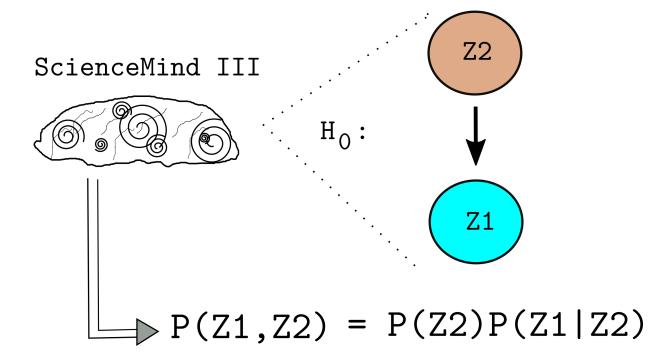
Probability

A Practical, Systems-Based Approach

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Z1={observed system behavior}
Z2={possible causally-linked data}



Preface

This book is dedeicated to hours of NOT understanding probability theory.

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Chapter 1

Introduction

Introduction to Probabi; lity Models, 10th edition by Sheldon Ross

1.1 Ross Examples

Suppose that n independent trials, each of which results in any of m possible outcomes with respective probabilities p1,..., pm, $\sum pi = 1$, are continually performed. Let X denote the number of trials needed until each outcome has occurred at least once. Define the equation for X=n and explain the reasoning.

Let outcomes be $1, \ldots, m$ with probabilities p_1, \ldots, p_m (independent across trials). For $n < m, \mathbb{P}(X = n) = 0$. For $n \ge m$

$$\mathbb{P}(X = n) = \sum_{\varnothing \neq S \subseteq \{1, ..., m\}} (-1)^{|S|+1} \Big(\sum_{j \in S} p_j\Big) \Big(1 - \sum_{j \in S} p_j\Big)^{n-1}$$

Why this is true (reasoning)

X=n means: after n-1 trials at least one outcome is missing, and on trial n the **last** missing type appears for the **first** time. Equivalently, pick any nonempty set S of outcomes and consider the event that **all** outcomes in S have been missing up to time n-1 and one of them appears at time n. The probability that no outcome from S appears in a given trial is $1-\sum_{j\in S}p_j$. So the chance they are all missing for the first n-1 trials and then one of them appears on trial n is

$$\left(1 - \sum_{j \in S} p_j\right)^{n-1} \left(\sum_{j \in S} p_j\right).$$

* But these events for different S overlap (inclusion–exclusion fixes the overcount), giving the alternating-sum formula above.

Coupon Collector Cheat Sheet (General Probabilities)

Problem: m outcomes with probabilities $p_1, \ldots, p_m > 0$, $\sum p_i = 1$. X = number of trials until all outcomes appear at least once. Define $q_S = \sum_{j \in S} p_j$ for subset $S \subseteq \{1, \ldots, m\}$.

Core Formulas

Quantity	Formula
$\overline{\mathrm{PMF}\ (n \geq m)}$	$\mathbb{P}(X=n) = \sum_{\varnothing \neq S} (-1)^{ S +1} q_S (1-q_S)^{n-1}$
CDF	$\mathbb{P}(X \le n) = \sum_{S}^{\infty \ne S} (-1)^{ S } (1 - q_S)^n$
Expectation	$\mathbb{E}[X] = \sum_{\varnothing \neq S} (-1)^{ S +1} \frac{1}{q_S}$
Minimum trials Sanity Check	$X_{\min} = m$ $\mathbb{P}(X = m) = m! \prod p_i; \sum_{n \ge m} \mathbb{P}(X = n) = 1$

Quick Insights

- Tail behavior dominated by largest $1 q_{\{i\}} = 1 \min p_i$; decays geometrically.
- Uniform probs: $p_i = 1/m \Rightarrow \mathbb{E}[X] = mH_m \approx m(\ln m + \gamma)$.
- Complexity: exact computation $O(2^m)$; feasible for $m \lesssim 20$.
- Simulation: use Monte Carlo for large m.
- First m trials all distinct probability: $m! \prod p_i$.

Algorithm (Bitmask / Combinations)

- 1. Enumerate all non-empty subsets S (bitmask or itertools.combinations).
- 2. Compute $q_S = \sum_{j \in S} p_j$.
- 3. Plug into PMF or expectation formulas.
- 4. Verify probabilities sum to 1 within tolerance.

Python Helper

```
from itertools import combinations
def coupon_collector_pmf(p, n):
    m = len(p)
    prob = 0.0
    idx = list(range(m))
    for r in range (1, m+1):
        for subset in combinations(idx, r):
            q = sum(p[i] for i in subset)
            prob += ((-1)**(r+1)) * q * (1-q)**(n-1)
    return prob
def coupon_collector_expectation(p):
    m = len(p)
    exp_val = 0.0
    idx = list(range(m))
    for r in range(1, m+1):
        for subset in combinations(idx, r):
            q = sum(p[i] for i in subset)
            exp_val += ((-1)**(r+1)) * (1/q)
    return exp_val
# Example usage:
p = [0.2, 0.3, 0.5]
for n in range (3, 8):
    print(f"P(X={n})_=_{coupon\_collector\_pmf(p,_n):.5f}")
print("E[X]_=", coupon_collector_expectation(p))
```

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