

## Chapter 2 Problems

2.2. Find a general inequality for  $H(X)$  vs  $H(Y)$ . let  $y = g(x)$ .  $p(y) = \sum_{x:y=g(x)} p(x)$  Meaning sum over all  $x$  given that  $y = g(x)$ . It is pretty clear that if there is a 1-1 mapping then  $p(x) = p(y)$ .

Consider a set of  $x$  that maps to a single  $y$ . For this set  $\sum_{x:y=g(x)} p(x) \log p(x) \leq \sum_{x:y=g(x)} p(x) \log p(y) = p(y) \log p(y)$  since  $p(y) \geq p(x)$ . Now extending the argument to the entire range of  $X$  and  $Y$  we obtain:

$$\begin{aligned} p(x) &\leq \sum_{x:y=g(x)} p(x) = p(y) \\ H(X) &= - \sum_x p(x) \log p(x) \\ &= - \sum_y \sum_{x:y=g(x)} p(x) \log p(x) \\ &\geq - \sum_y p(y) \log p(y) \\ &= H(Y) \end{aligned}$$

a)  $H(X)$  vs  $H(Y = 2^X)$

$Y$  and  $X$  are 1-1 so we have the special case of equality.

b)  $H(X)$  vs  $H(Y = \cos(X))$

There are multiple values of  $X$  that could map to single values of  $Y$  with this function. Thus we can only guarantee the general case where  $H(X) \geq H(Y)$

2.4. Show that the entropy of a function of  $X$  is  $\leq$  than the entropy of  $X$  by justifying the following steps.

a)  $H(X, g(X)) = H(X) + H(g(X)|X)$

This is justified by the chain rule.

b)  $= H(X)$

$H(g(X)|X) = 0$  because  $g(X)$  is known for any value of  $X$ . Thus  $H(g(X)|X) = \sum_x p(x) H(g(X)|X = x) = \sum_x 0 = 0$  leaving  $H(X)$

c)  $H(X, g(X)) = H(g(X)) + H(X|g(X))$

Just a different application of the chain rule.

d)  $\geq H(g(X))$

$H(X|g(X)) \geq 0$  and is equal to 0 when  $g(X)$  gives an unambiguous mapping to one  $X$ . Otherwise it will be greater than 0. Thus  $H(g(X)) \leq H(X, g(X))$

2.10. We have a set

$$X = \begin{cases} X_1 & \text{with probability } \alpha \\ X_2 & \text{with probability } (1 - \alpha) \end{cases}$$

a) What is  $H(X)$  in term of  $H(X_1), H(X_2), \alpha$ ?

$$\theta = f(X) = \begin{cases} 1 & \text{when } X = X_1 \\ 2 & \text{when } X = X_2 \end{cases}$$

$$\begin{aligned} H(X) &= H(X, f(X)) = H(\theta) + H(X|\theta) \\ &= H(\theta) + p(\theta = 1) H(X|\theta = 1) + p(\theta = 2) H(X|\theta = 2) \\ &= H(\alpha) + \alpha H(X_1) + (1 - \alpha) H(X_2) \end{aligned}$$

where  $H(\alpha) = -\alpha \log \alpha - (1 - \alpha) \log (1 - \alpha)$

2.12. We are given that  $p(x, y) = [\frac{1}{3}, \frac{1}{3}, 0, \frac{1}{3}]$  Find the following:

a)  $H(X), H(Y)$ 

$$\begin{aligned}
 H(X) &= H\left(\frac{2}{3}, \frac{1}{3}\right) \\
 &= \frac{2}{3} \log\left(\frac{3}{2}\right) + \frac{1}{3} \log(3) \\
 &\approx 0.9183 \text{ bits} \\
 H(Y) &= H\left(\frac{2}{3}, \frac{1}{3}\right) \\
 &\approx 0.9183 \text{ bits}
 \end{aligned}$$

b)  $H(X|Y), H(Y|X)$ 

$$\begin{aligned}
 H(X|Y) &= \frac{1}{3} H(X|Y=0) + \frac{2}{3} H(X|Y=1) \\
 &= \frac{1}{3} H\left(\frac{1}{2}, \frac{1}{2}\right) + \frac{2}{3} H(0, 1) \\
 &= \frac{1}{3} 1 + \frac{2}{3} 0 \\
 &= \frac{1}{3} \text{ bits} \\
 H(Y|X) &= \frac{2}{3} H(Y|X=0) + \frac{1}{3} H(Y|X=1) \\
 &= \frac{2}{3} H\left(\frac{1}{2}, \frac{1}{2}\right) + \frac{1}{3} H(0, 1) \\
 &= \frac{2}{3} \text{ bits}
 \end{aligned}$$

c)  $H(X, Y)$ 

$$\begin{aligned}
 H(X, Y) &= H(X) + H(Y|X) \\
 &\approx 1.585 \text{ bits}
 \end{aligned}$$

d)  $H(Y) - H(Y|X)$ 

$$\begin{aligned}
 H(Y) - H(Y|X) \\
 &\approx 0.251 \text{ bits}
 \end{aligned}$$

e)  $I(X; Y)$ 

$$\begin{aligned}
 I(X; Y) &= H(Y) - H(Y|X) \\
 &\approx 0.251 \text{ bits}
 \end{aligned}$$

f) See [Figure 1](#) which shows the above equations on a Venn diagram.2.15. Markov chains:  $X_1 \rightarrow X_2 \rightarrow \dots \rightarrow X_n$ , have joint probabilities of the following form:

$$p(x_1, x_2, \dots, x_n) = p(x_1)p(x_2|x_1) \cdots p(x_n|x_{n-1})$$

This means that for mutual information we can do some nice reduction.

$$\begin{aligned}
 I(X_1; X_2, \dots, X_n) &= \\
 \text{let } X_3, \dots, X_n &= Y \\
 I(X_1; X_2, Y) &= I(X_1; X_2) + I(X_1; Y|X_2)
 \end{aligned}$$

Now since this is a markov chain, and we know that  $X_1$  and  $Y$  are completely independent we can say that  $I(X_1; Y|X_2) = 0$  since  $Y|X_2$  is just a subset of  $Y$  which as noted previously is entirely independent of  $X_1$ . Thus the whole equation simplifies to  $I(X_1; X_2)$

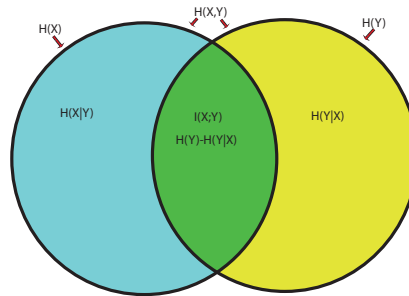


Figure 1: 2.12 Part e Venn Diagram

- 2.26. a)  $\ln(x) \leq x - 1$  for  $0 \leq x \leq \infty$  As  $x$  approaches 0  $x - 1$  approaches -1 while  $\ln(x)$  approaches  $-\infty$ .  $x - 1$  grows linearly while  $\ln(x)$  grows much slower, logarithmically, so for large numbers it is given that  $x - 1$  is larger.

From the derivative of  $x - 1$  and  $\ln(x)$  we know that  $x - 1$  has a slope of 1 for all  $x$ , while  $\ln(x)$  has slope  $\frac{1}{x}$ . This means that when  $x < 1$ ,  $\ln(x)$  has an increasingly steep negative slope, more steeply negative than 1, and when  $x > 1$  it has an increasingly shallow positive slope. The only point at which their slopes are equal is  $x = 1$ , and since  $\ln(1) = 0 = 1 - 1$  we can say that  $(x - 1) \geq \ln(x)$ .

- b) Since  $D(p||q) = \sum_x p(x) \ln \left( \frac{p(x)}{q(x)} \right)$  that means that  $-D(p||q) = -\sum_x p(x) \ln \left( \frac{p(x)}{q(x)} \right)$  which is equal to the reciprocal of the log, giving us:  $\sum_x p(x) \ln \left( \frac{q(x)}{p(x)} \right)$

Let  $\frac{q(x)}{p(x)}$  be  $x$ , as shown before  $x - 1 \geq \ln(x)$  which justifies the second step.

The final step is fairly straight forward as well. In the equation  $\sum_x p(x) \left( \frac{q(x)}{p(x)} - 1 \right)$  the  $p(x)$  terms cancel leaving  $\sum_x q(x) - 1$ . Since  $q(x)$  is a probability, we know that over all  $x$  it will sum to 1 meaning that the equation simplifies to 0.

- c) To equal 0, I have just shown that  $\sum_x p(x) \left( \frac{q(x)}{p(x)} - 1 \right) = 0$  additionally, in my argument for part a, I showed that  $\ln(x) = x - 1$  only when  $x = 1$ . Thus when  $\frac{q(x)}{p(x)} = 1$  (when  $q(x) = p(x)$ ) the value of  $D$  will be 0.

- 2.32. See attached sheet.