

/.

(a)

$$S(\beta) = (y - X\beta)^T (y - X\beta)$$

$$\frac{\partial S(\beta)}{\partial \beta} = \frac{\partial}{\partial \beta} \{ (y - X\beta)^T (y - X\beta) \}$$

$$= \frac{\partial}{\partial \beta} (y^T y - \beta^T X^T y - y^T X \beta + \beta^T X^T X \beta)$$

$$= \frac{\partial}{\partial \beta} (-2y^T X \beta + \beta^T X^T X \beta)$$

$$= -2X^T y + 2X^T X \beta \stackrel{!}{=} 0$$

$$X^T X \beta = X^T y$$

$$\Rightarrow (X^T X)^{-1} (X^T X) \beta = (X^T X)^{-1} X^T y$$

$$\therefore \hat{\beta} = (X^T X)^{-1} X^T y$$

(b)

$$\|y - \bar{y}\|^2 = \|y - \hat{y} + \hat{y} - \bar{y}\|^2$$

$$= \|y - \hat{y}\|^2 + \|\hat{y} - \bar{y}\|^2 + 2(y - \hat{y})^T (\hat{y} - \bar{y})$$

$$(y - \hat{y})^T (\hat{y} - \bar{y}) = (y - \hat{y})^T (X\hat{\beta} - \bar{y})$$

$$= (y - \hat{y})^T X \hat{\beta} - (y - \hat{y})^T \bar{y} \quad \text{vector}$$

$$= (y - X(X^T X)^{-1} X^T y)^T X \hat{\beta} - \bar{y} (y - \hat{y})^T \mathbf{1} \quad \text{scalar} \quad \text{vector}$$

$$= y^T (I - X(X^T X)^{-1} X^T) X \hat{\beta} - \bar{y} r^T \mathbf{1}$$

$$= y^T (X^T - X^T X (X^T X)^{-1} X^T)^T \hat{\beta} - \bar{y} \sum_{i=1}^n r_i$$

$$= y^T (X^T - X^T)^T \hat{\beta}$$

$$= 0$$

$$\bar{y} = \bar{y} \mathbf{1}$$

$$\therefore \|y - \bar{y}\|^2 = \|y - \hat{y}\|^2 + \|\hat{y} - \bar{y}\|^2$$

2.

(a)

$$\begin{aligned}
 E(\hat{\beta}_1) &= \frac{1}{S_x^2} E\left[\sum (x_i - \bar{x}_n) y_i - \cancel{\sum (x_i - \bar{x}_n) \bar{y}}^0\right] \\
 &= \frac{1}{S_x^2} E\left[\sum (x_i - \bar{x}_n) (\beta_0 + \beta_1 x_i + \varepsilon_i)\right] \\
 &= \frac{1}{S_x^2} \left[\cancel{\sum (x_i - \bar{x}_n) \beta_0}^0 + \beta_1 \sum (x_i - \bar{x}_n) x_i + \cancel{\sum (x_i - \bar{x}_n) E(\varepsilon_i)}^0\right] \\
 &= \frac{\beta_1}{S_x^2} \sum (x_i - \bar{x}_n) (x_i - \bar{x}_n + \bar{x}_n) \\
 &= \frac{\beta_1}{S_x^2} \left[\sum (x_i - \bar{x}_n)^2 + \cancel{\bar{x}_n \sum (x_i - \bar{x}_n)}^0\right] \\
 &= \beta_1
 \end{aligned}$$

$$\begin{aligned}
 \text{Var}(\hat{\beta}_1) &= \frac{1}{(S_x^2)^2} \text{Var}\left[\sum (x_i - \bar{x}_n) y_i - \cancel{\sum (x_i - \bar{x}_n) \bar{y}}^0\right] \\
 &= \frac{1}{(S_x^2)^2} \text{Var}\left[\sum (x_i - \bar{x}_n) (\beta_0 + \beta_1 x_i + \varepsilon_i)\right] \\
 &= \frac{1}{(S_x^2)^2} \text{Var}\left[\sum (x_i - \bar{x}_n) \varepsilon_i\right] \\
 &= \frac{1}{(S_x^2)^2} \sum (x_i - \bar{x}_n)^2 \text{Var}(\varepsilon_i) \quad \leftarrow \text{independent} \\
 &= \frac{\sigma^2}{S_x^2} = \frac{\sigma^2}{\sum (x_i - \bar{x}_n)^2}
 \end{aligned}$$

(b)

$$\hat{\beta}_1 = \frac{\sum (x_i - \bar{x}_n) \varepsilon_i}{S_x^2} + \frac{\sum (x_i - \bar{x}_n) (\beta_0 + \beta_1 x_i) - \sum (x_i - \bar{x}_n) \bar{y}_n}{S_x^2} \text{으로 정리할 수 있고}$$

ε_i 에 대한 선형 결합으로 볼 수 있다.

$\varepsilon_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$ 이고 정규분포의 선형 결합은 정규분포이므로 $\hat{\beta}_1$ 은 정규분포를 따른다.

3.

$$\frac{dA_n(c)}{dc} = \frac{1}{n} \sum_{i=1}^n \text{sign}(y_i - c) \stackrel{\text{set}}{=} 0$$

$\frac{dA_n(c)}{dc}$ 가 0이 되려면 $\text{sign}(y_i - c) = 1$ 인 y_i 와 $\text{sign}(y_i - c) = -1$ 인 y_i 의 개수가 같아야 한다.

y_1, \dots, y_n 을 정렬한 것은 $s_1 \leq s_2 \leq \dots \leq s_n$ 이라 하면

i) n 이 홀수인 때, $\text{sign}(y_i - c) = 1$ 인 y_i 와 $\text{sign}(y_i - c) = -1$ 인 y_i 의 개수가 같으므로

$$\sum_{i=1}^{\frac{n-1}{2}} \{ \text{sign}(s_i - c) + \text{sign}(s_{n+1-i} - c) \} = 0 \text{ 이고}$$

나머지 남은 수 $s_{\frac{n+1}{2}}$ 에 대하여 $\text{sign}(s_{\frac{n+1}{2}} - c) = 0$ 이므로 $c = s_{\frac{n+1}{2}} \Rightarrow \text{median}$ 이다.

ii) n 이 짝수인 때, $\text{sign}(y_i - c) = 1$ 인 y_i 와 $\text{sign}(y_i - c) = -1$ 인 y_i 의 개수가 같으므로

$$\sum_{i=1}^{\frac{n}{2}-1} \{ \text{sign}(s_i - c) + \text{sign}(s_{n+1-i} - c) \} = 0 \text{ 이고}$$

나머지 남은 수 $s_{\frac{n}{2}}, s_{\frac{n}{2}+1}$ 에 대하여 $\text{sign}\left(\frac{s_{\frac{n}{2}} + s_{\frac{n}{2}+1}}{2} - c\right) = 0$ 이므로 $c = \frac{s_{\frac{n}{2}} + s_{\frac{n}{2}+1}}{2} \Rightarrow \text{median}$ 이다.

4.

(a)

$$L(\theta) = \binom{n_1}{x_1} \theta^{x_1} (1-\theta)^{n_1-x_1} \times \binom{n_2}{x_2} \theta^{x_2} (1-\theta)^{n_2-x_2}$$

$$l(\theta) = \log L(\theta)$$

$$= \log \binom{n_1}{x_1} + \log \binom{n_2}{x_2} + (x_1 + x_2) \log \theta + (n_1 - x_1 + n_2 - x_2) \log (1-\theta)$$

(b)

$$\frac{\partial l(\theta)}{\partial \theta} = \frac{x_1 + x_2}{\theta} - \frac{n_1 - x_1 + n_2 - x_2}{1-\theta} \stackrel{\text{set}}{=} 0$$

$$\Rightarrow x_1 + x_2 - \theta(x_1 + x_2) = \theta(n_1 + n_2) - \theta(x_1 + x_2)$$

$$\Rightarrow x_1 + x_2 = \theta(n_1 + n_2)$$

$$\therefore \hat{\theta} = \frac{x_1 + x_2}{n_1 + n_2}$$

(c)

$$\frac{\partial^2 l(\theta)}{\partial \theta^2} = - \frac{x_1 + x_2}{\theta^2} - \frac{n_1 - x_1 + n_2 - x_2}{(1-\theta)^2}$$

$$I(\theta) = -E \left[\frac{\partial^2 l(\theta)}{\partial \theta^2} \right]$$

$$= E \left[\frac{x_1 + x_2}{\theta^2} + \frac{n_1 - x_1 + n_2 - x_2}{(1-\theta)^2} \right]$$

$$= \frac{n_1 \theta + n_2 \theta}{\theta^2} + \frac{n_1 (1-\theta) + n_2 (1-\theta)}{(1-\theta)^2}$$

$$= \frac{n_1 + n_2}{\theta} + \frac{n_1 + n_2}{1-\theta} = \frac{n_1 + n_2}{\theta(1-\theta)}$$