Supplement: Disaster Response on September 11, 2001 Through the Lens of Statistical Network Analysis

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We are interested in maximising the lower bound on $\log P(\mathbf{Y} = \mathbf{y})$ given by (23). We note that, despite of the fact that the family of auxiliary distributions is fully factorised, the lower bound on $\log P(\mathbf{Y} = \mathbf{y})$ is intractable, because

$$\int \left[\log \frac{P(\mathbf{Y} = \mathbf{y} \mid \boldsymbol{\delta}) p(\boldsymbol{\delta})}{q(\boldsymbol{\delta} \mid \boldsymbol{\vartheta})} \right] q(\boldsymbol{\delta} \mid \boldsymbol{\vartheta}) d\boldsymbol{\delta} = \sum_{i < j}^{n} E_{\boldsymbol{\vartheta}} \left[\log P(Y_{ij} = y_{ij} \mid \boldsymbol{\delta}) \right] + E_{\boldsymbol{\vartheta}} \left[\log p(\boldsymbol{\delta}) \right] - E_{\boldsymbol{\vartheta}} \left[\log q(\boldsymbol{\delta} \mid \boldsymbol{\vartheta}) \right], \tag{27}$$

where

$$E_{\mathbf{\vartheta}}[\log P(Y_{ij} = y_{ij} | \boldsymbol{\delta})] = \left[\sum_{k=1}^{K} \vartheta_{Z_{ik}} \vartheta_{\theta_{k}, 1} + \sum_{l=1}^{K} \vartheta_{Z_{jl}} \vartheta_{\theta_{l}, 1} \right] y_{ij}$$

$$+ \sum_{k=1}^{K} \sum_{l=1}^{K} \vartheta_{Z_{ik}} \vartheta_{Z_{jl}} E_{\mathbf{\vartheta}} \left\{ -\log \left[1 + \exp(\theta_{k} + \theta_{l}) \right] \right\}$$

$$(28)$$

includes the intractable expectations $E_{\vartheta} \{ -\log [1 + \exp(\theta_k + \theta_l)] \}$. To obtain a tractable lower bound on $\log P(\mathbf{Y} = \mathbf{y})$, we lower bound $E_{\vartheta} \{ -\log [1 + \exp(\theta_k + \theta_l)] \}$ by Jensen's inequality:

$$E_{\vartheta} \left\{ -\log \left[1 + \exp(\theta_k + \theta_l) \right] \right\} \ge -\log E_{\vartheta} \left[1 + \exp(\theta_k + \theta_l) \right]$$

$$= -\log \left[1 + \exp(\vartheta_{\theta_k, 1} + \frac{1}{2}\vartheta_{\theta_k, 2} + \vartheta_{\theta_l, 1} + \frac{1}{2}\vartheta_{\theta_l, 2}) \right].$$
(29)

The lower bound on $\log P(\mathbf{Y} = \mathbf{y})$ given by (27), (28), and (29) can be maximised with respect to $\boldsymbol{\vartheta}$ by coordinate ascent algorithms.

The lower bound on $\log P(\mathbf{Y} = \mathbf{y})$ can be maximised with respect to $\vartheta_{\alpha,1}$, $\vartheta_{\alpha,2}$, $\vartheta_{\mu,1}$, $\vartheta_{\mu,2}$, $\vartheta_{\sigma^{-2},1}$, $\vartheta_{\sigma^{-2},2}$, $\vartheta_{V_k,1}$, $\vartheta_{V_k,2}$, $\vartheta_{Z_{ik}}$ as follows:

$$\vartheta_{\alpha,1} = A_1 + K - 1, \tag{30}$$

$$\vartheta_{\alpha,2} = B_1 - \sum_{k=1}^{K-1} \left[\Psi(\vartheta_{V_k,2}) - \Psi(\vartheta_{V_k,1} + \vartheta_{V_k,2}) \right], \tag{31}$$

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$$\vartheta_{\mu,1} = \frac{MS^{-2} + E_{\vartheta}[\sigma^{-2}] \sum_{k=1}^{K} E_{\vartheta}[\theta_{k}]}{S^{-2} + K E_{\vartheta}[\sigma^{-2}]},$$
(32)

$$\vartheta_{\mu,2} = \frac{1}{S^{-2} + K E_{\eta}[\sigma^{-2}]},$$
 (33)

$$\vartheta_{\sigma^{-2},1} = A_2 + \frac{K}{2}, \tag{34}$$

$$\vartheta_{\sigma^{-2},2} = B_2 + \sum_{k=1}^{K} \frac{\vartheta_{\theta_k,1}^2 - 2\,\vartheta_{\theta_k,1}\vartheta_{\mu,1} + \vartheta_{\theta_k,2}}{2} + \frac{K(\vartheta_{\mu,1}^2 + \vartheta_{\mu,2})}{2},\tag{35}$$

$$\vartheta_{V_k,1} = 1 + \sum_{i=1}^n \vartheta_{Z_{ik}}, \ k = 1, \dots, K - 1,$$
 (36)

$$\vartheta_{V_k,2} = E_{\vartheta}[\alpha] + \sum_{i=1}^{n} \sum_{l=k+1}^{K} \vartheta_{Z_{il}}, \ k = 1, \dots, K-1,$$
 (37)

$$\log \vartheta_{Z_{ik}} \propto \Psi(\vartheta_{V_k,1}) - \Psi(\vartheta_{V_k,1} + \vartheta_{V_k,2}) + \sum_{l=1}^{k-1} \left[\Psi(\vartheta_{V_l,2}) - \Psi(\vartheta_{V_l,1} + \vartheta_{V_l,2}) \right]$$

$$+ \sum_{j\neq i}^{n} \vartheta_{\theta_k,1} y_{ij} - \sum_{j\neq i}^{n} \sum_{l=1}^{K} \vartheta_{Z_{jl}} \log \left[1 + \exp(\vartheta_{\theta_k,1} + \frac{1}{2}\vartheta_{\theta_k,2} + \vartheta_{\theta_l,1} + \frac{1}{2}\vartheta_{\theta_l,2}) \right],$$

$$(38)$$

where $\Psi(.)$ denotes the digamma function and the expectations $E_{\vartheta}[\alpha]$, $E_{\vartheta}[\theta_k]$, and $E_{\vartheta}[\sigma^{-2}]$ under the auxiliary distribution indexed by ϑ are tractable.

The lower bound on log $P(\mathbf{Y} = \mathbf{y})$ with respect to $\theta_{\theta_k,1}$ and $\theta_{\theta_k,2}$ can be maximised by maximising

$$-\sum_{k=1}^{K} \frac{\vartheta_{\sigma^{-2},1}}{2\vartheta_{\sigma^{-2},2}} \left[\vartheta_{\theta_{k},1}^{2} - 2 \vartheta_{\theta_{k},1} \vartheta_{\mu,1} + \vartheta_{\theta_{k},2} \right] + \frac{1}{2} \sum_{k=1}^{K} \log \vartheta_{\theta_{k},2}$$

$$+\sum_{i < j}^{n} \left[\sum_{k=1}^{K} \vartheta_{Z_{ik}} \vartheta_{\theta_{k},1} + \sum_{l=1}^{K} \vartheta_{Z_{jl}} \vartheta_{\theta_{l},1} \right] y_{ij}$$

$$-\sum_{i < j}^{n} \sum_{k=1}^{K} \sum_{l=1}^{K} \vartheta_{Z_{ik}} \vartheta_{Z_{jl}} \log \left[1 + \exp(\vartheta_{\theta_{k},1} + \frac{1}{2} \vartheta_{\theta_{k},2} + \vartheta_{\theta_{l},1} + \frac{1}{2} \vartheta_{\theta_{l},2}) \right].$$
(39)

Since $\vartheta_{\theta_k,2}$ are the precision parameters of the marginal auxiliary distributions of θ_k , the constraints $\vartheta_{\theta_k,2} > 0$ must be respected (k = 1, ..., K). Two possible approaches are constrained maximisation subject to the constraints $\vartheta_{\theta_k,2} > 0$ and unconstrained maximisation by reparameterising $\rho_{\theta_k,2} : \vartheta_{\theta_k,2} \mapsto \log \vartheta_{\theta_k,2}$ and maximising with respect to $\rho_{\theta_k,2}$ (k = 1, ..., K).