Probability and Estimation

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Probability Overview

- Events
 - discrete random variables, continuous random variables, compound events
- Axioms of probability
 - What defines a reasonable theory of uncertainty
- Independent events
- Conditional probabilities
- Bayes rule and beliefs
- Joint probability distribution
- Expectations
- Independence, Conditional independence

Random Variables

- Informally, A is a random variable if
 - A denotes something about which we are uncertain
 - perhaps the outcome of a randomized experiment
- Examples
 - A=True if a randomly drawn person from our class is female
 - A=The hometown of a randomly drawn person from our class
 - A=True if two randomly drawn persons from our class have same birthday
- Define P(A) as "the fraction of possible worlds in which A is true" or "the fraction of times A holds, in repeated runs of the random experiment"
 - The set of possible worlds is called the sample space, S
 - A random variable A is a function defined over S

A:
$$S \to \{0,1\}$$

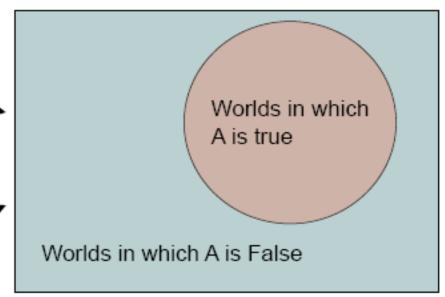
A little formalism

- A sample space S (e.g., set of students in our class)
 - aka the set of possible worlds
- A random variable is a function defined over the sample space
 - Gender: $S \rightarrow \{m,f\}$, Height: $S \rightarrow Reals$
- An event is a subset of S
 - e.g., the subset of S for which Gender=f
 - e.g., the subset of S for which (Gender=m) and (eyecolor=blue)
- We are often interested in probabilities of specific events
- And of specific events conditioned on other specific events

Visualizing A

Sample space of all possible worlds

Its area is 1

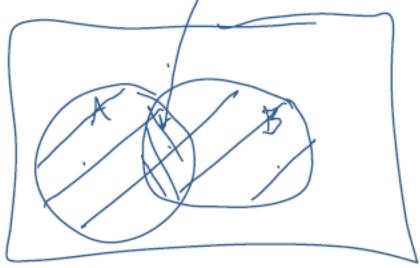


P(A) = Area of reddish oval

The Axioms of Probability

- $0 \le P(A) \le 1$
- P(True)=1
- P(False)=0

• P(A or B)=P(A)+P(B)-P(A and B)



Interpreting the axioms

- The area of A can't get any smaller than 0
- And a zero area would mean no world could ever have A true
- The area of A can't get any larger than 1
- And an area of 1 would mean all worlds will have A true

Theorems from the Axioms

- $0 \le P(A) \le 1$, P(True) = 1, P(False) = 0
- P(A or B)=P(A)+P(B)-P(A and B)

$$\rightarrow$$
 P(not A) = P(\sim A) = 1-P(A)

$$P(A \text{ or } \sim A) = 1$$
 $P(A \text{ and } \sim A) = 0$
 $P(A \text{ or } \sim A) = P(A) + P(\sim A) - P(A \text{ and } \sim A)$
 $1 = P(A) + P(\sim A) - 0$

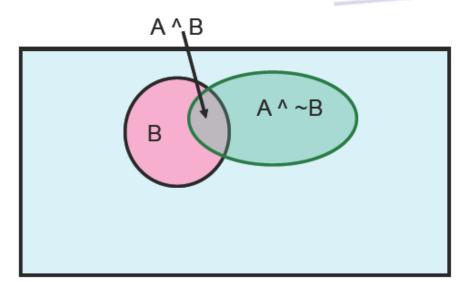
Elementary Probability in Pictures

- $P(\sim A) + P(A) = 1$
- $P(A)=P(A \land B) + P(A \land \sim B)$

A = A and $(B \text{ or } \sim B) = (A \text{ and } B) \text{ or } (A \text{ and } \sim B)$

$$P(A) = P(A \text{ and } B) + P(A \text{ and } \sim B) - P((A \text{ and } B) \text{ and } (A \text{ and } \sim B))$$

P(A) = P(A and B) + P(A and B) - P(A and A and B and B)



Multivalue Discrete Random Variables

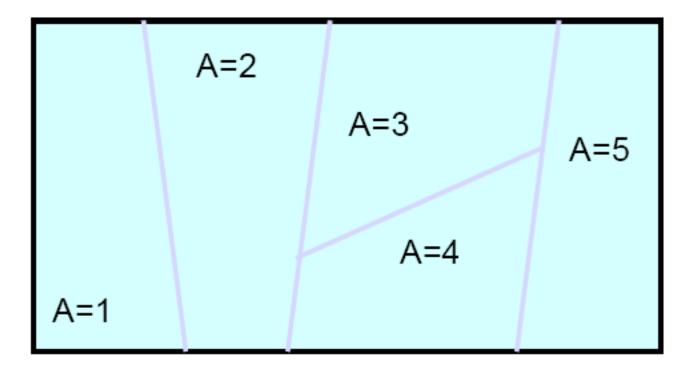
- Suppose A can take more than 2 values
- A is a random variable with arity k if it can take on exactly one value out of $\{v_1, v_2, ..., v_k\}$
- Thus...

$$P(A = v_i \land A = v_j) = 0 \text{ if } i \neq j$$

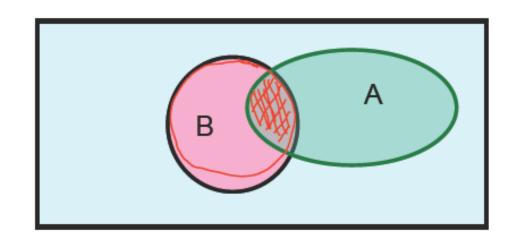
 $P(A = v_1 \lor A = v_2 \lor ... \lor A = v_k) = 1$

Elementary Probability in Pictures

$$\sum_{j=1}^k P(A = v_j) = 1$$



Definition of Conditional Probability



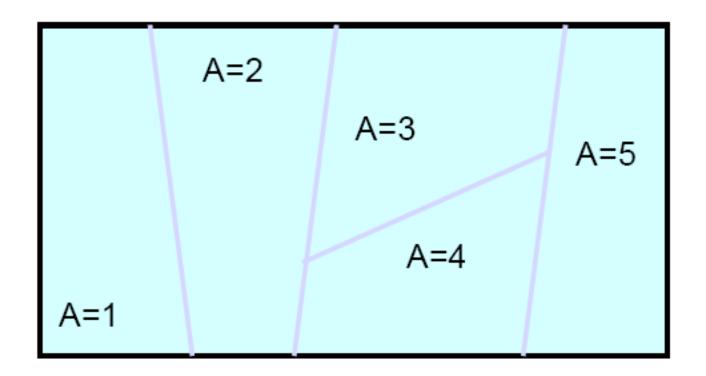
Corollary: The Chain Rule

$$P(A \land B) = P(A|B) P(B)$$

$$P(C \land A \land B) = P(C | A \land B)P(A | B)P(B)$$

Conditional Probability in Pictures

picture: P(B|A=2)



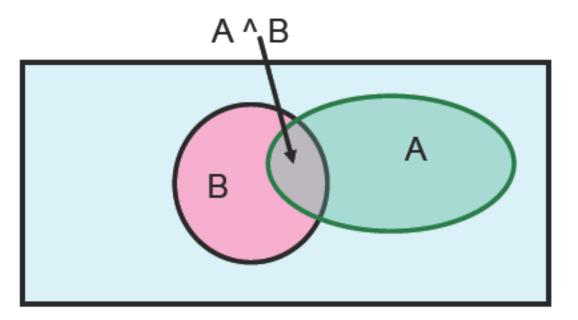
Independent Events

 Definition: two events A and B are independent if Pr(A and B)=Pr(A)*Pr(B)

• Intuition: knowing A tells us nothing about the value of B (and vice versa)

Elementary Probability in Pictures

• Let's write 2 expressions for P(A ^ B)



Bayes Rule

$$P(A|B) = \frac{P(B|A) * P(A)}{P(B)}$$
 Bayes' rule

we call P(A) the "prior"

and P(A|B) the "posterior"

Bayes, Thomas (1763) An essay towards solving a problem in the doctrine of chances. Philosophical Transactions of the Royal Society of London, 53:370-418

...by no means merely a curious speculation in the doctrine of chances, but necessary to be solved in order to a sure foundation for all our reasonings concerning past facts, and what is likely to be hereafter.... necessary to be considered by any that would give a clear account of the strength of analogical or inductive reasoning...

Other Forms of Bayes Rule

$$P(A | B) = \frac{P(B | A)P(A)}{P(B | A)P(A) + P(B | \sim A)P(\sim A)}$$

$$P(A|B \land X) = \frac{P(B|A \land X)P(A \land X)}{P(B \land X)}$$

Other Forms of Bayes Rule

$$P(A|BCD) = \frac{P(ABCD)}{P(BCD)}$$
(1)

$$P(A|BCD) = \frac{P(BCD|A)P(A)}{P(BCD)}$$
(2)

$$P(A|BCD) = \frac{P(B|ACD)P(A|CD)}{P(B|CD)}$$
(3)

$$P(A|BCD) = \frac{P(BC|AD)P(A|D)}{P(BC|D)}$$
(4)

Applying Bayes Rule

$$P(A \mid B) = \frac{P(B \mid A)P(A)}{P(B \mid A)P(A) + P(B \mid \sim A)P(\sim A)}$$

A=you have the flu, B=you just coughed

Assume:

$$P(A) = 0.05$$

$$P(B|A)=0.80$$

$$P(B|\sim A)=0.2$$

What does all this have to do with function Approximation??

 $f: X \rightarrow Y$

P(Y|X)

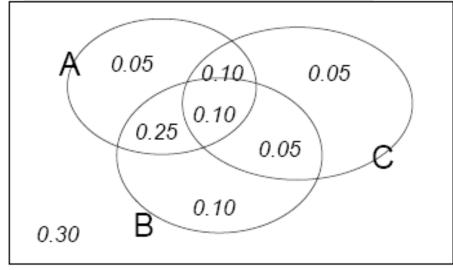
The Joint Distribution

Example: Boolean variables A, B, C

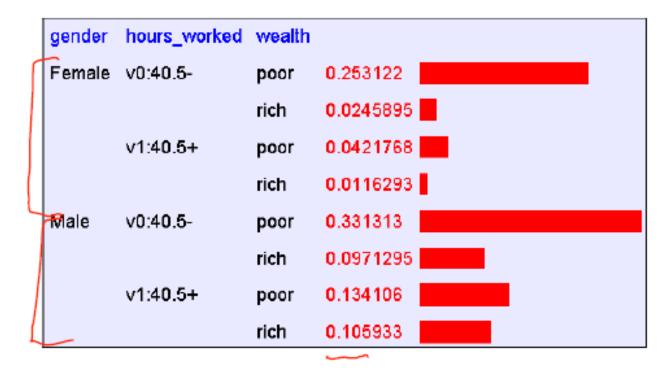
Recipe for making a joint distribution of M variables:

- Make a truth table listing all combinations of values of your variables (if there are M Boolean variables then the table will have 2^M rows).
- For each combination of values, say how probable it is.
- If you subscribe to the axioms of probability, those numbers must sum to 1.

Α	В	C	Prob
0	0	0	0.30
0	0	1	0.05
0	1	0	0.10
0	1	1	0.05
1	0	0	0.05
1	0	1	0.10
1	1	0	0.25
1	1	1	0.10



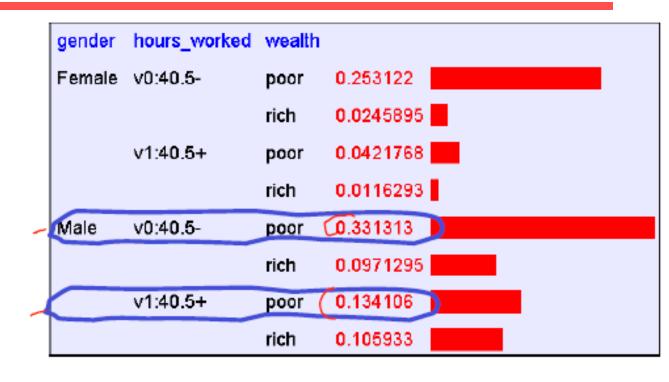
Using the Joint



One you have the JD you can ask for the probability of any logical expression involving your attribute

$$P(E) = \sum_{\text{rows matching } E} P(\text{row})$$

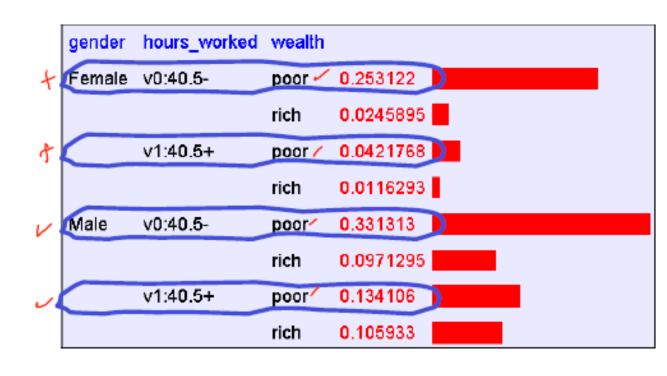
Using the Joint



$$P(Poor Male) = 0.4654$$

$$P(E) = \sum_{\text{rows matching } E} P(\text{row})$$

Using the Joint



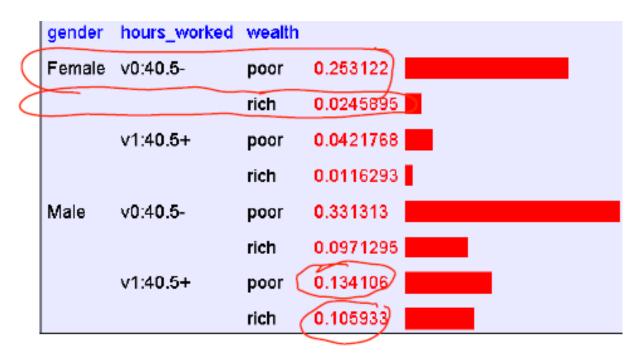
$$P(Poor) = 0.7604$$

$$P(E) = \sum_{\text{rows matching } E} P(\text{row})$$

Inference with the Joint

$$P(E_1 \mid E_2) = \frac{P(E_1 \land E_2)}{P(E_2)} = \frac{\sum_{\substack{rows \text{ matching } E_1 \text{ and } E_2 \\ rows \text{ matching } E_2}} \frac{P(row)}{P(row)}$$

Learning and the Joint Distribution



Suppose we want to learn the function f: <G, H>

W

Equivalently, P(W | G, H)

Solution: learn joint distribution from data, calculate P(W | G, H)

e.g., P(W=rich | G = female, H = 40.5-) =
$$\frac{.024}{.924 + .25}$$

Sounds like the solution to learning $F: X \rightarrow Y$, or P(Y|X).

Are we done?

Your first consulting job

- A billionaire asks you a questions:
 - He says: I have thumbtack, if I flip it, what's the probability it will fall with the nail up?
 - You say: please flip it a few times:



- -You say: The probability is: 0.6!
- -He says: Why???

Thumbtack – Binomial Distribution

■ P(Heads) =
$$\theta$$
, P(Tails) = $1-\theta$

Of tails outcomes

Of tails outcomes

At heads extreme

P(D| θ)= θ · θ · $(1-\theta)$ · θ · $(1-\theta)$ = θ · $(1-\theta)$

Flips produce data set D with α_H heads and α_T tails

- -Flips are independent, identically distributed 1's and 0's (Bernoulli)
- α_H and α_T are counts that sum these outcome (Binomial)

$$P(D \mid \theta) = P(\alpha_H, \alpha_T \mid \theta) = \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$$

Maximum Likelihood Estimation

- Data: Observed set D of α_H Heads and α_T Tails
- Hypothesis: Binomial distribution
- Learning θ is an optimization problem
 - What is the objective function?
- MLE: Choosing θ that maximizes the probability of observed data:

$$\hat{\theta} = \underset{\theta}{\operatorname{arg max}} P(D \mid \theta)$$

$$= \underset{\theta}{\operatorname{arg max}} \ln P(D \mid \theta)$$

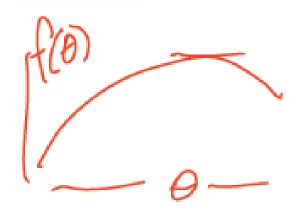
Maximum Likelihood Estimate for θ



$$\hat{\theta} = \arg\max_{\theta} P(D \mid \theta)$$

$$= \arg \max_{\theta} \ln P(D \mid \theta)$$

$$= \arg \max_{\theta} \ln \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$$



Set derivative to zero:

$$rac{d}{d heta}$$
 In $P(\mathcal{D} \mid heta) = 0$

$$\ln \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$$

$$= \alpha_H \ln \theta + \alpha_T \ln(1 - \theta)$$

$$\frac{d}{d\theta} \ln P(D \mid \theta) = \frac{d}{d\theta} (\alpha_H \ln \theta + \alpha_T \ln(1 - \theta))$$

$$=\alpha_H \frac{1}{\theta} - \alpha_T \frac{1}{1 - \theta} = 0$$

$$\hat{\theta} = \frac{\alpha_H}{\alpha_H + \alpha_T}$$

MLE & MAP

Bayesian Learning

MLE = arsmax P(DID)

Use Bayes rule:

argman
$$P(\theta \mid \mathcal{D})$$

$$= \frac{P(\mathcal{D} \mid \theta)P(\theta)}{P(\mathcal{D})}$$

$$= \frac{P(\mathcal{D} \mid \theta)P(\theta)}{P(\mathcal{D})}$$

$$= \frac{P(\mathcal{D} \mid \theta)P(\theta)}{P(\mathcal{D})}$$

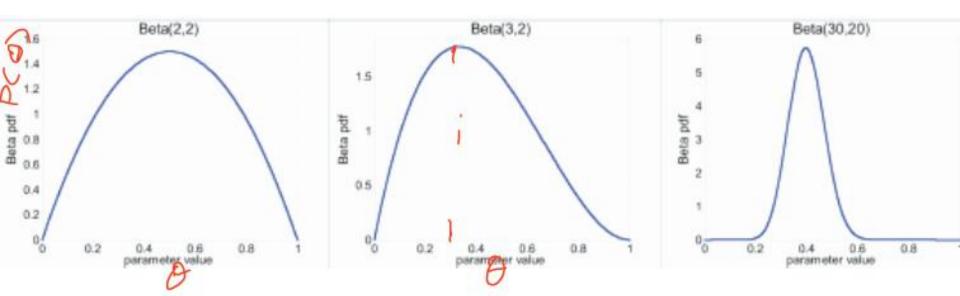
$$= \frac{P(\mathcal{D} \mid \theta)P(\theta)}{P(\mathcal{D})}$$

Or equivalently:

$$P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta)P(\theta)$$

Beta prior distribution – $P(\theta)$

$$P(\theta) = \frac{\theta^{\beta_H - 1} (1 - \theta)^{\beta_T - 1}}{B(\beta_H, \beta_T)} \sim Beta(\beta_H, \beta_T)$$



Beta prior distribution – $P(\theta)$

$$P(\theta) = \frac{\theta^{\beta_H - 1} (1 - \theta)^{\beta_T - 1}}{B(\beta_H, \beta_T)} \sim Beta(\beta_H, \beta_T)$$

Likelihood function:
$$P(D \mid \theta) = \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$$

Posterior:
$$P(\theta \mid D) \propto P(D \mid \theta) P(\theta)$$

$$\theta^{\alpha_{H}} (1-\theta)^{\alpha_{T}} \frac{\theta^{\beta_{H}-1} (1-\theta)^{\beta_{T}-1}}{B(\beta_{H}, \beta_{T})}$$

$$= \frac{\theta^{\alpha_{H}+\beta_{H}-1} (1-\theta)^{\alpha_{T}+\beta_{T}-1}}{B()}$$

MAP

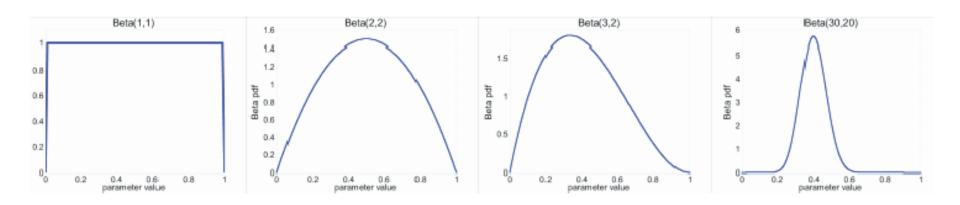
$$\arg\max_{\theta} P(\theta \mid D) = \arg\max_{\theta} \frac{\theta^{\alpha_H + \beta_H - 1} (1 - \theta)^{\alpha_T + \beta_T - 1}}{B()}$$

$$\hat{\theta}_{MAP} = \frac{\alpha_H + \beta_H - 1}{\alpha_H + \beta_H - 1 + \alpha_T + \beta_T - 1}$$

Posterior distribution

- Prior: $Beta(\beta_H, \beta_T)$
- Data: α_H heads and α_T tails
- Posterior distribution:

$$P(\theta \mid \mathcal{D}) \sim Beta(\beta_H + \alpha_H, \beta_T + \alpha_T)$$



MAP for Beta distribution

$$P(\theta \mid \mathcal{D}) = \frac{\theta^{\beta_H + \alpha_H - 1} (1 - \theta)^{\beta_T + \alpha_T - 1}}{B(\beta_H + \alpha_H, \beta_T + \alpha_T)} \sim Beta(\beta_H + \alpha_H, \beta_T + \alpha_T)$$

MAP: use most likely parameter: $\theta_{MAP} = rgmax_{ heta} P(\theta \mid D)$

- Beta prior equivalent to extra thumbtack flips
- As $N \to \infty$, prior is "forgotten"
- But for small sample size, prior is important!!

Conjugate priors

$P(\theta)$ and $P(\theta|D)$ have the same form

Eg.1 Coin flip problem

Likelihood is ~Binomial $P(D \mid \theta) = \theta^{\alpha_H} (1 - \theta)^{\alpha_T}$

If Prior is Beta distribution,
$$P(\theta) = \frac{\theta^{\beta_H - 1} (1 - \theta)^{\beta_T - 1}}{B(\beta_H, \beta_T)} \sim Beta(\beta_H, \beta_T)$$

Then posterior is Beta distribution

$$P(\theta \mid D) = Beta(\beta_H + \alpha_H, \beta_T + \alpha_T)$$

For Binomial, conjugate prior is Beta distribution.

Conjugate priors

$P(\theta)$ and $P(\theta|D)$ have the same form

Eg.2 Dice roll problem (6 outcomes instead of 2)

Likelihood is ~Multinomial $P(D \mid \theta) = \theta_1^{\alpha_1} \theta_2^{\alpha_2} ... \theta_k^{\alpha_k}$

If Prior is Dirichlet distribution,
$$P(\theta) = \frac{\prod_{i=1}^{k} \theta_i^{\beta_i - 1}}{B(\beta_1, \beta_2, ..., \beta_k)} \sim Dirichlet(\beta_1, \beta_2, ..., \beta_k)$$

Then posterior is Dirichlet distribution

$$P(\theta \mid D) = Dirichlet(\beta_1 + \alpha_1, ..., \beta_k + \alpha_k)$$

For Multinomial, conjugate prior is Dirichlet distribution.

Estimating Parameters

 Maximum Likelihood Estimate (MLE): choose θ that maximizes probability of observed data D

$$\hat{\theta} = \arg\max_{\theta} P(D \mid \theta)$$

• Maximum a Posterior (MAP): choose θ that is most probable given prior probability and the data $\hat{\theta} = \arg\max P(\theta \mid D)$

$$= \arg \max_{\theta} P(D \mid \theta) P(\theta)$$

Dirichelet distribution

- Number of heads in N flips of a two-sided coin
 - Follows a binomial distribution
 - Beta is a good prior (conjugate prior for binomial)
- What it's not two-sided, but k-sided?
 - Follows a multinomial distribution
 - Dirichlet distribution is the conjugate prior

$$P(\theta_1, \theta_2, ..., \theta_k) = \frac{1}{B(\alpha)} \prod_{i}^{K} \theta_i^{(\alpha_i - 1)}$$

You should know

Probability basics

- Random variables, events, sample space, conditional probs. ..
- Independence of random variables
- Bayes rule
- Joint probability distributions
- Calculating probabilities from the joint distribution

• Estimating parameters from data

- Maximum likelihood estimates
- Maximum a posterior estimates
- Distributions-binomial, Beta, Dirichlet, ...
- Conjugate priors

Expected values

 Given discrete random variable X, the expected value of X, written E[X] is

$$E[X] = \sum_{x \in X} xP(X = x)$$

 We also can talk about the expected value of functions of X

$$E[f(X)] = \sum_{x \in X} f(x)P(X = x)$$

Covariance

• Given two discrete r.v.'s X and Y, we define the covariance of X and Y as

$$Cov(X,Y) = E[(X - E(X))(Y - E(Y))]$$

e.g., X=gender, Y=playsFootball

Summary

• Questions?

• Thank You!