The model we are going to talk about is used to describe the results of full (complete) diallel experiments, where we have crosses + reciprocals + selfs. If you are not sure what a diallel experiment is, we suggest you go back to one of our previous posts on this sequence, where we give some preliminary information for beginners. Otherwise, we can proceed to the motivating example.

The example

In this post we will use the same example as provided in the original Hayman's paper (Hayman, 1954), relating to a complete diallel experiment with eight parental lines. The R dataset is included in the 'ImDiallel' package; in the box below we load the data, after installing (if necessary) and loading the 'ImDiallel' package (see box below).

The Hayman's model type 2

The Hayman's model type 2 is derived from type 1 (see our previous post), by partitioning the tSCA effect in three additive components. The equation is:

where $\(\)$ is the expected value (the overall mean, in the case of fully orthogonal designs), $\(\)$ is the number of parentals and $\(\)$ is the residual random error term. All the other terms correspond to genetic effects, namely:

- 1. the \(\textrm{g}_i\) and \(\textrm{g}_j\) terms are the **general combining abilities** (GCAs) of the \(i^{th}\) and \(j^{th}\) parents (see here).
- 2. The \(rg^a_i\) and \(rg^b_j\) terms are the **reciprocal general combining abilities** (RGCAs) for the \(i^{th}\) and \(j^{th}\) parents (see here).
- 3. The \(m\) term relates to the difference between the average value of all observations and the average values of crosses (**Mean Dominance Deviation**; MDD).
- 4. The \(d_i\) and \(df_j\) terms relate to the differences between the yield of each selfed parent (\(Y_{ij}\), with \(i = j\)) and the average yield of all selfed parents (**dominance deviation** for the \(i^{th}\) parent; DD).
- 5. The term \(s \{ij\\) is the SCA effect for the combination \(ij\\).
- 6. The \(rs \{ij\\) term is the reciprocal specific combining ability (RSCA) for a specific

\(ij\) combination, that is the discrepancy between the performances of the two reciprocals (e.g, A \(\times\) B vs. B \(\times\) A)(see here).

Similarly to type 1, the Hayman's model type 2 considers the genetical effects as differences with respect to the intercept \(\\mu\\), that is the mean of all observations (when the design is orthogonal). However, with respect to type 1, this latter model permits the estimation of a higher number of genetic effects (GCAs, RGCAs, MDD, DDs, SCAs and RSCAs) and provides an approach to quantify heterotic effects. We should consider that, due to unbalance (the number of crosses is never equal to the number of selfs), it is necessary to introduce some coefficients (i.e. $\(n-1\)$) and $\(n-2\)$ in Equation 1), which do not have an obvious meaning. In future posts we will see that other diallel models were proposed, which account for heterotic effects in a different manner (Gardner and Eberhart, 1966).

Model fitting with R

Let's assume that all effects are fixed, apart from the residual error effect. Consequently, Equation 1 is a specific parameterisation of a general linear model, which we can fit by the usual lm() function and related methods. However, we need to exploit some of the facilities in our new 'ImDiallel' extension package, which consist of the GCA(), MDD(), DD(), SCA(), RGCA() and RSCA() functions (see the box below). The resulting lm object can be explored by the usual R methods, such as summary() and anova() (the output of the summary() method is partly hidden, for brevity)

```
contrasts(hayman54$Block) <- "contr.sum"</pre>
dMod <- lm(Ftime ~ Block + GCA(Par1, Par2) + MDD(Par1, Par2) +</pre>
            DD(Par1, Par2) + SCA(Par1, Par2) +
            RGCA(Par1, Par2) + RSCA(Par1, Par2), data = hayman54)
summary(dMod)$coef[1:6,]
##
                     Estimate Std. Error
                                            t value
                                                        Pr(>|t|)
                    162.898437 1.804567 90.2700843 2.381071e-68
## (Intercept)
## Block1
                    -1.054688 1.804567 -0.5844545 5.610017e-01
## GCA(Par1, Par2)g A 46.195312 3.376036 13.6832990 1.558468e-20
## GCA(Par1, Par2)q B -24.585938 3.376036 -7.2824864 6.421946e-10
## GCA(Par1, Par2)g C 49.632812 3.376036 14.7015049 4.900927e-22
## GCA(Par1, Par2)g D 18.351563 3.376036 5.4358311 9.415231e-07
# ...
# ...
anova (dMod)
## Analysis of Variance Table
##
## Response: Ftime
##
                 Df Sum Sq Mean Sq F value
                                             Pr(>F)
                                142 0.3416 0.56100
                   1 142
## Block
## GCA(Par1, Par2) 7 277717 39674 95.1805 < 2.2e-16 ***
## MDD(Par1, Par2) 1 30797 30797 73.8840 3.259e-12 ***
## DD(Par1, Par2) 7 34153 4879 11.7050 1.957e-09 ***
                              1864 4.4729 2.560e-06 ***
## SCA(Par1, Par2) 20 37289
## RGCA(Par1, Par2) 7 6739
                                963 2.3097 0.03671 *
## RSCA(Par1, Par2) 21 12373
                                589 1.4135 0.14668
## Residuals 63 26260
                                417
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

For the sake of simplicity, we also built a wrapper function named lm.diallel(), which can be used in the very same fashion as lm(). The syntax is:

```
lm.diallel(formula, Block, Env, data, fct)
```

where 'formula' specifies the response variable and the two variables for parentals (e.g., Yield ~ Par1 + Par2) and the two arguments 'Block' and 'Env' are used to specify optional variables, coding for blocks and environments, respectively. The argument 'data' is a 'dataframe' where to look for the explanatory variables and, finally, 'fct' is a string variable coding for the selected model ("HAYMAN2", for this example; see below).

```
dMod2 <- lm.diallel(Ftime ~ Par1 + Par2, Block = Block,</pre>
                  data = hayman54, fct = "HAYMAN2")
# summary(dMod2)
anova (dMod2)
## Analysis of Variance Table
## Response: Ftime
## Df Sum Sq Mean Sq F value Pr(>F)
           1 142 142 0.3416 0.56100
## Block
## MDD
           1 30797 30797 73.8840 3.259e-12 ***
           7 277717 39674 95.1805 < 2.2e-16 ***
## GCA
           7 34153
                      4879 11.7050 1.957e-09 ***
## DD
## SCA 20 37289 1864 4.4729 2.560e-06 ***
                      963 2.3097 0.03671 *
589 1.4135 0.14668
## RGCA
          7 6739
## RSCA 21 12373
## Residuals 63 26260
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

The above function works very much like the lm() function and makes use of the general purpose linear model solver lm.fit(). Apart from simplicity, another advantage is that the call to lm.diallel() returns an object of both 'lm' and 'diallel' classes. For this latter class, we built several specific S3 methods, such as the usual anova(), summary() and model.matrix() methods, partly shown in the box above.

Considering that diallel models are usually fitted to determine genetical parameters, we also built the <code>glht.diallelMod()</code> method and the <code>diallel.eff()</code> function, which can be used with the 'multcomp' package, to retrieve the complete list of genetical parameters. An excerpt is shown in the box below.

```
library(multcomp)
gh <- glht(linfct = diallel.eff(dMod2))</pre>
# summary(gh, test = adjusted(type = "none"))
    Simultaneous Tests for General Linear Hypotheses
# Linear Hypotheses:
              Estimate Std. Error t value Pr(>|t|)
\# m == 0
              -5.8627
                       0.6821 -8.596 4.48e-09 ***
                        3.3760 13.683 2.17e-13 ***
\# g A == 0
              46.1953
# g B == 0
             -24.5859
                        3.3760 -7.282 9.83e-08 ***
# g_C == 0 49.6328 3.3760 14.702 4.13e-14 ***
```

```
\# g D == 0
               18.3516
                           3.3760 5.436 1.07e-05 ***
\# g E == 0
               -20.9297
                           3.3760 -6.199 1.47e-06 ***
\# g F == 0
                2.4453
                          3.3760 0.724 0.475340
\# g G == 0
               -44.7109
                          3.3760 -13.244 4.57e-13 ***
# g H == 0
                          3.3760 -7.819 2.71e-08 ***
              -26.3984
               1.2213 1.9492 0.627 0.536380
\# d A == 0
\# d B == 0
                          1.9492 -1.345 0.190113
               -2.6224
# ...
# ...
```

In a previous post (see here) we have shown that, when diallel models are fitted to the genotype means (and thus we have no replicates), an appropriate estimate of residual mean square and degrees of freedom can be passed as arguments to the summary(), anova() and diallel.eff() methods.

Estimation of variance components (random genetic effects)

If we intend to regard the genetic effects as random and to estimate variance components, we can use the mmer() function in the 'sommer' package (Covarrubias-Pazaran, 2016), although we need to code a bunch of dummy variables. In order to make things simpler for routine experiments, we have coded the mmer.diallel() wrapper using the same syntax as the lm.diallel() function. The exemplary code is given in the box below.

```
# Random genetic effects
mod1m <- mmer.diallel(Ftime ~ Par1 + Par2, Block = Block,</pre>
                    data = hayman54,
                    fct = "HAYMAN2")
mod1m
##
              VarComp VarCompSE
             0.00000
                        9.188298
## Block
          1783.96081 3118.893561
## MDD
## GCA
          1005.92052 574.893353
           17.97898 19.920016
## RGCA
## DD
           659.53567 468.205470
           351.74035 144.688653
## SCA
         32.02325 46.361581
## RSCA
## Residuals 412.54051 73.506382
```

That's all about the Hayman's models; you may have noted that both models (type 1 and 2) were devised for full diallel experiments, which are not so widespread in 'genetical' literature. A few years later, in 1956, the Australian scientist B. Griffing made the relevant effort of creating a comprehensive set of models which can be fitted to all types of diallel experiments. We will talk about these models in a future post.

Thanks for reading (and happy 2021!)