

Assignment 2: Data Visualization

Due beginning of class: Monday September 25

6. **Graduate students** (bonus undergraduates): Suppose we have n -dimensional real and linearly independent vectors $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_p$ and \mathbf{y} . The vector \mathbf{y} is the sum of two n -dimensional real vectors μ and \mathbf{r}

$$\mathbf{y} = \mu + \mathbf{r}$$

where μ is restricted to be a linear combination of the vectors $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_p$. That is

$$\mu = \theta_1 \times \mathbf{x}_1 + \theta_2 \times \mathbf{x}_2 + \dots + \theta_p \times \mathbf{x}_p$$

for some unknown real constants $\theta_1, \theta_2, \dots, \theta_p$, or equivalently

$$\mu = \mathbf{X}\theta$$

where $\mathbf{X} = [\mathbf{x}_1, \dots, \mathbf{x}_p]$ is an $n \times p$ matrix and $\theta = (\theta_1, \theta_2, \dots, \theta_p)^T$ is a $p \times 1$ vector.

- (a) (5 marks) For any \mathbf{y} , neither μ nor \mathbf{r} are uniquely defined. Suppose we choose particular vectors $\hat{\mu}$, and $\hat{\mathbf{r}}$ (with $\mathbf{y} = \hat{\mu} + \hat{\mathbf{r}}$) to be such that they are orthogonal to one another (whatever values any θ_i take). That is, $\hat{\mu}^T \hat{\mathbf{r}} = 0$.

Prove that this additional constraint implies that

$$\hat{\mu} = \mathbf{P}\mathbf{y}$$

where $\mathbf{P} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$ and hence show that $\hat{\mathbf{r}} = (\mathbf{I}_n - \mathbf{P})\mathbf{y}$.

$\hat{\mu}^T \hat{\mathbf{r}} = 0$ (whatever values any θ_i take) $\implies \hat{\mathbf{r}}$ is orthogonal to $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_p$. Hence:

$$\begin{aligned} \mathbf{X}^T \hat{\mathbf{r}} &= \begin{bmatrix} \mathbf{x}_1^T \hat{\mathbf{r}} \\ \dots \\ \mathbf{x}_p^T \hat{\mathbf{r}} \end{bmatrix} \\ &= \mathbf{0} \\ \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \hat{\mathbf{r}} &= \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{0} \\ \mathbf{P} \hat{\mathbf{r}} &= \mathbf{0} \\ \mathbf{P}\mathbf{y} - \mathbf{P} \hat{\mathbf{r}} &= \mathbf{P}\mathbf{y} \\ \mathbf{P} \hat{\mu} &= \mathbf{P}\mathbf{y} \end{aligned}$$

But we can simplify $\mathbf{P}\mathbf{u}$ to:

$$\begin{aligned} \mathbf{P}\mu &= \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mu \\ &= \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{X} \theta \\ &= \mathbf{X} \mathbf{I}_p \theta \\ &= \mathbf{X} \theta \\ &= \hat{\mu} \end{aligned}$$

$\therefore \hat{\mu} = \mathbf{P}\mathbf{y}$. For the second proof:

$$\begin{aligned} \mathbf{P}\mathbf{y} &= \hat{\mu} \\ &= \mathbf{y} - \hat{\mathbf{r}} \\ \hat{\mathbf{r}} &= \mathbf{y} - \mathbf{P}\mathbf{y} \\ &= (\mathbf{I}_n - \mathbf{P})\mathbf{y} \end{aligned}$$

- (b) (2 marks) Show that \mathbf{P} is an idempotent matrix, that is that $\mathbf{P}^2 = \mathbf{P}$.

$$\begin{aligned} \mathbf{P}^2 &= \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \\ &= \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{I}_p \mathbf{X}^T \\ &= \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \\ &= \mathbf{P} \end{aligned}$$

$\therefore \mathbf{P}$ is an idempotent matrix.

(c) (2 marks) Show that if \mathbf{P} is an idempotent matrix, then so must be $(\mathbf{I}_n - \mathbf{P})$.
 Assume that \mathbf{P} is an idempotent matrix.

$$\begin{aligned}
 (\mathbf{I}_n - \mathbf{P})^2 &= (\mathbf{I}_n - \mathbf{P})(\mathbf{I}_n - \mathbf{P}) \\
 &= \mathbf{I}_n^2 - \mathbf{I}_n\mathbf{P} - \mathbf{P}\mathbf{I}_n + \mathbf{P}^2 \\
 &= \mathbf{I}_n - \mathbf{P} - \mathbf{P} + \mathbf{P} \quad \text{since } \mathbf{P} \text{ is idempotent} \\
 &= \mathbf{I}_n - \mathbf{P}
 \end{aligned}$$

$\therefore (\mathbf{I}_n - \mathbf{P})$ is an idempotent matrix.

(d) (2 marks) Show that $\hat{\mathbf{r}}$ is in fact orthogonal to $\hat{\mu}$.

We first show that \mathbf{P} is symmetric:

$$\begin{aligned}
 \mathbf{P}^T &= (\mathbf{X}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T)^T \\
 &= (\mathbf{X}^T)^T(\mathbf{X}(\mathbf{X}^T\mathbf{X})^{-1})^T \\
 &= \mathbf{X}((\mathbf{X}^T\mathbf{X})^{-1})^T\mathbf{X}^T \\
 &= \mathbf{X}((\mathbf{X}^T\mathbf{X})^T)^{-1}\mathbf{X}^T \\
 &= \mathbf{X}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T \\
 &= \mathbf{P}
 \end{aligned}$$

$\therefore \mathbf{P}$ is symmetric.

$$\begin{aligned}
 \hat{\mu}^T \hat{\mathbf{r}} &= \hat{\mu} \cdot \hat{\mathbf{r}} \\
 &= \mathbf{P}\mathbf{y} \cdot (\mathbf{I}_n - \mathbf{P})\mathbf{y} \\
 &= \mathbf{y} \cdot \mathbf{P}(\mathbf{I}_n - \mathbf{P})\mathbf{y} \\
 &= \mathbf{y}^T(\mathbf{P} - \mathbf{P}^2)\mathbf{y} \\
 &= \mathbf{y}^T\mathbf{0}\mathbf{y} \\
 &= \mathbf{0}
 \end{aligned}$$

$\therefore \hat{\mu}$ is orthogonal to $\hat{\mathbf{r}}$.

(e) (5 marks) Show that $\hat{\mu} = \mathbf{P}\mathbf{y}$ is the choice of μ which minimizes the squared length of \mathbf{r} . That is it minimizes $\mathbf{r}^T\mathbf{r} = (\mathbf{y} - \mu)^T(\mathbf{y} - \mu)$.