

Multiscale occupancy model that accommodates false detections

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1 DESCRIPTION

A 3-level multiscale occupancy model that accommodates false detections. A secondary “negative control” data set is used for estimating the probability of a false detection.

2 IMPLEMENTATION

The file `occ.multiscale.fp.sim.R` simulates data according to the model statement below, and `occ.multiscale.fp.mcmc.R` contains the MCMC algorithm for parameter estimation.

3 MODEL STATEMENT

Let y_{ijk} be a binary observation representing a detection/non-detection, where $i = 1, \dots, N$ indexes primary sample unit (e.g., individual or site), $j = 1, \dots, J_i$ indexes subunits nested within the primary sample unit, and $k = 1, \dots, K_{ij}$ indexes replicate observations within a subunit. Also let v be the number of detections out of M trials obtained in the “negative control” dataset.

$$\begin{aligned} y_{ijk} &\sim \begin{cases} (1 - \phi) \text{Bern}(p_{ijk}) + \phi 1_{\{y_{ijk}=1\}}, & a_{ij} = 1 \\ \text{Bern}(\phi), & a_{ij} = 0 \end{cases} \\ a_{ij} &\sim \begin{cases} \text{Bern}(\theta_{ij}), & z_i = 1 \\ 0, & z_i = 0 \end{cases} \\ z_i &\sim \text{Bern}(\psi_i) \\ \psi_i &= \text{logit}^{-1}(\mathbf{x}'_i \boldsymbol{\beta}) \\ \theta_{ij} &= \text{logit}^{-1}(\mathbf{u}'_{ij} \boldsymbol{\gamma}) \\ p_{ijk} &= \text{logit}^{-1}(\mathbf{w}'_{ijk} \boldsymbol{\alpha}) \\ \boldsymbol{\beta} &\sim \mathcal{N}(\boldsymbol{\mu}_\beta, \sigma_\beta^2 \mathbf{I}) \\ \boldsymbol{\gamma} &\sim \mathcal{N}(\boldsymbol{\mu}_\gamma, \sigma_\gamma^2 \mathbf{I}) \\ \boldsymbol{\alpha} &\sim \mathcal{N}(\boldsymbol{\mu}_\alpha, \sigma_\alpha^2 \mathbf{I}) \\ v &\sim \text{Binom}(M, \phi) \\ \phi &\sim \text{Beta}(a, b) \end{aligned}$$

4 FULL-CONDITIONAL DISTRIBUTIONS

4.1 Occupancy state (z_i)

$$\begin{aligned} [z_i \mid \cdot] &\propto \prod_{j=1}^{J_i} [a_{ij} \mid \theta_{ij}, z_i] [z_i] \\ &\propto \prod_{j=1}^{J_i} \left(\theta_{ij}^{a_{ij}} (1 - \theta_{ij})^{1-a_{ij}} \right)^{z_i} \left(1_{\{a_{ij}=0\}}^{1-z_i} \right) \psi_i^{z_i} (1 - \psi_i)^{1-z_i} \\ &\propto \prod_{j=1}^{J_i} \left(\psi_i \theta_{ij}^{a_{ij}} (1 - \theta_{ij})^{1-a_{ij}} \right)^{z_i} \left((1 - \psi_i) 1_{\{a_{ij}=0\}} \right)^{1-z_i} \\ &= \text{Bern}(\tilde{\psi}_i), \end{aligned}$$

where,

$$\tilde{\psi}_i = \frac{\psi_i \prod_{j=1}^{J_i} \theta_{ij}^{a_{ij}} (1 - \theta_{ij})^{1-a_{ij}}}{\psi_i \prod_{j=1}^{J_i} \theta_{ij}^{a_{ij}} (1 - \theta_{ij})^{1-a_{ij}} + (1 - \psi_i) \prod_{j=1}^{J_i} 1_{\{a_{ij}=0\}}}.$$

4.2 “Use” state (a_{ij})

Note that the mixture specification for a_{ij} in the model statement above is equivalent to $a_{ij} \sim \text{Bern}(z_i \theta_{ij})$, an alternate specification that simplifies the update for a_{ij} .

$$\begin{aligned} [a_{ij} \mid \cdot] &\propto \prod_{k=1}^{K_{ij}} [y_{ijk} \mid p_{ijk}, a_{ij}, \phi] [a_{ij}] \\ &\propto \prod_{k=1}^{K_{ij}} \left((1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right)^{a_{ij}} \left(\phi^{y_{ijk}} (1 - \phi)^{1-y_{ijk}} \right)^{1-a_{ij}} (z_i \theta_{ij})^{a_{ij}} (1 - z_i \theta_{ij})^{1-a_{ij}} \\ &\propto \prod_{k=1}^{K_{ij}} \left[(z_i \theta_{ij}) \left((1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right) \right]^{a_{ij}} \left((1 - z_i \theta_{ij}) \phi^{y_{ijk}} (1 - \phi)^{1-y_{ijk}} \right)^{1-a_{ij}} \\ &= \text{Bern}(\tilde{\theta}_{ij}), \end{aligned}$$

where,

$$\tilde{\theta}_{ij} = \frac{z_i \theta_{ij} \prod_{k=1}^{K_{ij}} \left[(1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right]}{z_i \theta_{ij} \prod_{k=1}^{K_{ij}} \left[(1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right] + (1 - z_i \theta_{ij}) \prod_{k=1}^{K_{ij}} \phi^{y_{ijk}} (1 - \phi)^{1-y_{ijk}}}.$$

4.3 Regression coefficients affecting occupancy probability (β)

$$\begin{aligned} [\beta \mid \cdot] &\propto \prod_{i=1}^N [z_i \mid \psi_i] [\beta] \\ &\propto \prod_{i=1}^N \text{Bern}(z_i \mid \psi_i) \mathcal{N}(\beta \mid \mu_\beta, \sigma_\beta^2 \mathbf{I}). \end{aligned}$$

The update for β proceeds using Metropolis-Hastings.

4.4 Regression coefficients affecting probability of use (γ)

$$\begin{aligned} [\gamma \mid \cdot] &\propto \prod_{i=1}^N \prod_{j=1}^{J_i} [a_{ij} \mid \theta_{ij}, z_i] [\gamma] \\ &\propto \prod_{i=1}^N \prod_{j=1}^{J_i} \text{Bern}(a_{ij} \mid \theta_{ij})^{z_i} \mathcal{N}(\gamma \mid \mu_\gamma, \sigma_\gamma^2 \mathbf{I}). \end{aligned}$$

The update for γ proceeds using Metropolis-Hastings. Note that the product over i only includes instances of i such that $z_i = 1$.

4.5 Regression coefficients affecting detection probability (α)

$$\begin{aligned} [\alpha \mid \cdot] &\propto \prod_{i=1}^N \prod_{j=1}^{J_i} \prod_{k=1}^{K_{ij}} [y_{ijk} \mid p_{ijk}, a_{ij}, \phi] [\alpha] \\ &\propto \prod_{i=1}^N \prod_{j=1}^{J_i} \prod_{k=1}^{K_{ij}} \left[(1 - \phi) p_{ijk}^{y_{ijk}} (1 - p_{ijk})^{1-y_{ijk}} + \phi 1_{\{y_{ijk}=1\}} \right]^{a_{ij}} \mathcal{N}(\alpha \mid \mu_\alpha, \sigma_\alpha^2 \mathbf{I}). \end{aligned}$$

The update for α proceeds using Metropolis-Hastings. Note that the product over i and j only includes instances of i and j such that $a_{ij} = 1$.

4.6 Probability of a false detection (ϕ)

$$\begin{aligned}
[\phi \mid \cdot] &\propto [v \mid M, \phi] [\phi] \\
&\propto \text{Binom}(v \mid M, \phi) \text{Beta}(a, b) \\
&\propto \phi^v (1 - \phi)^{M-v} \phi^{a-1} (1 - \phi)^{b-1} \\
&\propto \phi^{v+a-1} (1 - \phi)^{M-v+b-1} \\
&= \text{Beta}(v + a, M - v + b)
\end{aligned}$$