

# Multiscale occupancy model that accommodates false detections

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## 1 DESCRIPTION

A 3-level multiscale occupancy model that accommodates false detections. A secondary “negative control” data set is used for estimating the probability of a false detection.

## 2 IMPLEMENTATION

The file `occ.multiscale.fp.sim.R` simulates data according to the model statement presented below, and `occ.multiscale.fp.mcmc.R` contains the MCMC algorithm for parameter estimation.

## 3 MODEL STATEMENT

Let  $y_{ijk}$  be a binary observation representing a detection/non-detection, where  $i = 1, \dots, N$  indexes primary sample unit (e.g., individual or site),  $j = 1, \dots, J_i$  indexes subunits (i.e., a secondary unit nested within the primary sample unit), and  $k = 1, \dots, K_{ij}$  indexes replicate observations within a subunit. Also let  $v$  be the number of detections out of  $n$  “trials” obtained in the “negative control” dataset.

$$\begin{aligned} y_{ijk} &\sim \begin{cases} (1 - \phi) \text{Bern}(p_{ijk}) + \phi 1_{\{y_{ijk}=1\}}, & a_{ij} = 1 \\ \text{Bern}(\phi), & a_{ij} = 0 \end{cases} \\ a_{ij} &\sim \begin{cases} \text{Bern}(\theta_{ij}), & z_i = 1 \\ 0, & z_i = 0 \end{cases} \\ z_i &\sim \text{Bern}(\psi_i) \\ \psi_i &= \text{logit}^{-1}(\mathbf{x}'_i \boldsymbol{\beta}) \\ \theta_{ij} &= \text{logit}^{-1}(\mathbf{u}'_{ij} \boldsymbol{\gamma}) \\ p_{ijk} &= \text{logit}^{-1}(\mathbf{w}'_{ijk} \boldsymbol{\alpha}) \\ \boldsymbol{\beta} &\sim \mathcal{N}(\boldsymbol{\mu}_\beta, \sigma_\beta^2 \mathbf{I}) \\ \boldsymbol{\gamma} &\sim \mathcal{N}(\boldsymbol{\mu}_\gamma, \sigma_\gamma^2 \mathbf{I}) \\ \boldsymbol{\alpha} &\sim \mathcal{N}(\boldsymbol{\mu}_\alpha, \sigma_\alpha^2 \mathbf{I}) \\ v &\sim \text{Binom}(M, \phi) \\ \phi &\sim \text{Beta}(a, b) \end{aligned}$$

## 4 FULL-CONDITIONAL DISTRIBUTIONS

### 4.1 Occupancy state ( $z_i$ )

$$\begin{aligned}
[z_i \mid \cdot] &\propto \prod_{j=1}^{J_i} [a_{ij} \mid \theta_{ij}, z_i] [z_i] \\
&\propto \prod_{j=1}^{J_i} \left( \theta_{ij}^{a_{ij}} (1 - \theta_{ij})^{1-a_{ij}} \right)^{z_i} \left( 1_{\{a_{ij}=0\}} \right) \psi_i^{z_i} (1 - \psi_i)^{1-z_i} \\
&\propto \prod_{j=1}^{J_i} \left( \psi_i \theta_{ij}^{a_{ij}} (1 - \theta_{ij})^{1-a_{ij}} \right)^{z_i} \left( (1 - \psi_i) 1_{\{a_{ij}=0\}} \right)^{1-z_i} \\
&= \text{Bern}(\tilde{\psi}_i),
\end{aligned}$$

where,

$$\tilde{\psi}_i = \frac{\psi_i \prod_{j=1}^{J_i} \theta_{ij}^{a_{ij}} (1 - \theta_{ij})^{1-a_{ij}}}{\psi_i \prod_{j=1}^{J_i} \theta_{ij}^{a_{ij}} (1 - \theta_{ij})^{1-a_{ij}} + (1 - \psi_i) \prod_{j=1}^{J_i} 1_{\{a_{ij}=0\}}}.$$

### 4.2 “Use” state ( $a_{ij}$ )

Note that the mixture specification for  $a_{ij}$  in the model statement above is equivalent to  $a_{ij} \sim \text{Bern}(z_i \theta_{ij})$ . The update for  $a_{ij}$  relies on this alternative specification.

$$\begin{aligned}
[a_{ij} \mid \cdot] &\propto \prod_{k=1}^{K_{ij}} [y_{ijk} \mid p_{ijk}, a_{ij}, \phi] [a_{ij}] \\
&\propto \prod_{k=1}^{K_{ij}} \left( (1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right)^{a_{ij}} \left( \phi^{y_{ijk}} (1 - \phi)^{1-y_{ijk}} \right) (z_i \theta_{ij})^{a_{ij}} (1 - z_i \theta_{ij})^{1-a_{ij}} \\
&\propto \prod_{k=1}^{K_{ij}} \left[ (z_i \theta_{ij}) \left( (1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right) \right]^{a_{ij}} \left( (1 - z_i \theta_{ij}) \phi^{y_{ijk}} (1 - \phi)^{1-y_{ijk}} \right)^{1-a_{ij}} \\
&= \text{Bern}(\tilde{\theta}_{ij}),
\end{aligned}$$

where,

$$\tilde{\theta}_{ij} = \frac{z_i \theta_{ij} \prod_{k=1}^{K_{ij}} \left[ (1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right]}{z_i \theta_{ij} \prod_{k=1}^{K_{ij}} \left[ (1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right] + (1 - z_i \theta_{ij}) \prod_{k=1}^{K_{ij}} \phi^{y_{ijk}} (1 - \phi)^{1-y_{ijk}}}.$$

### 4.3 Regression coefficients affecting occupancy probability ( $\beta$ )

$$\begin{aligned}
[\beta \mid \cdot] &\propto \prod_{i=1}^N [z_i \mid \psi_i] [\beta] \\
&\propto \prod_{i=1}^N \text{Bern}(z_i \mid \psi_i) \mathcal{N}(\beta \mid \mu_\beta, \sigma_\beta^2 \mathbf{I}).
\end{aligned}$$

The update for  $\beta$  proceeds using Metropolis-Hastings.

#### 4.4 Regression coefficients affecting probability of use ( $\gamma$ )

$$\begin{aligned}
[\gamma \mid \cdot] &\propto \prod_{i=1}^N \prod_{j=1}^{J_i} [a_{ij} \mid \theta_{ij}, z_i] [\gamma] \\
&\propto \prod_{i=1}^N \prod_{j=1}^{J_i} \text{Bern}(a_{ij} \mid \theta_{ij})^{z_i} \mathcal{N}(\gamma \mid \mu_\gamma, \sigma_\gamma^2 \mathbf{I}).
\end{aligned}$$

The update for  $\gamma$  proceeds using Metropolis-Hastings. Note that, in effect, the product over  $i$  only includes instances of  $i$  such that  $z_i = 1$ .

#### 4.5 Regression coefficients affecting detection probability ( $\alpha$ )

$$\begin{aligned}
[\alpha \mid \cdot] &\propto \prod_{i=1}^N \prod_{j=1}^{J_i} \prod_{k=1}^{K_{ij}} [y_{ijk} \mid p_{ijk}, a_{ij}] [\alpha] \\
&\propto \prod_{i=1}^N \prod_{j=1}^{J_i} \prod_{k=1}^{K_{ij}} \left[ (1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1 - y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right]^{a_{ij}} \mathcal{N}(\alpha \mid \mu_\alpha, \sigma_\alpha^2 \mathbf{I}).
\end{aligned}$$

The update for  $\alpha$  proceeds using Metropolis-Hastings. Note that, in effect, the product over  $i$  and  $j$  only includes instances of  $i$  and  $j$  such that  $a_{ij} = 1$ .

#### 4.6 Probability of a false detection ( $\phi$ )

$$\begin{aligned}
[\phi \mid \cdot] &\propto [v \mid M, \phi] [\phi] \\
&\propto \text{Binom}(v \mid M, \phi) \text{Beta}(a, b) \\
&\propto \phi^v (1 - \phi)^{M-v} \phi^{a-1} (1 - \phi)^{b-1} \\
&\propto \phi^{v+a-1} (1 - \phi)^{M-v+b-1} \\
&= \text{Beta}(v + a, M - v + b)
\end{aligned}$$