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1 Objective: Bootstrap Resampling Method

We use statistical sampling to determine the value of a parameter of a population. We sample a population, measure a statistic of this sample, and then use this statistic to say something about the corresponding parameter of the population. Bootstrapping is a resampling method with replacement from the initial sample. Each iteration of bootstrap improves the order of accuracy for the parameter in question.

The name bootstrap is a metaphor for "a self-sustaining process that proceeds without external help"

Our goal is to produce graphs like those of Mammen's paper, distributed in class, when sampling from two populations:

2 Key Words

rchisq, quintile, sample, t, apply, cbind, pnorm, quantiles, plain bootstrap, wild bootstrap, bias.

3 Brief Description

Consider the following two populations:

- POP1 : CHI 8: X has chi-squared distribution, df = 8, so E[X] = 8, V[X] = 16. where df is degrees of freedom, E[X] = df is expected value, and V[X] = 2(df) is variance.
- POP2: EXP 1: X has exponential distribution, $\lambda = 1$, E[X] = 1, V[X] = 1. where λ is the rate, $E[X] = \lambda^{-1}$ is expected value, and $V[X] = \lambda^{-2}$ is variance.

The statistic we'll look at is **sample mean**. There are **two pivots** of interest (notation is given below):

Difference $Q = \widehat{M} - M$ and studentized difference $Q_s = (\widehat{M} - M)/\widehat{D}$, and the corresponding bootstrap resampling pivots Q^* and Q_s^* . The claim of bootstrap is BEHAVIOR OF Q^* MIMICS THAT OF Q. Mammen's graphs are an attempt to prove this claim.

Read the notation and definitions as needed from Mammen's.

4 Notation

• POP: P; mean: M = M(P); var: V = V(P); dev: $D = \sqrt{V(P)}$.

SAMPLE: $S; n = \text{ns} = \text{number in sample}; \text{mean: } \widehat{M} = M(S); \text{var: } \widehat{V} = V(S) \neq V; \text{ dev: } \widehat{D} = D(\widehat{M}) = \sqrt{V(S)/n} \neq \sqrt{V/n}$

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• BOOTSTRAP: given $S = \{X_1, \dots, X_n\}$ (assumes independence)

SAMPLE: S^* : simple random sample, with replacement, of size ns from S; mean: $M^* = M(S^*)$; dev: $D^* = \sqrt{V(S^*)/n}$.

• WILDBOOTSTRAP: given $S = \{X_1, \ldots, X_n\}$. Wildbootstrap resamples from S differently, by choosing $a_1 \ x_1$'s, $a_2 \ x_2$'s, ... where a_1, a_2, \ldots are independent Poisson random variables (assumes heteroscedasticity):

```
SAMPLE: a_k = \text{ak iid POISSON}(1); MW = \text{mean}(a_k(X_k - \widehat{M})); DW = \sqrt{\text{Var}(a_k(X_k - \widehat{M}))}.
```

[A] QUANTILES: The first step is to use MC = Monte Carlo simulation to estimate the sequence of 5% quantiles uk and vk for the two pivots:

$$P\left[\widehat{M}-M\leq \mathsf{uk}\right]=k/20, \qquad \qquad P\left[(\widehat{M}-M)/\widehat{D}\leq \mathsf{vk}\right]=k/20, \quad k \ \ \text{in} \ \ 1:19$$

respectively.

[1] Fix a function q.chi8 with input ns and nq = number of quantile simulations, and output a vector of 5% quantiles:

[2] Fix a function qstu.chi8 like q.chi8 but with df replaced by dfs and

```
dfs[k]<-(mean(s) - 8)/(sqrt(var(s)/ns))
```

- [3] Fix two more functions q.expl and qstu.expl that work for the population expl.
- [4] Below we'll need uk.8 <- q.chi8(20,1000), and similarly for uks.8 <- qstu.chi8(20,1000). In the same fashion, generate uk.1, and uks.1.

Note: Mammen used nq = 16,000 rather than 1000.

- [B] PLAIN BOOTSTRAP: The second step is to use MC to estimate BIAS and MSE.
 - [1] Fix a function bs.p.chi8 with inputs: ns, nr = number of resamples from S, nb = number of bootstrap samples of different S and vector of quantiles uu, and output: matrix of bias and mse = expected squared error. These involve

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```
bs.p.chi8 <- function(ns, nr, nb, uu)
s <- double(ns)
                                                         # S
 ss <- double(ns)
                                                         # S*
# observed pk(S) across S's
for(i in 1:nb) {
        s <- rchisq(ns, 8)
                                                         # get S
        ms <- mean(s)</pre>
                                                         # M^
        fr <- rep(0, 19)
                                        # place for frequencies, replicates 0 19-times
        for (j in 1:nr) {
                ss <- sample(s, ns, rep = T)</pre>
                                                         # get S*
                dm <- mean(ss) - ms
                                                         # get M* - M^
                fr <- fr + (dm < uu)
                                                         # frequencies: WHY WORK?
        pp[i, ] <- fr / nr</pre>
                                                         \# pk(S), k = 1:19
}
err < - t(pp) - 1:19/20
                                        # t:=transpose func., pk (S)-k/20, k=1:19
bias <- apply(err, 1, mean)</pre>
mse <- apply(err^2, 1, mean)</pre>
cbind(bias, mse)
}
```

[2] Use the function to get plot data

```
bsp8 <- bs.p.chi8(20,50,100,uk.8) # uk.8 from [A4]
```

NOTE: Mammen used nr = 1,000 & nb = 10,000 rather than 50 & 100.

- [3] Make plots of bias and mse on y-axis versus xax<-1:19/20 on x-axis. They should look like Mammen. SHOW ME SCREEN OUTPUT FOR LAB CREDIT.
- [C] NORMAL APPROXIMATION: By the CLT. Distribution $(\widehat{M}-M)/\widehat{D}\neq N(0,1),\ Z:=N(0,1)$ random variable, so $k/20=P(\widehat{M}-M\leq \mathsf{uk})\neq P(\widehat{D}\ Z\leq \mathsf{uk})=P(Z\leq \mathsf{uk}/\widehat{D})$

Modify the outer loop and omit the inner loop of bs.p.chi8 to give normal approximation estimates bias & mse . Fix the function bs.na.chi8:

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