

Multivariate Statistics

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Slides adapted from Jhonson & Winchern

1 Principal Component Analysis

- A principal component analysis (PCA) is concerned with explaining the variance-covariance structure of a set of variables through a few linear combinations of these variables.
- Objectives
 - Data reduction
 - Data interpretation

- By PCA we select k principal components from a set for $p(\geq k)$ initial variables such that the total system variability is retained as much as possible.
- Data-set of size $(n \times p) \xrightarrow{PCA} \text{Data-set of size } (n \times k)$
- Note
 - To retain the total system variability, we need to retain all the p principal components

Population Principal Components I

- Principal components are particular linear combinations of the p random features/variables X_1, X_2, \dots, X_p .
 - These linear combinations represents selection of new coordinate system obtained by rotating the original system with X_1, X_2, \dots, X_p as the coordinate axes
 - the new axes represent the direction with maximum variability and
 - provide a simpler and more parsimonious description of the covariance structure
- Note:
 - Principal components depends solely on the covariance matrix Σ of X_1, X_2, \dots, X_p .
 - Their development does not require a multivariate normal assumption.
 - However, standard results on inference can be used if the samples are assumed to be coming from normal population

Population Principal Components II

- Formal definition
 - First principal component
 - is the linear combination of $\mathbf{a}_1'\mathbf{X}$ that maximizes $Var(\mathbf{a}_1'\mathbf{X})$ subject to $\mathbf{a}_1'\mathbf{a}_1 = 1$
 - Second principal component
 - is the linear combination of $\mathbf{a}_2'\mathbf{X}$ that maximizes $Var(\mathbf{a}_2'\mathbf{X})$ subject to $\mathbf{a}_2'\mathbf{a}_2 = 1$ and $Cov(\mathbf{a}_2'\mathbf{X}, \mathbf{a}_1'\mathbf{X}) = 0$
 -
 - i th principal component
 - is the linear combination of $\mathbf{a}_i'\mathbf{X}$ that maximizes $Var(\mathbf{a}_i'\mathbf{X})$ subject to $\mathbf{a}_i'\mathbf{a}_i = 1$ and $Cov(\mathbf{a}_i'\mathbf{X}, \mathbf{a}_k'\mathbf{X}) = 0$ for all $k < i$

Population Principal Components III

- Result: Let Σ be the covariance matrix associated with random vector $\mathbf{X} = [X_1, X_2, \dots, X_p]'$. Let Σ have the eigenvalue-eigenvector pairs $(\lambda_1, \mathbf{e}_1), (\lambda_2, \mathbf{e}_2), \dots, (\lambda_p, \mathbf{e}_p)$, where $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$. Then the i th principal component is given by

$$Y_i = \mathbf{e}_i' \mathbf{X} = e_{i1}X_1 + e_{i2}X_2 + \dots + e_{ip}X_p, \text{ for } i = 1, \dots, p$$

- With these choices,

$$\text{Var}(Y_i) = \mathbf{e}_i' \Sigma \mathbf{e}_i = \lambda_i, \text{ for } i = 1, \dots, p$$

$$\text{Cov}(Y_i, Y_k) = \mathbf{e}_i' \Sigma \mathbf{e}_k = 0, \text{ for } i \neq k$$

- Note: If some λ_i are equal then the choices of the corresponding coefficient vectors \mathbf{e}_i , and hence Y_i , are not unique.

Population Principal Components IV

- Sketch of proof:

To get the first principal component, we need

$$\max_{\mathbf{a}} (\text{Var}(\mathbf{a}'\mathbf{X})) \text{ s. t. } \mathbf{a}'\mathbf{a} = 1 \Rightarrow \max_{\mathbf{a}} \left(\frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right)$$

Thus (*Lemma: Maximization of Quadratic Forms for Points on the Unit Sphere*),

$$\max_{\mathbf{a}} \left(\frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right) = \lambda_1$$

and maximum is attained at $\mathbf{a} = \mathbf{e}_1$.

Hence, $Y_1 = \mathbf{e}_1'\mathbf{X}$ and $\text{Var}(Y_1) = \mathbf{e}_1'\Sigma\mathbf{e}_1 = \lambda_1$

Population Principal Components V

- Sketch of proof (contd.):

To get the i th principal component, we need

$$\max_{\mathbf{a}} (\text{Var}(\mathbf{a}'\mathbf{X})) \text{ s. t. } \mathbf{a}'\mathbf{a} = 1 \text{ and } \text{Cov}(\mathbf{a}'\mathbf{X}, \mathbf{a}_k'\mathbf{X}) = 0 \text{ for all } k < i$$

$$\Rightarrow \max_{\mathbf{a}} \left(\frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right) \text{ s. t. } \text{Cov}(\mathbf{a}'\mathbf{X}, \mathbf{e}_k'\mathbf{X}) = 0 \text{ for all } k < i$$

$$\Rightarrow \max_{\mathbf{a} \perp \mathbf{e}_1, \dots, \mathbf{e}_{i-1}} \left(\frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right), [\text{since } \mathbf{a}'\Sigma\mathbf{e}_k = \mathbf{a}'\lambda_k\mathbf{e}_k = 0 \Rightarrow \mathbf{a} \perp \mathbf{e}_k]$$

Thus (*Lemma: Maximization of Quadratic Forms for Points on the Unit Sphere*),

$$\max_{\mathbf{a} \perp \mathbf{e}_1, \dots, \mathbf{e}_{i-1}} \left(\frac{\mathbf{a}'\Sigma\mathbf{a}}{\mathbf{a}'\mathbf{a}} \right) = \lambda_i$$

and maximum is attained at $\mathbf{a} = \mathbf{e}_i$.

Hence, $Y_i = \mathbf{e}_i'\mathbf{X}$ and $\text{Var}(Y_i) = \mathbf{e}_i'\Sigma\mathbf{e}_i = \lambda_i$

Also,

$$\text{Cov}(Y_i, Y_k) = \text{Cov}(\mathbf{e}_i'\mathbf{X}, \mathbf{e}_k'\mathbf{X}) = \mathbf{e}_i'\Sigma\mathbf{e}_k = 0$$

Population Principal Components VI

- Result: Let the random vector $\mathbf{X} = [X_1, X_2, \dots, X_p]'$ have covariance matrix Σ , with the eigenvalue- eigenvector pairs $(\lambda_1, \mathbf{e}_1), (\lambda_2, \mathbf{e}_2), \dots, (\lambda_p, \mathbf{e}_p)$, where $\lambda_1 \geq \lambda_2 \geq \dots \lambda_p \geq 0$. Let $Y_1 = \mathbf{e}_1' \mathbf{X}$, $Y_2 = \mathbf{e}_2' \mathbf{X}$, \dots , $Y_p = \mathbf{e}_p' \mathbf{X}$ be the principal components. Then

$$\begin{aligned} \sum_{i=1}^p \text{Var}(X_i) &= \sigma_{11} + \sigma_{22} + \dots + \sigma_{pp} \\ &= \text{tr}(\Sigma) \\ &= \lambda_1 + \lambda_2 + \dots + \lambda_p \\ &= \sum_{i=1}^p \text{Var}(Y_i). \end{aligned}$$

Population Principal Components VII

- Proportion of total population variance explained by k th principal component:

$$\frac{\lambda_k}{\lambda_1 + \dots + \lambda_k + \dots + \lambda_p}$$

- Proportion of total population variance explained by first k principal components:

$$\frac{\lambda_1 + \dots + \lambda_k}{\lambda_1 + \dots + \lambda_k + \dots + \lambda_p}$$

Population Principal Components VIII

- Result: If $Y_1 = \mathbf{e}_1' \mathbf{X}$, $Y_2 = \mathbf{e}_2' \mathbf{X}$, \dots , $Y_p = \mathbf{e}_p' \mathbf{X}$ are the principal components obtained from the covariance matrix Σ , then

$$\rho_{Y_i, X_k} = \frac{e_{ik} \sqrt{\lambda_i}}{\sqrt{\sigma_{kk}}}, \text{ for } i, k = 1, 2, \dots, p$$

are the correlation coefficients between the components Y_i and the variables X_k . Here $(\lambda_1, \mathbf{e}_1), (\lambda_2, \mathbf{e}_2), \dots, (\lambda_p, \mathbf{e}_p)$ are the eigenvalue- eigenvector pairs for Σ .

- The magnitude of e_{ik} measures the importance of k th variable (X_k) to the i th principal component (Y_i)

Population Principal Components IX

- Sketch of proof:

$$\begin{aligned}\rho_{Y_i, X_k} &= \text{Cor}(Y_i, X_k) \\ &= \frac{\text{Cov}(\mathbf{e}_i' \mathbf{X}, [0 \ 0 \ \dots \underset{k^{th}}{1} \ \dots \ 0] \mathbf{X})}{\sqrt{\text{Var}(Y_i) \text{Var}(X_k)}} \\ &= \frac{[0 \ 0 \ \dots \underset{k^{th}}{1} \ \dots \ 0] \Sigma \mathbf{e}_i}{\sqrt{\lambda_i \sigma_{kk}}} \\ &= \frac{\lambda_i \mathbf{e}_{ik}}{\sqrt{\lambda_i \sigma_{kk}}} = \frac{\sqrt{\lambda_i} \mathbf{e}_{ik}}{\sqrt{\sigma_{kk}}}\end{aligned}$$

Principal Components on Standardized Variable I

- Given the vector \mathbf{X} , the standardized vector can be obtained as

$$\mathbf{Z} = \mathbf{V}^{-\frac{1}{2}}(\mathbf{X} - \boldsymbol{\mu}),$$

recall $\mathbf{V} = \begin{bmatrix} \sigma_{11} & \dots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \dots & \sigma_{pp} \end{bmatrix}.$

- Note:

- $E(\mathbf{Z}) = \mathbf{0} = [0 \dots 0]'$

- $Cov(\mathbf{Z}) = \boldsymbol{\rho} = \begin{bmatrix} 1 & \rho_{12} & \dots & \rho_{1p} \\ \vdots & \ddots & \vdots & \\ \rho_{1p} & \rho_{2p} & \dots & 1 \end{bmatrix}.$

Principal Components on Standardized Variable II

- Result: The i th principal component of the standardized variables $\mathbf{Z} = [Z_1 \ Z_2 \ \dots \ Z_p]'$ with $\text{Cov}(\mathbf{Z}) = \rho$, is given by

$$Y_i = \mathbf{e}_i' \mathbf{Z}, \text{ for } i = 1, 2, \dots, p.$$

Moreover,

$$\sum_{i=1}^p \text{Var}(Y_i) = \sum_{i=1}^p \text{Var}(Z_i) = p.$$

In this case, $(\lambda_1, \mathbf{e}_1), (\lambda_2, \mathbf{e}_2), \dots, (\lambda_p, \mathbf{e}_p)$ are the eigenvalue-eigenvector pairs for ρ , with $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p \geq 0$.

Principal Components on Standardized Variable III

- Proportion of total population variance explained by k th principal component:

$$\frac{\lambda_k}{p}$$

- Proportion of total population variance explained by first k principal components:

$$\frac{\lambda_1 + \dots + \lambda_k}{p}$$

Summarizing Sample Variations by Principal Components I

- Result: Let \mathbf{X} be the observation on the variables X_1, X_2, \dots, X_p with the corresponding sample covariance matrix $S_{p \times p}$. Then the i th sample principal component is given by

$$\hat{Y}_i = \hat{\mathbf{e}}_i' \mathbf{X} = \hat{e}_{i1} X_1 + \dots + \hat{e}_{ip} X_p \text{ for } i = 1, 2, \dots, p,$$

where $(\hat{\lambda}_1, \hat{\mathbf{e}}_1), (\hat{\lambda}_2, \hat{\mathbf{e}}_2), \dots, (\hat{\lambda}_p, \hat{\mathbf{e}}_p)$ are the eigenvalue-eigenvector pairs for S , with $\hat{\lambda}_1 \geq \hat{\lambda}_2 \geq \dots \geq \hat{\lambda}_p \geq 0$. Also,

$$\text{Var}(\hat{Y}_i) = \hat{\lambda}_i, \text{ for } i = 1, 2, \dots, p$$

and

$$\text{Cov}(\hat{Y}_i, \hat{Y}_k) = 0, \text{ for } i \neq k.$$

In addition,

$$\text{Total Sample Variance} = \sum_{i=1}^p s_{ii} = \sum_{i=1}^p \hat{\lambda}_i$$

Summarizing Sample Variations by Principal Components II

- Result: Let \mathbf{Z} be the observation on the variables $Z_i \left(= \frac{X_i - \bar{X}_i}{\sqrt{s_{ii}}} \right)$, $i = 1, \dots, p$, with the corresponding sample covariance matrix $R_{p \times p}$. Then the i th sample principal component is given by

$$\hat{Y}_i = \hat{\mathbf{e}}_i' \mathbf{Z} = \hat{e}_{i1} Z_1 + \dots + \hat{e}_{ip} Z_p \text{ for } i = 1, 2, \dots, p,$$

where $(\hat{\lambda}_1, \hat{\mathbf{e}}_1), (\hat{\lambda}_2, \hat{\mathbf{e}}_2), \dots, (\hat{\lambda}_p, \hat{\mathbf{e}}_p)$ are the eigenvalue-eigenvector pairs for R , with $\hat{\lambda}_1 \geq \hat{\lambda}_2 \geq \dots \geq \hat{\lambda}_p \geq 0$. Also,

$$\text{Var}(\hat{Y}_i) = \hat{\lambda}_i, \text{ for } i = 1, 2, \dots, p$$

and

$$\text{Cov}(\hat{Y}_i, \hat{Y}_k) = 0, \text{ for } i \neq k.$$

In addition,

$$\text{Total Sample Variance} = \sum_{i=1}^p \hat{\lambda}_i = p$$

Summarizing Sample Variations by Principal Components III

- How many principal components to be retained?
- No definite answer.
- Subjectively, we decide on
 - the relative size of the eigenvalues and the amount of sample variation explained
 - subject-matter interpretations of the components is also important
- Visual aid: Scree Plot
 - Plot of $\hat{\lambda}_i$ vs i
 - To determine the appropriate number of components we look for an elbow (bend) in the scree plot
 - The number of components is taken to be the point at which the remaining eigenvalues are relatively small and all about the same size.