

Variance

Lecture 5a (STAT 24400 F24)

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Variance of random variables

The **variance** of a random variable X is defined as

$$\text{Var}(X) = \mathbb{E}\left((X - \mathbb{E}(X))^2\right)$$

- Variance measures how much the r.v. typically varies from its mean.
- The definition is valid if and only if the expectation exists, and $\mathbb{E}(X)$ exists.
- Variance is often denoted as σ^2 or σ_X^2 .
- σ or σ_X is called the **standard deviation** of X .

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Variance of discrete and continuous r.v.'s

For a discrete random variable, we may calculate the variance as

$$\text{Var}(X) = \sum_x p_X(x) \cdot (x - \mu_X)^2$$

For a continuous random variable,

$$\text{Var}(X) = \int_{x=-\infty}^{\infty} f_X(x) \cdot (x - \mu_X)^2 dx$$

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Variance and measures of variation

Why use $\mathbb{E}\left((X - \mathbb{E}(X))^2\right)$ rather than, say, $\mathbb{E}\left(|X - \mathbb{E}(X)|\right)$?

- Intuitively, both measure the same thing:
how much variability is in the distribution of X ?
- However, $\mathbb{E}\left((X - \mathbb{E}(X))^2\right)$ has many convenient mathematical and statistical properties (because mathematically x^2 is a nicer function than $|x|$).

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Properties of variance

1. A simpler formula.

Variance can be calculated with a simpler formula:

$$\text{Var}(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2$$

(Note: this also shows that $\mathbb{E}(g(X)) \neq g(\mathbb{E}(X))$ in general.)

Proof. Write $\mu_X = \mathbb{E}(X)$,

$$\begin{aligned} \text{Var}(X) &= \mathbb{E}((X - \mu_X)^2) = \mathbb{E}(X^2 - 2\mu_X \cdot X + \mu_X^2) \\ &= \mathbb{E}(X^2) - 2\mu_X \mathbb{E}(X) + \mu_X^2 = \mathbb{E}(X^2) - \mu_X^2 \end{aligned}$$

by linearity of $\mathbb{E}(\cdot)$

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Properties of variance

2. The degenerate case.

$\text{Var}(X) = 0$ if and only if $\mathbb{P}(X = \mu_X) = 1$.

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Properties of variance

3. Chebyshev's inequality.

For any random variable X and any $t > 0$,

$$\mathbb{P}(|X - \mu_X| \geq t) \leq \frac{\sigma_X^2}{t^2}$$

Proof.

Recall Markov's inequality, for $X^* \geq 0$, $\mathbb{P}(X^* \geq t) \leq \frac{\mathbb{E}(X^*)}{t}$.

Let $Y = (X - \mu_X)^2$, so then $\mathbb{E}(Y) = \sigma_X^2$.

$$\mathbb{P}(|X - \mu_X| \geq t) = \mathbb{P}(Y \geq t^2) \leq \frac{\mathbb{E}(Y)}{t^2}$$

by Markov's inequality applied to Y

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Properties of variance

4. Linear transformations.

For any random variable X and any constants a, b ,

$$\text{Var}(a + bX) = b^2 \text{Var}(X)$$

Proof.

Using linearity of $\mathbb{E}(\cdot)$,

$$\begin{aligned} \text{Var}(a + bX) &= \mathbb{E}[(a + bX - \mu_{a+bX})^2] \\ &= \mathbb{E}[(a + bX - (a + b\mu_X))^2] \\ &= \mathbb{E}[(b(X - \mu_X))^2] = b^2 \mathbb{E}[(X - \mu_X)^2] = b^2 \text{Var}(X) \end{aligned}$$

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Examples (Normal, by formula)

Suppose $Y \sim N(\mu, \sigma^2)$. Formally, we may use the formulae to calculate

$$\mathbb{E}(Y) = \int_{-\infty}^{\infty} y f_Y(y) dy, \quad \text{Var}(Y) = \mathbb{E}(Y - \mu_Y)^2 = \int_{-\infty}^{\infty} (y - \mu)^2 f_Y(y) dy$$

Applying variable substitution $z = \frac{y-\mu}{\sigma}$ and Integration-by-parts,

$$\mathbb{E}(Y) = \int_{-\infty}^{\infty} y \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y-\mu)^2}{2\sigma^2}} dy = \dots = \mu$$

$$\text{Var}(Y) = \int_{-\infty}^{\infty} (y - \mu)^2 \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y-\mu)^2}{2\sigma^2}} dy = \dots = \sigma^2$$

(Side notes: Properties of odd and even functions.)

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Examples (Normal, by linearity)

Alternatively:

The standard normal distribution $N(0, 1)$ has mean 0 and variance 1 (assumption).

Let $X \sim N(0, 1)$. Then $Y = \mu + \sigma \cdot X$ is also of normal distribution (assumption).

Verify: $Y \sim N(\mu, \sigma^2)$.

Using the linearity property of expected value $\mathbb{E}(\cdot)$:

$$\mathbb{E}(Y) = \mathbb{E}(\mu + \sigma \cdot X) = \mu + \sigma \cdot \mathbb{E}(X) = \mu$$

$$\text{Var}(Y) = \text{Var}(\mu + \sigma \cdot X) = \sigma^2 \cdot \text{Var}(X) = \sigma^2$$

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Examples (Bernoulli)

Suppose $X \sim \text{Bernoulli}(p)$. What is $\text{Var}(X)$?

We know that

$$\mathbb{E}(X) = p$$

Compute

$$\mathbb{E}(X^2) = 1^2 \cdot p + 0^2 \cdot (1 - p) = p$$

So,

$$\text{Var}(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 = p - p^2 = p(1 - p)$$

Remark: Note the useful symmetry in the variance.

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Examples (Exponential)

Suppose $X \sim \text{Exponential}(\lambda)$. What is $\text{Var}(X)$?

We know that

$$\mathbb{E}(X) = 1/\lambda$$

Applying Integration-by-parts twice,

$$\mathbb{E}(X^2) = \int_{x=0}^{\infty} x^2 \cdot \lambda e^{-\lambda x} dx = \left[-x^2 e^{-\lambda x} - \frac{2}{\lambda} x e^{-\lambda x} - \frac{2}{\lambda^2} e^{-\lambda x} \right]_{x=0}^{\infty} = \frac{2}{\lambda^2}$$

So,

$$\text{Var}(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda} \right)^2 = \frac{1}{\lambda^2}$$

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Example: sampling effects

We perform and compare two experiments:

- Toss a fair coin 2 times, and let $X = \#$ of Heads
- A bag has 4 tickets (2 red, 2 blue). Draw 2 tickets, and let $Y = \#$ red

Questions

Are $\mathbb{E}(X)$ and $\mathbb{E}(Y)$ the same?

Are $\text{Var}(X)$ and $\text{Var}(Y)$ the same?

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Examples (sampling effects, cont.)

Calculate the probability mass functions of X and Y :

	0	1	2	
p_X	$1/4$	$1/2$	$1/4$	← using Binomial PMF
p_Y	$1/6$	$2/3$	$1/6$	← by counting

$$\mathbb{E}(X) = \frac{1}{4} \cdot 0 + \frac{1}{2} \cdot 1 + \frac{1}{4} \cdot 2 = 1$$

$$\mathbb{E}(Y) = \frac{1}{6} \cdot 0 + \frac{2}{3} \cdot 1 + \frac{1}{6} \cdot 2 = 1$$

$$\mathbb{E}(X^2) = \frac{1}{4} \cdot 0^2 + \frac{1}{2} \cdot 1^2 + \frac{1}{4} \cdot 2^2 = \frac{3}{2} \quad \rightsquigarrow \quad \text{Var}(X) = \frac{3}{2} - 1^2 = \frac{1}{2}$$

$$\mathbb{E}(Y^2) = \frac{1}{6} \cdot 0^2 + \frac{2}{3} \cdot 1^2 + \frac{1}{6} \cdot 2^2 = \frac{4}{3} \quad \rightsquigarrow \quad \text{Var}(Y) = \frac{4}{3} - 1^2 = \frac{1}{3}$$

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Examples (remarks)

Remarks

Notice the differences in the two experiments:

- Independent vs not-independent sampling
- without replace vs. with replacement
- infinite population vs finite population

⇒ The effects on variance

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