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Fertility, Agricultural labour Supply and Production____ Instrumental Variable Evidence from Uganda

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ABSTRACT

Human F[ertility] is likely to affect agricultural production through its effect on the supply of agricultural labourlabor. Using the fact that in traditional, patriarchal societies sons are often preferred to daughtersgirls, we isolated exogenous variation in the number of children born to athe mother and related it to agricultural labourlabor supply and pro-duction outcomes in Uganda—a country that combines a dominant agricultural sector with one of the highest fertility rates in the world. We founded that fertility has a seizable negative effect on household labourlabor allocation to subsistence agriculture. Households with lower fertility devote significantly more time to land preparation and weeding, while larger households grow less matooke and sweet potatoes. We founded no significant effect on agricultural productivity.

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INTRODUCTION

At the most basic level, subsistence farmers in rural Africa combine natural with human resources to make a living. They use mainly household <code>labourlabor</code> on their own small plots to produce <code>food</code> for <code>their</code> own consumption. As such, the quantity of family <code>laboumembers.r</code> available for agricultural <code>labor.e</code> is an important determinant of <code>well-beingwellbeing</code>. More children means mothers, and to a lesser extented fathers, will need to spend more time <code>caringfor children</code> on <code>reproductive chores</code>, meaning less time will be available to spend on agricultural activities. Since women are known to <code>supply most</code> of the agricultural <code>labourlabor</code>, the loss in time needed for <code>reproductive activities-may</code> hurt subsistence households disproportionately.

Uganda has one of the highest fertility rates in the world. Even in the context of large reductions in child mortality rates, total fertility rates remain stubbornly high. On average, Ugandan women in rural areas bear 6.8 children over the course of their reproductive lives (UBOS, 2 2013). At the same time, a substantial part of the population lives in rural areas making a living out of semi-subsistence agriculture. Ugandan agriculture accounts for about 35 percent of GDP and employs about 73 percent of the of the active labourlabor force. Virtually all households that reside in rural areas are engaged in farming, and about 80 percent are small;—scale, semi-subsistence farmers. The question of how fertility affects well-beingwellbeing through its effect on household labourlabor supply and agricultural production is therefore relevant. For example, knowledge of how fertility affects time allocation by different categories within the household is important to gender-stream efforts related to crop intensification and commercialization.

In this study, we will investigated the effect of fertility on agricultural production at the household level. In particular, we will investigated the effect of the number of biological children on household member labourlabor input in agriculture (further categorized asin land preparation, weeding, input application, and harvest—ing). We will also looked at the effect of fertility on crop portfolio, area cultivated, production, and productivity for the five most important crops. However, fertility is a choice variable to agricultural households. For in—stance, mothers who work long hours in the field may try to avoid becoming pregnant because this would only increase their hardship. If fertility, agricultural labourlabor allocation, and agricultural production were is jointly deter—mined, just looking at correlations wouldill be misleading, and so weene neededs to find a way to separate the exogenous variation from the part that is jointly determined.

Our identification strategy <u>wasis</u> a simple but powerful quasi-experimental approach inspired by the work of Angrist and Evans (1998). We use<u>d</u> the fact that, in conservative, patriarchal societies such

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Uganda's, male off-springs are generally preferred to female children. This preference and the random nature of the newborns' the sex determination of a newborn gives rise to particular fertility patterns. For example, households that have a girl as the first born irstborn are likely to have more children (Jayachandran and Kuziemko, 2 2011). In other words, we used the sex of off-springs as an instrumental variable (IV) to determine the exogenous component of fertility at the household level, that We expected that Ssuch a Two_-Stage Least Squares (2SLS) approach willis expected to would yield consistent estimates for the causal effect of fertility on agricultural labourlabor supply and associated agricultural production within the household.

There is an active debate among scholars in the field of labour economicsists on the relationship between fertility and labourlabor supply. Angrist and Evans (1998) used the fact that American couples prefer to have children of different sexes, and and that they are likely to keep trying if the first two children are of the same sex, as a source of exogenous variation. We will argue that in a developing country context, the sex of the (first—born) child makes more sense as an instrument. Indeed, this instrumentation strategy has been used in such a context. Gupta and Dubey (2006) used the sex of the first two children to predict exogenous variation in fertility in India and its effect on well-beingwellbeing. We thought feel it wasis too_ambitious to relate fertility directly to poverty and related measures of consumption, because the sex of the first two children may directly affect consumption, violating the exclusion restriction. We restricted ourselves to the agricultural labour allocation of adult household members, area planted, and production. In addition, most studies that look at the effects of fertility on labourlabor allocation in a developing country context use data from Asian countries. The high incidence of selective abortion in these countries areas may mean the sex of the first child or children becomes endogenous as well. This wasie likely to be much less of a problem in our application, which is to our knowledge, the first such application to an African country.

We found that the sex of the first-bornfirstborn, the sex of the first two children born, and as well as the percentage of girls as a share of the total number of children all significantly explain observed fertility, measured as the gap between actual children born and a theoretical maximal fecundity for each age cohort. Fertility has a strong negative effect on the number of days the mother worksed by the mother in the field. We also find some evidence of a negative effect on fer the father, but the size of the effect is only half of that one of the women. Households with lower fertility devote significantly more time to land preparation and weeding. We also found that smaller households grow and produce more matooke. This effect holds to a lesser extent for sweet potatoes. We figured no impact on yields.

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This article is organized as follows. The next section gives a brief overview of the most prominent

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papers that are related to our study. Then, we make our case for the use of the sex of the first-bernfirstborn as an instrument, using literature that documents child gender and reproductive behaviour. We then present the data we will used in our application, and describe the our main variables we will used in the analysis. Next, we present the results. In this section, we first take a close look at the first stage regression. We then look at the effect of fertility on household labourlabor supply, considering differential effects depending on specific agricultural labourlabor activities. We then turn to the effect of household size on aspects of agricultural production and productivity. A final section concludes.

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RELATED STUDIES

Fertility and the related concept of household size <u>effect_impacts</u> household <u>well-beingwellbeing</u> through consumption and production. Lanjouw and Ravallion (1995) focused on the consumption side effects of household size in a developing country context. They noted the contradiction between widely—held views that larger households are often poorer (due to increased competition for—a given food—steck) and scale economies in consumption. They fouind that, if economies of scale are accounted for within households, the negative correlation between household size and consumption expenditure disappears. On the production side of the farm household, the effect of household size is equally ambiguous. Some may argue that larger households <u>have means that</u> more <u>labourlabor</u> is available within the household. The additional advantage of this <u>labourlabor</u> is that it is not subject to the moral hazard effects often attributed to hired <u>labourlabor</u>.

But at the same time, more dependents within the household means more time needs to be allocated to <u>caring for them</u>reproductive activities. Also, agricultural <u>labourlabor</u> and agricultural production may be subject to diminishing returns.

The relationship between fertility and household labourlabor supply has been studied most carefully in the field of labourlabor economics. Since this literature is so extensive, we only mention two of the most influential works here. The first is Angrist and Evans (1998), who attempted to quantify the effect of fertility on labourlabor supply in the US. They dealt with the endogeneity of the number of a woman's children women have by exploiting the fact that Americans tend to prefer two mixed sibling genders-sex in their households. They argued that parents of same-sex siblings are significantly and substantially more likely to go on to have an additional child. They fouind that more children does indeed the reduce women's labour force participation in the labor force, of women, but that the effect is less pronounced than previous studies had suggested. They fouind no effect on the labourlabor for participation

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supply byof the fathers.

Another paper that trieds to answer the same question is Rosenzweig and Wolpin's (1980a). $\underline{\text{These two}} \underline{\text{-In this}}$

⁴For instance, Foder (1985) argues this may the reason why small farms appear to be more efficient than large farms

paper, obtained the exogenous variation in the number of childrechildren n is obtained by using the occurrence of multiple births (twins) at first birth as an instrument. The authors argued that the comparison between women who giave birth to a singleton child at first birth and women who gave birth to twins at first birth allowed them some to identify the causal effect of an extra child on an outcome (in their case labourlabor supply). Since the occurrence of twins is exogenous, there wasis no danger that heterogeneity in women preferences contaminateds the esti-mated coefficients. The study founds that household size reduces female labourlabor supply, but that the effect is only temporary².

Gupta and Dubey (2006) use $\underline{\underline{d}}$ the sex of the first two children as a natural experiment and found that

household size increaseds poverty in India. They used essentially the same argument as we will make in the next section. However, welfare, and the related concept of poverty, relyies on consumption per capita. Consumption per capita as the independent variable wasis likely to be problematic in our the two-stage least squares setting. There wasis a real danger that the instrumental variable would affects the outcome variable directly, instead of only through its influence on family size. For instance, if boys consume on average more than girls, the exclusion restriction would be is violated. There is also some evidence fromen Indonesia. Kim et al. (2009) lookeds at the relation between consumption and fertility. In-Kim and Aassve (2006), fertility is related fertility to the allocation of labourlabor within households. However, they moved away from the direct instrumental variable approach that is standard and instead estimated a reproduction function taking into account endogenous contraceptive choice.

All the above studies employed data from Southe—East Asia. It is well_known that fertility at birth is already skewed in many Asian countries. For instance India, from which Gupta and Dubey (2006) drewawn their sample, is particularly known for selective abortion of girls (Jha et al., 22011). This non-random distribution of sex of children opens the door to potential correlation between the instrument and the error term of the structural equation. One example would be that less educated, poorer households that depend heavily on agriculture engage more in the abortion of female fetuses. In the context of weak instruments, such correlation can seriously bias estimates (Bound et al., 1995).

In this paper, we will try to address some of the above challenges. We <u>will useused</u> the sex of the <u>first-bernfirstborn</u> and variations thereof as our instrumental variables. We will relate fertility to agricultural <u>labourlabor</u> supply and agricultural production, since there is a direct link between these three variables. We will also concentrate on Africa. Here, while there is a boy preference, reproduction rate norms are high and the cost of raising children is low. This means that selective abortion is much less of a concern.

¹For instance, Feder (1985) argues this may the reason why small farms appear to be more efficient than large farms.

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²In addition to studies that investigate the causal effect of household size on (female) labourlabor supply, there are also a range of papers that test Becker's quantity-quality fertility model (Becker, 1960; Becker and Lewis, 1973). Many of these articles also use twins (Rosenzweig and Wolpin, 1980b; Black et al., 2 2005) and/or sibling sex composition (Conley, 2 2000; Angrist et al., 2 2010).

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BOY PREFERENCE AND FERTILITY

There is quite a bit of some evidence that parents prefer boys over girls in many developing countries³. For instance there is a large body of literature that looks at correlations between sex and variables related to well-beingwellbeing or quality of children. Significant differences in these outcomes are then considered proof of sex bias. Das Gupta (1987) and Sen (1990) looked at excess mortality among female infants in India. Chen et al. (1981) and Pande (2003) investigated differential access to health, in Bangladesh and India respectively. Behrman (1988) and Hazarika (2000) founind a correlation between sex and nutrition and Behrman et al. (1982), Davies and Zhang (1995), and Alderman and King (1998) all investigated correlations between the gender of children and education.

However, at a more basic level, boy preference is already revealed by parents who, if asked in-fer example surveys, for example, often state clearly that they prefer boys to girls. Such preferences lead to a particular decision rule with respect to fertility, where the likelihood that children are added to the household is positively correlated to the number of surviving girls in the household. The preference for boys over girls results in what Jayachandran and Kuziemko (2011) refer to as the "stop-after-a-son" fertility pattern. There are indeed many studies that show empirically that in settings with characterized by son preference, a couple that has just had a son is more likely to stop having children (Das, 1987) or wait longer to have the next child (Trussell et al., 1985; Arnold et al., 1998; Clark, 22000; Drèze and Murthi, 22001).

Jayachandran and Kuziemko (2011) _argued that son preference leads mothers to breastfeed daughters and children without brothers for a relatively shorter timeless long. Since breastfeeding is an effective birth control method, this observed behaviour also explains why couples with a son seem to wait longer before they have the next child. In addition, this underlying consequence of sex bias may partly explain a range of outcomes observed in the area of health, mortality, and possibly even educational attainment. The Their model that Jayachandran and Kuziemko (2011) developed shows that even when parents want both boys and girls to have the same health and education, disparities can arise passively because of fertility preferences. The model y shows that a "try until you have a boy" fertility rule results in girls having on average more siblings, leading to more competition for resources within the household.

The occurrence of boy preference is explained by various cultural and economic factors documented

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³In developed countries, there is a preference for a mix of sexes among children, as shown in, for example Angrist and Evans (1998).

in the anthropological and demographic literature. In countries where no formal, risk_-free old age insurance (such as a pensions) is available, parents may choose to invest more in children that who will have a higher chance of being able to support them inat old age (Behrman et al., 1982). Anthropological and demographic evidence emphasize the dominant role of males in traditional patrilineal societies where descent and inheritance are transmitted through the male line. Furthermore, male children strengthen the relationship between the wife and her husband's kin (by guaranteeing the continuation of his lineage) and secure (the mother's) access to seto residence and inheritance and a place to live upon the husband's death. Older women have power through their sons and rule over their daughters-in-law (Kandiyoti, 1988). The spread of primary schooling in sub-Saharan-Africa south of the Sahara has also affected fertility patterns (Lloyd et al., 2 2000). Since boys are more likely to be sent (and kept) in school than girls, the extra cost associated with primary schooling will be higher in families with more boys. This, in turn, will encourage families who already have boys to reduce fertility.

Most of the evidence on the existence of boy preference comes from Asian countries. There are relatively few inquiries into sex preference in Sub-Saharan-Africa south of the Sahara. Even more, it is often assumed that gender preferences are much lower or even absent_therein Sub-Saharan Africa. This is surprising, since as many of the cultural and economic factors that are observed in Asia equally apply to Africa. One study that documents significant gender bias in Africa is Anderson and Ray's (2010), who found skewed sex ratios at older age in favor of men. Another study of in a small community in en Nigeria found reports that almost 90 percent of surveyed respondents reported male sex preference (Eguavoen et al., 22007). What is different from the Asian context, though, is that the vast majority of women are not just missing from the household labor supply when their children are born; at birth, they are missing but throughout their entire age spectrum4. Milazzo (2012) arguede that gender bias is likely not to be found at birth in the African context, where high fertility is culturally valued and less costly for families that still rely on the support from the extended family system. In Uganda, preference for boys has been extensively documented in Beyeza-Kashesya et al. (2010).

Even in Western societies, preference for the first-bern to be a sons, rather than daughters, has been observed. For example, Marleau and Saucier (2002) reported an extensive list of studies that found ind men and/or women prefer a boy rather than a girl as their first-bern firstborn. Even in the United States, the Angrist and Evans (1998) found provides some evidence of an association between having a male child and reduced childbearing at higher parities in addition to the mixed-child preference. As such,

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we fe<u>ltel</u> that the sex of the <u>first-born</u> (or closely related indicators) <u>would</u> provide a valid instrument for fertility at the household level in Uganda.

⁴The large effects documented in Anderson and Ray (2010) have recently been attenuated in Klasen and Vollmer (2013), who confirmed that only missing-young adult women were missing from the household labor force.

DATA AND DESCRIPTIVE STATISTICS

We will used the Uganda National Household Survey (UNHS) 2005—2006. While being somewhat dated, we have chosen this survey because it has much more information abouten agriculture than the more recent UNHS of 2009—2010, or that of 2012—2013. The 2005–2006 UNHS we chose was structured with a standard Living Standards Measurement Study-Survey - Integrated Surveys on Agriculture (LSMS-ISA) in mind, the UNHS 2005/06 it collected detailed information on a sample of almost 43,000 people in 7,500 households in Uganda.

Ideally, we would have it is a sample of households, where all desired children were born. The fact that we wereare working with a cross-section of households, where households are at different stages in their reproductive lives, createds some problems. Assume a couple that has just formed and is entering their reproductive stage. In our sample, such households will showed up with a smaller-than, average household size. Now, if the first-bornfirstborn happeneds to be a girl, this could have may mistakenly been interpreted as running against our hypothesis that households where the first-bornfirstborn wasis a girl would have higher fertility. On the other hand, if the first child wasis a boy, this could have may ledael us to put too much confidence in our hypothesis, as the smaller household size wasis not only due to the fact that the first-bornfirstborn wasis a son, but also wasis entered its their reproductive stage. The fact that we were are working with a cross_-section of households rather than historical data on all births by women who that had we reached the end of their reproductive lives wasis also reflected in the average number of children. This wasis only 3.13 children, while women bear almost 7 children over their entire reproductive period.

To deal with this problem, we worked with the difference in the actual number of children reported and the maximum reproductive capacity for a woman at a certain age, rather than simply working with the number of children in the household. We will referred to this measure of fertility as the gap or shortfall in fertility. To get the latter, number of children in the household we would have needed to estimate the average age at menarche within the population, and then simply divide into age the pregnancy period plus post pregnancy lactation in-fecund period. In addition, weene would also have needed to corporate the maternal mortality ratio for "censoring" the lives of women's lives who have had too many children and thus increasedeing her mortality rate (and exit from the sample). Instead, we took the 95th percentile of total fertility rates per age from the Demographic and Health Survey of Uganda (2011). This is probably a good approximation of the upper bound by age of fertility in the population.

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If it refers to the gap or shortfall in fertility, I'm confused about why you were not able to get it when you have something to refer to

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Our The selection of children wasis based on the household roster of the UNHS 2005/_2006. In particular, we

⁵Alternatively, one could use the number of children within the household and control for the age of the mother. We have also run the analysis using this strategy and came to virtually the same conclusions.

will selected individuals-who that were are indicated as son or daughter of the household head. There is another potential problem when someone usesing a cross sectional survey such as the UNHS that only looks at reported dependents currently living within the household to calculate the difference between actual and theoretical fertility. Older women may have been living in households where some of the older children hadve already left the household. Thus, at around the age of thirty, the gap between reported children in the household and theoretical fertility would have starteds to increase more rapidly because of children growing old enough to start households of their own. More troubling, the reported gender of the oldest son or daughter living in the household may not have been the gender of first-bernfirstborn. To overcome this problem, we restricted our sample to households where the mother wasis between 16 and 32. We chose Tthe cut-off age of 32 is because at this age, the first-born of the mother's firstborn or will turns 16, which is our entry age into the sample of mothers. Restricting our sample in this way hade a second advantage. For some of the indirect outcome variables we will use used, such as productivity, there was a risk ene may argue that the gender of the first-bernfirstborn hads a direct effect on the outcome, instead of only through fertility. Restricting our sample to households with only young mothers meants that the children wereare also likely to be younger, and thus less likely to engage in agricultural production, making a violation of the exclusion restriction less likely.

Looking at the sex of the first-bernfirstborn is only one possible strategy. One may argue that the sex of the first born is not very relevant in a context where women bear on average almost 7 children. Indeed, it is likely that households will get a second child irrespective of the sex of the first. This is supported by Jayachandran and Kuziemko (2011), who foundind that the difference in breastfeeding duration between boys and girls is largest near target family size, when gender is most predictive of subsequent fertility. Therefore, we_will not only used the sex of the first bernfirstborn child as an instrument for fertility, but also experimented with alternative instruments such as an indicator that the first two children would be_are girls or a variable that expresses the share of female children in the total family size. The next section presents some preliminary statistics that suggest how gender patterns in the household are related to fertility.

Gender of Offspring and Fertility

In t_This section, we makes a case for the different instruments we will use used in the analysis. While the next section we will runs a first stage regression in the next section, this one we present some simple descriptive statistics here to show that gender of the first few children, as well as the share of male children

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in a woman's the total number of children, affects fertility. Table 0.1 summarizes our findings. The first two columns in the top panel checks if households that

have a daughter daughter as first-born firstborn are more likely to have extra children. We simply calculated the percentage of households that had more than one child conditional on their first-born being a son or a girl. In other words, we calculated the probability that a household hads at least one additional child (prop +1). We found that in the sub-sample where the first child is a boy, about 37.46 percent of households would ill have at least one more child. However, if the first child happened to be a girl, almost 40 percent of households would ill have at least one more child. This confirmeds our hypothesis that households have a higher chance of adding children if the firstborn is a girl.

We also looked at the the effect of the firstborn's gender of the first born on shortfall in actual fertility to theoret-ical fertility. In t The first two columns in the bottom panel, we report this fertility gap for these two groups of households. We findfound that households that have a boy as the first born firstborn child have an average fertility gap of about 2.46 children. Consistent with the proposition that households with that have a first-born girl are likely to have more children conditional on age, we findfound that the gap is smaller than when the first-born is a girl (2.26). In other words, households where the first-born is a girl are will be-closer to the theoretical maximal fertility than those households whose firstborn withat had a is a boy-as the first-born. The difference in the fertility gap between the group of households with a first-born boy and the group with a first born-girl is significant (p=0.003).

The third and fourth column present the same statistics, but now looks at the sex of the first two children. We now looked at three possible scenarios. If the first two children wereare both boys, we expected that the chance that the household would have ey have extra children would!! be lowest. We findfound the probability of adding children in this case to be just over 13 percent (top panel). If the first was a girl and the second was also a girl, we expected the probability that the household would have additional children to be highest. In this case, there wasie indeed an almost 14 percent chance that a couple would they add at least one child to the household. For those households that had a first girl first but whose their second child was a boy, we expected the probability of increasing household size to lie between the two, which is indeed turned out to be the case. The lower panel shows that the gap between actual and potential household size is also largest when the first two children are boys. The gap is smallest when the first two children are both girls. All this again confirms our proposition that boy preference affects fertility.

Finally, in-columns 5 and 6, we propose the share of female children in the total number of a household's children as a potential instrument. As already stated above, because of boy preference, female children were likely to live in larger families and so we expected a positive correlation between this measure and household size. For the timebeing, new, we simply divided the sample in two, conditional on

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Should it say..." on the shortfall that exists when actual fertility is compared with theoretical fertility.

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whether more or less than half of the children <u>wereare</u> female. We <u>findfound</u> that the average number of children <u>wasie</u> indeed smaller in the sub-sample where less than half

of the children are were girls as opposed to the sub-sample where the majority were girls (top panel, 2.78 children as opposed to 2.88). We also find ound a difference in the fertility gap that wasis significant with an associated p-value of 0.021 (bottom panel). Households where girls were are in a majority were are closer to the theoretical maximal householdsize.

Agricultural <u>Labour</u><u>Labor</u> Supply

In—<u>T</u>this section we will briefly looks briefly—at some descriptive statistics on labourlabor supply in agriculture, one of the prime pathways through which fertility is likely to affect productivity and well-beingwellbeing. Most of Uganda has two cropping seasons. The first runs from January to June. The second cropping season-is from July to December. The UNHS 2005/—2006 interviewed households twice over the course of one year to capture this feature. It visited households in the beginning of 2005 to capture the second 2004 cropping season in 2004—(which runs that is, running from July to December 2004.—). Researchers revisited Hthe households were revisited at the end of 2005 to record information from the first 2005 cropping season in 2005—(which runs from January to June 2005). In—Qeur study, we will only considers the 2004 July to December cropping season, as data for labourlabor allocation in agriculture was unavailable for the 2005 cropping season.

most

of the work, and child labourlabor wasis restricted to about one day per crop and activity during the entire season. This already indicateds that the trade-off between the time lost by the mother because of rearing the children and the time gained by extra hands wasis likely to work against agricultural production. Typical

Figure 0.1 shows time reported eiin the fields along different dimensions⁶. Women seemed to do

for Uganda is the short-lew amount of time spent on applying inputs. Farmers in Uganda use very limited amounts of fertilizer and other inputs, so also the time spente on applying them is short-small. There is also some heterogeneity in the time spente on different crops. For instance, matooke is allocated less time than maize, both for men and for women. However, there are also differences between the sexes. For instance,

women spend much more time cultivating sweet potatoes, and to a lesser extent beans, than men.

The dimensions are mother, father or child; land preparation, input application, weeding and harvesting; and crop. The crops are the five most widely grown crops in Uganda.

⁷While the relationship between fertility and child labourlabor is an important research question, we <u>do will</u> not consider this in the present study. Reported child labourlabor seems lowsmall in Uganda. More importantly, the instruments we propose in this study (gender of first-born child/_children and sex composition of children within the household) are likely to directly influence the number of days the children worked in agriculture by the children, instead of only through fertility, risking to violate the exclusion restriction.

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Production

We also investigated fertility's the effect of fertility on production of some of the most important crops. More in particular, we looked at the fertility's effect of fertility on the likelihood that a household would cultivates each of the five most important crops. The first row in table 0.2 reports on the percentage of households that greew each of the crops. Over 50 percent reported growing maize, beans and cassava. We also looked at the impact on area cultivated, measured in acres. Households on average allocated about half an acre to maize, while. The least space wasis reserved for sweet potatoes. We also expressed area cultivated as a share of total area under cultivation. We findfound that about 17 percent of total land area wasis allocated to maize, while only 8 percent wasis allocated to sweet potatoes. The next line reports average production in kilograms at the household level. This may seem low, but this is because households that reported that they did not to produce the crop were are also part of the average. We also divided by household size. Finally, table 0.2 we reportst yields for the five crops, defined as the amount of each crop harvested per unit of land (per acre).

We also aggregated the different corps by weighing them to average prices. We used prices from FoodNet. In particular, we averaged prices observed in Kampala's Nakawa market over the July_-to_De-cember 2004 period in 2004. Doing so, we findfound that the average total value derived from these five crops wasis about UGX98_500, which translates to about UGX45,000 per capita. About 40 percent of the households in their reproductive age did_dees not cultivate any of these five crops. On average, about 0.69 acres wasis allocated to these five crops. The yield per acre wasis about UGX220,220.

RESULTS

In tThis section, we presents the results of our two-stage least squares estimates that looked at the causal impact of fertility on various agricultureal_related outcomes. The section we starts by presenting the first stage regression of that regresses the our proposed instruments on the fertility gap. Welt_then gives a detailed description of the second stage regression that focuses on the fertility's effect of fertility on agricultural labourlabor supply. This section we also explores the fertility's effect of fertility on area planted, production, and productivity.

 8 UGX stands for Ugandan Shillings, the national currency. At the time of the survey, USD1 = UGX1,780.

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Commented [JM52]: I've never heard this expression ...to weigh something to something

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The first stage regression

Table 0.3 reports the Rresults for the first stage regression that linked the sex of first child/children to fertility. are reported in table 0.3. The dependent variable, as explained above, wasis the difference between the maximum number of children of a typical woman at her age and the actual number of children the woman boreborn from the mother⁹. We referred to this as the fertility gap (fgap) or fertility shortfall. It was actually is the reverse of fertility, because the higher the gap, the lower the number of children in the household in a given age cohort. Apart from the exogenous variable that weis excluded from the second stage regression elaborated in the next sections, we included a series of control variables that were are clearly exogenous to fertility in all 4four specifications of the first stage. The first exogenous control variable, femhead, wasie a-an indicator variable that tookakes the value of 1 if the household head wais female. The second, urban, wasis an indicator variable that took takes the value of 1 if the household residede in an urban area¹⁰. Next, we included three dummies to account for the education level of the mother. The first, -mprim tookake the value of 1 if the mother hads completed primary education. The second, msec, iwass the additional effect of having completed secondary education. The third, mthird wasis the additional effect of the mother having completed tertiary education. The comparison category wasare therefore households where the mother did not complete at least a -primary education. We also added two community variables that which awere likely to influence household size. These werare school which iwas a dummy variable that took takes the value of 1 one if there wasis a school in the village, and health, which iwas a dummy that tookakes the value of 1-one if there iwas a public health centere or clinic in the community. Finally, we also added an indicator (cdied) that tookakes the value of one if a son or daughter of the mother hads died in the past.

We experimented with 4four different possible excluded instruments. Model (1) useds an indicator that tookakes the value of one if the first-bernfirstborn in the household iwas a girl as an excluded instrument (oldestgirl). The coefficient wasie significant at the 1ene percent level and hads the expected sign. Having a girl as the first-bernfirstborn offspring reduceds that fertility gap by about 0.2 children. In other words, households that hadve a girl as a first-bernfirstborn tended to be closer to maximal fecundity. For the controls, we findfound that households where females awere the head hadve a significantly larger fertility gap. The effect wasie very large, suggesting that such mothers have more than 1 child less. Also, in urban areas, households seemed to have significantly fewerless children. Schooling of the mother seemeds to reduce the number of children only at the secondary and tertiary levels. There seemeds

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 $^{^{9}\}mbox{The maximum number of children has been estimated from the DHS and is actually the 95th percentile.}$

¹⁰In some specifications where we expect regional variation in the outcome variable to be important, such as for production and yields for certain crops, we also include dummies for the four regions in both the first and second stage equations. This addition did not significantly change other estimated parameters in the first stage.

to be some indication that mothers who had that completed primary education had a slightly smaller fertility gap than mothers who that dihad not even completed primary education. The community variables dide not seem to have an effect on the fertility gap. Finally, having lost a child in the past leftaves a significant additional fertility gap compared withte households that had never lost a child. However, the additional gap iwas much lower than 1 ene, suggesting a substantial replacement effect in Ugandan fertility. The constant indicateds that the overall average fertility gap wasis about 2 two children.

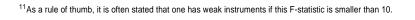
Model (2) useds an indicator that equalleds 1ene if the first two children born to the mother in the household were are both girls as excluded instrument (20ldestgirls). Using this instrument only madkes sense if we confined ourselves to households that hadve at least two children, hence the reduction in the sample size. As in (1) the parameter on the excluded instrument wasis significantly negative, in line with our hypothesis. The control variables were are very close to the what they had been were in model (1). Model (3) went goes one step further and considereds the first three children. In this case, the indicator, 30ldestgirls wasis one only if the 3three first children awere all girls. This again only make sense for in households that hadve at least three children, further reducing the sample size. The coefficient estimate wasis again negative, but this time it wasis not significant anymore. We assumed that the reduced sample size in this model might ay have reduced the power of the t-test too much.

Model (4) use<u>de</u> a continuous variable as an excluded instrument (*percentfemales*). We calculated the share of girls among children as a share of the total number of children in the household. Again, the coefficient on this instrument hade the expected <u>sign</u>. A higher share of females within the households <u>wasje</u> associated with a smaller fertility gap. This <u>wasje</u> consistent with Jayachandran and Kuziemko (2011), who observed <u>that</u> the "try until you have a boy" fertility rule leads to an outcome where larger households have on average more girls. Again, the other variables <u>wereare</u> similar to the previous models. We <u>findfound</u> that a <u>daughters-onlyn</u> all female <u>siblings</u>-household (*percentfemales* = 1) w<u>ouldill</u> be on average 0.28 children larger than an <u>sons-only</u> all boys <u>siblings</u>-household (*percentfemales* = 0).

While most of our instruments wereare significant and hadve the expected sign, they explained only a small part of the variance in the outcome. When all exogenous controls wereare included, the R-squared iwas indeed rather low. If we ranun partial regressions, regressing the excluded instruments one by one on the dependent variable, the R-square droppeds below 1 percent. The F-value of a regression with only excluded instruments—another important indicator of the strength of the instruments according to Bound et al. (1995)—also droppeds to 9.46¹¹. In other words, we hadve-serious concerns that our instruments wereare weak. We therefore used inference

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that is robust to weak instruments. In particular, we reliedy on the Anderson-Rubin test statistic to gauge the significance of the endogenous variable in all subsequent regressions (Staiger and Stock. 1997).

Household labourlabor supply

This section We new turns to <u>fertility's</u> the effect of <u>fertility</u> on total household adult <u>labourlabor</u> supply (table 0.4). We <u>will It</u> also looks at <u>labourlabor</u> supply separately for the mother and the father (table 0.5), and . We <u>will also looks</u> at <u>labourlabor</u> supply by activity (table 0.6).

In tTable 0.4, we-investigates the effect of our main variable of interest, the fertility shortfall, on the number of days worked in agriculture (land preparation, input application, weeding and harvesting)¹². The first column of the table reports the result without taking into account endogeneity of number of children. It reports Ordinary Least Squares (OLS) estimates that explain the number of days when adults reported that they had to have worked on the household farm in the 2004 agricultural season. Agricultural work wasie defined as work related to land preparation, input application, weeding and harvesting. We sawee that there wasie no significant correlation between the number of days worked and fertility as measured by the fertility gap. We did find o find significant and negative effects when of the household wasbeing headed by a female (femhead) and the household was being located in an urban area (urban). Primary and secondary education of mothers (mprim and msec) didees not seem to be-systematically related to the number of days worked in agriculture, but mothers who had that finished tertiary education (mthird) appeared to work less in agriculture. The OLS estimates also showed positive correlations between a school in the community (school) and days worked in agriculture and between a deceased child in the past (cdead) and days worked. There wasie also some indication of a positive correlation between health centers in the community (health) and days worked.

Models (2) to (4) estimated the same models, but instrumented the fertility gap with a single excluded instrument. In model (2), the instrument wasis an indicator taking the value of 1 one if the first-bernfirstborn iwas a girl. The coefficient on the fertility gap then new becaemes positive and significant at a 10 percent level, implying that higher fertility (and hence a shrinking fertility gap) causeds a reduction in the number of days worked on the family fields. Model (3) useds the sex of the first and second born as instruments for the fertility gap. Model

(4) use<u>d</u>s the share of daughters as an instrument. The estimate of the fertility effect bec<u>aemes</u> higher, and is now significant even at a 1 percent significance level.

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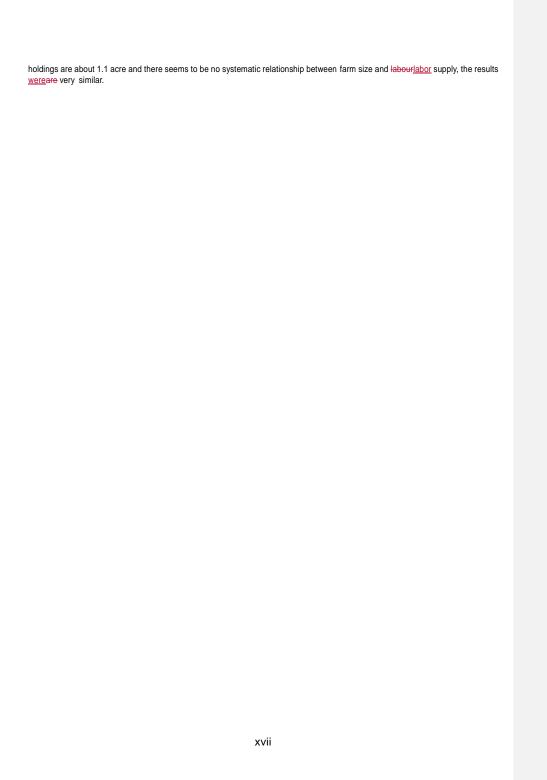
Commented [JM57]: Worked by whom? The mother? Both parents? The entire household?

Commented [JM58]: Worked by whom?

Commented [JM59]: Just checking...is instrument a verb here?

Commented [JM60]: Worked by whom? Whatever the answer is, I would insert it by saying, "a reduction in the number of days so-and-so worked on the family fields.

¹²We have also done this analysis using days worked per acre of land held by the household. However, since average land



Finally, model (5) useds both the gender of the first-bornfirstborn and the share of daughters as the ex-cluded instruments¹³. According to the Hansen-J statistic, our model that useds multiple instruments wasie valid (Hansen-J=0.849; p-val=0.357). We thus assumed this iwas the best specification. Each additional child causeds a reduction of about 66 days of labourlabor in agriculture. With respect to the other variables in the regression, we fiound some signs that households with women who that hadve finished tertiary education appeared to be less engaged in agriculture.

In Table 0.5, we differentiates between work done by the mother and by the father. For the sake of space, itwo only shows the coefficient on the fertility gap, but we also added the exogenous control variables that were also included in the first stage regression. Full results can be found in the appendix. The top panel in table 0.5 shows the effect of fertility on time worked in agriculture by the mother. The OLS estimate is not significant (model (1)). Accounting for endogeneity of fertility using the exogenous variation caused by the sex of the first-born rendereds the fertility gap significant at a 5 percent level (model (2)). An increase in the fertility gap per age cohort by one child leads a mother to work almost 30 days more in subsistence farming. Cycling through the results with the alternative instruments, the results changed little with respect to significance. In all, an additional child seemeds to reduce the number of days worked by the mother worked in agricultural production by about 40 days. Full results are reported in table 0.9.

In t_The first column of the second panel, we_reports the same OLS regression but with the number of days adult males worked by adult males as the dependent variable. As wasis the case with female labourlabor, the fertility gap dides not seem to be-correlated to male layout supply when we dide not take into account the endogenous nature of fertility choices. Table 0.10 in the appendix gives full results and the OLS results are in the first column. We findfound that living residing in urban areas ledads to farmers reporting fewerless days worked ein the field.—Win households headed by females, we also findfound a large negative effect on the number of days that men of female headedness on days worked on agriculture-related activities, by the male. This is because, in most cases, households are headed by females because the male head is missing, leading to fewerless days reported inen the field. We also findfound some evidence of males working less if the mother hadas higher education. This is most likely because higher educated men with higher educations choose women with higher educations around.

Judged by the instrumental variable models from (2) to (5), the effect of fertility on labourlabor supply by the father wasie less clear_cut. When we useding the sex of the first-born (model 2) and the

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Commented [JM62]: Is this word possibly supposed to be "labor"?

Commented [JM63]: I'm confused. How can farmers and their fields exist in urban areas?

sex of the first two children born (model 3) as instruments, the coefficient $\frac{1}{2}$ was positive but not significantly different from zero. If

 $^{^{13}\}mbox{We use}\underline{\mbox{d}}$ Limited Information Maximum Likelihood (LIML), as this is known to have better small sample properties than 2SLS in over-identified models with weak instruments (Angrist and Krueger, 2 2001).

we instrumented the fertility gap using the percentage of females, we findfound some indication that more children mightay reduce time allocated to working ein the field my males. The effect, however, is only half the size of the reductions we found for women. The over-identified model in model (5) showeds a significant effect at the 10 percent level only. These findings were are similar to what others have found. For instance, in their study on labourlabor supply response to fertility in the United States, Angrist and Evans (1998) also findfound that women work less while men dide not alter their labourlabor supply in response to having more children. Kim and Aassve (2006) findfound that Indonesian women reduced their working days in response to the higher fecundity in both rural and urban areas in Indonesia.

Table 0.6 looks at reported labourlabor by activity instead of by sex. Again, the results reported in table

0.6 only show the coefficients on the fertility gap. Full results are in the appendix. Model (1) in the top panel presents OLS results for number of days worked on land preparation. There wereare no effects from fertility in this specification. Again, as expected, households living residing in urban areas spented significantly less time preparing land. Female—headed households also allocated less time to land preparation. There wasis also some indication that women who that had tertiary education wereare less engaged in land preparation.

Model (2) presents the same model, but instruments fertility with the indicator for the first child being a girl. When we ran this model Fthe fertility effect new becaemes positive, but wasis not significantly different from zero. In model (3), which ere we looks at the sex of the first two children, the fertility gap_effect becaemes_ significant. The effect remaineds significant when we instrumented the fertility shortfall by the percentage of females born (model (4)) and in the over-identified model (5), but the effect size shrankreduces. An additional child reduceds time allocated to land preparation by about 25 days.

The second panel repeateds the same five models, but useds days spent on input application as the dependent variable. In none of the five specifications, did fertility seems to have a significant impact. Overall, time spented on input application wasis very limited anyway, as can be seen in figure 0.1. In all, households spented only about one day applying inputs (including planting). The only significant effect we findfound is that households where the mother has at least primary education allocated more time to input application (table 0.12).

The third panel presents results for time spentd on weeding. The results were similar to the ones for land preparation, but the effects were smaller. Each extra child reduceds time allocated to weeding by about 20 days. Full results in the appendix (table 0.13) show significant negative effects for

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Commented [JM65]: Apologies, but I don't understand....what higher fecundity?

Commented [JM66]: I wonder if it might be wise to briefly – in a phrase, describe Ugandan urban areas, because apparently people can farm in such areas, but they can't do that in American urban areas, (suburban, maybe, but not urban) and American or European readers will have difficulty picturing someone farming in a city

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 $\label{eq:communities} female_-headed \ households \ and \ for \ households \ in \ urban \ areas. \ We \ also \ \frac{find \underline{found}}{found} \ that \ communities \ that \ have \ a \ health \ center \ \frac{are}{find found} \ that \ communities \ that \ have \ a \ health \ center \ \frac{are}{find found} \ that \ cent$

spentding fewerless days on weeding. Finally, the last panel looks at the effects of fertility on days worked for harvesting. There iswas no significant positive association between the number of children in the family and the number of days spentd harvesting if we used only our binary instrument. We findfound a positive effect if we instrumented the fertility gap by the share of girls among siblings, but the effect iwas small compared to the other effects.

The above suggests that fertility affects time allocated to land preparation and weeding in a negative way. Harvesting seems to be less related to family size. Probably, when crops are ready to be harvested, farmers are more likely to put in the extra effort. This seems to be less evident for work that has an uncertain payoff in the future, such as weeding. The reductions of time allocated to land preparation and weeding may reduce both area planted and agricultural productivity. We will turn to this in tThe next section looks at this question.

Area planted, Production and Productivity

This section We will now looks at the fertility's effect of fertility on production and productivity. We will lt looks at productivity defined as kilograms harvested per acre of the five most important products separately. We will lt also looks at the value of total production, the value of production per acre, and the value of production per capita.

Table 0.7 reports on the second stage regressions of different aspects of production for the five most important crops. The table only reports the results for the coefficient on the fertility gap for the instrumental variable regression that uses the share of girls as excluded instrument. The regressions include the same control variables as in the previous sections. However, we now also added regional dummies, as some crops are grown more in some regions than in others. When the dependent variable wasis binary or censored, we estimated a tobit or probit is estimated using the methods described in Newey (1987).

The first row looks at the probability that a household grows the respective crop. For instance, the first entry in the first row tells us that the fertility gap didees not affect the probability that households cultivate maize. The third entry, however, shows that households that hadve a higher fertility gap wereare more likely to grow beans. Similarly, we findfound that higher fertility significantly reduceds the probability that matooke would be in grown. The next row looks at the total area reported to be used to grow each crop, measured in acres. We findfound a positive effect of the fertility gap on the area used to grow matooke. Fertility seemeds to be unrelated to the area used to grow any of the other crops. However, smaller households that greew more matooke mightay simply have had larger land holdings. Therefore, it was will

Commented [JM68]: Now a switch to present tense seems appropriate; the word "suggests" indicates you're talking about the implications of your findings.

be useful to also relate fertility to the share of each crop in total in terms of land size. This giaves an idea of the relative importance of each crop within the

household. This is presented in the next row. In this case, it seems that households with more children allocated less land as a share of total land to sweet potatoes. The next row looks at the value of production in kilograms. Only for matooke, larger households seemed to obtain a significantly lower quantity of matooke. The next row looks at production per capita. The lower production of matooke persisted if we accounted for household size. Finally, for none of the products did, the fertility gap haves a significant effect on yield.

Finally, in table 0.8 we presents results for en total production and productivity, using the prices for the different crops. Again we used We present again five different models. The first one wasis again regression that didees not take endogeneity into account. While in the previous regressions this was typically OLS, this might ay now have changed to a probit or tobit regression, depending on the nature of the dependent variable. The second regression instrumenteds the fertility gap with the sex of the first bernfirstborn. The third model useds the sex of the first two children born to the mother and the fourth useds the share of women amongstamong the children. As before, the fourth model instrumenteds the fertility difference by two instruments: the sex of the first-bernfirstborn and the share of girls among the children.

The first row gives results for the change in production. There seemeds to be no detectable effect from fertility on the total value of the production of the five crops. The second row expresses this production in per capita terms. The OLS estimates showed a positive effect of an increase in the fertility gap. However, if we confined to the exogenous part of fertility in the IV regressions, the effect disappeareds. The next row looks at a change in the total area allocated to the five crops. There is It shows again no significant effect from fertility. The final row, which looks at productivity defined as the total value of the five crops divided by the total area allocated to these five crops, also shows finds no causal impact from family size.

CONCLUSION

In this paper, wwe looked at the effect of fertility, defined as the number of biological children born to a enthe mother, on agricultural production and its determinants. One of the most evident determinants wasis household agricultural labourlabor. The identification strategy we used relies on the premise that, in partilineal societies, boys are preferred to girls in terms of offspring. Households that have a girl as the first child will have a higher propensity to add more children to the household. The fact that the sex of the first child is exogenous can be used to identify the causal impact of additional children on other variables such as labourlabor supply and productivity. Similarly, the fertility rule whereby one is more likely to stop

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Commented [JM70]: Is a noun missing here—is this perhaps supposed to say, "...we confined ourselves to the exogenous part etc. etc."? Or maybe, "...we confined it to the exogenous part etc. etc."?

Commented [JM71]: Present tense because it will always rely..

having children after a boy means that, on average, larger households consist of more girls. Therefore, the share of females in the total

number of children can also be used as an instrument.

Our first stage regression performede reasonably well. We findfound a significant negative effect of an indicator variable that the first bernfirstborn wasis a girl on a variable that measures the shortfall from fecundity. We equally findfound a negative effect of an indicator that the first two children wereare female. Finally, we also findfound that households with a relatively higher share of girls wereare negatively related to the fertility gap. While our instruments wereare significant and hadve the correct sign, explanatory power as measured by the partial R-squared wasis low. We therefore used inference methods in the second stage that wereare robust to weak instruments.

In the second stage regression, we <u>findfound</u> that fertility affects both time women and men allocated to agricultural production. However, most of the <u>labourlabor</u> time lost as a consequence of an exogenous increase in children <u>wasie</u> borne by the woman. <u>Especially IL</u> and preparation and weeding, <u>especially</u>, <u>were are</u> activities that seemed to suffer from excessive fertility. When we looked at crops, we <u>findfound</u> that only matooke and sweet potatoes <u>were are</u> significantly affected by fertility.

Matooke is the most important stable crop_in Uganda, providing 18 percent of caloric intake (Haggblade and Dewina, 2_2010). The finding that young households that have higher fertility wereare reducing the most important source of calories suggests that higher fertility also causes under-nutrition. Sweet potatoes are also a typical food security crop, with a low return but also low risk (Dercon, 1996). It is also a crop that is mostly under the control of the womean, who does much of the work on the field.

That said, the fact that we reliedy on a cross-section of households also limitse to what the extent to which our conclusions can be generalized. It may well be that households that are at a later stage in their life cycle profit much more from larger household size. For instance, in households where the mother has reduced fertility, she may have more time to work in agricultural activities. In addition, children may provide cheap and flexible labourlabor at a later age. Therefore, we want to stress that our results only hold for the subset of "young" households, where the womaen is between 16 and 32.

There are different ways in which the negative effect of fertility on labourlabor and production can be influenced. First, our analysis reconfirms the need for fertility:—reducing policies. Apart from known fertility—reducing policies such as women's education and improved maternal health care, the most promising policies wshould try to work on the root cause of increased fertility. This should be done by reducing the propensity of households' propensity to have higher fertility if the first-bornfirstborn is a girl. We can think of a host of policies that would do this by pushing go against the patrilineal nature of these societies. For example, Uganda may consider changing its land act to make it similar to what Kenya recently did and give

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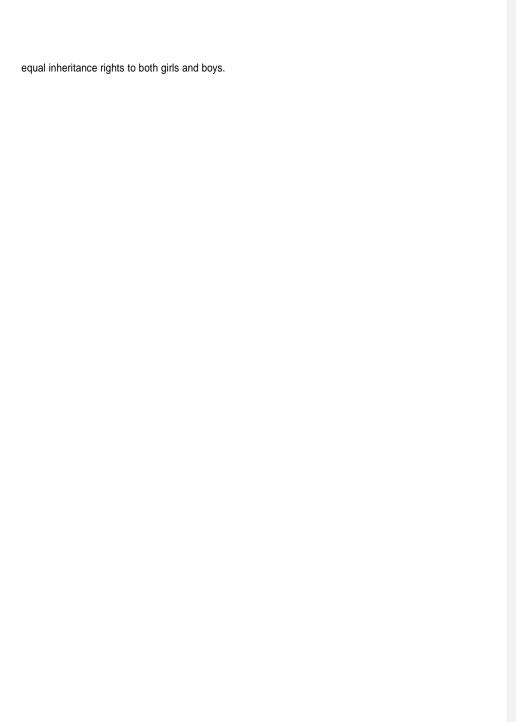
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The above policy response involves addressing cultural issues related to high fertility, some of which may face considerable resistance. Changing a set of cultural values is likely to be a very slow process. Meanwhile, the government of Uganda should support the nutritional needs of young families. It should also consider introducing agricultural technologies that save on agricultural labourlabor, especially for women.

Commented [JM76]: By "above policy response" do you mean having Uganda change its land act so that it's more like Kenya's?

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TABLES

Table 0.1—g Gender and fertility

	prob +1		prob +1		av fertility
1st=boy	0.375	1st=boy, 2nd=boy	0.132	% daughters<0.5	2.78
1st=girl	0.393	1st=girl, 2nd=boy	0.134	% daughters>0.5	2.88
		1st=girl, 2nd=girl	0.139		
	gap		gap		gap
1st=boy	2.46	1st=boy, 2nd=boy	2.46	% daughters<0.5	2.41
1st=girl	2.26	1st=girl, 2nd=boy	2.38	% daughters>0.5	2.32
-		1st=girl, 2nd=girl	2.06	-	

Table 0.2_—Descriptive statistics for crop production

	maize	beans	potatoes	cassava	matooke
growing crop (% of households)	59.1	52.6	38.8	51.8	43.1
crop area (acre)	0.473	0.263	0.165	0.301	0.264
crop area (% of total area)	17.1	12.6	8.2	12.0	10.2
production (kg)	38.7	12.8	85.1	83.2	348.9
production per capita (kg)	18.1	6.4	38.7	38.4	168.6
yield (kg per acre)	358.2	128.5	1096.3	1030.8	2067.3

Table 0.3_—First stage regression - OLS estimation of fertility gap

oldestgirl	-0.203** (0.067)			
2oldestgirls	,	-0.190* (0.082)		
3oldestgirls		, ,	-0.147 (0.117)	
percentfmales			,	-0.278** (0.094)
femhead	1.186** (0.098)	1.168** (0.105)	1.201** (0.118)	1.186**
urban	0.322**	0.273**	0.077 (0.115)	0.325**
mprim	-0.155 [*]	-0.025	-0.009	-0.159*
msec	(0.075) 0.259*	(0.082) 0.220+	(0.094) 0.193	(0.075) 0.257*
mthird	(0.101) 1.058**	(0.121) 0.914**	(0.150) 0.755+	(0.101) 1.060**
health	(0.192) 0.095	(0.250) 0.107	(0.403) 0.151	(0.192) 0.090
school	(0.124) 0.040	(0.146) -0.005	(0.171)	(0.124) 0.043
cdied	(0.070) 0.284**	(0.078) 0.204+	(0.090) 0.117	(0.070) 0.285**
cons	(0.100) 2.172** (0.074)	(0.108) 1.946** (0.075)	(0.127) 1.782** (0.079)	(0.100) 2.209** (0.081)
r2 N	0.091 2656	0.075 2036	0.065 1391	0.091 2656

Table 0.4_—Effect of fertility on total time worked in agriculture

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	-0.122	46.799+	64.297+	69.227**	66.349**
	(1.245)	(32.112)	(48.865)	(35.208)	(36.319)
femhead	-38.213**	-95.268*	-116.211+	-122.540**	-119.040**
	(4.793)	(39.623)	(59.496)	(44.167)	(45.268)
urban	-23.411**	-35.338**	-40.541**	-41.038**	-40.307**
	(5.780)	(11.352)	(13.796)	(13.622)	(13.516)
mprim	1.392	8.132	6.058	11.354	10.940
	(4.897)	(7.802)	(8.092)	(9.128)	(9.065)
msec	-6.121	-10.595	-7.207	-12.734	-12.459
	(6.516)	(8.899)	(11.617)	(11.167)	(10.850)
mthird	-20.792*	-70.642+	-83.864	-94.470*	-91.412*
	(9.946)	(39.814)	(54.795)	(45.224)	(46.122)
health	-15.118*	-21.809*	-18.573	-25.008+	-24.597+
	(7.694)	(11.099)	(14.923)	(13.928)	(13.544)
school	12.570**	7.633	9.887	5.274	5.576
	(4.793)	(6.914)	(9.145)	(7.881)	(7.879)
cdead	13.573*	3.161	1.904	-1.816	-1.177
	(6.273)	(10.788)	(12.672)	(12.954)	(12.801)
cons	87.193**	-7.527	-29.950	-52.803	-46.992
	(4.718)	(64.960)	(90.557)	(70.927)	(73.206)
N	2016	2016	1620	2016	2016
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.5_—2SLS estimates of household labourlabor supply

	(1)	(2)	(3)	(4)	(5)
	OLS	2ŠĹS	2ŠĹS	2ŠĹS	LÌML
			days worked <u>by</u> mo	other	
	0.533	29.890*	54.070**	40.841**	38.773**
	(0.735)	(17.769)	(33.364)	(19.353)	(19.085)
			days worked fath	ner	
	-0.620	10.928	5.915	22.327*	20.076+
	(0.668)	(13.983)	(25.595)	(13.580)	(14.756)
N	2016	2016	1620	2016	2016
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.6_—2SLS estimates of household labourlabor allocation

	(1)	(2)	(3)	(4)	(5)					
	OLS	2SLS	2SLS	2SLS	LIML					
	time allocated to land preparation									
	-0.048	15.876	41.127**	26.028**	25.278*					
	(0.528)	(12.737)	(25.884)	(13.720)	(14.761)					
		time a	allocated to input a	pplication						
	0.051	1.812	0.849	2.372	2.224					
	(0.136)	(1.632)	(1.686)	(1.829)	(1.741)					
		tiı	me allocated to we	eding						
	0.026	17.708*	25.959*	21.385*	20.492*					
	(0.468)	(11.724)	(17.541)	(11.730)	(11.253)					
		tim	ne allocated to harv	esting						
	-0.082	5.315	-3.591	13.852+	12.485+					
	(0.468)	(10.597)	(20.563)	(8.987)	(10.379)					
			. ,	•	,					
N	2015	2015	1619	2015	2015					
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl					

Table 0.7_—2SLS estimates of effect of fertility on crop mix, area, production and yield

	maize	beans	s pot	cassava	Matooke
		504.10	0 00.		
growing	0.431	-0.193	0.556+	-0.003	0.622+
0 0	(0.359)	(0.332)	(0.401)	(0.302)	(0.428)
total area	0.432	-0.119	0.187	0.105	0.603+
	(0.425)	(0.199)	(0.254)	(0.289)	(0.425)
area share	0.028	-0.085	0.169+	-0.027	0.043
	(0.087)	(0.080)	(0.115)	(0.081)	(0.081)
production	56.799	-4.677	182.878	29.714	1931.785+
	(71.795)	(22.600)	(176.010)	(189.684)	(1255.565)
production per capita	26.593	-1.079	55.706	-11.725	2026.613*
	(38.247)	(12.727)	(87.191)	(103.129)	(1135.716)
yield	-22.060	41.329	69.834	-480.421	-383.889
	(169.179)	(54.371)	(694.248)	(728.459)	(882.517)

Note: Huber-White standard errors in parentheses, +, * and ** denote significance at the 10, 5 and 1 percent level respectively. All regressions use the share of female children in total number of children as instrument.

Table 0.8_—2SLS estimates of total production

	(1)	(2)	(3)	(4)	(5)
	ÒĽS	2ŠĹS	2ŠĹS	2ŠĹS	LÌML
production	-3.411	21.056	-20.275	16.988	19.334
(x UGX1000)	(2.697)	(46.911)	(57.977)	(48.493)	(44.172)
production/capita	4.133**	10.501	-9.251	14.518	12.340
(x UGX1000)	(1.527)	(26.009)	(24.107)	(27.183)	(24.605)
area	-0.033	0.120	0.042	0.145	0.132
	(0.023)	(0.343)	(0.443)	(0.359)	(0.325)
yield	-1.697	43.088	-96.861	0.013	8.818
	(2.862)	(81.451)	(140.870)	(61.118)	(69.674)
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

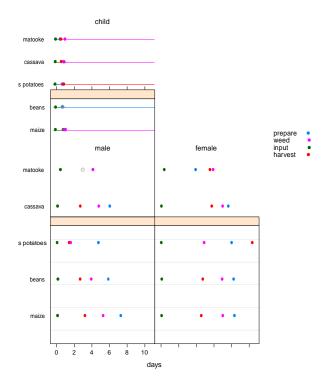
instrument: - 1st = girl 1st & 2nd = girl % girl 1st = girl & % girl

Note: Huber-White standard errors in parentheses, +, * and ** denote significance at the 10, 5 and 1 percent level respectively.

FIGURES

Figure 0.1_—Average number of days worked

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APPENDIX

Table 0.9_—Effect on days worked by mother (full results)

	(1)	(2)	(2)	(4)	(E)
	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	0.533	29.890*	54.070**	40.841**	38.773**
	(0.735)	(17.769)	(33.364)	(19.353)	(19.085)
femhead	-7.906*	-43.577*	-71.806+	-56.883*	-54.370*
	(3.715)	(21.981)	(40.831)	(24.354)	(23.883)
urban	-13.317**	-20.758**	-25.334*	-23.534**	-23.009**
	(3.536)	(6.910)	(10.247)	(8.041)	(7.811)
mprim	0.023	4.177	3.472	5.726	5.434
	(2.754)	(4.566)	(5.864)	(5.194)	(5.085)
msec	1.115	-1.624	2.608	-2.646	-2.453
	(4.706)	(5.869)	(9.177)	(6.913)	(6.676)
mthird	-15.001*	-46.192*	-70.087+	-57.827*	-55.630*
	(6.961)	(23.251)	(38.987)	(25.869)	(25.530)
health	-7.274	-11.634+	-9.296	-13.260	-12.953
	(4.607)	(6.967)	(11.406)	(8.326)	(8.044)
school	6.740*	3.705	3.073	2.573	2.786
	(2.703)	(3.837)	(6.100)	(4.461)	(4.334)
cdied	6.810+	0.209	-2.949	-2.253	-1.788
	(3.874)	(6.592)	(9.328)	(7.510)	(7.333)
cons	49.054**	-10.209	-48.668	-32.314	-28.140
	(2.815)	(36.011)	(61.945)	(39.017)	(38.525)
N	2016	2016	1620	2016	2016
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.10_—Effect on days worked by father (full results)

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	-0.620	10.928	5.915	22.327*	20.076+
	(0.668)	(13.983)	(25.595)	(13.580)	(14.756)
femhead	-30.349**	-44.381**	-39.124	-58.231**	-55.496**
	(2.216)	(16.903)	(30.436)	(16.952)	(18.155)
urban	-10.103**	-13.030**	-14.318**	-15.919**	-15.349**
	(2.902)	(4.347)	(5.249)	(5.056)	(5.016)
mprim	1.374	3.008	2.241	4.620	4.302
-	(2.717)	(3.271)	(2.967)	(3.769)	(3.687)
msec	-7.239*	-8.317*	-9.670 [*]	-9.381*	-9.171 [*]
	(3.120)	(3.454)	(3.770)	(4.404)	(4.175)
mthird	-5.829	-18.098	-9.522	-30.209+	-27.817
	(4.926)	(16.387)	(26.183)	(17.049)	(17.973)
health	-7.849*	-9.564*	-9.534*	-11.257*	-10.922*
	(3.589)	(4.045)	(4.444)	(5.168)	(4.913)
school	5.826*	4.632	7.310	3.454	3.687
	(2.843)	(3.630)	(4.782)	(3.620)	(3.742)
cdied	6.756*	4.159	5.314	1.596	2.102
	(3.272)	(4.374)	(5.030)	(5.209)	(5.089)
cons	38.067**	14.756	26.774	-8.254	-3.711
	(2.377)	(28.289)	(47.345)	(27.361)	(29.750)
N	2016	2016	1620	2016	2016
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.11_—Effect on days spend on preparing fields (full results)

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LÌML
fgap	-0.048	15.876	41.127**	26.028**	25.278*
	(0.528)	(12.737)	(25.884)	(13.720)	(14.761)
femhead	-15.173**	-34.568*	-64.891*	-46.932**	-46.018*
	(2.023)	(15.704)	(31.846)	(17.259)	(18.412)
urban	-6.940**	-11.013*	-17.002*	-13.610*	-13.418*
	(2.623)	(4.481)	(7.827)	(5.417)	(5.493)
mprim 2	-2.251	-0.069	-0.597	1.322	1.219
	(2.209)	(3.307)	(4.513)	(3.585)	(3.703)
msec	-2.980	-4.482	-2.924	-5.439	-5.368
	(2.832)	(3.537)	(6.729)	(4.377)	(4.336)
mthird	-11.214**	-28.140+	-51.405+	-38.931*	-38.133*
	(3.993)	(15.545)	(29.529)	(17.591)	(18.553)
health	-3.914	-6.215	-6.296	-7.682	-7.573
	(3.233)	(4.293)	(8.266)	(5.446)	(5.379)
school	2.919	1.343	0.511	0.338	0.412
	(1.932)	(2.546)	(4.654)	(3.079)	(3.073)
cdead	5.262+	1.632	-2.547	-0.682	-0.511
	(2.719)	(4.291)	(7.133)	(5.142)	(5.203)
cons	35.817**	3.741	-39.561	-16.708	-15.197
	(2.261)	(26.072)	(47.770)	(27.523)	(29.771)
N	2015	2015	1619	2015	2015
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.12_—Effect on days spend on input application (full results)

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	0.051	1.812	0.849	2.372	2.224
	(0.136)	(1.632)	(1.686)	(1.829)	(1.741)
femhead	0.349	-1.795	-0.493	-2.478	-2.297
	(0.747)	(1.611)	(1.756)	(1.717)	(1.628)
urban	-0.514	-0.964	-0.674	-1.107	-1.069
	(0.386)	(0.713)	(0.615)	(0.788)	(0.763)
mprim	0.466*	0.707*	0.657**	0.784*	0.764*
	(0.182)	(0.325)	(0.246)	(0.376)	(0.358)
msec	0.978	0.812	1.158	0.759	0.773
	(0.881)	(0.814)	(1.126)	(0.802)	(0.804)
mthird	-0.684	-2.555	-1.897	-3.151	-2.993
	(1.202)	(2.575)	(2.550)	(2.835)	(2.745)
health	0.198	-0.057	-0.088	-0.138	-0.116
	(0.489)	(0.611)	(0.613)	(0.687)	(0.663)
school	-0.045	-0.219	-0.108	-0.275	-0.260
	(0.289)	(0.302)	(0.339)	(0.302)	(0.300)
cdead	-0.142	-0.543	-0.235	-0.671	-0.637
	(0.222)	(0.513)	(0.401)	(0.591)	(0.564)
cons	0.526	-3.021	-0.935	-4.150	-3.851 [°]
	(0.448)	(3.370)	(3.238)	(3.793)	(3.609)
N	2015	2015	1619 [°]	2015	2015
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.13_—Effect on days spend on weeding (full results)

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LÌML
fgap	0.026	17.708*	25.959*	21.385*	20.492*
	(0.468)	(11.724)	(17.541)	(11.730)	(11.253)
femhead	-13.623**	-35.158*	-45.639*	-39.637**	-38.550**
	(1.728)	(14.648)	(21.645)	(14.779)	(14.180)
urban	-9.339**	-13.862**	-15.649**	-14.802**	-14.574**
	(2.340)	(4.495)	(5.509)	(4.756)	(4.613)
mprim	-0.489	1.934	1.010	2.438	2.315
	(1.772)	(2.772)	(3.137)	(2.968)	(2.880)
msec	-2.274	-3.942	-2.715	-4.289	-4.205
	(2.594)	(3.488)	(4.834)	(3.904)	(3.775)
mthird	-5.350	-24.143	-30.226	-28.052+	-27.103+
	(4.259)	(14.960)	(20.546)	(15.193)	(14.766)
health	-8.375**	-10.930**	-10.259+	-11.461**	-11.332**
	(2.059)	(3.844)	(5.363)	(4.270)	(4.136)
school	4.598**	2.848	3.290	2.484	2.572
	(1.751)	(2.336)	(3.165)	(2.470)	(2.417)
cdead	5.040*	1.010	0.656	0.172	0.375
	(2.323)	(4.315)	(5.183)	(4.525)	(4.412)
cons	29.534**	-6.083	-17.387	-13.490	-11.692
	(1.783)	(23.464)	(32.493)	(23.516)	(22.524)
N	`2015 [´]	` 2015 [′]	` 1619 [′]	` 2015 [′]	` 2015 [^]
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.14_—Effect on days spend on harvesting (full results)

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	-0.082	5.315	-3.591	13.852+	12.485+
	(0.468)	(10.597)	(20.563)	(8.987)	(10.379)
femhead	-10.650**	-17.223	-6.270	-27.620*	-25.955*
	(1.771)	(12.609)	(24.287)	(11.128)	(12.624)
urban	-6.876**	-8.257**	-8.124*	-10.440**	-10.091**
	(1.873)	(2.977)	(3.852)	(3.362)	(3.402)
mprim	2.741	3.481	3.569+	4.650+	4.463+
	(1.906)	(2.131)	(2.042)	(2.511)	(2.448)
msec	-1.835	-2.344	-2.761	-3.149	-3.020
	(2.343)	(2.335)	(2.849)	(2.969)	(2.822)
mthird	-3.720	-9.457	-0.912	-18.530+	-17.078
	(3.366)	(11.938)	(20.598)	(11.097)	(12.290)
health	-2.340	-3.120	-0.215	-4.353	-4.156
	(3.167)	(3.283)	(4.032)	(4.004)	(3.853)
school	5.394*	4.860+	6.889+	4.015	4.150
	(2.150)	(2.803)	(3.685)	(2.604)	(2.760)
cdead	3.501	2.271	3.443	0.325	0.637
	(2.351)	(2.914)	(3.565)	(3.384)	(3.345)
cons	22.551**	11.680	29.707	-5.516	-2.763
	(1.469)	(21.379)	(38.106)	(18.133)	(20.921)
N	2015	2015	1619	2015	2015
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.15_—Total production (full tobit results)

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	-3.411	21.056	-20.275	16.988	19.334
	(2.697)	(46.911)	(57.977)	(48.493)	(44.172)
femhead	-69.276**	-97.608+	-41.725	-92.870	-95.599+
	(13.700)	(55.994)	(67.508)	(57.696)	(52.888)
urban	-191.662**	-199.111**	-183.271**	-197.891**	-198.597**
	(13.527)	(18.843)	(19.978)	(19.209)	(18.219)
mprim	31.542**	35.247**	40.222**	34.604**	34.970**
	(9.652)	(12.321)	(11.070)	(12.382)	(12.050)
msec	-14.660	-20.572	-17.281	-19.563	-20.150
	(14.527)	(17.768)	(20.052)	(17.924)	(17.324)
mthird	29.624	2.888	78.249	7.286	4.749
	(35.753)	(58.152)	(63.402)	(59.720)	(55.538)
health	-12.603	-15.259	-10.523	-14.869	-15.089
	(15.952)	(16.802)	(19.526)	(16.816)	(16.688)
school	37.105**	36.348**	39.394**	36.483**	36.411**
	(9.238)	(9.316)	(10.397)	(9.273)	(9.279)
cdied	15.566	8.472	9.039	9.642	8.970
	(12.878)	(18.682)	(18.569)	(18.990)	(18.093)
cons	153.527**	100.065	197.728	108.964	103.831
	(12.388)	(103.164)	(121.588)	(106.568)	(97.211)
N	2637	2637	2020	2637	2637
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.16_—Total production per capita (full tobit results)

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	4.133**	10.501	-9.251	14.517	12.340
	(1.527)	(26.009)	(24.107)	(27.183)	(24.605)
femhead	-39.851**	-47.223	-10.658	-51.859	-49.346+
	(7.938)	(31.039)	(28.068)	(32.335)	(29.453)
substrat	-98.207**	-100.147**	-67.933**	-101.377**	-100.709**
	(8.560)	(10.451)	(8.290)	(10.764)	(10.150)
mprim	20.087**	21.052**	17.671**	21.646**	21.324**
	(5.220)	(6.835)	(4.606)	(6.946)	(6.716)
msec	-3.727	-5.265	-4.161	-6.221	-5.707
	(8.632)	(9.846)	(8.330)	(10.040)	(9.644)
mthird	20.268	13.311	55.245*	8.899	11.294
	(24.284)	(32.189)	(26.298)	(33.422)	(30.880)
health	-8.756	-9.447	-8.607	-9.909	-9.653
	(8.925)	(9.326)	(8.131)	(9.440)	(9.308)
school	12.227*	12.030*	11.697**	11.910*	11.976*
	(4.969)	(5.167)	(4.325)	(5.203)	(5.172)
cdied	5.932	4.085	2.685	2.916	3.551
	(7.121)	(10.360)	(7.726)	(10.649)	(10.082)
cons	49.910**	35.996	68.721	27.226	31.979
	(6.365)	(57.199)	(50.558)	(59.739)	(54.151)
N	2637	2637	2020	2637	2637
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl
mon annon.		iot – giii	101 G 211G - gill	70 giii	15t = giii & 70 giii

Table 0.17_—Total area (full tobit results)

-	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	-0.033	0.120	0.042	0.145	0.132
	(0.023)	(0.343)	(0.443)	(0.359)	(0.325)
femhead	-0.511**	-0.689+	-0.530	-0.718+	-0.703+
	(0.097)	(0.410)	(0.516)	(0.427)	(0.389)
urban	-1.351**	-1.398**	-1.345**	-1.406**	-1.401**
	(0.111)	(0.138)	(0.153)	(0.142)	(0.134)
mprim	0.172*	0.196*	0.250**	0.199*	0.197*
	(0.071)	(0.090)	(0.085)	(0.091)	(0.089)
msec	-0.175	-0.212	-0.215	-0.218	-0.215+
	(0.111)	(0.130)	(0.153)	(0.133)	(0.127)
mthird	-0.035	-0.202	0.088	-0.230	-0.215
	(0.210)	(0.428)	(0.487)	(0.443)	(0.411)
health	-0.133	-0.150	-0.158	-0.153	-0.151
	(0.108)	(0.123)	(0.150)	(0.125)	(0.123)
school	0.310**	0.306**	0.331**	0.305**	0.305**
	(0.070)	(0.068)	(0.080)	(0.069)	(0.068)
cdied	0.103	0.058	0.012	0.051	0.055
	(0.093)	(0.137)	(0.142)	(0.140)	(0.133)
cons	0.793**	0.459	0.630	0.403	0.432
	(0.080)	(0.755)	(0.930)	(0.788)	(0.715)
N	2637	2637	2020	2637	2637
	2037				
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

Table 0.18_—Total yield (x1000 UGX per acre)

	(1)	(2)	(3)	(4)	(5)
	OLS	2SLS	2SLS	2SLS	LIML
fgap	-1.697	43.088	-96.861	0.013	8.818
	(2.862)	(81.451)	(140.870)	(61.118)	(69.674)
femhead	-32.827*	-82.434	62.869	-34.721	-44.474
	(13.620)	(91.141)	(149.672)	(69.314)	(78.576)
urban	-19.120	-31.562	-17.771	-19.596	-22.042
	(17.498)	(29.340)	(36.155)	(24.900)	(26.586)
mprim	9.668	12.719	17.577	9.785	10.384
	(11.889)	(14.153)	(21.954)	(12.962)	(13.267)
msec	34.954+	32.322	24.785	34.854+	34.336+
	(20.106)	(22.005)	(27.654)	(20.229)	(20.478)
mthird	11.594	-41.631	115.467	9.562	-0.903
	(38.115)	(99.945)	(156.775)	(78.774)	(87.257)
health	-9.314	-18.619	0.558	-9.669	-11.498
	(17.760)	(26.018)	(27.844)	(21.578)	(22.827)
school	10.851	5.932	28.583	10.663	9.696
	(11.524)	(15.622)	(23.309)	(13.193)	(13.852)
cdied	1.373	-14.487	22.702	0.767	-2.351
	(16.065)	(30.019)	(38.702)	(25.321)	(27.246)
cons	178.153**	154.719	449.143	244.882+	226.452
	(16.668)	(171.814)	(283.600)	(130.707)	(148.357)
	,	•		•	,
N	1567	1567	1278	1567	1567
instrument:	-	1st = girl	1st & 2nd = girl	% girl	1st = girl & % girl

