

# Random Graph theory

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## 1 Basics of Graph Theory

- A graph  $G$  is a pair of sets  $G = (V, E)$  where  $V$  is a set of vertices,  $E$  is a set of edges, and  $e \in E$  can be written as  $e = (x, y)$  with  $x, y \in V$ .
- It is common to represent a graph by *drawing*. Each vertex  $v \in V$  is represented as a point in the plane, while edges are lines connecting pairs of points.

There are a number of special graphs, which we will only mention.

- A graph with  $n$  nodes is **complete** (denoted by  $K_n$ ) if every node forms an edge with every other node.
- A **cycle graph** (denoted by  $C_n$ ) is a graph that consists of nodes connected in a closed chain. The degree of each vertex is 2.
- A **tree** is a connected graph with no cycles.

The following theorem will get you started with the basics of graph theory.

### Theorem 1.1 First theorem of graph theory

A finite graph  $G$  has an even number of vertices with odd **degree** (i.e. the number of edges incident to it).

*Proof.* Since each edge connects 2 nodes,

$$2|E| = \sum_{v \in V} \deg(v) = \sum_{\substack{v \in V \\ \deg(v) \text{ even}}} \deg(v) + \sum_{\substack{u \in V \\ \deg(u) \text{ odd}}} \deg(u) \implies \left( \sum_{\substack{u \in V \\ \deg(u) \text{ odd}}} \deg(u) \right) \text{ is even.}$$

If the sum is even, and each summand is odd, then there must be an even number of summands. □

## 2 Sharp Threshold for Connectivity in Erdos-Renyi Graph Model

Most materials for this section note is taken from [1, 4]. First some background:

- We use  $G(n, p)$  to denote an undirected (Erdos-Renyi) graph with  $n$  nodes and probability of forming an edge  $p(n)$ .
- Each edge forms with probability  $p \in (0, 1)$  **independently** of other edges.
- An graph is **connected** if there is a path between any 2 pairs of nodes.
- When  $p = p(n)$  is a function of  $n$ , we may be interested in the behavior of  $G(n, p(n))$  as  $n \rightarrow \infty$ .

## 2.1 Warm-up

**Q1. What is the probability that a vertex is isolated in  $G(n, p)$ ?** **Ans:** A given node  $i$  cannot form an edge with each of the remaining  $n - 1$  nodes. Thus the probability is  $(1 - p)^{n-1}$ .

**Q2. What is the probability that node 1 and node 2 are both isolated?** **Ans:** Let  $I_1, I_2$  be the indicator that node 1 and node 2 are isolated. Then  $P(I_1 \cap I_2) = P(I_1)P(I_2 | I_1) = (1 - p)^{n-1} * (1 - p)^{n-2} = (1 - p)^{2n-3}$ .

**Q3. What is the probability that a group of  $k$  nodes do not connect to the rest of the  $n - k$  nodes?** **Ans:** There are  $\binom{n}{k}$  number of ways to choose  $k$  vertices. Each of these cannot form an edge with the remaining  $n - k$  nodes independently with probability  $(1 - p)^{n-k}$ . So overall we have  $(1 - p)^{(n-k)k}$ .

### Theorem 2.1 Erdos-Renyi 1961

Consider a graph  $g \sim G(n, p(n))$  where  $p(n) = \lambda \frac{\ln(n)}{n}$ . Then as  $n \rightarrow \infty$ ,

$$P(g \text{ connected}) \rightarrow 0 \quad \text{if } \lambda < 1$$

$$P(g \text{ connected}) \rightarrow 1 \quad \text{if } \lambda > 1$$

*Proof.* Suppose  $\lambda < 1$ . Since  $P(\text{connected}) = 1 - P(\text{disconnected})$ , we will show  $P(\text{disconnected}) \rightarrow 1$  by showing that **there is at least 1 isolated node**. Define

- $X_n$  to be a random variable that counts the number of isolated nodes
- $I_i$  to be a (Bernoulli) indicator random variable such that  $I_i = 1$  when node  $i$  is isolated and is 0 otherwise
- Let  $p = p(n)$  and  $q = q(n) = (1 - p(n))^{n-1}$  be the probability of a node being isolated

We want to show  $P(X_n > 0) \rightarrow 1$ , or equivalently,  $P(X_n = 0) \rightarrow 0$ . To get a bound on  $P(X_n = 0)$ , we observe:

$$\begin{aligned} \text{Var}(X_n) &= E(X_n - E(X_n))^2 \\ &= P(X_n = 0)(0 - E(X_n))^2 + P(X_n = 1)(1 - E(X_n))^2 + \dots \\ &\geq P(X_n = 0)E(X_n)^2. \end{aligned}$$

Thus

$$\frac{\text{Var}(X_n)}{E(X_n)^2} \geq P(X_n = 0). \quad (2.1)$$

We will now calculate  $\text{Var}(X_n)$  and  $E(X_n)$  explicitly to show that the left hand side of (2.1) goes to 0. By

linearity of expectation and applying definition of indicators,

$$E(X_n) = E\left(\sum_{i=1}^n I_i\right) = \sum_{i=1}^n E(I_i) = \sum_{i=1}^n P(I_i) = nq.$$

Since indicators  $I_i$  are **not independent** (why?), we use equation (1.10) in your book [3]:

$$\begin{aligned} \text{Var}(X_n) &= \text{Var}\left(\sum_{i=1}^n I_i\right) = \sum_{i=1}^n \text{Var}(I_i) + \sum_{i=1}^n \sum_{j \neq i} \text{Cov}(I_i, I_j) \\ &= \sum_{i=1}^n q(1-q) + \sum_{i=1}^n \sum_{j \neq i} [E(I_i I_j) - E(I_i)E(I_j)] \quad (\text{since } \text{Var}(\text{Bernoulli}) = p(1-p)) \\ &= nq(1-q) + \sum_{i=1}^n \sum_{j \neq i} [P(I_i \cap I_j) - P(I_i)P(I_j)] \\ &= nq(1-q) + \sum_{i=1}^n \sum_{j \neq i} [(1-p)^{n-1}(1-p)^{n-2} - (1-p)^{n-1}(1-p)^{n-1}] \\ &= nq(1-q) + \sum_{i=1}^n \sum_{j \neq i} \left[ \frac{q^2}{1-p} - q^2 \right] \\ &= nq(1-q) + n(n-1)q^2 \frac{p}{1-p}. \end{aligned}$$

Thus

$$\frac{\text{Var}(X_n)}{E(X_n)^2} = \frac{nq(1-q) + n(n-1)q^2 \frac{p}{1-p}}{(nq)^2} = \frac{1-q}{nq} + \frac{n-1}{n} \frac{p}{1-p}.$$

We will now show these terms approach 0 as  $n \rightarrow \infty$ , then eq (2.1) will give us what we need. The first term is dominated by  $nq$ , and

$$\begin{aligned} \lim_{n \rightarrow \infty} nq &= \lim_{n \rightarrow \infty} n(1-p)^{n-1} = \lim_{n \rightarrow \infty} \exp\{\ln(n) + (n-1)\ln(1-p)\} \\ &= \lim_{n \rightarrow \infty} \exp\left\{\ln(n) + (n-1)\ln\left(1 - \frac{\lambda \ln(n)}{n}\right)\right\} \\ &\approx \lim_{n \rightarrow \infty} \exp\left\{\ln(n) - \lambda \frac{n-1}{n} \ln(n)\right\} \quad \left(\ln(1+x) = x - \frac{x^2}{2} + \frac{x^3}{3} - \dots \approx x + O(x^2) \text{ for small } x\right) \\ &= \lim_{n \rightarrow \infty} \exp\left\{\ln(n) \left(1 - \lambda \frac{n-1}{n}\right)\right\} \\ &= \infty \quad (\text{since } \lambda < 1 \text{ and } n \rightarrow \infty) \end{aligned}$$

For the second term, observe that  $p = \lambda \frac{\ln(n)}{n} \rightarrow 0$  as  $n \rightarrow \infty$ . So  $\frac{p}{1-p} \rightarrow 0$  as well. This completes the case for  $\lambda < 1$ .

**Part II.** Now suppose  $\lambda > 1$ . We want to show  $P(\text{connected}) \rightarrow 1$ , or equivalently  $P(\text{disconnected}) \rightarrow 0$ . A graph is disconnected if there is a subgraph of  $k$  nodes that does not connect to any of the other  $n-k$  nodes

(draw a picture). By symmetry, we only have to consider  $k \in \{1, 2, \dots, \lfloor n/2 \rfloor\}$ . So

$$\begin{aligned}
P(\text{disconnected}) &= \bigcup_{k=1}^{\lfloor n/2 \rfloor} P(\text{some set of } k \text{ nodes not connected to the rest}) \\
&\leq \sum_{k=1}^{\lfloor n/2 \rfloor} P(\text{some set of } k \text{ nodes not connected to the rest}) \quad (\text{inclusion-exclusion picture}) \\
&= \sum_{k=1}^{\lfloor n/2 \rfloor} \binom{n}{k} \left[ (1-p)^{(n-k)} \right]^k \\
&\leq \sum_{k=1}^{\lfloor n/2 \rfloor} \binom{n}{k} e^{-p(n-k)k} \quad \left( e^{-x} = 1 - x + \frac{x^2}{2} - \dots \approx 1 - x + O(x^2) \text{ for small } x \right) \\
&= \sum_{k=1}^{\lfloor n/2 \rfloor} \binom{n}{k} \exp \left\{ \frac{-\lambda \ln(n)(n-k)k}{n} \right\} \\
&= \sum_{k=1}^{\lfloor n/2 \rfloor} \binom{n}{k} n^{\frac{-\lambda}{n}(n-k)k} \\
&= \sum_{k=1}^{n^*} \binom{n}{k} n^{\frac{-\lambda}{n}(n-k)k} + \sum_{k=n^*+1}^{\lfloor n/2 \rfloor} \binom{n}{k} n^{\frac{-\lambda}{n}(n-k)k} \quad \left( \text{Choose } n^* \text{ s.t. } \frac{\lambda(n-n^*)}{n} > 1 \iff n^* = \lfloor n(1 - \frac{1}{\lambda}) \rfloor \right)
\end{aligned}$$

For the first term,

$$\begin{aligned}
\sum_{k=1}^{n^*} \binom{n}{k} n^{\frac{-\lambda}{n}(n-k)k} &\leq \sum_{k=1}^{n^*} n^k n^{\frac{-\lambda}{n}(n-k)k} = \sum_{k=1}^{n^*} \left[ n^{1 - \frac{\lambda}{n}(n-k)} \right]^k \\
&\leq \sum_{k=1}^{n^*} \left[ n^{1 - \frac{\lambda}{n}(n-n^*)} \right]^k \quad (\text{judiciously bound inner } k \text{ with something bigger}) \\
&= \sum_{k=1}^{n^*} r^k \quad \left( \text{define } r = n^{1 - \frac{\lambda}{n}(n-n^*)} \right) \\
&= \left( \sum_{k=0}^{n^*} r^k \right) - 1 \\
&= \frac{r}{1-r} \quad \left( \text{geometric series. } r < 1 \text{ since } 1 - \frac{\lambda}{n}(n-n^*) < 0 \right) \\
&= \frac{1}{n^{\frac{\lambda}{n}(n-n^*)-1} - 1} \\
&\longrightarrow 0 \quad (\text{since } n \rightarrow \infty \text{ and exponent} > 0)
\end{aligned}$$

For the second term, we use a better bound than before (see homework):

$$\binom{n}{k} < \left( \frac{ek}{k} \right)^k.$$

Thus

$$\begin{aligned}
\sum_{k=n^*+1}^{\lfloor n/2 \rfloor} \binom{n}{k} n^{\frac{-\lambda}{n}(n-k)k} &< \sum_{k=n^*+1}^{\lfloor n/2 \rfloor} \left(\frac{en}{k}\right)^k n^{\frac{-\lambda(n-k)k}{n}} = \sum_{k=n^*+1}^{\lfloor n/2 \rfloor} \left[ \frac{en^{1-\frac{\lambda(n-k)}{n}}}{k} \right]^k \\
&\leq \sum_{k=n^*+1}^{\lfloor n/2 \rfloor} \left[ \frac{en^{1-\frac{\lambda(n-\frac{n}{2})}{n}}}{n^*+1} \right]^k \quad (\text{bound inner } k \text{ with something from above}) \\
&= \sum_{k=n^*+1}^{\lfloor n/2 \rfloor} \left[ \frac{en^{1-\frac{\lambda}{2}}}{n(1-\frac{1}{\lambda})+1} \right]^k \leq \sum_{k=n^*+1}^{\lfloor n/2 \rfloor} \left[ \frac{en^{\frac{-\lambda}{2}}}{1-\frac{1}{\lambda}} \right]^k \\
&\leq \sum_{k=n^*+1}^{\lfloor n/2 \rfloor} r^k \quad \left(r = \frac{en^{\frac{-\lambda}{2}}}{1-\frac{1}{\lambda}}, 0 < r < 1 \text{ for large } n\right) \\
&\leq \sum_{k=n^*+1}^{\infty} r^k = \sum_{k=0}^{\infty} r^k - \sum_{k=n^*+1}^{n^*} r^k \\
&= \frac{1}{1-r} - \frac{1-r^{n^*+1}}{1-r} \quad \left(\text{finite geometric series } \sum_{k=0}^m r^k = \frac{1-r^{m+1}}{1-r}\right) \\
&= \frac{r^{n^*+1}}{1-r} \rightarrow 0 \quad \text{since } n^* \rightarrow \infty.
\end{aligned}$$

□

### 3 Clustering graphs

Sometimes it is useful to **cluster** a graph, which lumps a graph's nodes into several groups so that there are much more edges within groups than between groups. There are many algorithms to do this (e.g. K-means, hierarchical), which is not our focus. Rather, we will combine random graph theory with (discrete time) Markov chains to define a new distance measure that can be used together with various clustering algorithms. Most material is based on [7].

#### 3.1 Euclidean distance for clustering in k-means algorithm

Clustering algorithms require some measures of similarity (i.e. distance) between 2 nodes. One common distance measure is the Euclidean distance:  $d(\mathbf{x}, \mathbf{y}) = \sqrt{\sum_i (x_i - y_i)^2}$ . This defines the traditional k-means and hierarchical clustering algorithms.

**(review) k-means algorithm:**

- (1) Randomly initialize  $k$  cluster centers  $c_1, \dots, c_k$ .
- (2) Assign each node to some cluster based on the smallest **Euclidean distance** to each cluster center  $c_i$ .
- (3) Update center location  $c_1, \dots, c_k$  by computing the means of the nodes in the cluster.
- (4) Repeat 2 and 3 until assignments no longer change

### 3.2 Random Walk and Euclidean Commute Time (ETC) Distance

We setup a Markov chain as follows:

- Consider a connected **weighted graph** with  $N$  nodes where each edge connecting nodes  $i$  and  $j$  has a weight  $w_{ij} > 0$  that is symmetric  $w_{ij} = w_{ji}$ .
- Let each node of a graph be a state in a Markov chain.
- For node  $i$ , the probability of jumping to an adjacent node  $j$  is  $p_{ij} = \frac{a_{ij}}{a_i}$  where  $a_i = \sum_j a_{ij}$ . Hence large  $w_{ij}$  values means easier communication through the edge.
- Connectivity implies the Markov chain is irreducible.

Based on this Markov chain, we define 2 important quantities:

- Starting at state  $i$ , the **average first passage time**

$$\begin{cases} m(k|k) = 0 \\ m(k|i) = 1 + \sum_{j=1}^N p_{ij} m(k|j) \quad \text{if } i \neq k. \end{cases} \quad (3.1)$$

is the average number of steps a random walker needs to enter state  $k$ .

- Starting at state  $i$ , the **average commute time**

$$n(i, j) = m(j|i) + m(i|j)$$

is the number of steps a random walker take to enter  $j \neq i$  for the first time, then go back to  $i$ . One can show that this function is a distance measure.

$\sqrt{n(i, j)}$  is called the **Euclidean Commute Time (ETC) Distance**. Intuitively,  $n(i, j)$  decreases when 2 nodes are highly connected, or when the length of any path decrease. The fact that it takes "connectivity" between nodes into account sets it apart from shortest path distances.

### 3.3 Computation of Euclidean Commute Time (ETC) Distance

#### 3.3.1 Closed form solution involving pseudoinverse

Define

- The **adjacency matrix**  $\mathbf{A} = (a_{ij})$  where  $a_{ij} = w_{ij}$  if node  $i$  is connected to  $j$ , otherwise  $a_{ij} = 0$
- $\mathbf{D} = \text{diag}(a_i)$  where  $a_i = \sum_j a_{ij}$
- The **Laplacian matrix** of the graph is  $\mathbf{L} = \mathbf{D} - \mathbf{A}$

Note the Laplacian  $\mathbf{L}$  is not full rank because  $\mathbf{1}$  (vector of 1s) is in its null space. Hence the following theorem involves a pseudoinverse  $\mathbf{L}^+$ :

### Theorem 3.1 Computaton of average commute time $n(i, j)$

We can compute average commute time between nodes  $i$  and  $j$  by:

$$n(i, j) = V_G(\mathbf{e}_i - \mathbf{e}_j)^t \mathbf{L}^+ (\mathbf{e}_i - \mathbf{e}_j)$$

where  $\mathbf{L}^+$  is the Moore-Penrose pseudoinverse of  $\mathbf{L}$ ,  $V_G = \sum_{i,j} a_{ij}$  is the volume of the graph, and  $\mathbf{e}_i$ 's are the standard basis vectors that is 0 everywhere and is 1 at position  $i$ .

*Proof.* See appendix of [2]. □

Observe that:

(1)  $\mathbf{L}^+$  is symmetric since  $\mathbf{L}$  is.

(2)  $\mathbf{L}^+$  is positive semidefinite, since  $\mathbf{L}$  psd  $\iff \mathbf{L}^+$  psd, and  $\mathbf{L}$  is psd since it is diagonally dominant.

Here (1) + (2) above implies that  $\mathbf{L}^+$  defines an inner product between  $\mathbf{x}$  and  $\mathbf{y}$  as  $\langle \mathbf{x}, \mathbf{y} \rangle = \mathbf{x}^t \mathbf{L}^+ \mathbf{y}$  in  $\mathbb{R}^n$ . This induces a norm:  $\|\mathbf{x}\| = (\mathbf{x}^t \mathbf{L}^+ \mathbf{x})^{1/2}$ . Therefore, the quantity  $[n(i, j)]^{1/2}$  is called the **Euclidean Commute Time (ETC) Distance**.

#### 3.3.2 Closed form formula for pseudoinverse

One way to compute the pseudoinverse is to do  $\mathbf{L}^+ = (\mathbf{L} - \mathbf{1}\mathbf{1}^t/n)^{-1} - \mathbf{1}\mathbf{1}^t/n$  via sparse (low rank) Cholesky factorization [2] (homework). For large graphs, [7] suggests Markov chain Monte Carlo.

#### 3.3.3 Iterative methods for ETC distance

An alternative is to use iterative methods. Using a more familiar notation, we observe that the recursive relation for  $m(k|i)$  in equation (3.1) can be interpreted as a matrix-matrix multiplication:

$$m_{ik} = \sum_{j=1}^N p_{ij} m_{jk} + 1 \iff \mathbf{M}^{(n+1)} = \mathbf{P}\mathbf{M}^{(n)} + (\mathbf{1}\mathbf{1}^t - \text{diag}(\mathbf{1})) \quad (3.2)$$

If the degree for each node is small, then  $\mathbf{P}$  is a sparse matrix, so the multiplication  $\mathbf{P}\mathbf{M}^{(n)}$  is linear in the number of non-zero entries. The programming illustrations below uses 50 iterations to estimate  $\mathbf{M}$ , which appears to be sufficient for the problems tested.

#### 3.4 Revised k-means clustering using ETC Distance

- (1) Connect each node to its  $h$  nearest neighbors, then add the edges of a minimum spanning tree. Then add weights  $w_{ij}$  using your favorite method (e.g. inverse Euclidean distance).
- (2) Choose number of clusters  $k$  and its respective cluster centers  $\mathbf{p}_1, \dots, \mathbf{p}_k$  where each  $\mathbf{p}_i$  is a node.
- (3) Assign each node  $\mathbf{x}_i$  to the nearest cluster  $C_l$  by finding which of  $\mathbf{p}_1, \dots, \mathbf{p}_k$  is closest to  $\mathbf{x}_i$ , where closeness is measured as a function of the ETC distance:  $\text{dist}(\mathbf{x}_i, \mathbf{p}_j) = n(i, j)^2$

(4) Recompute new cluster centers  $\mathbf{p}_1, \dots, \mathbf{p}_k$  by minimizing the within-cluster distance:

$$\mathbf{p}_l = \operatorname{argmin}_{\mathbf{x}_j} \left\{ \sum_{\mathbf{x}_k \in C_l} n(k, j)^2 \right\}$$

(5) Repeat (2) and (3) until convergence.

### 3.5 Some simulations

It is well known that K-means algorithm assumes Gaussian clusters with the same variance. Below illustrates 2 scenarios, one where k-means is expected to work, another where k-means is expected to fail. Here the graph is formed by connecting each node to 3 of its nearest neighbors with weight equal to the inverse Euclidean distance. Each node  $i$  is connected to node  $i + 1$  except for the transition node between the 2 clusters. For a better alternative to guarantee connectedness, [7] suggests to add edges from a **minimal spanning tree** (homework).

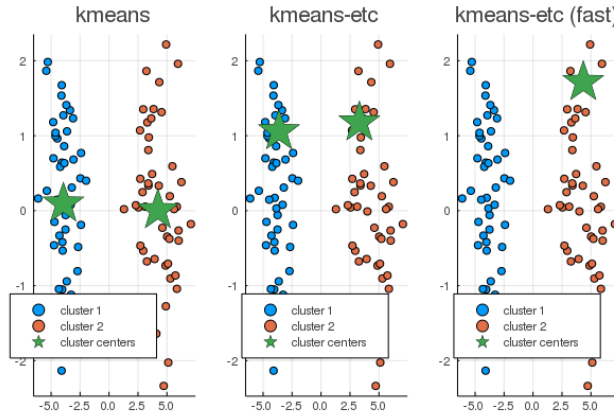


Figure 1: Two unit variance Gaussian clusters with mean -4 and 4.

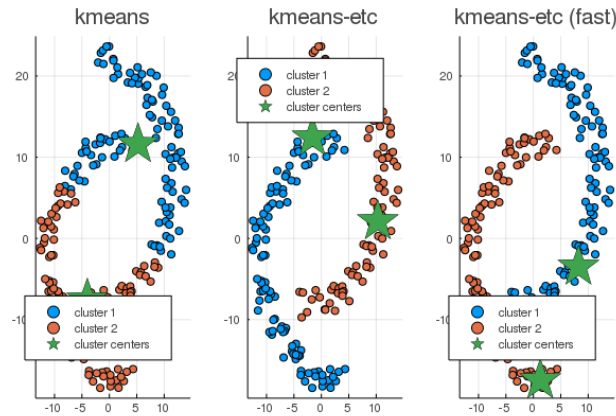
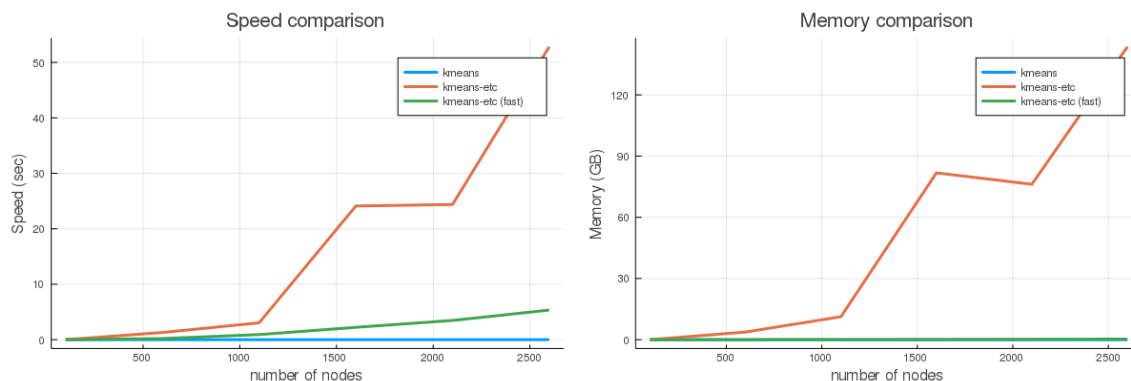


Figure 2: Two non-convex clusters.



Below compares scalability of the k-means, k-means using closed form pseudoinverse of ETC distance (kmeans-etc), and k-means using the iterative algorithm to estimate ETC distance (kmeans-etc (fast)). One should not take it too seriously because I did not try very hard to optimize this code.



## 4 Problems

In honor of Ken Lange, you are required to do 2 problems. If you do more I will grade your top 2 problems. Every problem is worth the same number of points. If a problem has subproblems, each subproblem is worth the same number of points.

### Problem 4.1 Bounds of binomial coefficients

For integers  $n$  and  $k$ , prove the following inequalities

$$\frac{n^k}{k^k} \leq \binom{n}{k} \leq \frac{n^k}{k!} < \left(\frac{ne}{k}\right)^k$$

which is used in part 2 of our sharp threshold proof. For the strict inequality, rewrite  $\frac{n^k}{k!} = \left(\frac{n}{k}\right)^k \frac{k^k}{k!}$  and use Taylor expansion on  $e^k$ .

### Problem 4.2 Verify this unproved claim

Prove that  $\mathbf{L}^+$  in theorem 3.1 can be computed via

$$\mathbf{L}^+ = \left( \mathbf{L} - \frac{\mathbf{1}\mathbf{1}^t}{n} \right)^{-1} + \frac{\mathbf{1}\mathbf{1}^t}{n}$$

where  $\mathbf{1}$  is a vector of 1s. This is equation (3) in [2], but it came without a proof.

### Problem 4.3 Mandatory computational problem

Starting with my code, do one of the computational problems below:

- Compute the pseudoinverse  $\mathbf{L}^+$  via sparse (low rank) Cholesky factorization. Recall Cholesky

for  $\mathbf{Ax} = \mathbf{b}$  is solved via back substitution:  $\mathbf{Ax} = \mathbf{b} \iff \mathbf{LL}'\mathbf{x} = \mathbf{b} \iff \mathbf{L}'\mathbf{x} = \mathbf{y}$  and  $\mathbf{Ly} = \mathbf{b}$ .

- In my code for computing the adjacency matrix  $\mathbf{A}$ , use your favorite algorithm to add graph edges corresponding to the **minimum spanning tree**. Hopefully this solves most of our stability problem.

#### Problem 4.4 Colorings of graphs

Let  $K_z$  be a **complete graph** where all  $z \in \mathbb{Z}_+$  nodes forms an edge with every other node. With equal probability, each edge is colored with red or green. Prove that  $z = 6$  is the minimal number of nodes needed to guarantee the existence of a **monochromatic triangle** (i.e. triangle with all edges the same color). This type of problem is what Ramsey theory studies, which we almost did.

We have just shown  $R(3,3) = 6$ . Similarly,  $R(3,4)$  is the minimal number of nodes to guarantee a red triangle or green rectangle. Function  $R$  obviously generalizes to more colors and shapes. Using Erdos' probabilistic method, Ramsey's theorem (see [6] or theorem 3.3 of [5]) says this number is finite but exponential. This takes us to Erdos' famous quote:

Suppose aliens invade earth and threaten to obliterate us within a year unless human beings can find  $R(5,5)$ . We could marshal the world's best minds and fastest computers, and within a year we could probably calculate the value. However, if the aliens demanded  $R(6,6)$ , we would have no choice but to launch a preemptive attack.

If you want to be famous, find  $R(5,5)$ .

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