

# Multivariate Statistics

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Source: Jhonson & Winchern

## 1 Principal Component Analysis

- A principal component analysis (PCA) is concerned with explaining the variance-covariance structure of a set of variables through a few linear combinations of these variables.
- Objectives
  - Data reduction
  - Data interpretation

- By PCA we select  $k$  principal components from a set for  $p(\geq k)$  initial variables such that the total system variability is retained as much as possible.
- Data-set of size  $(n \times p) \xrightarrow{PCA} \text{Data-set of size } (n \times k)$
- Note
  - To retain the total system variability, we need to retain all the  $p$  principal components

# Population Principal Components I

- Principal components are particular linear combinations of the  $p$  random features/variables  $X_1, X_2, \dots, X_p$ .
  - These linear combinations represents selection of new coordinate system obtained by rotating the original system with  $X_1, X_2, \dots, X_p$  as the coordinate axes
    - the new axes represent the direction with maximum variability and
    - provide a simpler and more parsimonious description of the covariance structure
- Note:
  - Principal components depends solely on the covariance matrix  $\Sigma$  of  $X_1, X_2, \dots, X_p$ .
  - Their development does not require a multivariate normal assumption.
  - However, standard results on inference can be used if the samples are assumed to be coming from normal population

# Population Principal Components II

- Formal definition
  - First principal component
    - is the linear combination of  $\mathbf{a}_1'\mathbf{X}$  that maximizes  $Var(\mathbf{a}_1'\mathbf{X})$  subject to  $\mathbf{a}_1'\mathbf{a}_1 = 1$
  - Second principal component
    - is the linear combination of  $\mathbf{a}_2'\mathbf{X}$  that maximizes  $Var(\mathbf{a}_2'\mathbf{X})$  subject to  $\mathbf{a}_2'\mathbf{a}_2 = 1$  and  $Cov(\mathbf{a}_2'\mathbf{X}, \mathbf{a}_1'\mathbf{X}) = 0$
  - . . . . .
  - $i$ th principal component
    - is the linear combination of  $\mathbf{a}_i'\mathbf{X}$  that maximizes  $Var(\mathbf{a}_i'\mathbf{X})$  subject to  $\mathbf{a}_i'\mathbf{a}_i = 1$  and  $Cov(\mathbf{a}_i'\mathbf{X}, \mathbf{a}_k'\mathbf{X}) = 0$  for all  $k < i$

# Population Principal Components III

- Result: Let  $\Sigma$  be the covariance matrix associated with random vector  $\mathbf{X} = [X_1, X_2, \dots, X_p]'$ . Let  $\Sigma$  have the eigenvalue-eigenvector pairs  $(\lambda_1, \mathbf{e}_1), (\lambda_2, \mathbf{e}_2), \dots, (\lambda_p, \mathbf{e}_p)$ , where  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$ . Then the  $i$ th principal component is given by

$$Y_i = \mathbf{e}_i' \mathbf{X} = e_{i1}X_1 + e_{i2}X_2 + \dots + e_{ip}X_p, \text{ for } i = 1, \dots, p$$

- With these choices,

$$\text{Var}(Y_i) = \mathbf{e}_i' \Sigma \mathbf{e}_i = \lambda_i, \text{ for } i = 1, \dots, p$$

$$\text{Cov}(Y_i, Y_k) = \mathbf{e}_i' \Sigma \mathbf{e}_k = 0, \text{ for } i \neq k$$

- Note: If some  $\lambda_i$  are equal then the choices of the corresponding coefficient vectors  $\mathbf{e}_i$ , and hence  $Y_i$ , are not unique.

# Population Principal Components IV

- Sketch of proof:

To get the first principal component, we need

$$\max_{\mathbf{a}} (\text{Var}(\mathbf{a}'\mathbf{X})) \text{ s. t. } \mathbf{a}'\mathbf{a} = 1 \Rightarrow \max_{\mathbf{a}} \left( \frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right)$$

Thus (*Lemma: Maximization of Quadratic Forms for Points on the Unit Sphere*),

$$\max_{\mathbf{a}} \left( \frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right) = \lambda_1$$

and maximum is attained at  $\mathbf{a} = \mathbf{e}_1$ .

Hence,  $Y_1 = \mathbf{e}_1'\mathbf{X}$  and  $\text{Var}(Y_1) = \mathbf{e}_1'\Sigma\mathbf{e}_1 = \lambda_1$



# Population Principal Components V

- Sketch of proof (contd.):

To get the  $i$ th principal component, we need

$$\max_{\mathbf{a}} (\text{Var}(\mathbf{a}'\mathbf{X})) \text{ s. t. } \mathbf{a}'\mathbf{a} = 1 \text{ and } \text{Cov}(\mathbf{a}'\mathbf{X}, \mathbf{a}_k'\mathbf{X}) = 0 \text{ for all } k < i$$

$$\Rightarrow \max_{\mathbf{a}} \left( \frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right) \text{ s. t. } \text{Cov}(\mathbf{a}'\mathbf{X}, \mathbf{e}_k'\mathbf{X}) = 0 \text{ for all } k < i$$

$$\Rightarrow \max_{\mathbf{a} \perp \mathbf{e}_1, \dots, \mathbf{e}_{i-1}} \left( \frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right), [\text{since } \mathbf{a}'\Sigma\mathbf{e}_k = \mathbf{a}'\lambda_k\mathbf{e}_k = 0 \Rightarrow \mathbf{a} \perp \mathbf{e}_k]$$

Thus (*Lemma: Maximization of Quadratic Forms for Points on the Unit Sphere*),

$$\max_{\mathbf{a} \perp \mathbf{e}_1, \dots, \mathbf{e}_{i-1}} \left( \frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right) = \lambda_i$$

and maximum is attained at  $\mathbf{a} = \mathbf{e}_i$ .

Hence,  $Y_i = \mathbf{e}_i'\mathbf{X}$  and  $\text{Var}(Y_i) = \mathbf{e}_i'\Sigma\mathbf{e}_i = \lambda_i$

Also,

$$\text{Cov}(Y_i, Y_k) = \text{Cov}(\mathbf{e}_i'\mathbf{X}, \mathbf{e}_k'\mathbf{X}) = \mathbf{e}_i'\Sigma\mathbf{e}_k = 0$$

# Population Principal Components VI

- Result: Let the random vector  $\mathbf{X} = [X_1, X_2, \dots, X_p]'$  have covariance matrix  $\Sigma$ , with the eigenvalue- eigenvector pairs  $(\lambda_1, \mathbf{e}_1), (\lambda_2, \mathbf{e}_2), \dots, (\lambda_p, \mathbf{e}_p)$ , where  $\lambda_1 \geq \lambda_2 \geq \dots \lambda_p \geq 0$ . Let  $Y_1 = \mathbf{e}_1' \mathbf{X}$ ,  $Y_2 = \mathbf{e}_2' \mathbf{X}$ ,  $\dots$ ,  $Y_p = \mathbf{e}_p' \mathbf{X}$  be the principal components. Then

$$\begin{aligned} \sum_{i=1}^p \text{Var}(X_i) &= \sigma_{11} + \sigma_{22} + \dots + \sigma_{pp} \\ &= \text{tr}(\Sigma) \\ &= \lambda_1 + \lambda_2 + \dots + \lambda_p \\ &= \sum_{i=1}^p \text{Var}(Y_i). \end{aligned}$$

# Population Principal Components VII

- Proportion of total population variance explained by  $k$ th principal component:

$$\frac{\lambda_k}{\lambda_1 + \dots + \lambda_k + \dots + \lambda_p}$$

- Proportion of total population variance explained by first  $k$  principal components:

$$\frac{\lambda_1 + \dots + \lambda_k}{\lambda_1 + \dots + \lambda_k + \dots + \lambda_p}$$

# Population Principal Components VIII

- Result: If  $Y_1 = \mathbf{e}_1' \mathbf{X}$ ,  $Y_2 = \mathbf{e}_2' \mathbf{X}$ ,  $\dots$ ,  $Y_p = \mathbf{e}_p' \mathbf{X}$  are the principal components obtained from the covariance matrix  $\Sigma$ , then

$$\rho_{Y_i, X_k} = \frac{e_{ik} \sqrt{\lambda_i}}{\sqrt{\sigma_{kk}}}, \text{ for } i, k = 1, 2, \dots, p$$

are the correlation coefficients between the components  $Y_i$  and the variables  $X_k$ . Here  $(\lambda_1, \mathbf{e}_1), (\lambda_2, \mathbf{e}_2), \dots, (\lambda_p, \mathbf{e}_p)$  are the eigenvalue- eigenvector pairs for  $\Sigma$ .

- The magnitude of  $e_{ik}$  measures the importance of  $k$ th variable ( $X_k$ ) to the  $i$ th principal component ( $Y_i$ )

# Population Principal Components IX

- Sketch of proof:

$$\begin{aligned}\rho_{Y_i, X_k} &= \frac{\text{Cor}(Y_i, X_k)}{\sqrt{\text{Var}(Y_i) \text{Var}(X_k)}} \\&= \frac{\text{Cov}(\mathbf{e}_i' \mathbf{X}, [0 \ 0 \ \dots \ 1 \ \dots \ 0] \mathbf{X})}{\sqrt{\lambda_i \sigma_{kk}}} \\&= \frac{[0 \ 0 \ \dots \ 1 \ \dots \ 0] \Sigma \mathbf{e}_i}{\sqrt{\lambda_i \sigma_{kk}}} \\&= \frac{\lambda_i \mathbf{e}_{ik}}{\sqrt{\lambda_i \sigma_{kk}}} = \frac{\sqrt{\lambda_i} \mathbf{e}_{ik}}{\sqrt{\sigma_{kk}}}\end{aligned}$$

Example 8.1 (Page: 434)

# Principal Components on Standardized Variable I

- Given the vector  $\mathbf{X}$ , the standardized vector can be obtained as

$$\mathbf{Z} = \mathbf{V}^{-\frac{1}{2}}(\mathbf{X} - \boldsymbol{\mu}),$$

recall  $\mathbf{V} = \begin{bmatrix} \sigma_{11} & \dots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \dots & \sigma_{pp} \end{bmatrix}.$

- Note:

- $E(\mathbf{Z}) = \mathbf{0} = [0 \dots 0]'$

- $Cov(\mathbf{Z}) = \boldsymbol{\rho} = \begin{bmatrix} 1 & \rho_{12} & \dots & \rho_{1p} \\ \vdots & \ddots & \vdots & \\ \rho_{1p} & \rho_{2p} & \dots & 1 \end{bmatrix}.$

# Principal Components on Standardized Variable II

- Result: The  $i$ th principal component of the standardized variables  $\mathbf{Z} = [Z_1 \ Z_2 \ \dots \ Z_p]'$  with  $\text{Cov}(\mathbf{Z}) = \rho$ , is given by

$$Y_i = \mathbf{e}_i' \mathbf{Z}, \text{ for } i = 1, 2, \dots, p.$$

Moreover,

$$\sum_{i=1}^p \text{Var}(Y_i) = \sum_{i=1}^p \text{Var}(Z_i) = p.$$

In this case,  $(\lambda_1, \mathbf{e}_1), (\lambda_2, \mathbf{e}_2), \dots, (\lambda_p, \mathbf{e}_p)$  are the eigenvalue-eigenvector pairs for  $\rho$ , with  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$ .

# Principal Components on Standardized Variable III

- Proportion of total population variance explained by  $k$ th principal component:

$$\frac{\lambda_k}{p}$$

- Proportion of total population variance explained by first  $k$  principal components:

$$\frac{\lambda_1 + \dots + \lambda_k}{p}$$

Example 8.2 (Page: 437)



# Summarizing Sample Variations by Principal Components I

- Result: Let  $\mathbf{X}$  be the observation on the variables  $X_1, X_2, \dots, X_p$  with the corresponding sample covariance matrix  $S_{p \times p}$ . Then the  $i$ th sample principal component is given by

$$\hat{Y}_i = \hat{\mathbf{e}}_i' \mathbf{X} = \hat{e}_{i1} X_1 + \dots + \hat{e}_{ip} X_p \text{ for } i = 1, 2, \dots, p,$$

where  $(\hat{\lambda}_1, \hat{\mathbf{e}}_1), (\hat{\lambda}_2, \hat{\mathbf{e}}_2), \dots, (\hat{\lambda}_p, \hat{\mathbf{e}}_p)$  are the eigenvalue-eigenvector pairs for  $S$ , with  $\hat{\lambda}_1 \geq \hat{\lambda}_2 \geq \dots \geq \hat{\lambda}_p \geq 0$ . Also,

$$\text{Var}(\hat{Y}_i) = \hat{\lambda}_i, \text{ for } i = 1, 2, \dots, p$$

and

$$\text{Cov}(\hat{Y}_i, \hat{Y}_k) = 0, \text{ for } i \neq k.$$

In addition,

$$\text{Total Sample Variance} = \sum_{i=1}^p s_{ii} = \sum_{i=1}^p \hat{\lambda}_i$$

# Summarizing Sample Variations by Principal Components II

- Result: Let  $\mathbf{Z}$  be the observation on the variables  $Z_i \left( = \frac{X_i - \bar{X}_i}{\sqrt{s_{ii}}} \right)$ ,  $i = 1, \dots, p$ , with the corresponding sample covariance matrix  $R_{p \times p}$ . Then the  $i$ th sample principal component is given by

$$\hat{Y}_i = \hat{\mathbf{e}}_i' \mathbf{Z} = \hat{e}_{i1} Z_1 + \dots + \hat{e}_{ip} Z_p \text{ for } i = 1, 2, \dots, p,$$

where  $(\hat{\lambda}_1, \hat{\mathbf{e}}_1), (\hat{\lambda}_2, \hat{\mathbf{e}}_2), \dots, (\hat{\lambda}_p, \hat{\mathbf{e}}_p)$  are the eigenvalue-eigenvector pairs for  $R$ , with  $\hat{\lambda}_1 \geq \hat{\lambda}_2 \geq \dots \geq \hat{\lambda}_p \geq 0$ . Also,

$$\text{Var}(\hat{Y}_i) = \hat{\lambda}_i, \text{ for } i = 1, 2, \dots, p$$

and

$$\text{Cov}(\hat{Y}_i, \hat{Y}_k) = 0, \text{ for } i \neq k.$$

In addition,

$$\text{Total Sample Variance} = \sum_{i=1}^p \hat{\lambda}_i = p$$

Example 8.3 (Page: 443)

# Summarizing Sample Variations by Principal Components III

- How many principal components to be retained?
- No definite answer.
- Subjectively, we decide on
  - the relative size of the eigenvalues and the amount of sample variation explained
  - subject-matter interpretations of the components is also important
- Visual aid: Scree Plot
  - Plot of  $\hat{\lambda}_i$  vs  $i$
  - To determine the appropriate number of components we look for an elbow (bend) in the scree plot
    - The number of components is taken to be the point at which the remaining eigenvalues are relatively small and all about the same size.

Example 8.4 (Page: 445)