Joint distributions (part 3)

Lecture 6b (STAT 24400 F24)

1/15

Conditional distribution

We often want to ask about probabilities of X based on some (partial) knowledge of Y, e.g. $\mathbb{P}(X \ge 3 \mid Y \ge 7)$.

We may also ask about the distribution of X given $\underline{\text{exact knowledge}}$ of Y, e.g. $\mathbb{P}(X \geq 3 \mid Y = 7)$.

this is the default meaning of "conditional distribution"

For example,

- If we draw a hand of 10 cards, what's the distribution of the number of Kings, given that the hand contains 4 red cards?
- If we sample a random person from the population, what's the distribution of their height, given that the age of the person chosen is 12.5 years?

Conditional distribution

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2/15

Conditional distribution: discrete case

If we know that Y = y, what is the distribution of X?

$$\mathbb{P}(X = x \mid Y = y) = \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(Y = y)} = \frac{p_{X,Y}(x,y)}{p_Y(y)} = \frac{p_{X,Y}(x,y)}{\sum_{x'} p_{X,Y}(x',y)}.$$

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We write this as $p_{X|Y}(x \mid y)$, the conditional PMF of X given Y = y.

3/15

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 $p_{X|Y}(\cdot \mid y)$ defines a valid PMF at any fixed y.

For example, it sums to 1:

$$\sum_{x} p_{X|Y}(x \mid y) = \sum_{x} \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(Y = y)} = \frac{\mathbb{P}(Y = y)}{\mathbb{P}(Y = y)} = 1.$$

3 / 15

Continuous case: problem with point-mass

The continuous case is harder...

$$\mathbb{P}(X = x \mid Y = y) = \frac{\mathbb{P}(X = x, Y = y)}{\mathbb{P}(Y = y)} = ????$$

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We will instead need to work with densities.

Continuous case: univariate density in a tiny interval

First we take another look at density in the univariate (not joint) case: For small $\epsilon>0$,

$$\mathbb{P}(x < X < x + \epsilon) = \int_{t=x}^{x+\epsilon} f(t) dt \approx \epsilon \cdot f(x) ,$$

So, $\mathbb{P}(X=x)$ is zero, but $\mathbb{P}(X\approx x)$ (x falls near x) is validly approximated via the density f(x).

 $5\,/\,15$

Continuous case: conditional density in a tiny region

For the joint case, we can't ask questions like $\mathbb{P}(X = x \mid Y = y)$, but $\mathbb{P}(X \approx x \mid Y \approx y)$ should be approximated via the *conditional density*.

$$\begin{split} \mathbb{P}(x < X < x + \epsilon \mid y < Y < y + \delta) &= \frac{\mathbb{P}(x < X < x + \epsilon, y < Y < y + \delta)}{\mathbb{P}(y < Y < y + \delta)} \\ &= \frac{\int_{u = x}^{x + \epsilon} \int_{v = y}^{y + \delta} f_{X,Y}(u, v) \, dv \, du}{\int_{v = v}^{y + \delta} f_{Y}(v) \, dv} \approx \frac{\epsilon \cdot \delta \cdot f_{X,Y}(x, y)}{\delta \cdot f_{Y}(y)} = \epsilon \cdot \frac{f_{X,Y}(x, y)}{f_{Y}(y)} \; . \end{split}$$

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 $6\,/\,15$

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$$\mathbb{P}(x < X < x + \epsilon \mid y < Y < y + \delta) = \frac{\mathbb{P}(x < X < x + \epsilon, y < Y < y + \delta)}{\mathbb{P}(y < Y < y + \delta)}$$

$$= \frac{\int_{u=x}^{x+\epsilon} \int_{v=y}^{y+\delta} f_{X,Y}(u,v) \, dv \, du}{\int_{v=y}^{y+\delta} f_{Y}(v) \, dv} \approx \frac{\epsilon \cdot \delta \cdot f_{X,Y}(x,y)}{\delta \cdot f_{Y}(y)} = \epsilon \cdot \frac{f_{X,Y}(x,y)}{f_{Y}(y)}.$$

So by analogy, we should define the conditional density as

$$f_{X|Y}(x \mid y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)}.$$

6 / 15

Independence (in PMF and PDF)

If (X, Y) is discrete,

 $X \perp \!\!\! \perp Y$ is equivalent to: $p_{X|Y}(x \mid y) = p_X(x)$ for all (x, y) (and, same for reversing X and Y).

If (X, Y) is continuous,

 $X \perp \!\!\! \perp Y$ is equivalent to: $f_{X|Y}(x \mid y) = f_X(x)$ for all (x, y) (and, same for reversing X and Y).

7/15

Law of total probability (in terms of conditional PMF & PDF)

For probabilities, given a partition B_1, B_2, \ldots , we showed

$$\mathbb{P}(A) = \sum_{i} \mathbb{P}(A \cap B_{i}) = \sum_{i} \mathbb{P}(A \mid B_{i}) \mathbb{P}(B_{i})$$

Similar rules hold for conditional distributions:

For discrete (X, Y),

$$p_Y(y) = \sum_{x} p_{X,Y}(x,y) = \sum_{x} p_{Y|X}(y \mid x) p_X(x)$$

For a continuous (X, Y),

$$f_Y(y) = \int_{x=-\infty}^{\infty} f_{X,Y}(x,y) \, \mathrm{d}x = \int_{x=-\infty}^{\infty} f_{Y\mid X}(y\mid x) f_X(x) \, \mathrm{d}x$$

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8 / 15

Example (uniform on unit disk)

(X,Y) is chosen uniformly at random from the unit disk, $\{x^2+y^2\leq 1\}$. Joint density:

$$f(x,y) = \begin{cases} \frac{1}{\pi}, & x^2 + y^2 \le 1, \\ 0, & \text{otherwise.} \end{cases}$$

Example (uniform on unit disk)

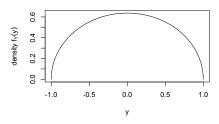
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Marginal density:

$$f_Y(y) = \int_{x=-\infty}^{\infty} f(x,y) \, \mathrm{d}x = \int_{x=-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{1}{\pi} \, \mathrm{d}x = \frac{2\sqrt{1-y^2}}{\pi} \quad \text{for } y \in [-1,1]$$

(& analogous for $f_X(x)$)



9 / 15

Example (cont.)

Conditional density:

$$f_{X|Y}(x \mid y) = \frac{f_{X,Y}(x,y)}{f_{Y}(y)} = \frac{\frac{1}{\pi}}{\frac{2\sqrt{1-y^2}}{\pi}} = \frac{1}{2\sqrt{1-y^2}}$$

10 / 15

Example (cont.)

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If Y = y for $y \in [-1, 1]$, possible values for X lie in $[-\sqrt{1 - y^2}, \sqrt{1 - y^2}]$.

Example (cont.)

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Equivalently:

$$X \mid Y = y \sim \mathsf{Uniform}[-\sqrt{1-y^2}, \sqrt{1-y^2}]$$

(& analogous for $Y \mid X$)

10 / 15

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Remarks : The *marginal* distribution of X (or of Y) is not uniform, but the *conditional* distribution of X|Y (or of Y|X) is uniform.

10 / 15

Functions of jointly distributed random variables

Expectations

If $Y = g(X_1, ..., X_n)$ where $X_1, ..., X_n$ have a joint PMF p, then

$$\mathbb{E}(Y) = \sum_{(x_1,\ldots,x_n)} g(x_1,\ldots,x_n) \cdot p(x_1,\ldots,x_n).$$

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If instead X_1, \ldots, X_n have a joint density f, then

$$\mathbb{E}(Y) = \int_{x_1} \dots \int_{x_n} g(x_1, \dots, x_n) \cdot f(x_1, \dots, x_n) dx_n \dots dx_1.$$

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(Alternatively, can compute the distribution of Y via the joint distribution of the X_i 's, and then compute $\mathbb{E}(Y)$, but this is generally less efficient.)

 $11\,/\,15$

Expectation of product function under independence

If $X \perp Y$, then $\mathbb{E}(X \cdot Y) = \mathbb{E}(X)\mathbb{E}(Y)$.

Proof: Let's prove for the continuous case:

the density factors since $X \perp \!\!\! \perp Y$

$$\mathbb{E}(X \cdot Y) = \int_{X} \int_{Y} x \cdot y \cdot f_{X,Y}(x,y) \, dy \, dx = \int_{X} \int_{Y} x \cdot y \cdot f_{X}(x) f_{Y}(y) \, dy \, dx$$
$$= \left(\int_{X} x \cdot f_{X}(x) \, dx \right) \cdot \left(\int_{Y} y \cdot f_{Y}(y) \, dy \right) = \mathbb{E}(X) \cdot \mathbb{E}(Y)$$

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12 / 15

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More generally:

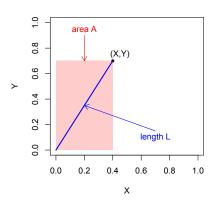
If $X \perp Y$, then for any functions $g, h, g(X) \perp h(Y)$ and so

$$\mathbb{E}(g(X)\cdot h(Y)) = \mathbb{E}(g(X))\cdot \mathbb{E}(h(Y))$$

Example (functions of (x, y) uniform on unit square)

The point (X, Y) is drawn uniformly from the unit square $[0, 1] \times [0, 1]$.

What is the expected value of A, the area of the rectangle? What is the expected value of L, the length of the segment?



Example: (unit square cont.)

Joint density of (X, Y):

$$f(x,y) = \begin{cases} 1, & (x,y) \in [0,1] \times [0,1], \\ 0, & (x,y) \notin [0,1] \times [0,1]. \end{cases}$$

13 / 15

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Equivalent:

$$f(x,y) = \mathbb{1}_{0 \le x \le 1} \cdot \mathbb{1}_{0 \le y \le 1}$$

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We can verify that marginally, X and Y are each $\mathsf{Uniform}[0,1]$

14 / 15

14 / 15

14 / 15

Example: (unit square cont.)

The area is $A = X \cdot Y$, so by independence,

$$\mathbb{E}(A) = \mathbb{E}(X \cdot Y) = \mathbb{E}(X) \cdot \mathbb{E}(Y) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}.$$

Example: (unit square cont.)

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The length is $L = \sqrt{X^2 + Y^2}$, so:

$$\mathbb{E}(L) = \iint \sqrt{x^2 + y^2} \cdot f(x, y) \, dy \, dx = \int_{x=0}^{1} \int_{y=0}^{1} \sqrt{x^2 + y^2} \, dy \, dx.$$

15 / 15

15 / 15