

**1 Algebra**

Absolute Value Inequalities:

$$|f(x)| < a \Rightarrow -a < f(x) < a$$

$$|f(x)| > a \Rightarrow f(x) > a \text{ or } f(x) < -a$$

**2 Calculus**

**2.1 Concavity in 1 dimension**

If  $g : I \rightarrow \mathbb{R}$  is twice differentiable in the interval  $I$ , i.e.  $g''(x)$  exists for all  $x \in I$ , then  $g$  is

- concave if and only if  $g''(x) \leq 0$  for all  $x \in I$ ;
- strictly concave if  $g''(x) < 0$  for all  $x \in I$ ;
- convex if and only if  $g''(x) \geq 0$  for all  $x \in I$ ;
- strictly convex if  $g''(x) > 0$  for all  $x \in I$ ;

**2.2 Multivariate Calculus**

**Gradient**

Let

$$f : \mathbb{R}^d \longrightarrow \mathbb{R} \theta = \theta = \begin{pmatrix} \theta_1 \\ \theta_2 \\ \vdots \\ \theta_d \end{pmatrix} \mapsto f(\theta)$$

denote a twice differentiable function, the Gradient  $\nabla$  of  $f$  is defined as:

$$\nabla f : \mathbb{R}^d \rightarrow \mathbb{R}^d$$

$$\theta = \begin{pmatrix} \theta_1 \\ \theta_2 \\ \vdots \\ \theta_d \end{pmatrix} \mapsto \begin{pmatrix} \frac{\partial f}{\partial \theta_1} \\ \frac{\partial f}{\partial \theta_2} \\ \vdots \\ \frac{\partial f}{\partial \theta_d} \end{pmatrix}$$

**Hessian**

The Hessian of  $f$  is the matrix

$$\mathbf{H} : \mathbb{R}^d \rightarrow \mathbb{R}^{d \times d}$$

whose entry in the  $i$ -th row and  $j$ -th column is defined by

$$(\mathbf{H}f)_{ij} := \frac{\partial^2}{\partial \theta_i \partial \theta_j} f, \quad 1 \leq i, j \leq d$$

**Semi-Definiteness**

A symmetric (real-valued)  $d \times d$  matrix  $\mathbf{A}$  is:

- Positive semi-definite if  $\mathbf{x}^T \mathbf{A} \mathbf{x} \geq 0$  for all  $\mathbf{x} \in \mathbb{R}^d$ .
- Positive definite if inequality above is strict  $\mathbf{x}^T \mathbf{A} \mathbf{x} > 0$  for all non-zero vectors  $\mathbf{x} \in \mathbb{R}^d$
- Negative semi-definite (resp. negative definite) if  $\mathbf{x}^T \mathbf{A} \mathbf{x}$  is non-positive (resp. negative) for all  $\mathbf{x} \in \mathbb{R}^d - \{0\}$ .

Positive (or negative) definiteness implies positive (or negative) semi-definiteness.

**3 Important probability distributions**

**Bernoulli**

Parameter  $p \in [0, 1]$ . Discrete, describes the success or failure in a single trial.

$$p_x(k) = \begin{cases} p, & \text{if } k = 1 \\ (1 - p), & \text{if } k = 0 \end{cases}$$

$$E[X] = p$$

$$Var(X) = p(1 - p)$$

**Exponential**

Parameter  $\lambda$ . Continuous

$$f_x(x) = \begin{cases} \lambda \exp(-\lambda x), & \text{if } x \geq 0 \\ 0, & \text{o.w.} \end{cases}$$

$$F_x(x) = \begin{cases} 1 - \exp(-\lambda x), & \text{if } x \geq 0 \\ 0, & \text{o.w.} \end{cases}$$

$$E[X] = \frac{1}{\lambda}$$

$$Var(X) = \frac{1}{\lambda^2}$$

**Normal (Gaussian)**

Parameters  $\mu$  and  $\sigma^2 > 0$ . Continuous

$$f(x) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

$$E[X] = \mu$$

$$Var(X) = \sigma^2$$

Useful properties:

**Poisson**

Parameter  $\lambda$ . Discrete, approximates the binomial PMF when  $n$  is large,  $p$  is small, and  $\lambda = np$ .

$$(p_{\mathbf{x}}(k) = \exp(-\lambda) \frac{\lambda^k}{k!} \text{ for } k = 0, 1, \dots,$$

$$E[X] = \lambda$$

$$Var(X) = \lambda$$

**Uniform**

**4 Expectation and Variance**

**Expectation**

**Variance**

**Covariance**

**Variance and expectation of mean of n iid random variables**

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} P_\mu$ , where  $E(X_i) = \mu$  and  $Var(X_i) = \sigma^2$  for all  $i = 1, 2, \dots, n$  and  $\overline{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ .

Variance of the Mean:

$$Var(\overline{X}_n) = \left(\frac{\sigma^2}{n}\right)^2 Var(X_1 + X_2, \dots, X_n) = \frac{\sigma^2}{n}.$$

Expectation of the mean:

$E[\overline{X}_n] = \frac{1}{n} E[X_1 + X_2, \dots, X_n] = \mu$ .

**5 LLN and CLT**

Let  $X_1, \dots, X_n \stackrel{iid}{\sim} P_\mu$ , where  $E(X_i) = \mu$  and  $Var(X_i) = \sigma^2$  for all  $i = 1, 2, \dots, n$

Weak and strong law of large numbers:

$$\overline{X}_n = \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow[n \rightarrow \infty]{P, a.s.} \mu.$$

$$\frac{1}{n} \sum_{i=1}^n g(X_i) \xrightarrow[n \rightarrow \infty]{P, a.s.} E[g(X)]$$

Central Limit Theorem:

$$\sqrt{(n)} \frac{\overline{X}_n - \mu}{\sqrt{(\sigma^2)}} \xrightarrow[n \rightarrow \infty]{(d)} N(0, 1)$$

$$\sqrt{(n)} (\overline{X}_n - \mu) \xrightarrow[n \rightarrow \infty]{(d)} N(0, \sigma^2)$$

**6 Statistical models**

**7 Estimators**

**8 Confidence intervals**

**Onesided**

**Twosided**

**Delta Method**

**9 Hypothesis tests**

**Onesided**

**Twosided**

**P-Value**

**10 Distance between distributions**

**Total variation**

The total variation distance TV between the propability measures  $P$  and  $Q$  with a sample space  $E$  is defined as:

$$TV(\mathbf{P}, \mathbf{Q}) = \max_{A \subseteq E} |P(A) - Q(A)|,$$

Calculation with  $f$  and  $g$ :

$$TV(\mathbf{P}, \mathbf{Q}) = \begin{cases} \frac{1}{2} \sum_{x \in E} |f(x) - g(x)|, & \text{discr} \\ \frac{1}{2} \int_{x \in E} |f(x) - g(x)| dx, & \text{cont} \end{cases}$$

Symmetry:  
 $d(\mathbf{P}, \mathbf{Q}) = d(\mathbf{Q}, \mathbf{P})$   
nonnegative:  
 $d(\mathbf{P}, \mathbf{Q}) \geq 0$   
definite:  
 $d(\mathbf{P}, \mathbf{Q}) = 0 \iff \mathbf{P} = \mathbf{Q}$   
triangle inequality:  
 $d(\mathbf{P}, \mathbf{V}) \leq d(\mathbf{P}, \mathbf{Q}) + d(\mathbf{Q}, \mathbf{V})$   
If the support of  $\mathbf{P}$  and  $\mathbf{Q}$  is disjoint:  
 $d(\mathbf{P}, \mathbf{V}) = 1$   
TV between continuous and discrete r.v:  
 $d(\mathbf{P}, \mathbf{V}) = 1$

**KL divergence**

the KL divergence (also known as relative entropy) KL between between the propability measures  $P$  and  $Q$  with the common sample space  $E$  and pmf/pdf functions  $f$  and  $g$  is defined as:

$$KL(\mathbf{P}, \mathbf{Q}) = \begin{cases} \sum_{x \in E} p(x) \ln \left( \frac{p(x)}{q(x)} \right), & \text{discr} \\ \int_{x \in E} p(x) \ln \left( \frac{p(x)}{q(x)} \right) dx, & \text{cont} \end{cases}$$

Not a distance!  
Asymetric in general:  
 $KL(\mathbf{P}, \mathbf{Q}) \neq KL(\mathbf{Q}, \mathbf{P})$   
Nonnegative:  
 $KL(\mathbf{P}, \mathbf{Q}) \geq 0$

Define:  
if  $\mathbf{P} = \mathbf{Q}$  then  $KL(\mathbf{P}, \mathbf{Q}) = 0$   
Does not satisfy triangle inequality in general:  
 $KL(\mathbf{P}, \mathbf{V}) \not\leq KL(\mathbf{P}, \mathbf{Q}) + KL(\mathbf{Q}, \mathbf{V})$

Estimator of KL divergence:  
 $KL(\mathbf{P}_{\theta_*}, \mathbf{P}_{\theta}) = const - E[\ln(p_{\theta}(X_i))]$

$\widehat{KL}(\mathbf{P}_{\theta_*}, \mathbf{P}_{\theta}) = const - \frac{1}{n} \sum_{i=1}^n \log(p_{\theta}(X_i))$

**11 Likelihood**

Let  $(E, \{P_{\theta}\}_{\theta \in \Theta})$  denote a discrete or continuous statistical model. Let  $p_{\theta}$  denote the pmf or pdf of  $P_{\theta}$ . Let  $X_1, \dots, X_n \stackrel{iid}{\sim} P_{\theta^*}$  where the parameter  $\theta^*$  is unknown. Then the likelihood is the function

$$L_n : E^n \times \Theta$$

$$(x_1, \dots, x_n, \theta)$$

$$L_n(x_1, \dots, x_n, \theta) = \prod_{i=1}^n P_{\theta}[X_i = x_i]$$

**Discrete Likelihood**

Bernoulli likelihood:

$$L_n(x_1, \dots, x_n, p) = p^{\sum_{i=1}^n x_i} (1 - p)^{n - \sum_{i=1}^n x_i}$$

$$= L_n = \prod_{i=1}^n (x_i p + (1 - x_i)(1 - p))$$

Poisson likelihood:

$$L_n(x_1, \dots, x_n, \lambda) = \prod_{i=1}^n \frac{\lambda^{\sum_{i=1}^n x_i}}{\prod_{i=1}^n x_i!} e^{n\lambda}$$

Poisson loglikelihood:

$$\log(L(x_1 \dots x_n; \lambda)) = -n\lambda + \log(\lambda)(\sum_{i=1}^n x_i) - \log(\prod_{i=1}^n x_i!)$$

**Continuous Likelihood**

Gaussian likelihood:

$$L(x_1 \dots x_n; \mu, \sigma^2) = \frac{1}{(\sigma \sqrt{2\pi})^n} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2\right)$$

Gaussian loglikelihood:

$$\log(L(x_1 \dots x_n; \mu, \sigma^2)) = -n \log(\sigma \sqrt{2\pi}) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$$

Gaussian Maximum-loglikelihood estimators:

$$\text{MLE estimator for } \sigma^2 = \tau:$$

$$\hat{\tau}_n^{MLE} = \frac{1}{n} \sum_{i=1}^n X_i^2$$

MLE estimators:

$$\hat{\mu}_n^{MLE} = \frac{1}{n} \sum_{i=1}^n (x_i)$$

Exponential likelihood:

$$L(x_1 \dots x_n; \lambda) = \lambda^n \exp\left(-\lambda \sum_{i=1}^n x_i\right)$$

Uniform:

$$L(x_1 \dots x_n; b) = \frac{1(\max_i(x_i \leq b))}{b^n}$$