N(0, 1)$

② $U \sim \chi_n^2$

③ $Z \perp U$

then the distribution of

$$\frac{Z}{\left(\frac{U}{n} \right)^{1/2}}$$

t-distribution with n -degrees of freedom.

The probability density function of such t -distribution is given by:

$$f(t) = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{n\pi} \Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{t^2}{n}\right)^{-\left(\frac{n+1}{2}\right)},$$

$$-\infty < t < \infty.$$

Definition:

① $U \perp V$

② $U \sim \chi_m^2$

$V \sim \chi_n^2$

Then

$$W = \frac{\left(\frac{U}{m}\right)}{\left(\frac{V}{n}\right)} \sim$$

F-distribution,
with
(m, n)
degrees of
freedom.

Notation: $F_{m,n}$

Sample mean and Sample Variance

Let (X_1, X_2, \dots, X_n) be independent $N(\mu, \sigma^2)$ random variables.

(sample from a normal distribution)

Then,

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

called SAMPLE mean

it is a random variable.

μ : Population mean

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

called SAMPLE variance

$(\sigma^2$: Population variance)

$$E \left(\sum_{i=1}^n (X_i - \bar{X})^2 \right) = \sum_{i=1}^n E \left(X_i^2 - 2 X_i \bar{X} + \bar{X}^2 \right) = (\star)$$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

$$E(\bar{X}) = \frac{1}{n} \sum_{i=1}^n E(X_i)$$

$$= \frac{1}{n} \sum_{i=1}^n \mu$$

$$= \frac{n\mu}{n}$$

$$= \mu$$

$$\bar{X}_1, \bar{X}_2, \dots \sim N(\mu, \sigma^2)$$

$$E(X_i^2) = E(X^2)$$

$i=1, 2, \dots, n$

$$E(\bar{X}^2) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right)^2$$

$$= \frac{1}{n^2} \cdot E\left(\sum_{i=1}^n \sum_{j=1}^n X_i X_j\right)$$

$$\left(\sum_{i=1}^n a_i\right)^2 = \sum_{i=1}^n \sum_{j=1}^n a_i a_j$$

So,

$$\sum_{i=1}^n E(X_i^2) = n \cdot \frac{1}{n^2} E\left(\sum_{i=1}^n \sum_{j=1}^n X_i X_j\right)$$

\uparrow
numbers

$$= \frac{1}{n} E\left(\sum_{i=1}^n \sum_{j=1}^n X_i X_j\right)$$

$$\sum_{i=1}^n E(X; \bar{X}) = \sum_{i=1}^n E\left(X; \frac{1}{n} \sum_{j=1}^n X_j\right) \quad (1)$$

$$= \frac{1}{n} \sum_{i=1}^n \sum_{j=1}^n E(X; X_j)$$

(*) gives.

$$E\left(\sum_{i=1}^n (X_i - \bar{X})^2\right)$$

$$= n \cdot E(X^2) - \frac{2}{n} \sum_{i=1}^n \sum_{j=1}^n E(X_i X_j)$$

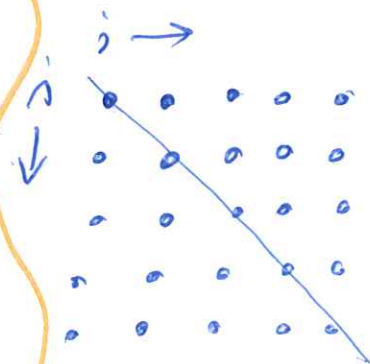
+ ~~$\frac{1}{n} \sum_{i=1}^n \sum_{j=1}^n E(X_i X_j)$~~

$$+ \frac{1}{n} \sum_{i=1}^n \sum_{j=1}^n E(X_i X_j)$$

$$= n E(X^2) - \frac{1}{n} \sum_{i=1}^n \sum_{j=1}^n E(X_i X_j)$$

$$= n E(X^2) - \frac{1}{n} \sum_{i=1}^n \sum_{\substack{j=1 \\ j \neq i}}^n E(X_i) E(X_j)$$

$$- \frac{1}{n} \sum_{i=1}^n E(X_i^2)$$



$$= n E(\bar{X}^2) - \frac{1}{n} \sum_{i=1}^n \sum_{\substack{j=1 \\ i \neq j}}^n x_i^2 - \frac{n}{n} E(X^2) \quad (12)$$

$$= n E(\bar{X}^2) - E(\bar{X}^2) - \frac{n^2}{n} (n^2 - n)$$

$$= n E(\bar{X}^2) - E(\bar{X}^2) - \mu^2 (n-1)$$

$$= (n-1) (E(\bar{X}^2) - \mu^2) = (n-1) \sigma^2$$

$$\text{So, } E\left(\sum_{i=1}^n (X_i - \bar{X})^2\right) = (n-1) \sigma^2$$

$$E\left(\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2\right) = \sigma^2$$

$$E(*) = \sigma^2$$

(Unbiased estimator of σ^2)

Sample Variance is an UNBIASED ESTIMATOR of population variance

However, we have already
seen

$$E(\bar{X}) = \mu$$

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

Sample mean is an
unbiased estimator of
population mean
