

Cochran Theorem – from *The distribution of quadratic forms in a normal system, with applications to the analysis of covariance* published in 1934 – is probably the most important one in a regression course. It is an application of a nice result on quadratic forms of Gaussian vectors. More precisely, we can prove that if $\mathbf{Y} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$ is a random vector with d independent standard normal variables then (i) if A is a (squared) idempotent matrix $\mathbf{Y}^{\top} A \mathbf{Y} \sim \chi^2_r$ where r is the rank of matrix A , and (ii) conversely, if $\mathbf{Y}^{\top} A \mathbf{Y} \sim \chi^2_r$ then A is an idempotent matrix of rank r . And just in case, A is an idempotent matrix means that $A^2 = A$, and a lot of results can be derived (for instance on the eigenvalues). The proof of that result (at least the (i) part) is nice: we diagonalize matrix A , so that $A = P \Delta P^{\top}$, with P orthonormal. Since A is an idempotent matrix observe that $A^2 = P \Delta P^{\top} P \Delta P^{\top} = P \Delta^2 P^{\top}$ where Δ is some diagonal matrix such that $\Delta^2 = \Delta$, so terms on the diagonal of Δ are either 0 or 1's. And because the rank of A (and Δ) is r then there should be r 1's and $d-r$ 0's. Now write $\mathbf{Y}^{\top} A \mathbf{Y} = \mathbf{Y}^{\top} P \Delta P^{\top} \mathbf{Y} = \mathbf{Z}^{\top} \Delta \mathbf{Z}$ where $\mathbf{Z} = P^{\top} \mathbf{Y}$ that satisfies $\mathbf{Z} \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)$. Thus $\mathbf{Z}^{\top} \Delta \mathbf{Z} = \sum_{i: \Delta_{ii}=1} Z_i^2 \sim \chi^2_r$. Nice, isn't it. And there is more (that will be strongly connected actually to Cochran theorem). Let $A = A_1 + \dots + A_k$, then the two following statements are equivalent (i) A is idempotent and $\text{rank}(A) = \text{rank}(A_1) + \dots + \text{rank}(A_k)$ (ii) A_i 's are idempotents, $A_i A_j = 0$ for all $i \neq j$.

Now, let us talk about projections. Let \mathbf{y} be a vector in \mathbb{R}^n . Its projection on the space $V(\mathbf{v}_1, \dots, \mathbf{v}_p)$ (generated by those p vectors) is the vector $\hat{\mathbf{y}} = \mathbf{V} \hat{\mathbf{a}}$ that minimizes $\|\mathbf{y} - \mathbf{V} \mathbf{a}\|$ (in \mathbf{V}). The solution is $\hat{\mathbf{a}} = (\mathbf{V}^{\top} \mathbf{V})^{-1} \mathbf{V}^{\top} \mathbf{y}$ and $\hat{\mathbf{y}} = \mathbf{V} \hat{\mathbf{a}}$. Matrix $P = \mathbf{V} (\mathbf{V}^{\top} \mathbf{V})^{-1} \mathbf{V}^{\top}$ is the orthogonal projection on $V(\mathbf{v}_1, \dots, \mathbf{v}_p)$ and $\hat{\mathbf{y}} = P \mathbf{y}$.

Now we can recall Cochran theorem. Let $\mathbf{Y} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}_d)$ for some $\sigma > 0$ and \mathbf{I}_d . Consider sub-vector orthogonal spaces F_1, \dots, F_m , with dimension d_i . Let P_{F_i} be the orthogonal projection matrix on F_i , then (i) vectors $P_{F_1} \mathbf{X}, \dots, P_{F_m} \mathbf{X}$ are independent, with respective distribution $\mathcal{N}(P_{F_i} \mathbf{0}, \sigma^2 \mathbf{I}_{d_i})$ and (ii) random variables $\|P_{F_i} \mathbf{X}\|^2 / \sigma^2$ are independent and $\chi^2_{d_i}$ distributed.

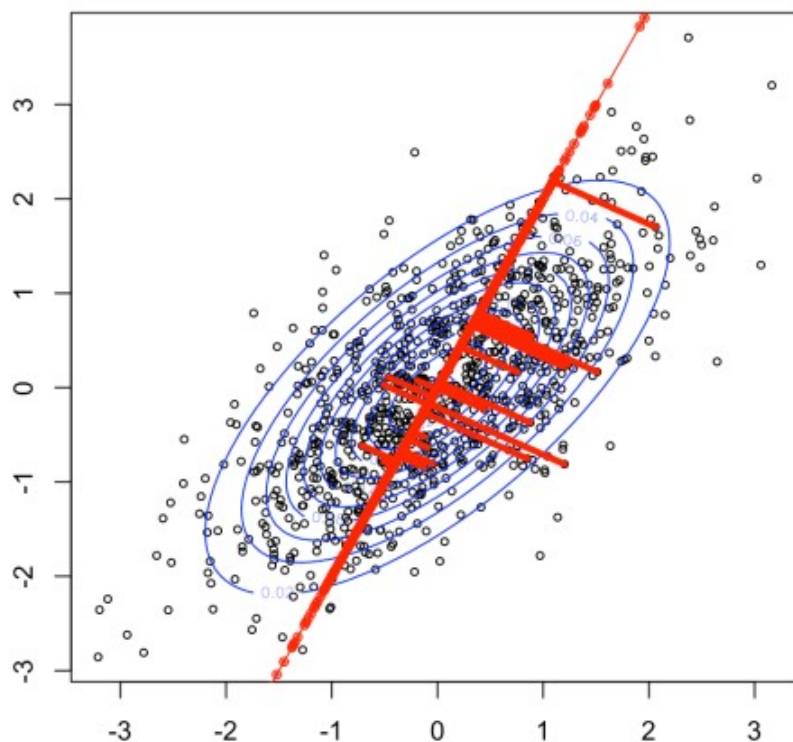
We can try to visualize those results. For instance, the orthogonal projection of a random vector has a Gaussian distribution. Consider a two-dimensional Gaussian vector

```
library(mnormt)
r = .7
s1 = 1
s2 = 1
Sig = matrix(c(s1^2, r*s1*s2, r*s1*s2, s2^2), 2, 2)
Sig
Y = rmnorm(n = 1000, mean=c(0,0), varcov = Sig)
plot(Y, cex=.6)
vu = seq(-4,4,length=101)
vz = outer(vu,vu,function(x,y) dmnorm(cbind(x,y),
mean=c(0,0), varcov = Sig))
contour(vu,vu,vz,add=TRUE,col='blue')
abline(a=0,b=2,col='red')
```

Consider now the projection of points $\mathbf{y} = (y_1, y_2)$ on the straight line with directional vector $\overrightarrow{\mathbf{u}}$ with slope a (say $a=2$). To get the projected point $\mathbf{x} = (x_1, x_2)$ recall that $x_2 = a y_1$ and $\overrightarrow{\mathbf{x}}, \mathbf{y} \perp \overrightarrow{\mathbf{u}}$

. Hence, the following code will give us the orthogonal projections

```
p = function(a) {
  x0=(Y[,1]+a*Y[,2]) / (1+a^2)
  y0=a*x0
  cbind(x0,y0)
}
```

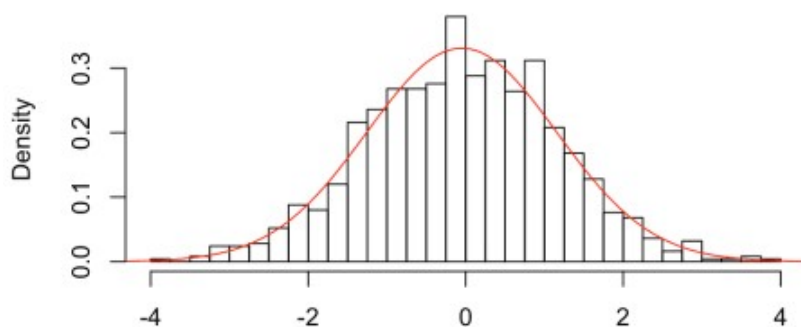


with

```
P = p(2)
for(i in 1:20) segments(Y[i,1],Y[i,2],P[i,1],P[i,2],lwd=4,col="red")
points(P[,1],P[,2],col="red",cex=.7)
```

Now, if we look at the distribution of points on that line, we get... a Gaussian distribution, as expected,

```
z = sqrt(P[,1]^2+P[,2]^2)*c(-1,+1) [(P[,1]>0)*1+1]
vu = seq(-6,6,length=601)
vv = dnorm(vu,mean(z),sd(z))
hist(z,probability = TRUE,breaks = seq(-4,4,by=.25))
lines(vu,vv,col="red")
```



Or course, we can use the matrix representation to get the projection on $\overrightarrow{\text{u}}$, or a normalized version of that vector actually

```

a=2
U = c(1,a)/sqrt(a^2+1)
U
[1] 0.4472136 0.8944272
matP = U %*% solve(t(U) %*% U) %*% t(U)
matP %*% Y[1,]
[,1]
[1,] -0.1120555
[2,] -0.2241110
P[1,]
x0 y0
-0.1120555 -0.2241110

```

(which is consistent with our manual computation). Now, in Cochran theorem, we start with independent random variables,

```
Y = rmnorm(n = 1000, mean=c(0,0), varcov = diag(c(1,1)))
```

Then we consider the projection on $\overrightarrow{\mathbf{u}}$ and $\overrightarrow{\mathbf{v}} = \overrightarrow{\mathbf{u}}^\perp$

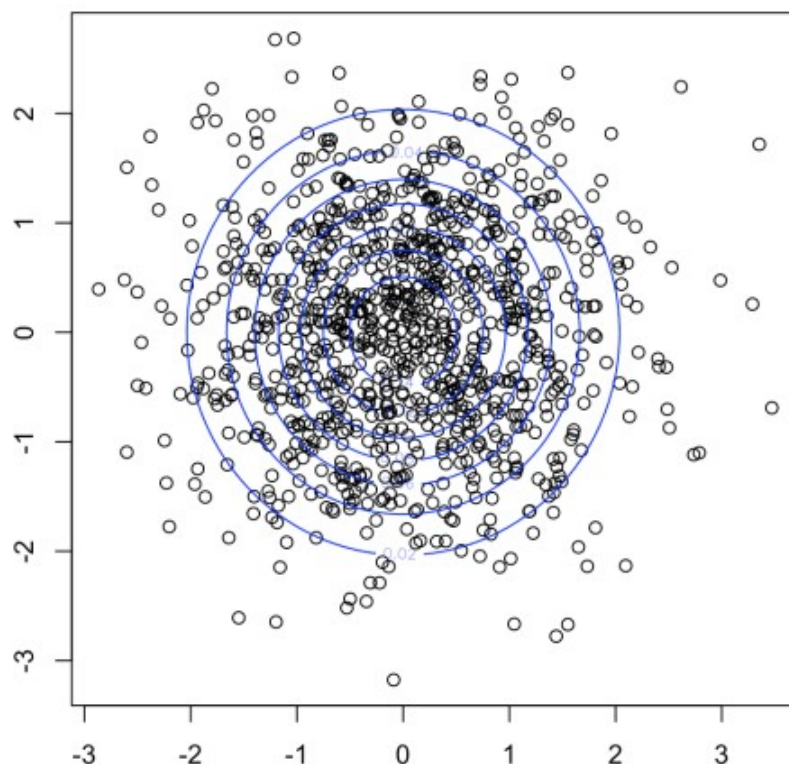
```

U = c(1,a)/sqrt(a^2+1)
matP1 = U %*% solve(t(U) %*% U) %*% t(U)
P1 = Y %*% matP1
z1 = sqrt(P1[,1]^2+P1[,2]^2)*c(-1,+1)[(P1[,1]>0)*1+1]
V = c(a,-1)/sqrt(a^2+1)
matP2 = V %*% solve(t(V) %*% V) %*% t(V)
P2 = Y %*% matP2
z2 = sqrt(P2[,1]^2+P2[,2]^2)*c(-1,+1)[(P2[,1]>0)*1+1]

```

We can plot those two projections

```
plot(z1,z2)
```



and observe that the two are indeed, independent Gaussian variables. And (of course) their squared norms

are $\chi^2_{\{1\}}$ distributed.