8: Eksperimenter II

 $\label{lem:videregaende} \mbox{Videregaende kvantitative metoder i studiet af politisk adfærd}$

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- 3 Clustering
- 4 Kovariater
 - Præ-treatment outcome
 - Andre kovariater
 - Blocking
- 5 Noncompliance
- 6 Case: Gerber & Green (2000)
- 7 Kig fremad

	Uge	Dato	Tema	Litteratur	Case
-	1	5/9	Introduktion til R	lmai kap 1	
	2	12/9	Regression I: OLS	GH kap 3, MM kap 2	Gilens & Page (2014
	3	26/9	Regression II: Paneldata	GH kap 11	Larsen et al. (2016)
	4	29/9	Regression III: Multileveldata, interaktioner	GH kap 12	Berkman & Plutzer
	5	3/10	Introduktion til kausal inferens	Hariri (2012), Samii (2016)	
	6	10/10	Matching	Justesen & Klemmensen (2014)	Ladd & Lenz (2009)
		17/10	*Efterårsferie*	` ,	, ,

Uge	Dato	Tema	Litteratur	Case
	17/10	*Efterårsferie*		
7	24/10	Eksperimenter I	MM kap 1, GG kap $1+2$	Gerber et al. (2008)
8	31/10	Eksperimenter II	GG kap 3+4+5	Gerber & Green (2000
9	14/11	Instrumentvariable	MM kap 3	Arunachalam & Watso
10	14/11	Regressionsdiskontinuitetsdesigns	MM kap 4	Eggers & Hainmueller
11	21/11	Difference-in-difference designs	MM kap 5	Enos (2016)
12	28/11	'Big data' og maskinlæring	Grimmer (2015), Varian (2014)	, ,
13	5/12	Scraping af data fra online-kilder	MRMN kap 9	
14	12/12	Tekst som data	Grimmer & Stewart (2013), Imai kap 5	

Formalia 000

Spørgsmål?

- mere om potential outcomes framework
- randomisering i praksis: sample(), set.seed()
- faldgruber ved eksperimentelle designs: excludability, noninterference
- case: GGL (2008)

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Spørgsmål?

Formlen for \widehat{ATE} 's standardfejl:

$$SE(\widehat{ATE}) = \sqrt{\frac{1}{N-1}} \left\{ \frac{mVar(Y_{i0})}{N-m} + \frac{(N-m)Var(Y_{i1})}{m} + 2Cov(Y_{i0}Y_{i1}) \right\}$$
 (1)

Kilder til mindre standardfejl:

- N ↑
- $Var(Y_{i0})$ eller $Var(Y_{i1}) \uparrow$
- Cov(Y_{i0} Y_{i1}) ↓

I nogle situationer måles *outcome* individuelt, men *assignment* sker på cluster-niveau

- fx. mediemarkeder, kommuner, klasseværelser
- grundlæggende implikation: ingen bias, men svækket præcision

 \widehat{ATE} 's standardfejl med k clusters:

$$SE(\widehat{ATE}) = \sqrt{\frac{1}{k-1}} \left\{ \frac{mVar(\overline{Y}_{j0})}{N-m} + \frac{(N-m)Var(\overline{Y}_{j1})}{m} + 2Cov(\overline{Y}_{j0}\overline{Y}_{j1}) \right\}$$
(2)

Kilder til mindre standardfejl:

- k ↑
- $Var(\overline{Y}_{j0})$ eller $Var(\overline{Y}_{j1})) \uparrow$

p-værdier kan (med eller uden cluster assignment) beregnes med randomiseringsinferens

- simulér alle tænkelige assignments af treatment
- for hver hypotetisk assignment, estimér ATE
- beregn p-værdi pba. faktisk ATE ift. fordelingen af estimater

Spørgsmål?

- outcome kan måles som ændring præ-post
- i st. for difference-in-means, difference-in-differences estimator (jf. u. 11)
- når præ-treatment kovariater korrelerer m. potential outcomes, stærkt øget præcision

Andre kovariater

Regression af Y_i på d_i og kovariat X_i :

$$Y_i = Y_{i0}(1 - d_i) + Y_{i1}d_i = a + bd_i + cX_i + (u_i - cX_i)$$
(3)

ightarrow prædiktive kovariater reducerer residualer ightarrow $\sigma_{\hat{b}} \downarrow$

Andre kovariater

Men: øger også 'researcher degrees of freedom'

»This type of analysis introduces an element of discretion in terms of what results are reported. Perhaps unconsciously, the researcher may settle on a regression model that makes the estimated ATE look impressive or interesting, a decision rule that jeopardizes the unbiasedness of the estimator.« (105)

Table I. Explaining support for socially protective policies with physiological reactions to threatning images. Results of ordinary least squares (CLS) regression with support for socially protective policies (possible range from 0 to 18), with higher numbers indicating attitudes more supportno of policies thought to protect the social unit regressed on five explanatory variables: gender (0 = male; 1 = female), age (in) years), education (gio, categories ranging from 3 and 3

Variable	Unstandardized coefficient (SE)	Standardized coefficient	
SCL	92.2* (29.03)	0.377	
Income	-0.395 (0.471)	-0.10	
Education	-1.63* (0.465)	-0.42	
Age	0.19 (0.10)	0.235	
Gender	-2.34 (1.3)	-0.20	
Constant	-353* (193)		
N	46		
Adi. R-square	0.37		

Table 2. Explaining support for socially protective policies with physiological reactions to nonthrest ening images. Results of regression (CDS) with support for socially protective policies regressed on five explanatory variables variables are the same as those described for Table 1 except that skin conductance (SCL) is the change in skin conductance occusioned by the viewing of nonthreasting regis. Descriptive statistics and further discussion of the regression techniques are available in the SCM #9 F old Sc hove-alled frace.

	Unstandardized	Standardized
Variable	coefficient (SE)	coefficient
SCL	-1.8 (35.08)	-0.007
Income	-0.438 (0.533)	-0.115
Education	-1.57* (0.53)	-0.408
Age	0.165 (0.11)	0.204
Gender	-2.23 (1.52)	-0.196
Constant	-304* (217)	
N	46	
Adj. R-square	0.21	

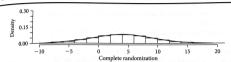
Blocking

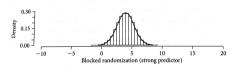
Blocking på kovariat X hjælper når:

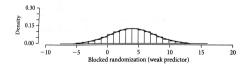
- N er relativt lille
- X er ubalanceret
- X er stærk prædiktor for Y

God pakke til block random assignment: randomizr

FIGURE 4.2
Comparison of sampling distributions based on completely randomized and block randomized designs







Formalia Opsamling Clustering Kovariater Noncompliance Case: Gerber & Green Kig fremad

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Blocking

Spørgsmål?

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Konceptuelt: to typer subjekter under ensidet noncompliance

- compliers: $d_i(z=1)=1$
- never-takers: $d_i(z=1)=0$
- \rightarrow tre grupper:
 - treatede compliers
 - non-treatede compliers
 - never-takers

For hvert subjekt *i* defineres:

$$ITT_{i,D} \equiv d_i(1) - d_i(0) \tag{4}$$

$$ITT_{i,Y} \equiv Y_i(1) - Y_i(0) \tag{5}$$

CACE er forholdet mellem $\overline{ITT_{i,Y}}$ og $\overline{ITT_{i,D}}$:

$$CACE = \frac{ITT}{ITT_D} \tag{6}$$

• ATE =
$$(2+6+4+2+4+8+3+3+4)/9 = 4$$

• ITT =
$$(2+0+4+2+4+0+3+3+0)/9 = 2$$

• CACE =
$$(2+4+2+4+3+3)/6 = 3$$

TABLE 5.1

Hypothetical schedule of potential outcomes assuming one-sided noncompliance

Observation	$Y_i(d=0)$	$Y_i(d=1)$	$d_i(z=0)$	$d_i(z=1)$	Туре
1	4	6	0	1	Complier
2	2	8	0	0	Never-Taker
3	1	5	0	1	Complier
4	5	7	0	1	Complier
5	6	10	0	1	Complier
6	2	10	0	0	Never-Taker
7	6	9	0	1	Complier
8	2	5	0	1	Complier
9	5	9	0	0	Never-Taker

Direkte sammenligning af treatede og nontreatede v. noncompliance estimerer flg.:

$$CACE + \{E[Y_i(d=0)|D_i(1)=1] - E[Y_i(d=0)|D_i(1)=0]\}(1 - ITT_D)$$
(7)

 \rightarrow bias hvis compliers og never-takers har uens untreated potential outcomes

Spørgsmål?

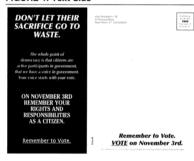
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Formalia Opsamling Clustering Kovariater Noncompliance Case: Gerber & Green Kig frema

FIGURE 2. Picture Side



FIGURE 1. Text Side



»to find the treatment effect, subtract the turnout rate of the control group from the turnout rate of the experimental group and divide this difference by the observed "contact rate," which is 28%. Using this formula, we find that personal contact raises the probability of turnout by 8.7 percentage points« (658)

»Despite limitations, this experiment provides important new clues in the ongoing mystery of why turnout has declined even as the average age and education of the population has risen. A certain segment of the electorate tends not to vote unless encouraged to do so through face-to-face contact. As voter mobilization grows more impersonal, fewer people receive this kind of encouragement. This point is of great practical significance for those who seek to reverse the declining trend in turnout. Many of the recent policy innovations designed to encourage voter participation (e.g., absentee balloting) focus on reducing the costs of voting. Our findings suggest the importance of focusing as well on the personal connection between voters and the electoral process.« (662)

Formalia

Næste gang:

- instrumentvariable
- pensum: MM kap. 3
- ullet case: Arunachalam & Watson o læs kun ift. argumentation om instrumentvaliditet

Tak for i dag!

Kig fremad ○●