

## 13.1 Histograms

Suppose  $X_1 \dots X_n$  are i.i.d p on  $[0, 1]^d$ . Divide the unit cube into  $B_1, \dots, B_n$  cubes each with side length  $h$ ,  $N \leq (\frac{1}{h})^d$ . Define

$$\hat{p}_h(x) = \sum_j \frac{\hat{\theta}_j}{h^d} I(x \in B_j) \quad \text{where} \quad \hat{\theta}_j = \frac{1}{n} \sum_{i=1}^n I(x_i \in B_j)$$

We want to bound

$$P(\|\hat{p}_h - p\|_\infty > \epsilon)$$

Consider :

$$\begin{aligned} & P(\max_j |\frac{\hat{\theta}_j}{h^d} - \frac{\theta_j}{h^d}| > \epsilon) \\ & (\theta_j = \int_{B_j} p(x) dx) \\ & = P(\max_j |\hat{\theta}_j - \theta_j| > \epsilon h^d) \end{aligned}$$

assume  $|p(x) - p(y)| \leq L\|x - y\|$  then  $p$  is bounded. It follows that

$$\begin{aligned} \theta_j &= \int_{B_j} p(u) du \leq c \int_{B_j} du = ch^d \\ &\rightarrow \theta_j(1 - \theta_j) \leq \theta_j \leq ch^d \end{aligned}$$

Now

$$\begin{aligned} P(|\hat{\theta}_j - \theta_j| > h^d \epsilon) &\leq 2 \exp(-\frac{1}{2} \frac{n \epsilon^2 h^{2d}}{\theta_j(1 - \theta_j) + \epsilon h^d / 3}) \\ &\leq 2 \exp(-\frac{1}{2} \frac{n \epsilon^2 h^{2d}}{ch^d + \epsilon h^d / 3}) \\ &\leq 2 \exp(-cn \epsilon^2 h^d) \\ &(c = \frac{1}{2(c + \frac{1}{3})}) \end{aligned}$$

$$P(\max_j |\hat{\theta}_j - \theta_j| > h^d \epsilon) \leq 2h^{-d} \exp(-cn \epsilon^2 h^d) = \Pi_n$$

Now

$$\begin{aligned}
p(x) - p_n(x) &= p(x) - \frac{\int_{B_j} p(u) du}{h^d} \\
&= \frac{1}{h^d} \int_{B_j} (p(x) - p(u)) du \\
\rightarrow |\cdot| &\leq \frac{1}{h^d} \int_{B_j} |p(x) - p(u)| du \\
&\leq \frac{1}{h^d} Lh\sqrt{d} \int du = Lh\sqrt{d}
\end{aligned}$$

wp  $1 - \Pi_n$

$$\begin{aligned}
\|\hat{p}_h - p\|_\infty &\leq \|\hat{p}_h - p_h\|_\infty + \|p_h - p\|_\infty \\
&\leq \epsilon + L\Pi h \\
\epsilon &= \sqrt{\frac{1}{cnh^d} \log\left(\frac{2}{\delta h^d}\right)} \\
\delta &= \Pi n
\end{aligned}$$

wp  $1 - \delta$ ,

$$\begin{aligned}
\|\hat{p}_h - p\|_\infty &\leq \sqrt{\frac{1}{cnh^d} \log\left(\frac{2}{\delta h^d}\right)} + L\sqrt{d}h \\
h &= \left(\frac{c}{n}\right)^{\frac{1}{2+d}}
\end{aligned}$$

wp  $1 - \Pi_n$ .

$$\|\hat{p}_h - p\|_\infty = O\left(\left[\frac{\log n}{n}\right]^{\frac{1}{2+d}}\right)$$

## 13.2 Martingales

Let  $(\Omega, \mathcal{F})$ . Define sequence  $\mathcal{F}_i$  of  $\sigma$ -algebras s.t.  $\mathcal{F}_i \subset \mathcal{F}_{i+1}$  and  $\mathcal{F}_0 = \sigma\{\Omega\}$ . We call  $((\mathcal{F}_i, (X_i)))$  a martingale w.r.t  $(\mathcal{F}_i)$  if  $X_i$  is  $\mathcal{F}_i$  measurable and  $E[X_i | \mathcal{F}_{i-1}] = X_{i-1}$  where  $\mathcal{F}_i = \sigma(X_1, X_2, \dots, X_i)$ . Note the following

- $E[X_i] = \mu$
- If  $E[X_i | \mathcal{F}_{i-1}] = 0$  call it a martingale difference.

**Example 13.1.**  $Y_0 = E[X_0], Y_1 = X_1 - X_0$

**Example 13.2.**  $\exists X$  s.t. can create  $\mathcal{F}_i \subset \sigma[X], X_i = E[X | \mathcal{F}_i]$ .

## 13.3 Azuma's inequality

Let  $((\mathcal{F}_i), (y_i))$  be a martingale difference s.t. for  $(a_i)$  and  $(b_i)$ ,  $P(Y_i \in [a_i, b_i]) = 1$  for all  $i$ . Then for  $\delta > 0$ ,

$$P\left(\frac{1}{n} \left| \sum Y_i \right| \geq \delta\right) \leq 2 \exp^{-\frac{n\delta^2}{\sum_{i=1}^n (b_i - a_i)^2}}$$

Pf: sketch

$$E e^{\lambda \sum_{i=1}^n Y_i} = E[E[e^{\lambda \sum_{i=1}^n Y_i} | \mathcal{F}_{n-1}]] = E[e^{\lambda \sum_{i=1}^n Y_i} E[e^{\lambda Y_n} | \mathcal{F}_{n-1}]]$$

Let  $Z = f(X_1, \dots, X_n)$ .

**Theorem 13.3** (McDiamid's). *If  $f$  satisfies "bounded differences" :*

$$\sup_{x_1, x_2, \dots, x_n, x' \in X^{n+1}} |f(x_1, \dots, x_n) - f(x_2, \dots, x_{i-1}, x, x_{i+1}, \dots, x_n)| < C_i$$

Then  $\delta > 0$ ,  $P(Z - E[Z] > \delta) \leq \exp^{-\frac{2\delta^2}{\sum c_i^2}}$

Pf : sketch

$$Y_0 = E[Z], Y_1 = E[Z|X_1 \dots X_i] = E[Z|X_1 \dots X_{i-1}].$$