

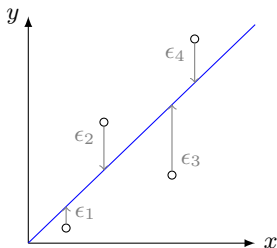
Exam 2 Review

Spring 2021

Announcements

- ▶ Exam 2 will be “take home”: 80 minutes from when you **begin** the exam.
- ▶ The exam *must* be completed by 5pm Central time on Friday 3/26.
- ▶ I will have additional office hours on Thursday, 8:30–9:20am for last-minute questions.

RMSE of a linear model



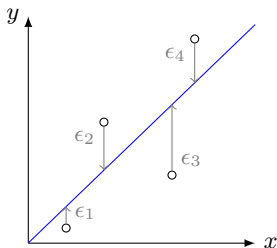
$$\text{— } y = \beta_0 + \beta_1 x$$

$$\text{RMSE} = \sqrt{\frac{1}{n} \sum_{i=1}^n \left(y_i^{\text{pred}} - y_i^{\text{true}} \right)^2}$$

- ▶ The RMSE measures the uncertainty in the model's predictions. Predictions should be reported as $y^{\text{pred}} \pm \text{RMSE}$.
- ▶ The 95% confidence interval of a model's predictions are

$$[y^{\text{pred}} - 2 \text{RMSE}, y^{\text{pred}} + 2 \text{RMSE}]$$

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0-, 1-, and 2-norm regularization

$$\beta = \begin{pmatrix} 0 \\ 1.2 \\ -0.3 \\ 0 \\ 1 \end{pmatrix}$$

$$\|\beta\|_0 = 0 + 1 + 1 + 0 + 1 = 3 \quad (3 \text{ nonzero entries})$$

$$\|\beta\|_1 = |0| + |1.2| + |-0.3| + |0| + |1| = 2.5$$

$$\|\beta\|_2 = \sqrt{0^2 + 1.2^2 + (-0.3)^2 + 0^2 + 1^2} = 2.53$$

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Regularization	Computation	Sparsity	Unique?
0-norm	Hard (combinatorial)	Sparse	No
1-norm	Easier (discontinuous derivative)	Mostly sparse	No
2-norm	Easy	Dense	Yes
Elastic Net (1&2-norm)	Easier (like 1-norm)	Some sparsity	Yes

Systems of linear inequalities

- ▶ A system of linear inequalities has the form $\mathbf{Ax} \leq \mathbf{b}$.
- ▶ This includes inequalities of the form $\mathbf{Ax} \geq \mathbf{b}$, since these can be transformed to $-\mathbf{Ax} \leq -\mathbf{b}$.
- ▶ Systems of inequalities often have infinitely many solutions. We have not discussed how to solve these systems (we let `MATLAB` solve the SVM problem).
- ▶ Also, note that we haven't discussed *strict* inequalities ($\mathbf{Ax} < \mathbf{b}$). These are much harder to solve since the solution set is open.
- ▶ Are the solution sets to linear inequalities always convex?

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- ▶ Are the solution sets to linear inequalities always convex?

Proof. Assume \mathbf{x}_1 and \mathbf{x}_2 are solutions to $\mathbf{Ax} \leq \mathbf{b}$. Then

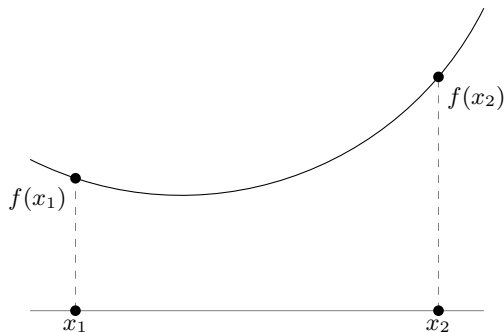
$$\begin{aligned}\mathbf{A}(\lambda \mathbf{x}_1 + (1 - \lambda) \mathbf{x}_2) &= \lambda \mathbf{Ax}_1 + (1 - \lambda) \mathbf{Ax}_2 \\ &\leq \lambda \mathbf{b} + (1 - \lambda) \mathbf{b} \\ &= \mathbf{b}\end{aligned}$$

Since all points on the line connecting \mathbf{x}_1 and \mathbf{x}_2 are also solutions, the solution space must be convex.

Convex functions

A function f is convex if and only if

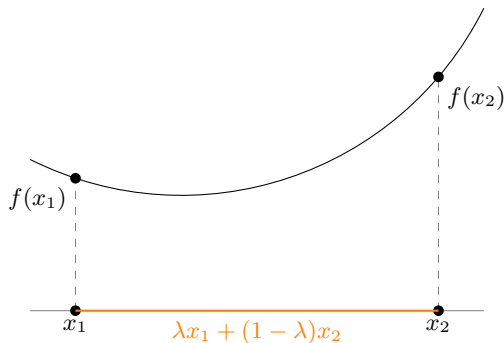
$$f(\lambda \mathbf{x}_1 + (1 - \lambda) \mathbf{x}_2) \leq \lambda f(\mathbf{x}_1) + (1 - \lambda) f(\mathbf{x}_2), \quad \lambda \in [0, 1]$$



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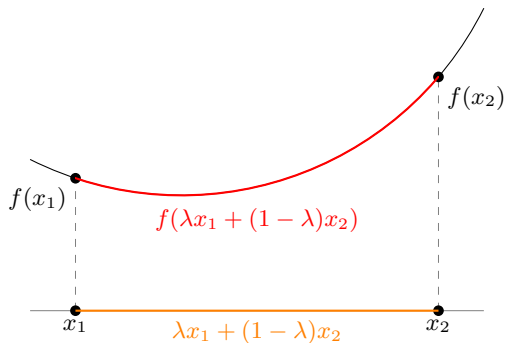
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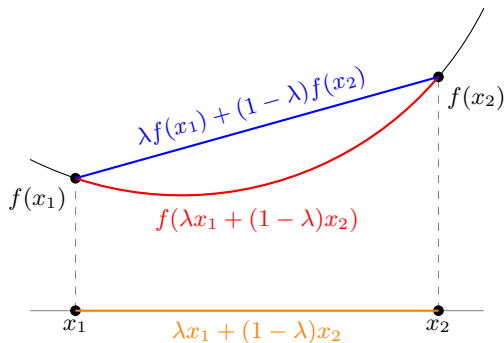
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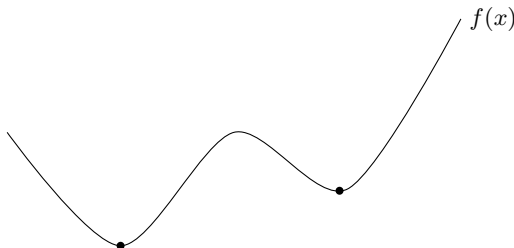


Convex functions with convex domains have only global minima.

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Here is a function with local minima. However, it isn't convex.

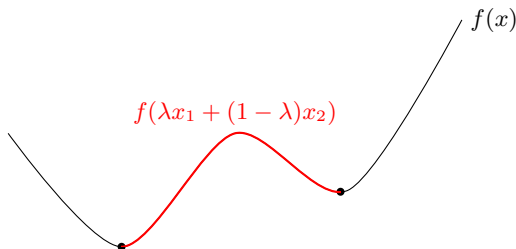


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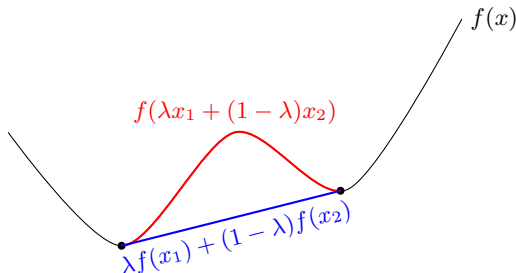


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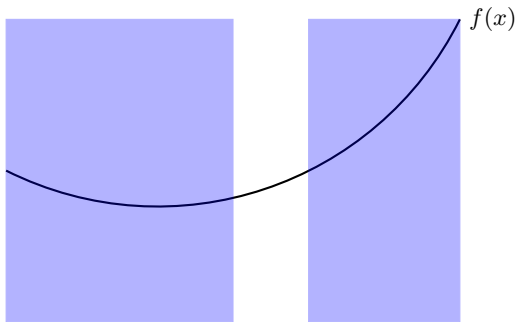
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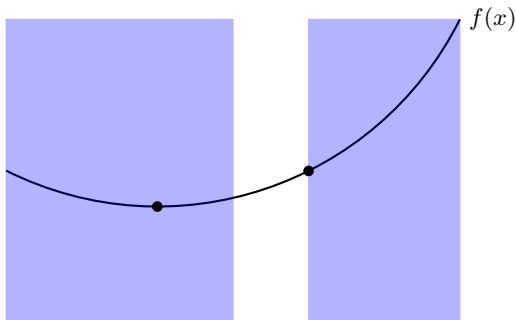
Why does the domain need to be convex?

Imagine minimizing a convex function over a discontinuous (non-convex) domain (shaded blue).



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Imagine minimizing a convex function over a discontinuous (non-convex) domain (shaded blue).



The boundaries create a second local (but not global) minimum of the convex function.

The Jacobian Matrix

Consider a multivariate function $\mathbf{g}(\mathbf{x}) = \begin{pmatrix} g_1(\mathbf{x}) \\ g_2(\mathbf{x}) \\ g_3(\mathbf{x}) \end{pmatrix}$

The Jacobian matrix of partial derivatives is

$$\mathbf{J}(\mathbf{x}) = \begin{pmatrix} \frac{\partial g_1}{\partial x_1} & \frac{\partial g_1}{\partial x_2} & \frac{\partial g_1}{\partial x_3} \\ \frac{\partial g_2}{\partial x_1} & \frac{\partial g_2}{\partial x_2} & \frac{\partial g_2}{\partial x_3} \\ \frac{\partial g_3}{\partial x_1} & \frac{\partial g_3}{\partial x_2} & \frac{\partial g_3}{\partial x_3} \end{pmatrix}$$

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$$\mathbf{g}(\mathbf{x}) = \begin{pmatrix} x_1 x_2 x_3 \\ x_2^2 - x_1 x_3 \\ 2x_1^3 \end{pmatrix} \Rightarrow \mathbf{J}(\mathbf{x}) = \begin{pmatrix} x_2 x_3 & x_1 x_3 & x_1 x_2 \\ -x_3 & 2x_2 & -x_1 \\ 6x_1^2 & 0 & 0 \end{pmatrix}$$

The pseudoinverse

Imagine a linear system $\mathbf{Ax} = \mathbf{y}$. If \mathbf{A} is square and full rank, then it has a true inverse \mathbf{A}^{-1} . We can use the inverse to solve for \mathbf{x} :

$$\mathbf{x} = \mathbf{A}^{-1}\mathbf{y}$$

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What if \mathbf{A} is not square or not full rank? It still has a pseudoinverse \mathbf{A}^+ that we can use to solve for \mathbf{x} :

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Final note: If the matrix $\mathbf{A}^T\mathbf{A}$ has full column rank, then $\mathbf{A}^+ = (\mathbf{A}^T\mathbf{A})^{-1}\mathbf{A}^T$. Otherwise, we find the pseudoinverse using the Singular Value Decomposition (as we will see in Part III).

Interpreting the output of fitlm

```
model2 =
```

```
Linear regression model:
```

```
y ~ 1 + x + x^2
```

```
Estimated Coefficients:
```

	Estimate	SE	tStat	pValue
	<hr/>	<hr/>	<hr/>	<hr/>
(Intercept)	0.33485	8.0944	0.041369	0.96709
x	1.3816	2.3069	0.59887	0.55065
x^2	1.0595	0.14057	7.537	2.5514e-11

```
Number of observations: 100, Error degrees of freedom: 97
```

```
Root Mean Squared Error: 20.9
```

```
R-squared: 0.932, Adjusted R-Squared 0.931
```

```
F-statistic vs. constant model: 667, p-value = 2.1e-57
```

- ▶ **Estimate:** The estimated values of the parameter (β_i).
- ▶ **SE:** The standard error of the estimate. Roughly, if $\beta \pm 2$ SE includes 0, the parameter is not significant, but we prefer judgements based on the p -value of a t -test (below).
- ▶ **tStat:** The t -statistic used to calculate the p -value. Not directly interpretable.
- ▶ **pValue:** The probability that a nonzero parameter estimate of this size could have occurred randomly. If $p < 0.05$, we say the parameter is significantly nonzero.

Why predict $\log(\text{odds})$ in logistic regression?

Remember that for logistic regression we use a linear model to predict the $\log(\text{odds})$, i.e.

$$\log(\text{odds}(y = 1)) = \beta_0 + \beta_1 x_1 + \cdots + \beta_n x_n$$

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Why log odds? Two reasons:

- ▶ The odds of an event is always non-negative, and we have no way of forcing our linear model to only make non-negative predictions. However, the log odds can take any value, with negative values indicating odds < 1 .
- ▶ When we pass the log odds through the logistic link function, the output becomes the probability $P(y = 1)$. This is easier to interpret and can be compared with binary data during model fitting.

Deriving estimators

1. The goal of model fitting is to find parameters that minimize the *loss*.
2. For example, the quadratic loss for a model is

$$L(\boldsymbol{\beta}) = \sum_{i=1}^n \left(y^{\text{pred}}(\boldsymbol{\beta}) - y^{\text{true}} \right)^2$$

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3. We substitute the model in for y^{pred} and minimize, usually by finding roots of the gradient with respect to $\boldsymbol{\beta}$.

Newton's Method vs. Gradient descent

We use Newton's method to find the roots of a multivariate function $\mathbf{g}(\mathbf{x})$.

1. Compute $\mathbf{J}(\mathbf{x})$, the Jacobian of \mathbf{g} .
2. Start with an initial guess $\mathbf{x}^{(0)}$.
3. Iterate:

$$\mathbf{x}^{(1)} = \mathbf{x}^{(0)} - \mathbf{J}^{-1}(\mathbf{x}^{(0)})\mathbf{g}(\mathbf{x}^{(0)}).$$

4. Repeat until

$$\mathbf{g}(\mathbf{x}^{(k)}) \approx 0$$

or

$$\mathbf{x}^{(k)} \approx \mathbf{x}^{(k-1)}.$$

We use gradient descent to minimize a scalar-valued function $f(\mathbf{x})$.

1. Compute $\mathbf{g}(\mathbf{x})$, the gradient of f .
2. Start with an initial guess $\mathbf{x}^{(0)}$.
3. Iterate:

$$\mathbf{x}^{(1)} = \mathbf{x}^{(0)} - \alpha \mathbf{g}(\mathbf{x}^{(0)}).$$

4. Repeat until

$$f(\mathbf{x}^{(k)}) \approx f(\mathbf{x}^{(k-1)})$$

or

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