Introduction



The story of the Donner-Reed party is one of the more tragic stories of the American Pioneers. The story is of a group of families and individuals who migrated to California on a wagon train from Midwestern America. The group initially traveled via the Oregon Trail but then opted to take the Hastings Cutoff- a route which was never traveled with wagon trains. While there was some elements of luck involved, there are some interesting characteristics which can be used to model the survival outcome for a given member of the party. While American history is by no means my domain knowledge, I got some more context from the Weird History channel's video and the History Channel's blog on the topic.

In this blog post I explore the <code>Donner</code> data set from the <code>vcdExtra</code> package and develop a logistic regression model. A initial model is proposed and refined by using backwards elimination and is diagnosed at every step. Finally we compare all the models made with the main effects model and come up with a model of choice.

Initial Questions

If there is no meaningful questions being posed before exploring our data set, it will be harder to come up with insights. This is why before getting into any EDA or modeling, I'm going to ask the following questions:

- 1. What is the relationship between a individual's sex, age and family with survival?
- 2. How does the interaction between these 3 variables relate to survival.

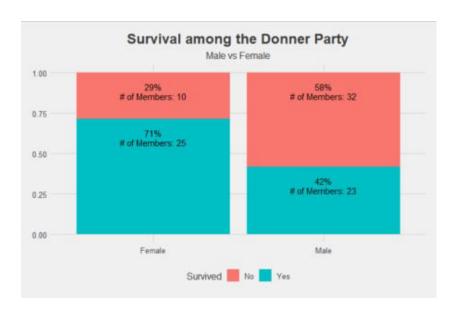
As basic as these questions are, these questions will guide the rest of the analysis into model development.

Exploratory data analysis

Because there is no major issue of missing data in this data set, the check I usually do for missing data will be ignored. Lets now look at the relationship of sex, age and family of a given individual in the Donner-Reed party and their survival. Because this is a blog post, all the R code will be included in-line with the output it produces

Individual Sex and Survival

```
library(vcdExtra)
library(tidyverse)
library(ggthemes)
sexSurvival <-Donner %>% group by(sex,survived) %>%
  summarize(count=n()) %>%
 mutate(pct=count/sum(count),Count=count)
ggplot(data=sexSurvival,aes(x=sex,y=pct,fill=as.factor(survived)))+
  theme fivethirtyeight()+
 geom_bar(stat="identity",position="fill")+
 geom text(mapping= aes(label=scales::percent(pct)),position="stack",
vjust=+2.5)+
  geom text(mapping= aes(label=paste("# of Members:",Count)),position="stack",
vjust=+3.8)+
  labs(title="Survival among the Donner Party", subtitle = "Male vs Female")+
  scale_fill_manual(name="Survived",
                    values=c("#F8766D","#00BFC4"),
                    labels=c("No","Yes"))+
  theme(plot.title = element text(hjust=0.5),plot.subtitle =
element text(hjust=0.5))
```



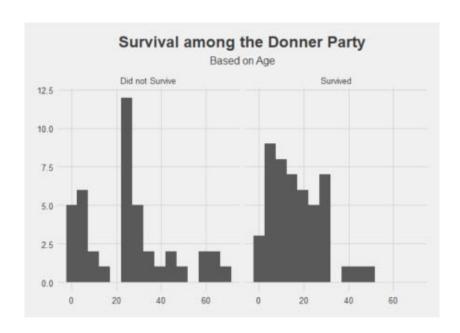
As expected for the time (chivalry didn't die yet) women had ha higher probability of survival than men in the Donner-Reed party.

Individual Age and Suvival

There are two ways age can be viewed, it can either be viewed as a continuous variable or categorically by splitting age in ranges. Lets first see what the raw data can tell us.

```
# Aesthetics
Donner$SurvivedFactor<-as.factor(Donner$survived)
levels(Donner$SurvivedFactor)<-c("Did not Survive","Survived")
# The plot
ggplot(data=Donner, aes(x=age))+
    theme_fivethirtyeight()+
    geom_histogram(bins=15)+
    facet wrap(~SurvivedFactor)+</pre>
```

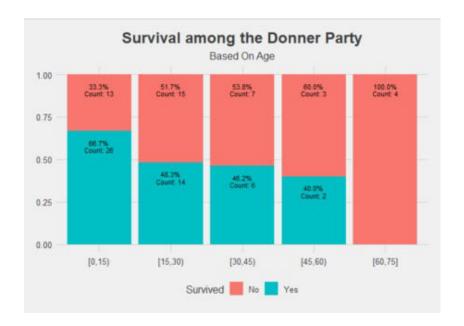
```
labs(title="Survival among the Donner Party", subtitle = "Based on Age") +
   theme(plot.title = element_text(hjust=0.5), plot.subtitle =
element_text(hjust=0.5))
```



While there preference for survival to be among individuals under 20, there does not seems to be a clear picture with this diagnostic in terms of exploratory data analysis. Lets try grouping our ages into 14 year ranges:

```
Donner$AgeRange<-cut(Donner$age,breaks=c(0,15,30,45,60,75),include.lowest =
TRUE,right=FALSE)</pre>
```

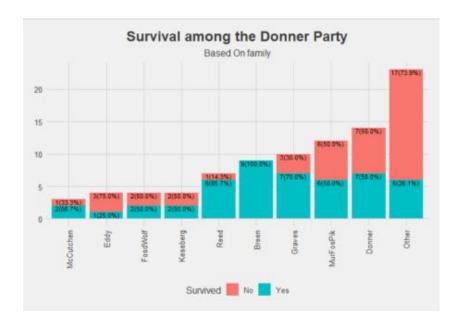
```
AgeRangeProp<-Donner %>% group by (AgeRange, survived) %>%
  summarize(count=n()) %>%
 mutate(pct=count/sum(count),Count=count)
ggplot(data=AgeRangeProp, mapping=aes(x=AgeRange, y=pct,
fill=as.factor(survived)))+
 theme fivethirtyeight()+
 geom bar(stat="identity",position="fill")+
  geom text(mapping= aes(label=scales::percent(pct)),position="stack",size=3,
vjust=+2.5)+
  geom text(mapping= aes(label=paste("Count:",Count)),position="stack",size=3,
vjust=+3.8)+
  scale fill manual(name="Survived",
                    values=c("#F8766D","#00BFC4"),
                    labels=c("No","Yes"))+
  labs(title="Survival among the Donner Party", subtitle = "Based On Age")+
  theme(plot.title = element text(hjust=0.5),plot.subtitle =
element text(hjust=0.5))
```



From looking at the data like this, we can see a clearer picture of the effect of age. A given individual in the Donner-Reed party who was under 15 seemed to have had the best chance of survival, with a jump down in chances of survival if they were 15 and older. Interpretation of age as a range thus provides a more intuitive insight.

Survival among Families

```
FamilyProp<-Donner %>% group by (family, survived) %>%
  summarize(count=n()) %>%
  mutate(pct=count/sum(count),Count=count) %>%
  group by (family, count)
FamilyOrder<-aggregate(FamilyProp$count,by=list(Category=FamilyProp$family),
FUN=sum) %>% group by(x)
FamilyProp<-left join(FamilyProp,FamilyOrder,by=c("family"="Category"))</pre>
# Taken from
FamilyProp$family<-reorder(FamilyProp$family,FamilyProp$x)</pre>
ggplot(data=FamilyProp, mapping=aes(x=family, y=count,
fill=as.factor(survived)))+
  theme fivethirtyeight()+
  geom bar(stat="identity", position="stack") +
  scale fill manual(name="Survived",
                    values=c("#F8766D","#00BFC4"),
                    labels=c("No","Yes"))+
  geom text(mapping= aes(label=paste0(Count,"(",scales::percent(pct),")")),
            position="stack", size=3, vjust=1)+
  labs(title="Survival among the Donner Party", subtitle = "Based On family")+
  theme(plot.title = element text(hjust=0.5),
        plot.subtitle = element text(hjust=0.5),
        axis.text.x = element text(angle = 90, vjust = 0.5, hjust=1))
```



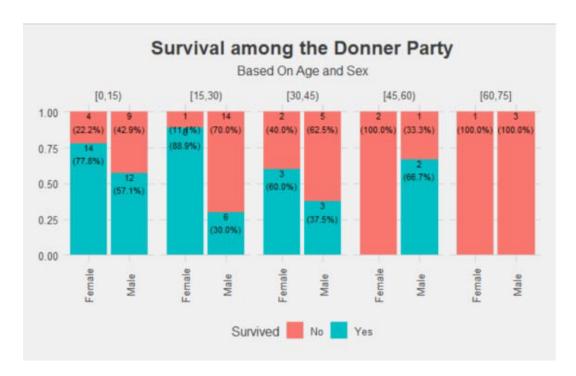
Besides for the Breen family, everyone in the Donner party lost at least one member of their family. In fact, its possible to assign probabilities of survival based on the family for a given individual of the Donner-Reed party. It would be a misstep not to consider this characteristic in a predictive model.

Interaction betweeen Sex and Age

There are two ways age can be viewed- in its raw form as a "continuous" variable or as a categorical variable. To offer a comparison we will make visuals of both of them.

Age as a Categorical Variable

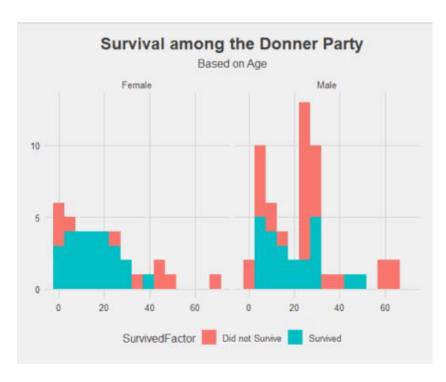
```
sexAgeRangeDf<- Donner %>% group by(sex,AgeRange,survived) %>%
  summarize(count=n()) %>%
  mutate(pct=count/sum(count),Count=count)
ggplot(data=sexAgeRangeDf,mapping=aes(x=sex,y=pct,fill=as.factor(survived)))+
  theme fivethirtyeight()+
  geom bar(stat="identity", position="fill") +
  scale fill manual(name="Survived",
                      values=c("#F8766D","#00BFC4"),
                      labels=c("No","Yes"))+
  {\tt geom\_text} \; ({\tt mapping= aes} \; ({\tt label=paste0} \; ({\tt Count,"} \\ {\tt n(",scales::percent(pct),")"}) \; ) \; , \\
             position="stack", size=3, vjust=1) +
  facet grid(~AgeRange)+
  labs(title="Survival among the Donner Party", subtitle = "Based On Age and
Sex") +
  theme(plot.title = element text(hjust=0.5),
        plot.subtitle = element text(hjust=0.5),
        axis.text.x = element text(angle = 90, vjust = 0.5))
```



When looking at age as a categorical variable, we note that women overall had a higher probability of survival than men. However, women who were 15-29 years old had the highest probability of survival. With men- those whom were under 15 had the highest probability of survival. This once again appears to push the idea that women and children generally were prioritized for survival.

Age as a Continuos variabe

```
ggplot(data=Donner, aes(x=age,fill=SurvivedFactor))+
  theme_fivethirtyeight()+
  geom_histogram(bins=15)+
  facet_wrap(~sex)+
  labs(title="Survival among the Donner Party",subtitle = "Based on Age")+
  theme(plot.title = element_text(hjust=0.5),plot.subtitle =
  element text(hjust=0.5))
```

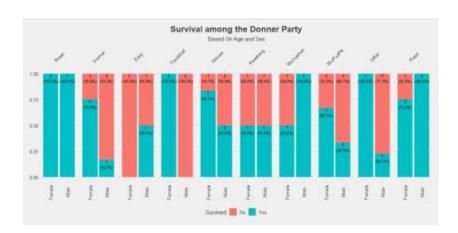


We see that same idea when viewing age as a continuous variable. Because age as a continuous variable

gives a finer control in the model, age will be treated as a continuous variable in the modelling step.

Interaction between Sex and family

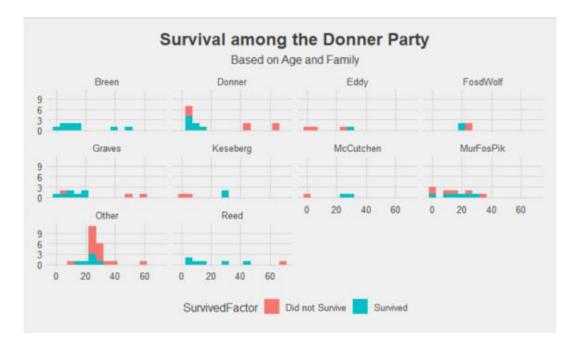
```
sexFamilyDf<- Donner %>% group by(sex,family,survived) %>%
  summarize(count=n()) %>%
 mutate(pct=count/sum(count),Count=count)
ggplot(data=sexFamilyDf, mapping=aes(x=sex, y=pct, fill=as.factor(survived)))+
 theme fivethirtyeight()+
 geom bar(stat="identity", position="fill") +
  scale fill manual(name="Survived",
                    values=c("#F8766D","#00BFC4"),
                    labels=c("No","Yes"))+
 geom text(mapping= aes(label=paste0(Count,"\n(",scales::percent(pct),")")),
            position="stack", size=3, vjust=1) +
  facet grid(~family)+
  labs(title="Survival among the Donner Party", subtitle = "Based On Age and
Sex")+
  theme(plot.title = element_text(hjust=0.5),
        plot.subtitle = element text(hjust=0.5),
        axis.text.x = element text(angle = 90, vjust = 0.5),
        strip.text.x = element text(angle=45))
```



The Eddy, Keseberg, McCutchen and families had more women die in their families than men, while the other families had more men die.

Interaction between age and family

```
ggplot(data=Donner, aes(x=age,fill=SurvivedFactor))+
   theme_fivethirtyeight()+
   geom_histogram(bins=15)+
   facet_wrap(~family)+
   labs(title="Survival among the Donner Party",subtitle = "Based on Age and Family")+
   theme(plot.title = element_text(hjust=0.5),plot.subtitle = element_text(hjust=0.5))
```

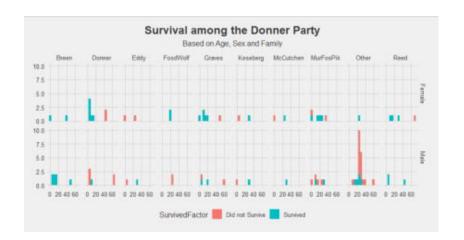


We see that overall, people who were in their "prime" (20-40) or younger generally had a better chance of survival than those who were older.

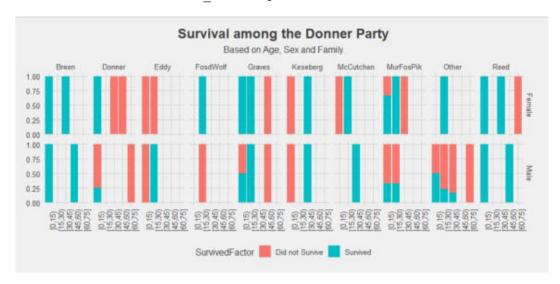
Interaction between sex, age and family.

Because the visual using age as a continuous variable is more unwieldy, I have included plots treat age as a continuous and categorical variable.

```
ggplot(data=Donner, aes(x=age,fill=SurvivedFactor))+
   theme_fivethirtyeight()+
   geom_histogram(bins=15)+
   facet_grid(sex~family)+
   labs(title="Survival among the Donner Party",subtitle = "Based on Age, Sex and Family")+
   theme(plot.title = element_text(hjust=0.5),plot.subtitle = element_text(hjust=0.5))
```



```
ggplot(data=Donner, aes(x=AgeRange,fill=SurvivedFactor))+
    theme_fivethirtyeight()+
    geom_bar(position="fill")+
    facet_grid(sex~family)+
    labs(title="Survival among the Donner Party",subtitle = "Based on Age, Sex and Family")+
    theme(plot.title = element_text(hjust=0.5),plot.subtitle = element_text(hjust=0.5),
```



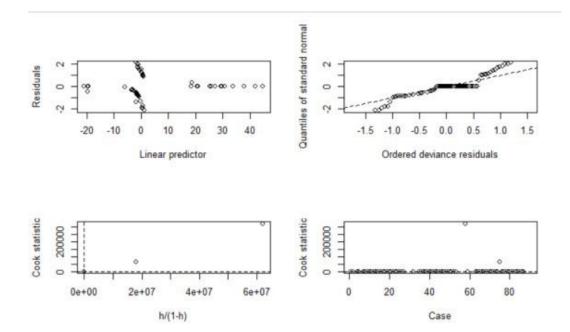
With the three way interaction of age, sex and family, its harder to see any particular trends present. In the modeling step we will consider the three way interaction model and preform some diagnostics to see if we want it.

Constructing a model

The model that we are choosing to use is (obviously) logistic regression. This is because the response we are trying to model is binary (i.e. Survival). Below are the models and their diagnostics.

Model 1: The full (3-way interaction) model.

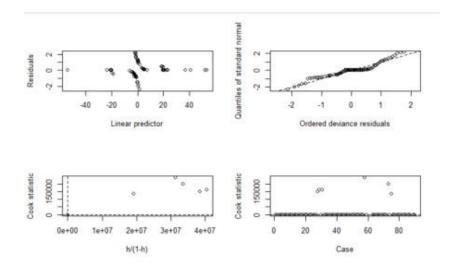
```
library(car)
library(boot)
fit<-glm(survived~family*age*sex, family=binomial(link="logit"),data=Donner)</pre>
fit diag <- glm.diag(fit)</pre>
Anova(fit)
## Analysis of Deviance Table (Type II tests)
##
## Response: survived
##
                LR Chisq Df Pr(>Chisq)
                 24.1364 9 0.004091 **
## family
                  5.6049 1 0.017910 *
## age
                 10.4845 1 0.001204 **
## sex
                 25.5736 9 0.002398 **
## family:age
## family:sex
                  2.5283 9 0.980112
## age:sex
                   2.7518 1
                               0.097142 .
## family:age:sex 1.0224 7 0.994459
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
glm.diag.plots(fit, fit diag)
```



This model considers all the effects of age family and sex. Based on the Anova output we can already see that the 3-way interaction is not significant. Based on residual (top left) and QQ-Plot (top right) we see that the ordered deviance residuals do not behave well. While the response is binary, clearly the model can be improved.

Model 2: the 2-way interaction model.

```
fit2<-glm(survived~family*age+family*sex+age*sex,</pre>
family=binomial(link="logit"),data=Donner)
fit2_diag <- glm.diag(fit2)</pre>
Anova(fit2)
## Analysis of Deviance Table (Type II tests)
##
## Response: survived
##
              LR Chisq Df Pr(>Chisq)
                24.1364
                         9
                             0.004091 **
## family
## age
                 5.6049
                         1
                             0.017910 *
## sex
                10.4845
                        1
                             0.001204 **
               25.5736
                             0.002398 **
## family:age
                         9
## family:sex
                 2.5283
                         9
                             0.980112
## age:sex
                 2.7518
                             0.097142 .
                         1
## ---
## Signif. codes:
                    0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
glm.diag.plots(fit2, fit2 diag)
```



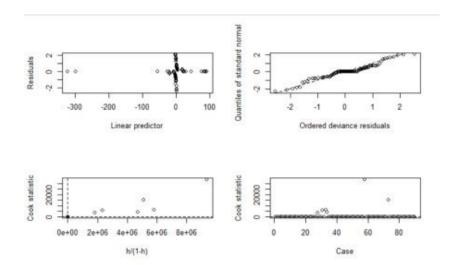
Looking at the 2-way interaction model, we see that there is an improvement in the residuals. However from the Anova model we see that there are still insignificant interaction variables that can be removed from the model.

Model 3: Considering only sigificant effects

While respecting the principal of marginality a model was selected with considering only the significant (p-value < 0.05) effects.

```
fit3<-glm(survived~family*age+sex, family=binomial(link="logit"),data=Donner)</pre>
fit3 diag <- glm.diag(fit3)</pre>
Anova(fit3)
## Analysis of Deviance Table (Type II tests)
##
## Response: survived
              LR Chisq Df Pr(>Chisq)
## family
                22.199 9
                            0.008268 **
## age
                 5.398 1
                            0.020162 *
                10.485 1
                            0.001204 **
## sex
                37.097 9 2.529e-05 ***
## family:age
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

glm.diag.plots(fit3, fit3 diag)

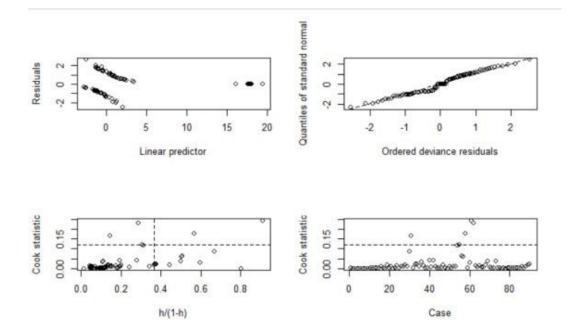


While the residuals look much better, there is still flatness present in the QQ-Plot. Lets try finally look at the main effects to see if the residuals will be imporved.

Model 4: The main-effects model.

Anova(fit4)

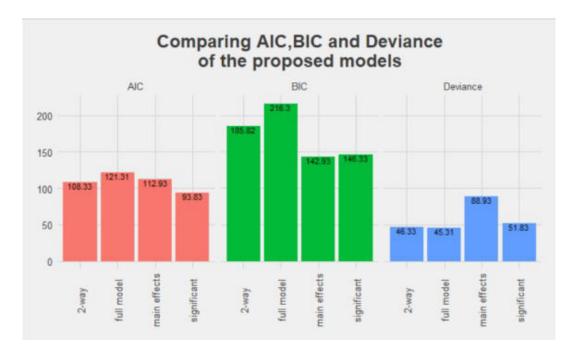
```
## Analysis of Deviance Table (Type II tests)
##
## Response: survived
         LR Chisq Df Pr(>Chisq)
## family 22.1993 9
                        0.008268 **
            5.3978 1
                        0.020162 *
## age
## sex
           5.2261
                   1
                        0.022250 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
glm.diag.plots(fit4, fit4 diag)
```



Out of all the models, the main effects model seems to preform the best in terms of residuals.

Comparing Deviance, AIC and BIC of models

```
library(reshape2)
diagnosticDf<-tibble(</pre>
  model=c("full model", "2-way", "significant", "main effects"),
 AIC = c(AIC(fit), AIC(fit2), AIC(fit3), AIC(fit4)),
  BIC = c(BIC(fit), BIC(fit2), BIC(fit3), BIC(fit4)),
  Deviance=c(deviance(fit), deviance(fit2), deviance(fit3), deviance(fit4))
                      ) %>% melt()
ggplot(data=diagnosticDf,mapping = aes(x=model,y=value,fill=variable))+
  theme_fivethirtyeight()+
  geom bar(position="dodge", stat="identity") +
  geom text(mapping= aes(label=round(value,2)),
            position="stack", size=3, vjust=1) +
  facet wrap(~variable)+
  labs(title="Comparing AIC,BIC and Deviance\nof the proposed models")+
  theme(plot.title = element text(hjust=0.5),
        legend.position = 'none',axis.text.x = element_text(angle = 90, vjust
= 0.5)
```



What the main effects model makes up with in terms of residuals, it suffers in terms of AIC and deviance. While it preforms better with AIC when compared to the full model. It does not do better than the 2-way (model 2) or significant effects models (model 3). In terms of deviance, the main effects model fairs the worst. But from the QQ and residual plots, the main effects model also seems to provide the best goodness-of-fit.

With all this information the best model out of the four proposed models is the main effects model (let me know if you disagree!).

Conclusion

From this blog its clear that there's more than meets the eye for constructing a logistic regression model. Even after asking pointed questions about the data- diagnostics need to be done to determine the validity of the model. It could be that the simplest model is what best describes the data- but determining that can only be done with applying some diagnostics.